## Appendix A

Abundance and Total Numbers of Chinook Salmon and Trout in the Chiwawa River basin, Washington, 2015

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TO: HCP Hatchery Committee
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Subject: Abundance and Total Numbers of Chinook Salmon and Trout in the Chiwawa River basin, Washington, 2015

The Chelan County Public Utility District (PUD) hatchery program is operated through a habitat conservation program (HCP) that was incorporated into the PUD's license in 2004. The HCP directed the signatories to develop a monitoring and evaluation plan within one year of the effective date. This resulted in the development of the Conceptual Approach to Monitoring and Evaluating the Chelan County Public Utility District Hatchery Programs (Murdoch and Peven 2005). In 2013, the Hatchery Committees updated the hatchery monitoring and evaluation plan (Hillman et al. 2013). This study will help the Hatchery Committees determine if it is meeting Objective 2 in the updated monitoring and evaluation plan.

Objective 2: Determine if the proportion of hatchery fish on the spawning grounds affects the freshwater productivity of supplemented stocks.

We estimated densities and total numbers of age-0 spring Chinook salmon Oncorhynchus tshawytscha, trout Oncorhynchus sp., and char Salvelinus sp. in the Chiwawa River basin, Washington, in August and September 2015. This was the $23^{\text {rd }}$ year of an ongoing study to assess the freshwater productivity (juveniles/redd) of Chinook salmon in the Chiwawa River basin. We used landscape classification to stratify streams in the basin that supported juvenile Chinook salmon (Hillman and Miller 2004). Classification "explained" most of the variability in fish numbers caused by geology, land type, valley bottom type, stream state condition, and habitat type. We identified ten reaches on the lower 31 miles ( 50 km ) of the Chiwawa River and one reach in each of Phelps, Rock, Chikamin, Big Meadow, Alder, Brush, Clear, Y, and Unnamed ${ }^{1}$ creeks (Figure 1). Each reach consisted of several combinations of state-type and habitat-type strata. We used classification to find reference areas for reaches in the Chiwawa River. We matched Reach 3 and Reach 8 of the Chiwawa River with a moderately-confined section of Nason Creek (RM 0.62-1.70) and an unconfined area of the Little Wenatchee River (RM 4.39-

[^0]8.55), respectively (Hillman and Miller 2004). Because of the supplementation program in Nason Creek, the use of Nason Creek as a reference for the Chiwawa River is no longer valid. However, as directed by the Hatchery Committee, we continue to sample sites in Nason Creek. Following methods described in Hillman and Miller (2004), we used underwater observations to estimate numbers of fish in 199 randomly selected sites.
During sampling in August 2015, discharge in the Chiwawa River averaged 108 cubic feet per second (cfs) and ranged from 89-137 cfs (Figure 2). Stream temperatures during the study period ranged from 8.0 to $20.0^{\circ} \mathrm{C}$. Fish species observed in the Chiwawa River basin and reference areas during the 1992-2015 survey period ${ }^{2}$ included: spring Chinook salmon, coho salmon O. kisutch, sockeye salmon $O$. nerka, steelhead/rainbow trout $O$. mykiss (hatchery rainbow were present only in 1992 and 1993), cutthroat trout $O$. clarki lewisi, bull trout $S$. confluentus, brook trout $S$. fontinalis, mountain whitefish Prosopium williamsoni, dace Rhinichthys sp., northern pikeminnow Ptychocheilus oregonensis, suckers Catostomus sp., and sculpin Cottus sp. The age0 spring Chinook that we observed in the Chiwawa River basin during the 2015 survey were produced from 485 redds counted in the fall of 2014 (Hillman et al. 2015). Assuming a mean fecundity of 4,045 eggs per female Chinook (from females collected for broodstock), and that no female produced more than one redd (Murdoch et al. 2009), we estimated that the Chiwawa River basin was seeded with $1,961,825$ eggs in 2014 (Appendix A).

In 2015, riffles made up the largest fraction of habitat types in reaches of the Chiwawa River basin (53\% of the total stream surface area) (Table 1). Pools (24\%), glides (7\%), and multiple channels ( $16 \%$ ) constituted the remaining $47 \%$ of the stream surface area. We found woody debris associated with most multiple-channel habitat.

## Chinook Salmon Abundance

Chinook salmon were the most abundant salmonid in the Chiwawa River basin. We estimated, based on surface area, that age-0 Chinook salmon numbered 111,224 ( $\pm 7 \%$ of the estimated total) in the Chiwawa River basin in August 2015 (Table 2). Extrapolating based on volume of habitat types, age-0 Chinook numbered $97,358( \pm 7 \%)$ in the Chiwawa River basin. About 7\% of the juvenile Chinook were in tributaries to the Chiwawa River. During the 1992-2015 surveys, numbers of age-0 Chinook ranged from 5,815 to 149,563 in the Chiwawa River basin (Figure 3; Appendix A and B). Most of the difference in juvenile numbers among years resulted from different seeding (stock) levels (Figure 4). Numbers of Chinook redds in the Chiwawa River basin during 1992-2014 ranged from 13 to 1,078, resulting in seeding levels of 66,248 to 4,984,672 eggs (Appendix A).
As in most years, age-0 Chinook in 2014 were distributed contagiously among reaches in the Chiwawa River (Table 2). In the Chiwawa River, densities of age-0 Chinook were highest in the upper reaches (Reaches 7-10). The highest densities in the Chiwawa River basin were in tributaries to the Chiwawa River (Table 2). Age-0 Chinook were most abundant in multiple channels and least abundant in glides and riffles. We found the majority of the Chinook

[^1]associated with woody debris in multiple channels (multiple channel use index $=2.80$ ) ${ }^{3}$. These sites (multiple channels) made up $16 \%$ of the total surface area of the Chiwawa River basin, but they provided habitat for $63 \%$ of all the age-0 Chinook in the basin in 2015 (Appendix C). In contrast, riffles made up $53 \%$ of the total surface area, but provided habitat for only $5 \%$ of all age-0 Chinook in the Chiwawa River basin (riffle use index $=0.25$ ). Pools made up $24 \%$ of the total surface area and provided habitat for $31 \%$ of all age-0 Chinook in the basin (pool use index $=1.58$ ). Few Chinook used glides that lacked woody debris (glide use index $=0.26$ ).

As noted earlier, we assumed that the Chiwawa River was seeded with 1,961,825 Chinook eggs ( 485 redds times 4,045 eggs/female) in fall, 2014, and that at least 111,224 of those survived to August 2015. This means that the egg-to-parr survival was at least $5.7 \%$ ( $95 \%$ confidence bound 5.2-6.1\%). During 1992-2015, egg-to-parr survival averaged $8.1 \%$ (range $2.7-19.1 \%$ ) in the Chiwawa River basin (Appendix A). This survival rate comports with those from other streams. For example, Mullan et al. (1992) estimated an egg-to-parr survival rate of $9.8 \%$ for spring Chinook salmon in Icicle Creek, a tributary of the Wenatchee River. Using a Beverton and Holt model, Hubble (1993) estimated that egg-to-parr survival of Chinook in the Chewuck River, a tributary to the Methow River, ranged between $13 \%$ and $32 \%$, depending on percent seeding level in the basin. Kiefer and Forster (1991) estimated a mean egg-to-parr survival rate of 5.5\% (range 5.1-6.7\%) for naturally-spawning spring Chinook salmon in the entire upper Salmon River. They also noted that egg-to-parr survival of natural spawners and adult outplants in the headwater streams of the upper Salmon River averaged 24.4\% (range 16.1-32.0\%). Petrosky (1990) reported an egg-to-parr survival range of 1.2-29.0\% for Chinook in the upper Salmon River, Idaho. Konopacky et al. (1986) estimated egg-to-parr survival of Chinook in Bear Valley Creek, Idaho, as 8.1-9.4\%. Work by Richards and Cernera (1987) in Bear Valley Creek indicated an egg-to-parr survival of $2.1 \%$.
Mean densities of age-0 Chinook salmon in two reaches of the Chiwawa River were generally less than those in corresponding reference areas (Figure 5). Within both the Chiwawa River and its reference areas, pools and multiple channels consistently had the highest densities of age-0 Chinook.

We estimated a total of 620 ( $\pm 43 \%$ of the estimated total) age- $1+$ Chinook salmon in the Chiwawa River basin in August 2015 (Table 3). In August 1992-2015, numbers of age-1+ Chinook ranged from 5 to 967 in the Chiwawa River basin (Figure 3; Appendix B). These fish occurred throughout the Chiwawa River. We found relatively few age-1+ Chinook in tributaries; although, numbers in Rock Creek were higher in 2015 than in past years. Age-1+ Chinook were most abundant in multiple channels and pools.

[^2]
## Juvenile Chinook Salmon Productivity (Fish/Redd)

Freshwater productivity of juvenile Chinook salmon was estimated as the number of parr (age-0 Chinook) per redd in the Chiwawa River basin. Theoretically, the relationship between number of parr and redds can be explained mathematically provided the relationship between the two parameters goes through the origin, increases monotonically at low spawning levels, and shows some level of density dependence at high spawning levels. We identified four alternative hypotheses that may explain the relationship between spawning level (redds) and numbers of age-0 Chinook:

1. The first hypothesis assumed that the number of juveniles increases constantly toward an asymptote as the number of redds increases. After the asymptote is reached, the number of juveniles neither increases nor decreases. The asymptote represents the maximum number of juveniles the system can support (i.e., carrying capacity for the system). This hypothesis was modeled with a Beverton-Holt curve that took the form:

$$
J=\frac{(\alpha R)}{(\beta+R)}
$$

where $\boldsymbol{J}$ is the number of juvenile (age- 0 ) Chinook, $\boldsymbol{R}$ is the number or redds, $\boldsymbol{\alpha}$ is the maximum number of juveniles produced, and $\boldsymbol{\beta}$ is the number of redds needed to produce (on average) juveniles equal to one-half the maximum number of juveniles.
2. The second hypothesis, like the first, assumed that the number of juveniles increases toward an asymptote (carrying capacity) as the number of redds increases. After the carrying capacity is reached, the number of juveniles neither increases nor decreases. The carrying capacity represents the maximum number of juveniles the system can support. This hypothesis was modeled with a smooth hockey stick function that took the form:

$$
J=J_{\infty}\left(1-e^{-\left(\frac{\alpha}{J_{\infty}}\right) R}\right)
$$

where $\boldsymbol{J}$ and $\boldsymbol{R}$ are as above, $\boldsymbol{\alpha}$ is the slope at the origin of the spawner-recruitment curve, and $\boldsymbol{J}_{\infty}$ is the carrying capacity of juveniles.
3. The third hypothesis assumed that the number of juveniles increases to a maximum and then declines as the number or redds increases. In this case, mortality rate of juveniles (or eggs) is proportional to the initial number of redds. Higher mortality rate is associated with density-dependent growth coupled with size-dependent predation. This hypothesis was modeled with a Ricker curve that took the form:

$$
J=\alpha R e^{-\beta R}
$$

where $\boldsymbol{J}$ and $\boldsymbol{R}$ are as above, $\boldsymbol{\alpha}$ is the number of juveniles per redd at low spawning levels, and $\boldsymbol{\beta}$ describes how quickly the juveniles per redd drop as the number of redds increases.
4. The fourth hypothesis, like the first, assumed that the number of juveniles increases constantly, but unlike the first, the number of juveniles does not reach an asymptote. Rather, the number of juveniles increases indefinitely, but at a slowing rate of increase. This hypothesis was modeled with both a Cushing curve and a Gamma function. The

Cushing curve took the form:

$$
\boldsymbol{J}=\boldsymbol{\alpha} \boldsymbol{R}^{\gamma}
$$

where $\boldsymbol{J}$ and $\boldsymbol{R}$ are as above, $\boldsymbol{\alpha}$ is the number of juveniles per redd at low spawning levels, and $\gamma$ describes the level of density dependence at high spawning levels. The Gamma function is a three-parameter model that has the form:

$$
J=\alpha R^{\gamma} e^{-\beta R}
$$

This is an un-normalized gamma function that is similar to the Cushing curve when $\beta=0$.
We used Akaike's Information Criterion for small sample size ( $\mathrm{AIC}_{\mathrm{c}}$ ) to determine which model(s) best explained the productivity of juvenile Chinook in the Chiwawa River basin. AIC ${ }_{c}$ was estimated as:

$$
A I C_{c}=-2 \log (£(\theta \mid \text { data }))+2 K+\left(\frac{2 K(K+1)}{n-K-1}\right)
$$

where $\log (\boldsymbol{£}(\boldsymbol{\theta} \mid$ data $))$ is the maximum likelihood estimate, $\boldsymbol{K}$ is the number of estimable parameters (structural parameters plus the residual variance parameter), and $\boldsymbol{n}$ is the sample size (Burnham and Anderson 2002). We used least-squares methods to estimate $\boldsymbol{\operatorname { l o g }}(\boldsymbol{£}(\boldsymbol{\theta} \mid \boldsymbol{d a t a})$ ), which was calculated as $\boldsymbol{\operatorname { l o g }}\left(\boldsymbol{\sigma}^{2}\right)$, where $\sigma^{2}=$ residual sum of squares divided by the sample size ( $\boldsymbol{\sigma}^{2}=\boldsymbol{R S S} / \boldsymbol{n}$ ). AIC $\mathrm{c}_{\mathrm{c}}$ assesses model fit in relation to model complexity (number of parameters). The model with the smallest $\mathrm{AIC}_{\mathrm{c}}$ value represents the "best approximating" model within the model set. Remaining models were ranked relative to the best model using AIC ${ }_{c}$ difference scores ( $\mathbf{\Delta} \mathbf{A I C} \mathbf{c}$ ), Akaike weights ( $\boldsymbol{w}_{\boldsymbol{i}}$ ), and evidence ratios. Models with $\boldsymbol{\Delta A I C} \mathbf{c}$ values less than 2 indicate that there is substantial support for these models as being the best-fitting models within the set (Burnham and Anderson 2002). Models with values greater than 2 have less support. Akaike weights are probabilities estimating the strength of the evidence supporting a particular model as being the best model within the model set. Models with small $\boldsymbol{w}_{\boldsymbol{i}}$ values are less plausible as competing models (Burnham and Anderson 2002). If no single model could be specified as the best model, a "best subset" of competing models was identified using (1) AIC ${ }_{c}$ differences to indicate the level of empirical support each model had as being the best model, (2) evidence ratios based on Akaike weights to indicate the relative probability that any model is the best model, and (3) coefficients of determination $\left(R^{2}\right)$ assessing the explanatory power of each model.

The use of $\mathrm{AIC}_{\mathrm{c}}$ indicated that the Beverton-Holt model best approximated the information in the juveniles/redd data (Table 4; Figure 6). The estimated structural parameters for this model were:

$$
\text { Juveniles }=\frac{(148,410 \times \text { Redds })}{(184+\text { Redds })}
$$

where the bootstrap estimated standard errors for the two parameters were 17,021 and 55 , respectively. The adjusted $R^{2}=0.84$. The second-best model was the smooth hockey stick model, which was $1.64 \mathrm{AIC}_{\mathrm{c}}$ units from the best model (Table 4; Figure 6). The estimated parameters for this model were:

$$
L N(\text { Juveniles })=11.6+L N\left(1-e^{-\left(\frac{723.8}{113,413}\right) R e d d s}\right)
$$

where the bootstrap estimated standard errors of the two parameters were 0.1 and 136, respectively, and the $R^{2}=0.83$. The $\mathrm{AIC}_{\mathrm{c}}$ difference scores, Akaike weights, and evidence ratios indicated that there was substantial support for both the Beverton-Holt and smooth hockey stick models (Table 4). There was less support for the remaining models (Ricker, Gamma ${ }^{4}$, and Cushing), which were $>2 \mathrm{AIC}_{c}$ units from the best models. This was further supported by the fact that, relative to the best models, the remaining models had evidence ratios greater than 10.
Although the Beverton-Holt, smooth hockey stick, and Ricker models have different biological assumptions, they all indicated a density-dependent relationship between spawning levels (redds) and juvenile Chinook production. This was not only evident in the best approximating models, but there was also a significant negative relationship between juveniles per redd and numbers of redds in the Chiwawa River basin (Figure 7). Although data at high seeding levels are lacking, the Beverton-Holt model would limit the capacity of juvenile Chinook to about 180,000 parr in the basin (bootstrap upper 95\% CI of $\boldsymbol{\alpha}$ in the Beverton-Holt model). This equates to about 1,621 Chinook parr per hectare. In contrast, the smooth hockey stick model, which fit the data as well as the Beverton-Holt model, would limit the carrying capacity for juvenile Chinook to about 140,000 parr (bootstrap upper $95 \%$ CI of $J_{\infty}$ in the smooth hockey stick model). This equates to about 1,261 Chinook parr per hectare. As a comparison, Thorson et al. (2013) estimated the carrying capacity for 15 populations of juvenile Chinook in the Snake River metapopulation as 5,000 juveniles per hectare. However, those authors noted that the estimate could be biased because of imperfect detectability and estimates of spawning numbers.

## Steelhead/Rainbow Abundance

Based on stream surface area, we estimated a total of 10,208 ( $\pm 11 \%$ of the estimated total) age-0 steelhead/rainbow ( $<4 \mathrm{in}$ ) in reaches of the Chiwawa River basin in August 2015 (Table 5). During the 1992-2015 survey period, numbers of age-0 steelhead/rainbow ranged from 1,410 to 45,727 in the Chiwawa River basin (Figure 8; Appendix B). In 1992-2015, numbers of age-0 steelhead/rainbow varied among reaches, but were typically highest in the lower reaches of the Chiwawa River. In all years they most often used riffle and multiple channel habitats in the Chiwawa River, although we also found them associated with woody debris in pool and glide habitat. In tributaries they were generally most abundant in small pools. Those that we observed in riffles selected stations in quiet water behind small and large boulders or occupied stations in quiet water along the stream margin. In pool and multiple-channel habitats, we found age-0 steelhead/rainbow using the same kinds of habitat as age-0 Chinook salmon.

We estimated that 754 ( $\pm 26 \%$ of the estimated total) age- $1+$ steelhead/rainbow ( $4-8$ in) lived in reaches of the Chiwawa River basin in August 2015 (Table 6). This was the lowest number of age- $1+$ steelhead/rainbow that we recorded during the more than 20 -year survey period. During the survey period 1992-2015, numbers of age-1+ steelhead/rainbow ranged from 754 to 22,130 (Figure 8; Appendix B). In most years we found these fish in nearly all reaches, but they were

[^3]typically most numerous in lower reaches of the Chiwawa River. We observed age-1+ steelhead/rainbow mostly in pool, riffle, and multiple-channel habitats. Those that we observed in pools were usually in deeper water than age- 0 steelhead/rainbow and Chinook. Like age-0 steelhead/rainbow, age-1+ steelhead/rainbow selected stations in quiet water behind boulders in riffles, but we generally did not find the two age groups together. Age-1+ steelhead/rainbow appeared to use deeper and faster water than did age-0 steelhead/rainbow.

We estimated that steelhead/rainbow larger than 8 inches numbered 18 ( $\pm 106 \%$ of the estimated total) in the Chiwawa River basin in August 2015 (Table 7). During the period 1992-2015, steelhead/rainbow numbers ranged from 8 to 1,869 (Appendix B). Steelhead/rainbow larger than 8 inches were most abundant in the lower Chiwawa River; however, in 1992 and 1993, they were most abundant near campgrounds in Reaches 8,9 , and 10 (these were mostly hatchery rainbow trout planted near the campgrounds). We found very few in tributaries. Most of the
steelhead/rainbow larger than 8 inches used deep pools ( $>5$ feet), and occupied stations near the bottom at the upstream end of pools.

## Bull Trout Abundance

We estimated, based on surface area that at least 239 ( $\pm 17 \%$ of the estimated total) juvenile (2-8 in) bull trout lived in reaches of the Chiwawa River basin in August 2015 (Table 8). We found most of these fish in the upper-most reaches of the Chiwawa River and in Rock Creek. During 1992-2015, numbers of juvenile bull trout ranged from 79 to 505 (Figure 9; Appendix B). These estimates and those for adult bull trout are incomplete because we did not sample the entire range of bull trout in all tributaries. That is, we did not extend our surveys into the headwaters of the Chiwawa River because there were no juvenile Chinook there. Areas beyond the distribution of juvenile Chinook salmon are known to support bull trout, steelhead/rainbow, and cutthroat trout (USFS 1993). In addition, our estimates of bull trout abundance were based on daytime snorkel surveys, which may underestimate the actual abundance of bull trout. ${ }^{5}$ Several studies (e.g., Goetz 1994; Thurow and Schill 1996; Hillman and Chapman 1996; Bonar et al. 1997) have found bull trout population estimates based on nighttime snorkeling to be in some cases more accurate than daytime snorkeling, especially for juvenile bull trout. Our estimates of adult bull trout numbers may be more accurate than those for juveniles.

In all years we found most juvenile bull trout in the upstream reaches of the Chiwawa River. Of the reaches we surveyed, they were most numerous in Reaches 7-10 on the Chiwawa River. In 2015, they occurred in Reaches 9-10 on the Chiwawa River. We found the majority of these fish in multiple channels, pools, and riffles, and few in glides. They consistently occupied stations close to the stream bottom over rubble and small boulder substrate or near woody debris. This is similar to the observation of Pratt (1984) in the upper Flathead River Basin in Montana. She found that juvenile bull trout lay close to instream cover and that they tended to conceal themselves. As a result, she found it difficult to accurately estimate their numbers. Although this implies that we underestimated numbers of juvenile bull trout in the Chiwawa River, the relative

[^4]distribution of juvenile bull trout is valid if we assume that we saw the same fraction of juveniles in all reaches (i.e., detection probability was the same across survey sites).

We estimated a total of 2,286 ( $\pm 14 \%$ of the estimated total) adult ( $>8 \mathrm{in}$ ) bull trout in reaches of the Chiwawa River basin in August 2015 (Table 9). This was the greatest number of adult bull trout that we recorded during the more than 20-year survey period. In previous years, numbers ranged from 76 to 900 (Figure 9; Appendix B). As with juvenile bull trout, we found most of the adult bull trout upstream from Reach 6; although they were found in all reaches on the Chiwawa River. We found few adult bull trout in tributaries of the Chiwawa River. Adult bull trout primarily used pools and multiple channel habitat, although most of the smaller adults ( $<10 \mathrm{in}$ ) used riffles.

## Abundance of Other Salmonids

In August 2015, we estimated that at least 28 brook trout, an exotic species closely related to the bull trout, occurred in the Chiwawa River, Chikamin Creek, Big Meadow Creek, Minnow Creek, and in the Little Wenatchee River survey areas. Brook trout occurred in the lower seven reaches on the Chiwawa River. In both the Chiwawa and Little Wenatchee rivers, brook trout usually used multiple channels. Few appeared to be bull trout/brook trout hybrids. In Chikamin, Minnow, and Big Meadow creeks, brook trout were most abundant in pools. Brook trout lengths ranged from 2-12 inches.

At least 294 westslope cutthroat trout occurred in the Chiwawa River, Phelps Creek, Nason Creek, and Little Wenatchee River survey areas in August 2015. These fish most often occurred in pools and multiple channel habitats. They ranged in size from 2-22 inches. Juvenile coho salmon were observed in Nason Creek and the Chiwawa River.

We observed both juvenile and adult mountain whitefish in the Chiwawa River, Phelps Creek, Rock Creek, Nason Creek, and the Little Wenatchee River survey areas. In sum, at least 6,861 adult and 2,145 juvenile whitefish lived in these streams in August 2015. We found few whitefish in most tributaries to the Chiwawa River.

## Conclusion

This was the $23^{\text {rd }}$ year of a study to monitor trends in juvenile spring Chinook production in the Chiwawa River basin. As shown in Figure 3, numbers of juvenile Chinook salmon in the Chiwawa River basin have fluctuated widely over the 23-year period. Numbers of juveniles in 2001, 2002, and 2009-2015 were some of the highest recorded, while numbers in the mid-1990s were some of the lowest. Interestingly, the highest spawning escapements (highest redd numbers) resulted in the lowest egg-parr survival rates (Appendix A). This is supported by the fact that the best approximating models clearly demonstrated a density-dependent relationship between seeding levels and juvenile production. Indeed, there was a significant negative relationship between parr per redd and numbers of redds in the Chiwawa River basin. This is an important observation because some of the hypotheses in the revised monitoring and evaluation plan (Hillman et al. 2013) are only valid when the supplemented population is below its carrying capacity.
The best fitting stock-recruitment models indicate that the capacity of the Chiwawa River basin
is between 140,000 to 180,000 spring Chinook parr. This equates to an overall density of about 1,300-1,600 parr per hectare. These densities can be achieved with about 470 redds. Assuming that a female Chinook produces only one redd (Murdoch et al. 2009), a spawning escapement of about 470 females is needed to fill the capacity of the Chiwawa River basin.
The proportion of hatchery-origin spawners ( $\mathrm{pHOS} \mathrm{)} \mathrm{within} \mathrm{the} \mathrm{Chiwawa} \mathrm{River} \mathrm{basin} \mathrm{during} \mathrm{the}$ survey period has ranged from 0 to $100 \%$. Thus, some of the variation in juvenile productivity may be related to pHOS . Although there appeared to be a negative relationship between juvenile productivity (parr/redd) and pHOS, the correlation was not significant (Figure 10). In addition, there was no relationship between juvenile productivity and pHOS after the effects of spawning escapement were removed from the analysis (Figure 10). This suggests that spawning escapement has a larger effect on juvenile productivity than does the presence of hatchery spawners.
The presence of density dependence in the early life stages of spring Chinook is not surprising. Rarely does density dependence appear in numbers of adult spring Chinook or on their spawning grounds. The Chiwawa River basin appears to have plenty of spawning habitat, as indicated by the large numbers of spawners and redds widely distributed throughout the basin during high spawning escapements. However, those large spawning escapements did not translate into large numbers of juveniles or smolts. Thus, density-dependent regulation appears to occur sometime during the early life stages of the fish, likely at the fry stage. It is possible that physical habitat (space) during higher flows when fry are emerging may limit juvenile Chinook production in the basin. Low nutrient levels and its effects on food webs may also be a limiting factor in the basin. If spawning escapements remain relatively high, marine-derived nutrients should increase in the basin, resulting in more food for juvenile Chinook salmon.

## References

Bonar, S. A., M. Divens, and B. Bolding. 1997. Methods for sampling the distribution and abundance of bull trout and Dolly Varden. Washington Department of Fish and Wildlife, Research Report No. RAD97-05. Olympia, WA.
Burnham, K. P. and D. R. Anderson. 2002. Model selection and multimodel inference: a practical information-theoretic approach. Second edition. Springer, New York, N.Y.

Goetz, F. A. 1994. Distribution and juvenile ecology of bull trout (Salvelinus confluentus) in the Cascade Mountains. Master's thesis. Oregon State University, Corvallis.

Hillman, T. W. and D. W. Chapman. 1996. Comparison of underwater methods and electrofishing for estimating fish populations in the upper Blackfoot River Basin. Report to Seven-Up Pete Joint Venture, Lincoln, MT.

Hillman, T. W. and M. D. Miller. 2004. Abundance and total numbers of Chinook salmon and trout in the Chiwawa River basin, Washington, 2004. Report to Chelan Public Utility District, Wenatchee, WA. BioAnalysts, Inc., Boise, ID.

Hillman, T, M. Miller, T. Miller, M. Tonseth, M. Hughes, A. Murdoch, L. Keller, and J. Murauskas. 2013. Monitoring and evaluation of the Chelan County PUD Hatchery programs: 2012 annual report. Report to the HCP Hatchery Committee, Wenatchee, WA.

Hillman, T., T. Kahler, G. Mackey, J. Murauskas, A. Murdoch, K. Murdoch, T. Pearsons, and M. Tonseth. 2013. Updated monitoring and evaluation plan for PUD hatchery programs. Report to the Hatchery Committees, Wenatchee, East Wenatchee, and Ephrata, WA.
Hillman, T., T. Kahler, G. Mackey, J. Murauskas, A. Murdoch, K. Murdoch, T. Pearsons, and M. Tonseth. 2013. Monitoring and evaluation plan for PUD hatchery programs; 2013 update. Report to the HCP and PRCC Hatchery Committees, Wenatchee, WA.

Hubble, J. 1993. Methow valley spring Chinook supplementation project. Yakima Indian Nation. Annual report to Douglas County Public Utility District, East Wenatchee, WA.
Kiefer, R. and K. Forster. 1991. Idaho habitat and natural production monitoring. Idaho Department of Fish and Game, Annual Report 1989, Project No. 83-7, Contract No. DE-BI79-84BP13381.

Konopacky, R. C., P. J. Cernera, and E. C. Bowles. 1986. Natural propagation and habitat improvement, Idaho: Salmon River habitat enhancement. Subproject I, Bear Valley Creek: inventory, 1984 and 1985. Shoshone-Bannock Tribes, Fort Hall, ID. Report to U.S. Department of Energy, Bonneville Power Administration, Division of Fish and Wildlife, Project No. 83-359, Contract No. DE-A179-84BP14383, Portland, OR.

Mullan, J. W., K. R. Williams, G. Rhodus, T. W. Hillman, and J. D. McIntyre. 1992. Production and habitat of salmonids in mid-Columbia River tributary streams. U.S. Fish and Wildlife Service, Monograph I. 489 p.

Murdoch, A. and C. Peven. 2005. Conceptual approach to monitoring and evaluating the Chelan County Public Utility District hatchery programs. Chelan County Public Utility District
and the Washington Department of Fish and Wildlife, Wenatchee, WA.
Murdoch, A., T. Pearsons, and T. Maitland. 2009. The number of redds constructed per female spring Chinook salmon in the Wenatchee River Basin. North American Journal of Fisheries Management 29:441-446.

Petrosky, C. E. 1990. Estimating spring Chinook parr and smolt abundance in wild and natural production areas. Pages 57-61 in: D. L. Park, editor. Status and future of spring Chinook salmon in the Columbia River basin--conservation and enhancement. Spring Chinook salmon workshop, U.S. Dept. Comm. NOAA Tech. Mem. NMFS F/NWD-187.
Pratt, K. L. 1984. Habitat use and species interactions of juvenile cutthroat Salmo clarki lewisi and bull trout Salvelinus confluentus in the upper Flathead River Basin. Master's thesis. University of Idaho, Moscow, ID. 95 p.

Quinn, T. P. 2004. The behavior and ecology of Pacific salmon and trout. University of Washington Press, Seattle, WA.
Richards, C. and P. J. Cernera. 1987. Salmon River habitat enhancement, annual report, 1986. Shoshone-Bannock Tribes, Fort Hall, ID. Report to Bonneville Power Administration, Project No. 83-359, Contract No. DE-A179-84BP14383, Portland, OR.
Thorson, J. T., M. D. Scheuerell, E. R. Buhle, and T. Copeland. 2013. Spatial variation buffers temporal fluctuations in early juvenile survival for an endangered Pacific salmon. Journal of Animal Ecology 2013:1-11.
Thurow, R. F. and D. J. Schill. 1996. Comparison of day snorkeling, night snorkeling, and electrofishing to estimate bull trout abundance and size structure in a second-order Idaho stream. North American Journal of Fisheries Management 16:314-323.

USFS (United States Forest Service). 1993. Upper Chiwawa River stream survey report. Wenatchee National Forest, Wenatchee, WA.


Figure 1. Location of study reaches on the Chiwawa River, and Chikamin, Rock, Big Meadow, Unnamed, Alder, Brush and Phelps creeks, Chelan County, Washington. Reach 2 on Nason Creek and Reach 2 on the Little Wenatchee River were matched with Reaches 3 and 8 on the Chiwawa River, respectively.

## Chiwawa River

 2015

Figure 2. Mean, minimum, and maximum monthly flows in the Chiwawa River for 2015.

## Chinook Salmon

Age-0


Age-1+


Figure 3. Numbers of age-0 and age-1+ Chinook salmon within the Chiwawa River basin in August 1992-2015; ND = no data.

## Chiwawa Spring Chinook



Figure 4. Relationship between total numbers of age-0 Chinook salmon (based on fish/ha) and numbers of eggs in the Chiwawa River basin. Vertical bars indicate $95 \%$ confidence bounds.


Figure 5. Comparison of the means ( $95 \% \mathrm{CI}$ ) of age-0 Chinook salmon densities (fish/ha) within state/habitat types in Reaches 3 and 8 of the Chiwawa River and their matched reference areas on Nason Creek and the Little Wenatchee River. There was no sampling in 2000 and no sampling in reference areas in 1992.


Figure 6. Relationship between numbers of juvenile (age-0) Chinook and redds in the Chiwawa River basin, 1992-2015 (no sampling occurred in 2000). Figures show the fit of the Beverton-Holt model, smooth hockey stick, Ricker model, and the Cushing model to the data. Gray lines indicate the upper and lower 95\% C.B.

## Chiwawa Spring Chinook



Figure 7. Relationship between parr/redd and numbers of redds (top figure) and natural log parr/redd and numbers of redds (bottom figure) in the Chiwawa River basin, 1992-2015. No sampling was conducted in 2000. Estimates for 1993-2015 included the Chiwawa River and its tributaries; the 1992 estimate included only the Chiwawa River. The linear relationship $\mathrm{LN}(\mathrm{P} / \mathrm{R})=6.38-0.002$ (Redds) was significant with $\mathrm{P}=0.0000 ; R^{2}=0.690$.


Figure 8. Numbers of age-0 ( $<4 \mathrm{in}$ ) and age-1+ ( $4-8 \mathrm{in}$ ) steelhead/rainbow within the Chiwawa River basin in August 1992-2015; ND = no data.


Figure 9. Numbers of juvenile ( $2-8$ inches) and adult ( $>8$ inches) bull trout within the Chiwawa River basin in August 1992-2015; ND = no data.

Chiwawa Spring Chinook


Chiwawa Spring Chinook


Figure 10. Relationship between juvenile productivity (parr/redd) and the proportion of hatchery-origin spawners ( pHOS ) (top figure) and the relationship between the residuals from the Beverton-Holt stock/recruitment relationship and pHOS (bottom figure).

Table 1. Description, location (river mile), and area (hectares) of land-class strata (reaches) used by age-0 Chinook salmon in the Chiwawa River basin, 2015. Reaches were classified according to geologic district, landtype association, valley-bottom type, stream state-type, and habitat type within the Cascade Ecoregion; MCV = moderately confined valley, $\mathrm{CC}=$ confined canyon, $\mathrm{UCV}=$ unconfined valley, $\mathrm{NC}=$ natural channel, $\mathrm{EB}=$ eroded banks, $\mathrm{S}=$ straight, $\mathrm{G}=$ glide, $\mathrm{P}=$ pool, $\mathrm{R}=$ riffle, and $\mathrm{MC}=$ multiple channel. See Hillman and Miller (2004) for definitions of stream state codes.

| Reach | RM | Gradient | Geologic district | Landtype association | Valley bottom type | Stream state type | Habitat type | Area (ha) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | Total | Sample |
| Chiwawa River |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-3.77 | 0.007 | Glacial Drift over Chumstick Formation | Glacial Valley | MCV <br> Alluvial | NC/EB | G | 0.49 | 0.49 |
|  |  |  |  |  |  | NC/EB | P | 1.17 | 0.88 |
|  |  |  |  |  |  | NC/EB | R | 16.60 | 1.57 |
| 2 | 3.77-5.51 | 0.010 | Glacial Drift over Chumstick Formation | Glacial Canyon | CC Fluvial | NC/EB | G | 0.29 | 0.25 |
|  |  |  |  |  |  | NC/EB | P | 0.70 | 0.24 |
|  |  |  |  |  |  | NC/EB | R | 6.08 | 0.58 |
| 3 | 5.51-7.88 | 0.009 | Glacial Drift over Chumstick Formation | Glacial Valley | MCV <br> Alluvial | NC/S | R | 4.45 | 0.70 |
|  |  |  |  |  |  | NC/EB | G | 0.11 | 0.11 |
|  |  |  |  |  |  | NC/EB | R | 4.13 | 0.48 |
|  |  |  |  |  |  | MC | MC | 0.38 | 0.38 |
| 4 | 7.88-8.90 | 0.007 | Glacial Drift over Chumstick Formation | Glacial Canyon | CC Fluvial | NC/EB | P | 0.34 | 0.26 |
|  |  |  |  |  |  | NC/EB | R | 2.34 | 0.33 |
|  |  |  |  |  |  | MC | MC | 0.39 | 0.39 |
| 5 | 8.90-10.83 | 0.011 | Glacial Drift over Chumstick Formation | Glacial Valley | MCV Alluvial | NC/EB | P | 0.13 | 0.13 |
|  |  |  |  |  |  | NC/EB | R | 7.63 | 0.92 |
| 6 | 10.83-11.80 | 0.008 | Glacial Drift over Chumstick Formation | Glacial Canyon | CC Fluvial | NC/EB | P | 0.35 | 0.35 |
|  |  |  |  |  |  | NC/EB | R | 3.72 | 0.93 |
|  |  |  |  |  |  | MC | MC | 0.36 | 0.36 |
| 7 | 11.80-20.03 | 0.001 | Glacial Drift over Chumstick Formation | Glacial Valley | $\begin{gathered} \text { UCV } \\ \text { Alluvial } \end{gathered}$ | NC | G | 1.89 | 0.44 |
|  |  |  |  |  |  | NC | P | 5.11 | 0.49 |
|  |  |  |  |  |  | NC | R | 0.71 | 0.17 |
|  |  |  |  |  |  | NC/EB | G | 2.30 | 1.20 |
|  |  |  |  |  |  | NC/EB | P | 5.83 | 1.50 |
|  |  |  |  |  |  | NC/EB | R | 4.20 | 0.47 |
|  |  |  |  |  |  | MC | MC | 4.05 | 1.77 |
| 8 | 20.03-25.42 | 0.003 | Glacial Drift over Swakane Gneiss | Glacial Valley | UCV <br> Alluvial | NC/EB | G | 2.09 | 0.85 |
|  |  |  |  |  |  | NC/EB | P | 7.01 | 2.02 |
|  |  |  |  |  |  | NC/EB | R | 4.46 | 0.81 |
|  |  |  |  |  |  | EB | P | 0.22 | 0.22 |
|  |  |  |  |  |  | EB | R | 0.34 | 0.34 |
|  |  |  |  |  |  | MC | MC | 5.90 | 2.34 |
| 9 | 25.42-28.81 | 0.007 | Glacial Drift over Swakane Gneiss | Glacial Valley | MCV <br> Alluvial | NC | P | 3.92 | 0.43 |
|  |  |  |  |  |  | NC | R | 2.20 | 0.47 |
|  |  |  |  |  |  | MC | MC | 2.58 | 1.10 |
| 10 | 28.81-31.11 | 0.011 | Pre-upper Jurassic Gneiss | Glacial Valley | $\begin{gathered} \text { MCV } \\ \text { Alluvial } \end{gathered}$ | NC | P | 0.47 | 0.24 |
|  |  |  |  |  |  | NC | R | 1.87 | 0.27 |
|  |  |  |  |  |  | MC | MC | 3.92 | 0.28 |

Table 1. Concluded.

| Reach | $\mathbf{R M}$ | Gradient | Geologic district | Landtype association | Valley bottom type | Stream state type | Habitat type | Area (ha) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | Total | Sampled |
| Trinity Side Channel |  |  |  |  |  |  |  |  |  |
| 10b | 0.00-0.75 | 0.011 | Pre-upper Jurassic Gneiss | Glacial Valley | MCV <br> Alluvial | NC | P | 0.40 | 0.08 |
|  |  |  |  |  |  | NC | R | 0.14 | 0.06 |
|  |  |  |  |  |  | NC | MC | 0.07 | 0.07 |
| Phelps Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.35 | 0.043 | Pre-upper Jurassic Gneiss | Glacial Valley | MCV Alluvial | NC | R | 0.00 | 0.00 |
|  |  |  |  |  |  | NC | MC | 0.14 | 0.14 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.94 | 0.013 | Glacial Drift over Chumstick Formation | Glacial Valley | UCV <br> Alluvial | NC | G | 0.05 | 0.05 |
|  |  |  |  |  |  | NC | P | 0.19 | 0.06 |
|  |  |  |  |  |  | NC | R | 0.32 | 0.10 |
|  |  |  |  |  |  | MC | MC | 0.14 | 0.14 |
| Rock Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.73 | 0.020 | Glacial Drift over Swakane Gneiss | Glacial Valley | UCVAlluvial | NC | P | 0.20 | 0.05 |
|  |  |  |  |  |  | NC | R | 0.37 | 0.08 |
|  |  |  |  |  |  | MC | MC | 0.10 | 0.10 |
| Unnamed Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.05 |  | Pre-upper Jurassic Gneiss | Glacial Valley | MCV Alluvial | NC | P | 0.00 | 0.00 |
|  |  |  |  |  |  | NC | R | 0.00 | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.35 | 0.025 | Glacial Drift over Chumstick Formation | Glacial Valley | MCV <br> Alluvial | NC | G | 0.02 | 0.02 |
|  |  |  |  |  |  | NC | P | 0.10 | 0.05 |
|  |  |  |  |  |  | NC | R | 0.07 | 0.02 |
|  |  |  |  |  |  | NC | MC | 0.00 | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.01 |  | Glacial Drift over Chumstick Formation | Glacial Valley | MCV <br> Alluvial | NC | P | 0.001 | 0.001 |
|  |  |  |  |  |  | NC | R | 0.006 | 0.006 |
| Brush Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.01 |  | Glacial Drift over Chumstick Formation | Glacial Valley | UCV <br> Alluvial | NC | P | 0.002 | 0.002 |
|  |  |  |  |  |  | NC | R | 0.003 | 0.003 |
| Clear Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.05 |  | Glacial Drift over Chumstick Formation | Glacial Valley | UCV Alluvial | NC | P | 0.003 | 0.003 |
|  |  |  |  |  |  | NC | R | 0.002 | 0.002 |
| Y Creek |  |  |  |  |  |  |  |  |  |
| 1 | 0.00-0.05 |  | Glacial Drift over Swakane Gneiss | Glacial Valley | UCV Alluvial | NC | P | 0.000 | 0.000 |
|  |  |  |  |  |  | NC | R | 0.000 | 0.000 |

[^5]Table 2. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of age-0 Chinook salmon in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 70.3 | 0.025 | 1,285 | $\pm 437$ | 0.34 | 1,250 | $\pm 384$ | 0.31 |
| 2 | 111.7 | 0.027 | 790 | $\pm 135$ | 0.17 | 690 | $\pm 75$ | 0.11 |
| 3 | 47.5 | 0.014 | 431 | $\pm 22$ | 0.05 | 471 | $\pm 21$ | 0.05 |
| 4 | 368.7 | 0.082 | 1,132 | $\pm 66$ | 0.06 | 1,137 | $\pm 89$ | 0.08 |
| 5 | 44.2 | 0.012 | 343 | $\pm 27$ | 0.08 | 377 | $\pm 21$ | 0.06 |
| 6 | 58.7 | 0.020 | 260 | $\pm 45$ | 0.17 | 252 | $\pm 37$ | 0.15 |
| 7 | 728.6 | 0.113 | 17,553 | $\pm 3,979$ | 0.23 | 15,333 | $\pm 2,998$ | 0.20 |
| 8 | 743.2 | 0.135 | 14,878 | $\pm 5,167$ | 0.35 | 13,792 | $\pm 4,405$ | 0.32 |
| 9 | 1,953.8 | 0.343 | 16,998 | $\pm 4,623$ | 0.27 | 14,448 | $\pm 1,710$ | 0.12 |
| 10 | 7,283.8 | 1.992 | 50,040 | $\pm 1,852$ | 0.04 | 41,690 | $\pm 2,185$ | 0.05 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 2,035.7 | 2.074 | 285 | $\pm 0$ | 0.00 | 285 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 2,738.6 | 1.947 | 1,917 | $\pm 626$ | 0.33 | 2,467 | $\pm 560$ | 0.23 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 6,205.9 | 2.524 | 4,158 | $\pm 564$ | 0.14 | 4,110 | $\pm 1,875$ | 0.46 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 5,446.2 | 2.688 | 1,013 | $\pm 545$ | 0.54 | 915 | $\pm 373$ | 0.41 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 10,142.9 | 11.270 | 71 | $\pm 0$ | 0.00 | 71 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 12,400.00 | 22.963 | 62 | $\pm 0$ | 0.00 | 62 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 1,600.0 | 1.404 | 8 | $\pm 0$ | 0.00 | 8 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 1,001.6 | 0.206 | 111,224 | $\pm 8,280$ | 0.07 | 97,358 | $\pm 6,342$ | 0.07 |

[^6]Table 3. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of age-1+ Chinook salmon in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 0.9 | 0.000 | 16 | $\pm 21$ | 1.31 | 15 | $\pm 10$ | 0.67 |
| 2 | 4.5 | 0.001 | 32 | $\pm 37$ | 1.16 | 26 | $\pm 24$ | 0.92 |
| 3 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 4 | 7.5 | 0.002 | 23 | $\pm 0$ | 0.00 | 23 | $\pm 0$ | 0.00 |
| 5 | 0.5 | 0.000 | 4 | $\pm 0$ | 0.00 | 3 | $\pm 0$ | 0.00 |
| 6 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 7 | 11.9 | 0.002 | 286 | $\pm 227$ | 0.79 | 244 | $\pm 140$ | 0.57 |
| 8 | 3.3 | 0.001 | 67 | $\pm 78$ | 1.16 | 61 | $\pm 53$ | 0.87 |
| 9 | 6.0 | 0.001 | 52 | $\pm 72$ | 1.38 | 42 | $\pm 53$ | 1.26 |
| 10 | 1.2 | 0.001 | 8 | $\pm 11$ | 1.38 | 10 | $\pm 6$ | 0.60 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 31.4 | 0.025 | 22 | $\pm 33$ | 0.00 | 32 | $\pm 33$ | 0.00 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 164.2 | 0.066 | 110 | $\pm 67$ | 0.61 | 108 | $\pm 144$ | 1.33 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 5.6 | 0.001 | 620 | $\pm 265$ | 0.43 | 564 | $\pm 218$ | 0.39 |

[^7]Table 4. Summary of the five productivity models of juvenile (age-0) Chinook salmon in the Chiwawa River basin. Models are shown, including the number of parameters ( $K$ ), $\mathrm{AIC}_{\mathrm{c}}$ values, $\mathrm{AIC}_{\mathrm{c}}$ difference scores $\left(\Delta_{\mathrm{i}}\right)$, the likelihood of the model given the data $\left(£\left(g_{i} \mid x\right)\right.$ ), Akaike weights ( $w_{i}$ ), and adjusted $R^{2}$ values. The sample size ( $n$ ) for all models was 23 . Models describe the relationship between juvenile Chinook numbers (dependent variable) and redd numbers (independent variable).

| Model | $K^{a}$ | $\mathrm{AIC}_{\text {c }}$ | $\Delta_{i}$ | $f\left(g_{i} \mid x\right)$ | $\boldsymbol{w}_{i}$ | Adj $\mathrm{R}^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Beverton-Holt | 3 | -123.272 | 0.000 | 1.000 | 0.663 | 0.838 |
| Smooth Hockey Stick | 3 | -121.632 | 1.640 | 0.440 | 0.292 | 0.826 |
| Gamma ${ }^{\text {b }}$ | 4 | -116.473 | 6.799 | 0.033 | 0.022 | 0.799 |
| Ricker | 3 | -115.227 | 8.046 | 0.018 | 0.012 | 0.778 |
| Cushing | 3 | -115.186 | 8.087 | 0.018 | 0.012 | 0.770 |

${ }^{\mathrm{a}} \boldsymbol{K}$ is the number of structural parameters in the model plus 1 for $\sigma^{2}$.
${ }^{\mathrm{b}}$ The $\gamma$ parameter in the Gamma model was greater than 0 , which means that this model is nearly identical to the Ricker model.

Table 5. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of age-0 ( $<4 \mathrm{in}$ ) steelhead/rainbow in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 52.2 | 0.017 | 953 | $\pm 90$ | 0.09 | 852 | $\pm 91$ | 0.11 |
| 2 | 58.0 | 0.015 | 410 | $\pm 114$ | 0.28 | 388 | $\pm 139$ | 0.36 |
| 3 | 102.4 | 0.030 | 929 | $\pm 17$ | 0.02 | 988 | $\pm 9$ | 0.01 |
| 4 | 61.6 | 0.014 | 189 | $\pm 42$ | 0.22 | 190 | $\pm 32$ | 0.17 |
| 5 | 46.1 | 0.013 | 358 | $\pm 38$ | 0.11 | 433 | $\pm 32$ | 0.07 |
| 6 | 18.7 | 0.006 | 83 | $\pm 15$ | 0.18 | 78 | $\pm 11$ | 0.14 |
| 7 | 65.4 | 0.011 | 1,575 | $\pm 689$ | 0.44 | 1,448 | $\pm 700$ | 0.48 |
| 8 | 1.7 | 0.000 | 35 | $\pm 42$ | 1.20 | 31 | $\pm 28$ | 0.90 |
| 9 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 10 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 2,841.4 | 1.913 | 1,989 | $\pm 585$ | 0.29 | 2,424 | $\pm 571$ | 0.24 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 2,500.0 | 1.064 | 1,675 | $\pm 391$ | 0.23 | 1,732 | $\pm 683$ | 0.39 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 9,467.7 | 4.780 | 1,761 | $\pm 446$ | 0.25 | 1,627 | $\pm 179$ | 0.11 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 24,285.7 | 26.984 | 170 | $\pm 0$ | 0.00 | 170 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 12,400.0 | 22.963 | 62 | $\pm 0$ | 0.00 | 62 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 3,800.0 | 3.333 | 19 | $\pm 0$ | 0.00 | 19 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 91.9 | 0.022 | 10,208 | $\pm 1,093$ | 0.11 | 10,442 | $\pm 1,160$ | 0.11 |

[^8]Table 6. Estimated mean densities (fish/hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of age-1+ (4-8 in) steelhead/rainbow in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 7.4 | 0.003 | 135 | $\pm 46$ | 0.34 | 123 | $\pm 38$ | 0.31 |
| 2 | 2.5 | 0.001 | 18 | $\pm 24$ | 1.33 | 16 | $\pm 19$ | 1.19 |
| 3 | 19.3 | 0.006 | 175 | $\pm 31$ | 0.18 | 206 | $\pm 22$ | 0.11 |
| 4 | 12.7 | 0.003 | 39 | $\pm 10$ | 0.26 | 39 | $\pm 7$ | 0.18 |
| 5 | 14.4 | 0.004 | 112 | $\pm 14$ | 0.13 | 130 | $\pm 9$ | 0.07 |
| 6 | 8.1 | 0.003 | 36 | $\pm 14$ | 0.39 | 33 | $\pm 11$ | 0.33 |
| 7 | 4.1 | 0.001 | 99 | $\pm 108$ | 1.09 | 95 | $\pm 118$ | 1.24 |
| 8 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 9 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 10 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 57.1 | 0.032 | 40 | $\pm 0$ | 0.00 | 40 | $\pm 0$ | 0.00 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 149.3 | 0.060 | 100 | $\pm 149$ | 1.49 | 98 | $\pm 178$ | 1.82 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand <br> Total | 6.8 | 0.002 | 754 | $\pm 195$ | 0.26 | 780 | $\pm 219$ | 0.28 |

[^9]Table 7. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of steelhead/rainbow larger than 8 inches in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | $\text { Volume }\left(\mathbf{m}^{3}\right)$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 0.7 | 0.000 | 13 | $\pm 19$ | 1.46 | 15 | $\pm 5$ | 0.33 |
| 2 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 3 | 0.1 | 0.000 | 1 | $\pm 0$ | 0.00 | 1 | $\pm 0$ | 0.00 |
| 4 | 0.3 | 0.000 | 1 | $\pm 0$ | 0.00 | 1 | $\pm 0$ | 0.00 |
| 5 | 0.4 | 0.000 | 3 | $\pm 0$ | 0.00 | 3 | $\pm 0$ | 0.00 |
| 6 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 7 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 8 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 9 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 10 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 0.2 | 0.000 | 18 | $\pm 19$ | 1.06 | 20 | $\pm 5$ | 0.25 |

[^10]Table 8. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of juvenile bull trout ( $2-8 \mathrm{in}$ ) in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 2 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 3 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 4 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 5 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 6 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 7 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 8 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| 9 | 6.9 | 0.001 | 60 | $\pm 34$ | 0.57 | 55 | $\pm 27$ | 0.49 |
| 10 | 21.8 | 0.006 | 150 | $\pm 21$ | 0.14 | 120 | $\pm 14$ | 0.12 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 35.7 | 0.036 | 5 | $\pm 0$ | 0.00 | 5 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 35.8 | 0.002 | 24 | $\pm 6$ | 0.25 | 24 | $\pm 17$ | 0.71 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 2.2 | 0.000 | 239 | $\pm 41$ | 0.17 | 204 | $\pm 35$ | 0.17 |

[^11]Table 9. Estimated mean densities (fish $/$ hectare and fish $/ \mathrm{m}^{3}$ ), total numbers, $95 \%$ confidence bounds on total numbers, and error of the estimated total number of adult bull trout ( $>8$ in) in reaches in the Chiwawa River basin, Washington, August 2015.

| Reach | Mean density |  | Surface area (ha) |  |  | Volume ( $\mathrm{m}^{3}$ ) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fish/ha | Fish/m ${ }^{3}$ | Total No. | 95\% C.B. | $\pm$ Error | Total No. | 95\% C.B. | $\pm$ Error |
| Chiwawa River |  |  |  |  |  |  |  |  |
| 1 | 1.3 | 0.000 | 24 | $\pm 16$ | 0.67 | 20 | $\pm 6$ | 0.30 |
| 2 | 7.1 | 0.002 | 50 | $\pm 21$ | 0.42 | 42 | $\pm 1$ | 0.02 |
| 3 | 0.1 | 0.000 | 1 | $\pm 0$ | 0.00 | 1 | $\pm 0$ | 0.00 |
| 4 | 3.9 | 0.001 | 12 | $\pm 10$ | 0.83 | 12 | $\pm 7$ | 0.58 |
| 5 | 1.4 | 0.000 | 11 | $\pm 0$ | 0.00 | 10 | $\pm 0$ | 0.00 |
| 6 | 1.1 | 0.000 | 5 | $\pm 0$ | 0.00 | 5 | $\pm 0$ | 0.00 |
| 7 | 16.3 | 0.003 | 392 | $\pm 204$ | 0.52 | 352 | $\pm 128$ | 0.36 |
| 8 | 9.2 | 0.002 | 185 | $\pm 159$ | 0.86 | 184 | $\pm 55$ | 0.30 |
| 9 | 37.4 | 0.007 | 325 | $\pm 64$ | 0.20 | 283 | $\pm 47$ | 0.17 |
| 10 | 185.4 | 0.051 | 1,274 | $\pm 169$ | 0.13 | 1,072 | $\pm 177$ | 0.17 |
| Phelps Creek |  |  |  |  |  |  |  |  |
| 1 | 42.9 | 0.044 | 6 | $\pm 0$ | 0.00 | 6 | $\pm 0$ | 0.00 |
| Chikamin Creek ${ }^{1}$ |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Rock Creek |  |  |  |  |  |  |  |  |
| 1 | 1.5 | 0.001 | 1 | $\pm 0$ | 0.00 | 1 | $\pm 0$ | 0.00 |
| Unnamed Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Big Meadow Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Alder Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Brush Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Clear Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Y Creek |  |  |  |  |  |  |  |  |
| 1 | 0.0 | 0.000 | 0 | $\pm 0$ | 0.00 | 0 | $\pm 0$ | 0.00 |
| Grand Total | 20.6 | 0.004 | 2,286 | $\pm 316$ | 0.14 | 1,988 | $\pm 230$ | 0.12 |

[^12]APPENDIX A. Numbers of redds, eggs, age-0 Chinook salmon, parr per redd, and percent egg-to-parr survival in the Chiwawa River basin, brood years 1991-2014; NS = not sampled. Numbers of eggs were calculated as the number of redds times the mean fecundity of females collected for broodstock.

| Brood Year | Chinook Salmon |  |  | Parr/Redd | Egg-to-parr survival (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Redds | Eggs | Age-0 (parr) |  |  |
| 1991 | 104 | 478,400 | 45,483 | 437 | 9.5 |
| 1992 | 302 | 1,570,098 | 79,113 | 262 | 5.0 |
| 1993 | 106 | 556,394 | 55,056 | 519 | 9.9 |
| 1994 | 82 | 485,686 | 55,240 | 674 | 11.4 |
| 1995 | 13 | 66,248 | 5,815 | 447 | 8.8 |
| 1996 | 23 | 106,835 | 16,066 | 699 | 15.0 |
| 1997 | 82 | 374,740 | 68,415 | 834 | 18.3 |
| 1998 | 41 | 218,325 | 41,629 | 1,015 | 19.1 |
| 1999 | 34 | 166,090 | NS | NS | NS |
| 2000 | 128 | 642,944 | 114,617 | 895 | 17.8 |
| 2001 | 1,078 | 4,984,672 | 134,874 | 125 | 2.7 |
| 2002 | 345 | 1,605,630 | 91,278 | 265 | 5.7 |
| 2003 | 111 | 648,684 | 45,177 | 407 | 7.0 |
| 2004 | 241 | 1,156,559 | 49,631 | 206 | 4.3 |
| 2005 | 332 | 1,436,564 | 79,902 | 241 | 5.6 |
| 2006 | 297 | 1,284,228 | 60,752 | 205 | 4.7 |
| 2007 | 283 | 1,256,803 | 82,351 | 291 | 6.6 |
| 2008 | 689 | 3,163,888 | 106,705 | 155 | 3.4 |
| 2009 | 421 | 1,925,233 | 128,220 | 305 | 6.7 |
| 2010 | 502 | 2,165,628 | 141,510 | 282 | 6.5 |
| 2011 | 492 | 2,157,420 | 103,940 | 211 | 4.8 |
| 2012 | 880 | 3,412,184 | 149,563 | 185 | 4.4 |
| 2013 | 714 | 3,367,224 | 121,240 | 170 | 3.6 |
| 2014 | 485 | 1,961,825 | 111,224 | 229 | 5.7 |
| Average | 324 | 1,466,346 | 82,078 | 244 | 8.1 |

APPENDIX B. Estimated numbers of salmonids (based on fish/ha) in the Chiwawa River basin, Washington, 1992-2015; NS = not sampled.

| Survey year | Chinook salmon |  | Steelhead/Rainbow |  |  | Bull trout |  | Cutthroat trout |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age-0 | Age-1+ | Age-0 | Age-1+ | $>8$ in $^{1}$ | 2-8 in | $>8$ in |  |
| $1992{ }^{2}$ | 45,483 | 563 | 4,927 | 2,533 | 1,869 | 299 | 208 | NS |
| 1993 | 79,113 | 174 | 4,004 | 2,860 | 768 | 158 | 156 | NS |
| 1994 | 55,056 | 18 | 1,410 | 5,856 | 67 | 90 | 76 | NS |
| 1995 | 55,241 | 13 | 7,357 | 9,517 | 140 | 97 | 664 | NS |
| 1996 | 5,815 | 22 | 4,245 | 11,849 | 78 | 79 | 343 | NS |
| 1997 | 16,066 | 5 | 8,823 | 6,905 | 48 | 220 | 472 | 56 |
| 1998 | 68,415 | 63 | 3,921 | 10,585 | 78 | 300 | 900 | 93 |
| 1999 | 41,629 | 41 | 5,838 | 22,130 | 33 | 130 | 423 | 80 |
| 2000 | NS | NS | NS | NS | NS | NS | NS | NS |
| 2001 | 114,617 | 69 | 45,727 | 10,623 | 420 | 505 | 542 | 108 |
| 2002 | 134,874 | 32 | 20,521 | 9,090 | 181 | 217 | 521 | 111 |
| 2003 | 91,278 | 134 | 18,020 | 6,179 | 49 | 196 | 282 | 52 |
| 2004 | 45,177 | 21 | 10,380 | 8,190 | 8 | 140 | 157 | 22 |
| 2005 | 49,631 | 79 | 11,463 | 6,188 | 48 | 125 | 346 | 23 |
| 2006 | 79,902 | 388 | 16,245 | 10,533 | 50 | 238 | 686 | 68 |
| 2007 | 60,752 | 41 | 14,073 | 8,448 | 77 | 95 | 520 | 47 |
| 2008 | 82,351 | 189 | 15,230 | 10,576 | 144 | 124 | 510 | 109 |
| 2009 | 106,705 | 54 | 17,179 | 5,629 | 85 | 82 | 618 | 128 |
| 2010 | 128,220 | 291 | 25,018 | 9,616 | 63 | 79 | 547 | 252 |
| 2011 | 141,510 | 967 | 39,446 | 14,903 | 65 | 86 | 621 | 240 |
| 2012 | 103,940 | 767 | 27,134 | 8,576 | 65 | 159 | 768 | 188 |
| 2013 | 149,563 | 852 | 21,682 | 7,253 | 76 | 299 | 820 | 358 |
| 2014 | 121,240 | 939 | 16,083 | 5,084 | 87 | 259 | 875 | 761 |
| 2015 | 111,224 | 620 | 10,208 | 754 | 18 | 239 | 2,286 | 292 |

${ }^{1}$ During 1992-1993, numbers of steelhead/rainbow greater than 8 inches included both hatchery and wild rainbow trout.
Thereafter, only wild trout were observed.
${ }^{2}$ Only the Chiwawa River was sampled in 1992. No tributaries were sampled in that year.

APPENDIX C. Proportion of total habitat available, fraction of all age-0 Chinook within each habitat type, and densities (fish/ha) and numbers of age-0 Chinook within each habitat type in the Chiwawa River basin, survey years 1992-2015; NS = not sampled.

| Habitat | 1992 | 1993 | 1994 | 1995 | 1996 | 1997 | 1998 | 1999 | 2000 | 2001 | 2002 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Proportion of total habitat available |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.10 | 0.09 | 0.10 | 0.10 | 0.10 | 0.09 | 0.09 | 0.09 | NS | 0.07 | 0.08 |
| Pool | 0.19 | 0.19 | 0.21 | 0.18 | 0.18 | 0.17 | 0.16 | 0.17 | NS | 0.15 | 0.16 |
| Riffle | 0.61 | 0.61 | 0.57 | 0.59 | 0.57 | 0.57 | 0.58 | 0.55 | NS | 0.49 | 0.48 |
| M. Chan | 0.10 | 0.11 | 0.12 | 0.14 | 0.14 | 0.17 | 0.17 | 0.19 | NS | 0.29 | 0.28 |
| Fraction of all age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.07 | 0.03 | 0.02 | 0.01 | 0.02 | 0.01 | 0.01 | 0.01 | NS | 0.03 | 0.01 |
| Pool | 0.30 | 0.28 | 0.22 | 0.21 | 0.30 | 0.16 | 0.17 | 0.14 | NS | 0.23 | 0.24 |
| Riffle | 0.19 | 0.16 | 0.12 | 0.11 | 0.43 | 0.23 | 0.08 | 0.11 | NS | 0.18 | 0.15 |
| M. Chan | 0.45 | 0.53 | 0.64 | 0.67 | 0.24 | 0.60 | 0.74 | 0.74 | NS | 0.57 | 0.60 |
| Densities of age-0 Chinook within habitat types (fish/ha) |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 254 | 251 | 93 | 55 | 11 | 12 | 78 | 13 | NS | 351 | 187 |
| Pool | 584 | 1,049 | 619 | 541 | 82 | 122 | 607 | 257 | NS | 1,392 | 1,468 |
| Riffle | 116 | 188 | 124 | 91 | 38 | 52 | 79 | 62 | NS | 336 | 300 |
| M. Chan | 1,710 | 3,408 | 2,985 | 2,328 | 84 | 449 | 2,620 | 1,201 | NS | 1,820 | 2,069 |
| Number of age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 2,967 | 2,458 | 857 | 623 | 137 | 130 | 837 | 157 | NS | 3,231 | 1,931 |
| Pool | 13,468 | 21,814 | 12,131 | 11,294 | 1,755 | 2,553 | 11,454 | 5,933 | NS | 25,890 | 32,612 |
| Riffle | 8,531 | 12,616 | 6,698 | 6,197 | 2,525 | 3,699 | 5,392 | 4,626 | NS | 20,629 | 19,754 |
| M. Chan | 20,517 | 42,225 | 35,370 | 36,965 | 1,396 | 9,682 | 50,728 | 30,912 | NS | 64,866 | 80,576 |

APPENDIX C. Continued.

| Habitat | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Proportion of total habitat available |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.07 | 0.07 | 0.08 | 0.08 | 0.07 | 0.09 | 0.08 | 0.08 | 0.08 | 0.07 | 0.07 |
| Pool | 0.17 | 0.16 | 0.16 | 0.16 | 0.17 | 0.23 | 0.22 | 0.23 | 0.18 | 0.23 | 0.23 |
| Riffle | 0.49 | 0.50 | 0.47 | 0.47 | 0.47 | 0.51 | 0.54 | 0.53 | 0.57 | 0.53 | 0.53 |
| M. Chan | 0.26 | 0.27 | 0.29 | 0.30 | 0.29 | 0.17 | 0.15 | 0.16 | 0.17 | 0.17 | 0.17 |
| Fraction of all age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.02 | 0.01 | 0.01 | 0.03 | 0.02 | 0.03 | 0.02 | 0.02 | 0.04 | 0.01 | 0.02 |
| Pool | 0.23 | 0.07 | 0.19 | 0.31 | 0.46 | 0.40 | 0.36 | 0.34 | 0.34 | 0.41 | 0.37 |
| Riffle | 0.15 | 0.14 | 0.07 | 0.12 | 0.12 | 0.11 | 0.11 | 0.11 | 0.19 | 0.15 | 0.13 |
| M. Chan | 0.60 | 0.77 | 0.73 | 0.54 | 0.40 | 0.45 | 0.51 | 0.53 | 0.43 | 0.43 | 0.48 |
| Densities of age-0 Chinook within habitat types (fish/ha) |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 200 | 58 | 49 | 237 | 113 | 238 | 230 | 286 | 526 | 173 | 321 |
| Pool | 951 | 155 | 492 | 1,240 | 1,211 | 1,210 | 1,453 | 1,436 | 1,805 | 1,360 | 1,890 |
| Riffle | 216 | 101 | 60 | 166 | 118 | 156 | 175 | 200 | 330 | 221 | 281 |
| M. Chan | 1,626 | 1,008 | 1,057 | 1,147 | 603 | 1,872 | 2,993 | 3,293 | 2,515 | 2,061 | 3,190 |
| Number of age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 1,884 | 540 | 442 | 2,498 | 1,120 | 2,668 | 2,371 | 3,164 | 6,122 | 1,535 | 2,822 |
| Pool | 21,091 | 3,183 | 9,626 | 26,754 | 28,851 | 34,314 | 39,382 | 44,765 | 48,846 | 42,209 | 55,651 |
| Riffle | 13,783 | 6,501 | 3,367 | 10,753 | 7,809 | 9,773 | 11,558 | 14,446 | 27,883 | 15,418 | 19,619 |
| M. Chan | 54,519 | 34,952 | 36,196 | 46,580 | 25,409 | 38,275 | 55,607 | 69,609 | 61,944 | 44,779 | 73,057 |

APPENDIX C. Concluded.

| Habitat | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | 2023 | Mean |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Proportion of total habitat available |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.07 | 0.07 |  |  |  |  |  |  |  |  | 0.08 |
| Pool | 0.22 | 0.24 |  |  |  |  |  |  |  |  | 0.19 |
| Riffle | 0.54 | 0.53 |  |  |  |  |  |  |  |  | 0.53 |
| M. Chan | 0.17 | 0.16 |  |  |  |  |  |  |  |  | 0.20 |
| Fraction of all age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 0.01 | 0.01 |  |  |  |  |  |  |  |  | 0.02 |
| Pool | 0.37 | 0.31 |  |  |  |  |  |  |  |  | 0.30 |
| Riffle | 0.11 | 0.05 |  |  |  |  |  |  |  |  | 0.13 |
| M. Chan | 0.51 | 0.63 |  |  |  |  |  |  |  |  | 0.55 |
| Densities of age-0 Chinook within habitat types (fish/ha) |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 133 | 66 |  |  |  |  |  |  |  |  | 171 |
| Pool | 1,569 | 1,300 |  |  |  |  |  |  |  |  | 1,048 |
| Riffle | 190 | 98 |  |  |  |  |  |  |  |  | 163 |
| M. Chan | 2,957 | 3,768 |  |  |  |  |  |  |  |  | 1,855 |
| Number of age-0 Chinook within habitat types |  |  |  |  |  |  |  |  |  |  |  |
| Glide | 1,120 | 518 |  |  |  |  |  |  |  |  | 1,745 |
| Pool | 44,321 | 34,993 |  |  |  |  |  |  |  |  | 24,908 |
| Riffle | 13,085 | 6,017 |  |  |  |  |  |  |  |  | 10,899 |
| M. Chan | 62,713 | 69,969 |  |  |  |  |  |  |  |  | 45,515 |

## Appendix B

Fish Trapping at the Chiwawa and Wenatchee Smolt Traps during 2015

# Monitoring Juvenile Salmonids in the Wenatchee River Subbasin: Activities in the Chiwawa River and Lower Wenatchee River during 2015 

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## INTRODUCTION

## Background

## Monitoring and Evaluation

Productivity indicators in the freshwater environment provide data essential to inform evolving salmon and steelhead hatchery programs. In the Wenatchee River subbasin, the Juvenile Monitoring Component of the Monitoring and Evaluation Plan for PUD Hatchery Programs gathers data directed at informing these productivity indicators (see Hillman et al. 2013). More specifically, this data directly addresses Objective 2 of the monitoring and evaluation framework:
"Determine if the proportion of hatchery fish on the spawning grounds affects the freshwater productivity of supplemented stocks."

## Objectives

The Washington Department of Fish and Wildlife monitors juvenile salmonids in the Wenatchee River subbasin with the primary objective of estimating: natural productivity, migration timing, and age with size at migration. This has occurred at the tributary level (Chiwawa River since 1991) and population level (Wenatchee River since 1997). Target species include spring Chinook Salmon (Oncorhynchus tshawytscha) and summer steelhead (O. mykiss) in the Chiwawa River, and is expanded to include sockeye Salmon (O. nerka) and summer Chinook Salmon ( 0. tshawytscha) in the mainstem Wenatchee River.

Monitoring has primarily been conducted with rotary screw traps that capture emigrating salmonids from spring through fall. In an effort to reduce biases in emigrant estimates, and to improve understanding of survival and movement during non-trapping periods (December through February), WDFW began remote sampling spring Chinook Salmon in the Chiwawa Basin in 2012.

## Study Area

## Chiwawa River

The Chiwawa River is a fourth-order river draining a $474-\mathrm{km}^{2}$ basin and has a mean annual discharge of 14.4 cubic meters per second (cms); contributing about $15 \%$ of the mean annual discharge of the Wenatchee River. The Chiwawa basin is dominated by the snow melt cycle with peak discharge occurring May through July with occasional fall freshets (Figure 1). The Chiwawa River originates in the North Cascades and flows southeast for 60 km before joining the Wenatchee River at river kilometer (rkm) 76, about 9 km downstream of Lake Wenatchee (Figure 2). The Chiwawa River basin is relatively natural, with $96 \%$ managed as part of the Wenatchee National Forest and the upper $32 \%$ designated wilderness.

Precipitation in the basin varies between 76 cm near the confluence and 356 cm at the peaks, while elevations range from 573 to $2,768 \mathrm{~m}$. The river is dynamic with generally shallow pool
riffle segments as it meanders through a U-shaped valley formed by ancient glaciers in the region. Gradients remain well under $1 \%$ for the majority of the river.


Figure 1. Discharge of the Chiwawa River at Plain, USGS gauge \# 12456500. Black line represents 2015 discharge and grey line represents mean discharge from 1990-2014.


Figure 2. Wenatchee River subbasin (with rotary screw trap locations).

## Wenatchee River

The Wenatchee River is a fourth-order river draining a $3,437-\mathrm{km}^{2}$ basin and has a mean annual discharge of 91.4 cms . The hydrograph is dominated by the snow melt cycle with peak discharge occurring May through July with occasional fall freshets (Figure 3). The mainstem originates at the outlet of Lake Wenatchee and flows southeast 84.5 km before joining the Columbia River, 753 km upstream of the Pacific Ocean (Figure 2). While most of the lowlands (17\%) are private, the majority ( $83 \%$ ) of basin is public land.

Precipitation in the basin varies from 22 cm near the Columbia River confluence to 381 cm at the crest of the Cascade Mountains with elevations ranging from 237 to $2,768 \mathrm{~m}$. The Wenatchee River has a relatively low gradient except from rkm 40-64 where the river flows through a bedrock canyon (Tumwater Canyon) and has a gradient of approximately 9.8 meters per kilometer.


Figure 3. Discharge of the Wenatchee River at Monitor, USGS gauge \# 12462500. Black line represents 2015 discharge and grey line represents mean discharge from 1990-2014.

## METHODS

## Rotary Screw Traps

## Trap Operations

The Chiwawa River trap consists of a single 2.4 m cone and has been operating since 1991 at its current location, 0.6 km upstream from the confluence with the Wenatchee River. Trap operations usually begin in late February and continue until ice suspends operations in late fall. The Lower Wenatchee trap consists of two 2.4 m cones and has been operating in its current location (rkm 12.5) since 2013. Trap operations usually begin in late January and continue until fall, when river conditions force its removal.

Operational procedures and techniques follow the standardized basin-wide monitoring plan developed by the Upper Columbia Regional Technical Team for the Upper Columbia Salmon Recovery Board (UCSRB; Hillman 2004), which was adapted from Murdoch and Petersen (2000). The traps remain in operation 24 hours a day unless environmental condition (high/low flow, extreme temperature, and high debris), hatchery releases, mechanical failure or human recreational activities halt operations. During periods of high recreational activities in the spring and summer the Lower Wenatchee trap is pulled during daylight hours to minimize human danger.

Fish Sampling
At a minimum of once a day, all fish collected at the traps were identified to genus or species, enumerated, weighed, and fork length (FL) measured. All salmonids were classified as hatchery, wild, or unknown and visually classified as fry, parr, transitional, or smolt. All hatchery salmonids in the basin are marked (adipose fin-clip, coded-wire tags, or Passive Integrated Transponder (PIT) with the exception of coho. Based on length subsamples of known hatchery coho at Leavenworth Fish Hatchery, all coho collected at the Lower Wenatchee smolt trap were considered wild if $<80 \mathrm{~mm}$ FL or unknown origin if $\geq 80 \mathrm{~mm}$ FL. All coho collected in the Chiwawa River were considered wild. Target species ( $\geq 65 \mathrm{~mm}$ FL) were tagged using 12.5 mm FDX PIT tags and all PIT tagging information was uploaded to a reginal PIT tag database (PTAGIS) maintained by the Pacific States Marine Fisheries Commission.

A combination of age and trap location was used to determine race (spring or summer) of captured juvenile Chinook Salmon. All Chinook Salmon captured in the Chiwawa River trap were considered spring Chinook, regardless of size since summer Chinook Salmon spawning has not been documented upstream of the trap. All yearling (age-1) Chinook captured at the Lower Wenatchee River trap during the spring migration period were considered spring Chinook Salmon because spring Chinook Salmon are yearling migrants and summer Chinook Salmon are typically subyearling migrants. All subyearling fry and parr (age-0) Chinook captured at the Lower Wenatchee River trap during spring were considered summer Chinook Salmon.

## Mark-Recapture Trials

Groups of marked juveniles were released during a range of stream discharges in order to determine trapping efficiencies under the varied flow regime. Natural origin fish were marked with a PIT tag if $\geq 65 \mathrm{~mm}$ FL or stained with Bismarck Brown dye if $<65 \mathrm{~mm}$ FL. Hatchery origin fish were marked using a caudal fin clip. All marked fish were released evenly upstream on both sides of the river between 1800 hours and 2000 hours. Marked fish from the Lower Wenatchee River trap were transported and released 14.5 km upstream of the trap site while fish from the Chiwawa River trap were released 2.6 km upstream. Each trial was conducted over a four-day ( 96 hour) period to allow time for passage or capture. Target mark group sizes were based on historical data, location and species, ranging from 100 to over 500 individual fish.

## Emigrant Estimates

All emigration estimates were calculated using estimated daily trap efficiency derived from the regression formula using trap efficiency (dependent variable) and discharge (independent variable). Trap efficiency models used a modified Bailey estimator (recaptures +1 ) in the calculation of efficiency as a method of bias correction. If a significant relationship ( $R^{2}>0.5$ and $\mathrm{P}<0.05$ ) could not be found a pooled trap efficiency estimate was used. All estimates of emigrating spring Chinook do not include fry due to the uncertainty that these fish were actively migrating to the ocean (UCRTT, 2001). See appendices A and B for detailed equations and information on how the point estimate, variance, and standard error were calculated.

During minor breaks in operation (less than seven days), the number of individual fish collected was estimated. This estimate was calculated using the mean number of fish captured two days prior and two days after the break in operation. For major breaks in operations (greater than seven days), an estimate based on historical run timing was developed. This estimate of daily capture was incorporated into the overall emigration estimate.

## Egg-to-emigrant Survival

The estimated total egg deposition (d) was calculated by multiplying the mean fecundity (f) of the brood spawners by the total number of redds (r) found during surveys (Hillman et al. 2014). Egg-to-emigrant survival (s) was calculated by dividing total emigrants (e) by estimated egg deposition (d).

## Backpack Electrofishing

## Sampling Procedure

From 2012 to present, WDFW has had a goal of PIT tagging 3,000 juvenile spring Chinook Salmon each year. In order to representatively tag the population throughout all reaches, the number of fish tagged in each reach was based on the reach specific abundance encountered during snorkeling surveys in late summer. See Appendix C for further explanation.

## Detections and Calculations

Detections occur at PIT tag interrogation sites in and out of the basin as well as rotary smolt traps downstream of the sampling reaches. Calculations of non-trapping emigrant estimates are based on a flow-detection efficiency regression developed using mark-groups previously released to test smolt trap efficiencies. The total number of tagged fish ( t ) divided by the estimated total parr abundance ( $p$ ), as based off of standard snorkeling techniques (Hillman et al. 2013), resulted in an overall tag rate ( $\mathrm{t}_{\mathrm{i}}$ ). See Appendix C for further explanation.

## RESULTS

## Rotary Screw Traps - Chiwawa

## Trap Operation

The Chiwawa trap operated between 25 February and 24 November 2015. During that time the trap was inoperable for 29 days as a result of low or high discharge, debris and hatchery fish releases. The trap was operated in two positions based on season (i.e., lower position through June 30 and upper position after July 1).

## Fish Sampling

A total of 60,302 individual fish were collected, with wild spring Chinook Salmon and steelhead comprising $62 \%$ and $5 \%$ of the total catch, respectively. Additionally, 7,162 hatchery spring Chinook, 3,151 hatchery steelhead, and 38 wild coho were collected. Throughout the sampling period 18,470 PIT tag were deployed into wild spring Chinook and steelhead (16,675 and 1,795 respectively). Spring Chinook mortality for the season totaled 42 yearling, 390 subyearling parr, and 31 fry ( $0.7 \%, 2.1 \%$, and $0.24 \%$, respectively). Mortality of steelhead throughout the season totaled 45 ( $1.38 \%$ ). The mean fork length (SD) of captured yearling and subyearling spring Chinook Salmon (fry excluded) was 93 (9.0) mm and 71 (10.7) mm, respectively (Table 1).

Table 1. Mean fork length (mm) and weight (g) of spring Chinook Salmon captured in the Chiwawa River smolt trap during 2015.

|  | Yearling transitional/smolts |  |  |  | Subyearling parr |  |  |  |
| :--- | ---: | ---: | ---: | :--- | :--- | :--- | :--- | :--- |
|  | Mean | SD | $N$ |  | Mean | SD | $N$ |  |
| Fork length | 92.5 | 9.0 | 6,304 |  | 71.1 | 10.7 | 15,241 |  |
| Weight | 8.8 | 2.9 | 6,244 |  |  | 4.2 | 1.7 | 14,660 |

## Yearling Spring Chinook (Brood Year 2013)

Wild yearling spring Chinook Salmon were primarily captured between 25 February and 14 June (Figure. 4). A total of 6,350 yearling Chinook Salmon were captured and an estimated 6,891 would have been captured if the trap had operated without interruption. Nine mark/recapture efficiency trials using PIT tags were conducted when the trap was in the lower position producing a mean trap efficiency of 19\%. In 2015, mark/recapture trials were conducted at all desired discharge levels but a statistically significant flow-efficiency regression model could not
be obtained ( $\mathrm{R}^{2}=0.22, \mathrm{P}<0.069$ ). Thus, a pooled estimate combining the 2014 and 2015 mark/recapture trials was developed. The estimated number ( $95 \%$ C.I.) of yearling spring Chinook Salmon that emigrated from the Chiwawa River in 2015 was $39,396( \pm 8,399)$.


Figure 4. Daily catch of yearling spring Chinook Salmon at the Chiwawa River rotary screw trap.

## Subyearling Spring Chinook (Brood Year 2014)

Wild subyearling spring Chinook Salmon were captured throughout the sampling period, with peak catches of parr in October and November and fry occurring in March and April (Figures 5 and 6, respectively). A total of 18,190 subyearling parr and 12,962 fry were captured with an estimated 19,435 subyearling parr and 13,936 fry had the trap operated without interruption. Four mark/recapture efficiency trials were conducted (three PIT and one Bismarck Brown) with a mean trap efficiency of 25.4\%. A combination of mark/recapture efficiency trials from 2014 and 2015 were used to create a regression model for the upper trap position ( $R^{2}=0.58, P=$ 0.002). Data from 2002, 2003, 2013 and 2015 were combined to create a regression model ( $\mathrm{R}^{2}$ $=0.83, P<0.001$ ) for subyearling Chinook captured at the lower trap position. In 2015, the estimated number of subyearling spring Chinook Salmon (excluding fry < 50 mm FL) emigrating from the Chiwawa River during the sampling period was $77,510( \pm 9,074)$.


Figure 5. Daily catch of wild spring Chinook subyearling parr at the Chiwawa River rotary screw trap.


Figure 6. Daily catch of wild spring Chinook fry at the Chiwawa River rotary screw trap.

## Summer Steelhead

During the trapping period, 259 steelhead transitional/smolts and 3,004 steelhead/rainbow parr and fry were captured. While collections occurred in moderate numbers throughout the year, peak collections occurred during October (Figure 7). The mean fork length (SD) of steelhead parr and transitional/smolts captured was 75.8 (23.1) and 167.1 (21.8) mm, respectively (Table. 2).


Figure 7. Daily catch of all wild steelhead at the Chiwawa River rotary screw trap.

Table 2. Mean fork length (mm) and weight (g) and of steelhead/rainbow captured in the Chiwawa River smolt trap during 2015.

|  | Transitional/smolts |  |  | Parr |  |  |  |
| :--- | ---: | ---: | ---: | :--- | ---: | ---: | :---: | :---: |
|  | Mean | SD | $N$ |  | Mean | SD | $N$ |
| Fork length | 167.1 | 21.8 | 256 |  | 75.8 | 23.1 | 2,570 |
| Weight | 50.1 | 19.2 | 252 |  | 6.0 | 7.88 | 2,557 |

## Egg-to-emigrant Survival

For BY 2013, 714 redds were counted in the Chiwawa River with an estimated 3,367,224 eggs being deposited. A total of 113,091 emigrants were estimated resulting in an egg-to-emigrant survival of $3.4 \%$ (Table 3). This is down slightly from a five year moving average of $3.8 \%$.

Table 3. Estimated egg deposition and egg-to-emigrant survival rates for Chiwawa River spring Chinook Salmon.

|  |  |  | Estimated number <br> Brood <br> Year |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number <br> of redds | Estimated <br> egg <br> deposition | Sub- <br> yearling | Non <br> trapping | Yearling | Total <br> emigrants |
| 1992 | 302 | $1,570,098$ | 25,818 | Egg-to- <br> emigrant <br> survival <br> $(\%)$ |  |  |
| 1993 | 106 | 556,394 | 14,036 | 39,723 | 65,541 | 4.2 |
| 1994 | 82 | 485,686 | 8,595 | 8,662 | 22,698 | 4.1 |
| 1995 | 13 | 66,248 | 2,121 | 16,472 | 25,067 | 5.2 |
| 1996 | 23 | 106,835 | 3,708 | 3,830 | 5,951 | 9.0 |
| 1997 | 82 | 374,740 | 16,228 | 15,475 | 19,183 | 18.0 |


| Brood Year | Number of redds | $\begin{aligned} & \text { Estimated } \\ & \quad \text { egg } \\ & \text { deposition } \end{aligned}$ | Estimated number |  |  |  | Egg-toemigrant survival (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Subyearling | Non trapping | Yearling | Total emigrants |  |
| 1998 | 41 | 207,675 | 2,855 |  | 23,068 | 25,923 | 11.9 |
| 1999 | 34 | 166,090 | 4,988 |  | 10,661 | 15,649 | 9.4 |
| 2000 | 128 | 642,944 | 14,854 |  | 40,831 | 55,685 | 8.7 |
| 2001 | 1,078 | 4,836,704 | 459,784 |  | 86,482 | 546,266 | 11.0 |
| 2002 | 345 | 1,605,630 | 93,331 |  | 90,948 | 184,279 | 11.5 |
| 2003 | 111 | 648,684 | 16,881 |  | 16,755 | 33,637 | 5.2 |
| 2004 | 241 | 1,156,559 | 44,079 |  | 72,080 | 116,158 | 10.0 |
| 2005 | 333 | 1,436,564 | 108,595 |  | 69,064 | 177,659 | 12.3 |
| 2006 | 297 | 1,284,228 | 62,922 |  | 45,050 | 107,972 | 8.4 |
| 2007 | 283 | 1,241,521 | 60,196 |  | 25,809 | 86,006 | 6.9 |
| 2008 | 689 | 3,163,199 | 85,161 |  | 35,023 | 120,184 | 3.8 |
| 2009 | 421 | 1,925,233 | 30,996 |  | 30,959 | 61,955 | 3.2 |
| $2010^{\text {a }}$ | 502 | 2,165,628 | 53,619 |  | 47,511 | 101,130 | 4.7 |
| $2011{ }^{\text {a }}$ | 492 | 2,157,420 | 67,982 | 3,665 | 37,185 | 108,832 | 5.0 |
| 2012 ${ }^{\text {a }}$ | 880 | 3,716,240 | 49,774 | 25,305 | 34,334 | 109,413 | 2.9 |
| $2013{ }^{\text {a }}$ | 714 | 3,367,224 | 73,695 | NA | 39,396 | 113,091 | 3.4 |
| $2014{ }^{\text {a }}$ | 485 | 1,961,825 | 77,510 | -- | -- | -- | -- |

${ }^{\text {a }}$ calculated with Bailey model

## Non-target Taxa

Bull trout (Salvelinus confluentus) also comprised a large proportion of incidental species captured. During the trapping period 298 bull trout ( $32 \geq 300 \mathrm{~mm} \mathrm{FL}$ and $266<300 \mathrm{~mm} \mathrm{FL}$ ) were captured. Additionally, a total of 72 western cutthroat trout ( $O$. clarki lewisi), 2 resident rainbow ( $O$. mykiss) and 8 Eastern brook trout (S. fontinalis) were collected. In all, 260 bull trout, and 65 western cutthroat trout were released with PIT tags. Monthly and annual totals of all fish captured are presented in Appendix D and Appendix E, respectively.

## Rotary Screw Traps - Lower Wenatchee

## Trap Operation

The Lower Wenatchee trap operated from 30 January through 27 June 2015. During this time the trap was inoperable for a total of 5 days due to high/low flows, high temperatures, heavy debris and major hatchery releases. Extreme river temperatures and low flows resulted in trapping operations being suspended for the season on 28 June. Throughout the season, the trap cones were operated in the lower position.

Fish Sampling

A total of 282,976 individual fish were collected, with wild summer Chinook Salmon comprising $89 \%$ of the total catch. Additionally, 1,559 wild yearling spring Chinook Salmon, 9,920 hatchery yearling Chinook Salmon, 4,178 wild sockeye, 331 wild steelhead, and 2,288 hatchery steelhead were captured. Throughout the sampling period 5,513 PIT tag were deployed into wild yearling spring Chinook, sockeye and steelhead (1,301; 3,922; and 290 respectively). Mortality for the season totaled 17 yearling spring Chinook, 282 subyearling summer Chinook, 64 sockeye, and 2 steelhead ( $1.1 \%, 0.1 \%, 1.5 \%$, and $0.6 \%$, respectively).

## Wild Yearling Spring Chinook (Brood Year 2013)

Wild yearling spring Chinook Salmon were primarily captured in March and April (Figure 8). Throughout the trapping period 1,559 spring Chinook were collected and an estimated 1,654 would have been collected had the trap operated without interruption. One mark/recapture efficiency trial was carried out using caudal fin clipped yearling hatchery spring Chinook Salmon. A combination of 2013, 2014, and 2015 trials were used to develop a significant relationship between discharge and trap efficiency ( $R^{2}=0.62, P=0.02$ ). This model was used to calculate an emigrant estimate of $58,595( \pm 6,731)$. The mean fork length (SD) of captured yearling Chinook was 96 (9.7) mm (Table 4).


Figure 8. Daily capture of wild yearling Chinook Salmon at the Lower Wenatchee smolt trap.
Table 4. Average length and weight for wild yearling spring Chinook Salmon sampled at the Lower Wenatchee trap.

|  | Mean | SD | N |
| :--- | :---: | :---: | :---: |
| Fork length | 96 | 9.8 | 1,491 |
| Weight | 9.4 | 3.7 | 1,473 |

Wild Subyearling Summer Chinook (Brood Year 2014)
Wild subyearling summer Chinook dominated the catch with 252,293 fish being processed,
most being collected in April and May (Figure 9). An estimated 274,346 would have been captured had the trap operated without interruption. Over the season, eight mark/recapture efficiency trials were carried out using Bismarck Brown during the 2015 trapping season. When combined with trials from the previous trapping season a significant discharge efficiency relationship was developed ( $\mathrm{R}^{2}=0.61, P<0.001$ ) and an emigrant estimate ( $95 \%$ C.I.) of $14,157,778( \pm 2,125,578)$ was calculated. The mean fork length (SD) for captured subyearling parr and fry summer Chinook was 63 (9.7) and 41 (3.3), respectively (Table 5). No PIT tags were deployed in summer Chinook.


Figure 9. Daily capture of wild summer Chinook Salmon at the Lower Wenatchee River trap.

Table 5. Fork length and weight of subyearling Summer Chinook Salmon sampled at the lower Wenatchee smolt trap.

|  | Transition / Smolt |  | Parr |  |  | Fry |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | SD | N | Mean | SD | N | Mean | SD | N |
| Fork length | 75.3 | 7.2 | 8 | 62.8 | 9.7 | 2,011 | 41.0 | 3.3 | 6,267 |
| Weight | 4.36 | 1.3 | 7 | 3.07 | 1.5 | 1,690 | 0.62 | 0.3 | 2,863 |

## Wild Sockeye

A total of 4,178 juvenile sockeye were collected in the 2015 season and an estimated 5,239 had the trap operated without interruption. Almost all of these fish (96\%) were collected in April (Figure 10). Three mark/recapture efficiency trials were carried out using PIT tagged juvenile sockeye Salmon. When combined with efficiency trials from the 2014 and 2013 season a significant discharge efficiency model ( $\mathrm{R}^{2}=0.52, P<0.043$ ) was developed. This model produced an estimate ( $95 \%$ C.I.) of the 2015 emigrant population of juvenile sockeye at $1,065,614( \pm 238,901)$. Smolt survival (SE) to McNary of those tagged fish was $45 \%$ ( $5 \%$ ) using a

Cormack Jolly Seber estimator. Over 90\% of sockeye in run year 2013 and 2014 migrated as Age $1+$ with the remaining being Age $2+$ (Table 6). Mean fork length (SD) for captured sockeye was 86 (9.4) mm (Table 7).


Figure 10. Daily capture of wild sockeye Salmon at the Lower Wenatchee River trap.

Table 6. Age structure and estimated number of wild sockeye smolts that emigrated from Lake Wenatchee in 2013-2015.

| Run year | Proportion of Wild Smolts |  |  | Total Wild Smolts |
| :---: | :---: | :---: | :---: | :---: |
|  | Age 1+ | Age 2+ | Age 3+ |  |
| 2013 | 0.932 | 0.068 | 0.000 | 873,096 |
| 2014 | 0.924 | 0.076 | 0.000 | $1,275,027$ |
| 2015 | NA | NA | NA | $1,065,614$ |

Table 7. Fork length and weight of wild sockeye Salmon smolts sampled at the Lower Wenatchee smolt trap.

|  | Mean | SD | N |
| :--- | :---: | :---: | :---: |
| Fork length | 86.0 | 9.4 | 4,067 |
| Weight | 5.37 | 3.0 | 4,049 |

## Wild Summer Steelhead

Capture of wild steelhead at the Lower Wenatchee site for all life stages was low, totaling 331 smolts, parr, and fry combined and an estimated 339 collected had the trap operated without interruption. Peak catches of steelhead occurred in May (Figure 11). Due to the low captures no mark/recapture trials were conducted in 2015. In 2014 however, two trials using hatchery steelhead transitional/smolts were piloted. Based on these two trials a pooled efficiency of
0.036 was used to estimate ( $95 \%$ C.I.) the emigrant population at $8,632( \pm 45,053)$ parr and smolt emigrant steelhead. However, due to the small number of trials, small sample sizes, use of hatchery transitional/smolts surrogates and the relationship not being significant, caution should be used in the interpretation and use of the estimate. Mean length (SE) of transitional/smolts and parr was 179 (24.8) and 94 (22.7) mm, respectively (Table 8).


Figure 11. Daily capture of wild steelhead at the Lower Wenatchee River trap.

Table 8. Fork length and weight of wild steelhead sampled at the lower Wenatchee smolt trap.

|  | Transitional/Smolt |  |  | Parr |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | SD | N | Mean | SD | N |
| Fork length | 179 | 24.8 | 227 | 94 | 22.7 | 74 |
| Weight | 60.24 | 25.6 | 226 | 10.39 | 9.4 | 71 |

## Survival

For BY 2013, 1,159 spring Chinook Salmon redds were surveyed in the Wenatchee Basin producing an estimated 5,512,204 eggs. An estimate of 58,595 emigrants results in an estimated egg-to-emigrant survival of $1.06 \%$. This is down from the last two year average of 1.65\% (Table 9).

Table 9. Estimated egg deposition and egg-to-smolt survival rates for Wenatchee Basin spring Chinook Salmon.

| Brood <br> Year | Number <br> of redds | Estimated egg <br> deposition | Estimated number  <br>  $c c o$Total <br> emigrants | Egg-to-emigrant <br> survival (\%) |
| :---: | :---: | :---: | :---: | :---: |
|  | 350 | $1,758,050$ | 76,643 | 4.36 |
| 2001 | 1,876 | $8,674,624$ | 243,516 | 2.81 |
| 2002 | 1,139 | $5,300,906$ | 165,116 | 3.11 |


| Brood <br> Year | Number <br> of redds | Estimated egg <br> deposition | Total <br> emigrants | Egg-to-emigrant <br> survival $(\%)$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 323 | $1,887,612$ | 70,738 | 3.75 |
| 2004 | 555 | $2,663,445$ | 55,619 | 2.09 |
| 2005 | 829 | $3,587,083$ | 302,116 | 8.42 |
| 2006 | 588 | $2,542,512$ | 85,558 | 3.37 |
| 2007 | 466 | $2,069,506$ | 60,219 | 2.91 |
| 2008 | 1,411 | $6,479,312$ | 82,137 | 1.27 |
| 2009 | -- | -- | -- | -- |
| 2010 | -- | -- | -- | -- |
| 2011 | 872 | $3,823,720$ | 89,917 | 2.35 |
| 2012 | 1,704 | $7,195,992$ | 67,973 | 0.94 |
| 2013 | 1,159 | $5,512,204$ | 58,595 | 1.06 |

For BY 2014, 3,458 summer Chinook Salmon redds were surveyed in the Wenatchee Basin, 95.9\% being upstream of the Lower Wenatchee smolt trap. After extrapolating by the proportion of redds above the trap a total emigrant population of 14,763,064 was estimated resulting in an egg-to-emigrant survival of $89.17 \%$. This is up from the last two year average of 80.73\% (Table 10).

Table 10. Estimated egg deposition and egg-to-emigrant survival rates for Wenatchee Basin summer Chinook Salmon.

|  |  |  | Redds | Estimated number |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Brood <br> year | Peak total <br> redd <br> above | Estimated egg <br> deposition <br> trap/ <br> total <br> redds | Trap estimate | Total <br> emigrants | Egg-to- <br> emigrant <br> survival <br> $(\%)$ |  |
| 1999 | 2,738 | $13,654,406$ | 0.988 | $9,572,392$ | $9,685,591$ | 70.93 |
| 2000 | 2,540 | $13,820,140$ | 0.983 | $1,299,476$ | $1,322,383$ | 9.57 |
| 2001 | 3,550 | $18,094,350$ | 0.987 | $8,229,920$ | $8,340,342$ | 46.09 |
| 2002 | 6,836 | $37,488,624$ | 0.977 | $13,167,855$ | $13,475,368$ | 35.95 |
| 2003 | 5,268 | $28,241,748$ | 0.996 | $20,336,968$ | $20,426,149$ | 72.33 |
| 2004 | 4,874 | $26,207,498$ | 0.989 | $14,764,141$ | $14,935,745$ | 56.99 |
| 2005 | 3,538 | $17,877,514$ | 0.993 | $11,612,939$ | $11,695,581$ | 65.42 |
| 2006 | 8,896 | $45,663,168$ | 0.979 | $9,397,044$ | $9,595,512$ | 21.01 |
| 2007 | 1,970 | $10,076,550$ | 0.983 | $4,470,672$ | $4,546,838$ | 45.12 |
| 2008 | 2,800 | $14,302,400$ | 0.978 | $4,309,496$ | $4,405,473$ | 30.8 |
| 2009 | 3,441 | $18,206,331$ | 0.983 | $6,695,977$ | $6,814,805$ | 37.43 |
| 2010 | 3,261 | $16,184,343$ | 0.957 | -- | -- | -- |
| 2011 | 3,078 | $15,122,214$ | 0.958 | -- | - | -- |
| 2012 | 2,504 | $12,021,704$ | 0.93 | $9,333,214$ | $10,034,508$ | 83.47 |


|  |  |  | Redds |  | Estimated number |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Brood <br> year | Peak total <br> redd <br> expansion | Estimated egg <br> deposition | trap <br> tral <br> total | Trap estimate | Total <br> emigrants | Egg-to- <br> emigrant <br> survival |
|  |  |  | redds |  |  | $(\%)$ |
| 2013 | 3,241 | $16,162,867$ | 0.947 | $11,936,928$ | $12,605,925$ | 77.99 |
| 2014 | 3,458 | $16,556,904$ | 0.959 | $14,157,778$ | $14,763,064$ | 89.17 |

## Non-target Taxa

One westslope cutthroat trout was sampled at the Lower Wenatchee site and no bull trout where sampled. No PIT tags were applied to non-target taxa. Monthly and annual totals of all fish captured are presented in Appendix F and Appendix G, respectively.

## Backpack Electrofishing

## Fish Sampling

Between 1 October and 17 November 2014, WDFW personnel sampled the Chiwawa River over a 13-day span for a total of 55,895 seconds. During this sampling 1,019 subyearling spring Chinook received a PIT tag. The majority of the sampling (95\%) occurred between rkm 35 and 55. The greatest concentration of juvenile Chinook occurred between rkm 50 and 53 which had a mean sample rate of one Chinook collected for every 53 seconds of sampling. Over the sample period 14 Chinook died resulting in a mortality rate of $1.3 \%$. Additionally, 121 juvenile bull trout and 94 steelhead were collected, with 67 bull trout and 23 steelhead receiving PIT tags. Highest catch rates for bull trout were between rkm 42 and 47 while the lowest site sampled (rkm 11) had the highest catch rate of steelhead. There was no mortality associated with bull trout or steelhead.

## Detections and Calculations

Between the non-trapping season of 18 November 2014 through 24 February 2015, a total of 16 detections of remotely tagged Chinook were recorded at the lower Chiwawa antenna array. During the trapping season of 17 October and 6 November 2014, and 13 March and 6 June 2015, the Chiwawa rotary smolt trap collected 17 and 47 remotely tagged Chinook, respectively. Due to uneven distribution of effort throughout the Chiwawa River and poor sample size, no emigrant estimate for the non-trapping period was calculated for the BY 2013.

## DISCUSSION

## Chiwawa River Smolt Trap

Over the last five years the Chiwawa River smolt trap has had an average installation date of 3 March. With the relatively mild spring in 2015, the smolt trap was installed almost a week earlier on 25 February. The 2015 trapping season provided relatively good trapping conditions with two minor stoppages in the spring (due to hatchery releases) and two minor stoppages in the fall (due to high discharge and debris). The Chiwawa River smolt trap is considered operable
between discharges of 90 and $1,500 \mathrm{cfs}$, and the only significant stoppages occurred between mid-September and mid-October when flow periodically dropped below 90 cfs.

A significant discharge efficiency model was produced for subyearling Chinook and a pooled estimate was used for yearling Chinook. Historically, emigrant estimates were calculated using the Peterson estimator of abundance (Seber 1982), however more accurate estimates currently utilize a modified Bailey estimator (Murdoch et al. 2012).

The total production estimate for brood year 2013 was 119,615 and comprises estimates of subyearling emigrants in 2014 and yearling emigrants in 2015. Unfortunately, high flows and the inability to electrofish the Chiwawa River due to spawning bull trout concerns resulted in an abbreviated sampling window and prevented the completion of 2014 remote tagging efforts. This resulted in no estimate being calculated for the 2014 non-trapping season and a known underestimate of the total brood year production. Protocols and field sampling will be continually adapted to fit within environmental and permit constraints and estimates will be improved upon when possible.

Abnormally low discharge levels also limited the number of mark/recapture trials that could be done at the Chiwawa River smolt trap and reliance on historical data was necessary. Further complicating estimates, emigrating yearling and subyearling Chinook were collected when the trap was operating at both the upper and lower cone positions. However, insufficient numbers were present to produce a trap efficiency model for both life stages at each cone positions. In an effort to expand operational condition and reduce the dependence on historic data, 2016 trap operations will eliminate the lower cone position and a single upper cone positon will be used.

## Lower Wenatchee River Smolt Trap

Historically, the smolt trap on the mainstem Wenatchee River has moved location numerous times due to poor trap efficiencies of target species and environmental factors causing abbreviated trapping seasons. At the lower Wenatchee site, the smolt trap has been able to operate into September in 2013 and October in 2014. This marks a relatively large increase in operational length over the old site (located 2.5 km downstream) which had an average trap removal date of 14 August. However, 2015 proved to be a difficult trapping season for the Lower Wenatchee trap. Up until late June the Lower Wenatchee trap only had three minor stoppages due to hatchery fish releases and debris. However, the Lower Wenatchee trap is considered operable between discharges of 1,300 and 10,000 cfs and summer proved to be a substantial departure from normal discharge and river temperature. From late June through July water temperatures at our Lower Wenatchee trapping site fluctuated between 18 and 26 degrees Celsius and discharge was about $25 \%$ of normal. The culmination of these factors resulted in trapping operations terminating at its earliest known date of 28 June.

The early removal of the lower Wenatchee trap proved to be the most difficult part of the 2015 trapping season. To account for the early removal of the trap, historical run timing was used to extrapolate what the catch would have been had the trap been able to operate as normal. Historical emigration timing showed no sockeye, and only a small percentage of spring and
summer Chinook emigrated after 28 June ( $0.4 \%$ and $3.5 \%$, respectively). Emigration estimates used these percentages to extrapolate to a total estimate of emigrants had the trap been able to operate further into the season.

Discharge efficiency models were obtained for three of the four target species at the lower Wenatchee trap during the 2015 trapping season (wild spring and summer Chinook Salmon and sockeye Salmon). Collections of wild steelhead continue to be inadequate for conducting a mark/recapture trial. In 2016, hatchery steelhead from the Chiwawa acclimation site will be used in mark/recapture trials in an effort to improve emigrant estimates of this target species. This approach requires the assumption that hatchery fish behave in a similar manner to wild fish, an assumption we will test over time as possible. While the new trap location has allowed for greater operational flexibility, it does require the development of new flow-efficiency models. While this can be accomplished relatively quickly with species that are relatively abundant (e.g., summer Chinook and sockeye), it may take several years for those in low abundance (e.g., steelhead). Fortunately, given similar operation parameters across time, we will be able to reexamine past abundance estimates when those models are fully developed.

## Backpack Electrofishing

Remote sampling in the Chiwawa Basin started in 2012. Some success occurred early on with PIT tag targets being met, however, there have been substantial obstacles since 2013. Permit restrictions limit field operations until bull trout spawning has concluded; which typically occurs early October. At this time, weather becomes increasingly unfavorable and elevated discharge and cold air and water temperatures hinder sampling efforts. In 2014, early high water events halted sampling efforts and limited not only the area that was sampled, but also the number of fish that were processed. Future investigations will look into alternative sampling techniques and the allocation of personnel to maximize sampling efforts in the basin.

## REFERENCES

Anderson, R. O., \& Neumann, R. M. (1996). Length, Weight, and Associated Structural Indices. In Murphy B. E. \& Willis D. W. (Eds.), Fisheries Techniques (2nd ed. pp. 461-480). American Fisheries Society.

Bailey, N.T.J. 1951. On estimating the size of mobile populations from capture-recapture data. Biometrika 38:293-306.

Hillman, T.W. 2004. Monitoring strategy for the Upper Columbia Basin: Draft report February 1, 2004. Prepared for Upper Columbia Regional Technical Team, Wenatchee, Washington.

Hillman, T., T. Kahler, G. Mackey, J. Murauskas, A. Murdoch, K. Murdoch, T. Pearsons, and M. Tonseth. 2013. Monitoring and Evaluation plan for PUD hatchery programs; 2013 update. Report to the HCP and PRCC Hatchery Committees, Wenatchee, WA.

Hillman, T., M. Miller, C. Moran, M. Tonseth, M. Hughes, A. Murdoch, L. Keller, C. Willard, B. Ishida, C. Kamphaus, T. Pearsons, and P. Graff. 2014. Monitoring and evaluation of the Chelan and Grant County PUDs hatchery programs: 2013 annual report. Report to the HCP and PRCC Hatchery Committees, Wenatchee, WA.

Murdoch, A., and K. Petersen. 2000. Freshwater Production and Emigration of Juvenile Spring Chinook from the Chiwawa River in 2000. Washington State Department of Fish and Wildlife

Murdoch, A.R., Miller, T.L., Truscott, B.L., Snow, C., Frady, C., Ryding, K., Arterburn, J.E., Hathaway, D. 2012. Upper Columbia Spring Chinook Salmon and Steelhead Juvenile and Adult Abundance, Productivity, and Spatial Structure Monitoring. BPA Project No. 2010-034-00. Bonneville Power Administration, Portland, Oregon.

Seber, G.A.F. 1982. The estimation of animal abundance and related parameters, 2nd edition. Macmillan Co., New York, New York, USA

UCRTT (Upper Columbia Regional Technical Team). 2001. A Strategy to Protect and Restore Salmonid Habitat in the Upper Columbia Region, a Discussion Draft Report. Upper Columbia Salmon Recovery Board.

## APPENDICES

## Appendix A. Peterson Population and Variance Equations.

Trap efficiency was calculated using the following formula:

$$
\text { Trap efficiency }=E_{i}=R / M i,
$$

Where $E_{i}$ is the trap efficiency during time period $i ; M_{i}$ is the number of marked fish released during time period $i$; and $R_{i}$ is the number of marked fish recaptured during time period $i$. The number of fish captured was expanded by the estimated daily trap efficiency ( $e$ ) to estimate the daily number of fish migrating past the trap using the following formula:

$$
\text { Estimated daily migration }=\hat{N}_{i}=C_{i} / \hat{e}_{i}
$$

where $N_{i}$ is the estimated number of fish passing the trap during time period $i ; C_{i}$ is the number of unmarked fish captured during time period $i$; and $e_{i}$ is the estimated trap efficiency for time period $i$ based on the regression equation.

The variance for the total daily number of fish migrating past the trap was calculated using the following formulas:

Variance of daily migration estimate $=$

$$
\operatorname{var}\left[\hat{N}_{i}\right]=\hat{N}_{i}^{2} \frac{\operatorname{MSE}\left(1+\frac{1}{n}+\frac{\left(X_{i}-\bar{X}\right)^{2}}{(n-1) s_{\mathrm{x}}^{2}}\right)}{\hat{e}_{i}^{2}}
$$

where $X_{i}$ is the discharge for time period $i$, and $n$ is the sample size. If a relationship between discharge and trap efficiency was not present (i.e., $P<0.05$; $r^{2}$ @0.5), a pooled trap efficiency was used to estimate daily emigration:

$$
\text { Pooled trap efficiency }=e_{p}=\sum R / \sum M
$$

The daily emigration estimate was calculated using the formula:

$$
\text { Daily emigration estimate }=\hat{N}_{i}=C_{i} / e_{p}
$$

The variance for daily emigration estimates using the pooled trap efficiency was calculated using the formula:

Variance for daily emigration estimate $=$

$$
\operatorname{var}\left[\hat{N}_{i}\right]=\hat{N}_{i}^{2} \frac{e_{p}\left(1-e_{p}\right) / \sum M}{e_{p}^{2}}
$$

The total emigration estimate and confidence interval was calculated using the following formulas:

$$
\begin{gathered}
\text { Total emigration estimate }=\sum \hat{N}_{i} \\
95 \% \text { confidence interval }=1.96 \times \sqrt{\sum \operatorname{var}}\left[\hat{N}_{i}\right]
\end{gathered}
$$

## Appendix B. Bailey Population and Variance Equations.

Trap efficiency was calculated using the following formula:

$$
\begin{gathered}
\qquad \text { Trap efficiency }=E_{i}=R+1 / \mathrm{Mi}, \\
\text { Estimated daily emigration }=\hat{N}_{i}=\frac{C_{i}+1}{\hat{e}_{i}}
\end{gathered}
$$

The variance of the total population abundance was calculated as follows:

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)=\underbrace{\sum_{i} \operatorname{Var}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}\right)}_{\text {Part } A}+\underbrace{\sum_{i} \sum_{j} \operatorname{Cov}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}, \frac{\left(C_{j}+1\right)}{\hat{e}_{j}}\right)}_{\text {Part } B}
$$

Part A is the variance of the daily estimates where $C_{i}$ is the number of fish caught in period $i, \mathrm{e}_{\mathrm{i}}$ is the estimated trap efficiency for period $i$, and Cov is the between day covariance for days that the same linear model is used (part B). For a more details and derivation of Peterson and Bailey estimation methods see Murdoch et al. (2012).

## Appendix C. Emigration during non-trapping periods.

A flow-efficiency regression model was developed for the lower Chiwawa River PIT tag interrogation site ( CHL ) using the same mark/recapture trials used for estimating efficiency at the smolt trap. This CHL model was used to calculate emigration outside of the trapping period by incorporating the tag rate into the Bailey estimator.

$$
\begin{aligned}
& \text { Estimated daily emigration }=\left(\hat{N}_{i}=\frac{C_{i}+1}{\hat{e}_{i}}\right) / t_{i} \\
& \text { Where } \mathrm{t}_{\mathrm{i}} \text { is equal to the tag rate }=t_{i}=\frac{t}{p}
\end{aligned}
$$

Appendix D. Monthly collection information for the Chiwawa River smolt trap.

|  |  | 2015 |  |  |  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Species/Origin | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Total |  |
| Chinook |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Wild yearling | -- | 55 | 1,839 | 3,072 | 1,277 | 94 | 13 | 0 | 0 | 0 | 0 | 6,350 |  |
| Wild subyearling | -- | 83 | 3,516 | 7,639 | 352 | 5,509 | 3,058 | 1,423 | 641 | 5,340 | 3,591 | 31,152 |  |
| $\quad$ Hatchery yearling | -- | 0 | 0 | 7,141 | 1 | 2 | 4 | 8 | 6 | 0 | 0 | 7,162 |  |
| Steelhead |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Wild |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Smolt | -- | 0 | 9 | 59 | 163 | 8 | 6 | 12 | 2 | 0 | 0 | 259 |  |
| $\quad$ Parr and fry | -- | 2 | 45 | 200 | 416 | 447 | 283 | 453 | 168 | 538 | 452 | 3,004 |  |
| $\quad$ Hatchery | -- | 0 | 1 | 630 | 2,433 | 63 | 4 | 12 | 3 | 4 | 1 | 3,151 |  |
| Coho |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Wild |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Smolt | -- | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |  |
| $\quad$ Parr and fry | -- | 0 | 1 | 2 | 8 | 22 | 3 | 2 | 0 | 0 | 0 | 38 |  |
| Hatchery | -- | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |  |
| Bull trout |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\quad$ Juvenile | -- | 0 | 9 | 1 | 4 | 7 | 18 | 13 | 14 | 147 | 53 | 266 |  |
| $\quad$ Adult | -- | 0 | 0 | 0 | 0 | 0 | 2 | 5 | 10 | 14 | 1 | 32 |  |
| Westslope cutthroat | -- | 0 | 3 | 0 | 6 | 8 | 22 | 24 | 8 | 0 | 1 | 72 |  |
| Eastern brook trout | -- | 0 | 0 | 1 | 4 | 1 | 0 | 0 | 0 | 1 | 1 | 8 |  |
| Rainbow trout | -- | 0 | 0 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 2 |  |
| Mountain whitefish | -- | 0 | 3 | 17 | 6 | 44 | 2,407 | 2,619 | 355 | 42 | 51 | 5,544 |  |
| Longnose dace | -- | 1 | 21 | 33 | 636 | 661 | 197 | 369 | 255 | 415 | 75 | 2,663 |  |
| Northern pikeminnow | -- | 0 | 0 | 0 | 1 | 16 | 157 | 150 | 7 | 0 | 0 | 331 |  |
| Sculpin spp. | -- | 0 | 8 | 0 | 13 | 40 | 48 | 23 | 13 | 58 | 22 | 225 |  |
| Sucker spp. | -- | 0 | 0 | 0 | 0 | 0 | 11 | 16 | 1 | 2 | 0 | 30 |  |
| Redside shiner | -- | 0 | 0 | 0 | 0 | 0 | 1 | 11 | 0 | 1 | 0 | 13 |  |


| Appendix E. Annual collection information from the Chiwawa River smolt trap. |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Species origin | 2015 | 2014 | 2013 | 2012 | 2011 | 2010 |
| Chinook |  |  |  |  |  |  |
| $\quad$ Wild yearling | 6,350 | 5,419 | 3,199 | 7,626 | 4,848 | 6,482 |
| $\quad$ Wild subyearling | 31,152 | 23,755 | 27,621 | 14,831 | 20,561 | 13,344 |
| $\quad$ Hatchery yearling | 7,162 | 5,293 | 15,909 | 30,751 | 25,620 | 22,481 |
| Steelhead |  |  |  |  |  |  |
| $\quad$ Wild |  |  |  |  |  |  |
| $\quad$ Smolt | 3,263 | 1,938 | 2,034 | 1,921 | 1,176 | 1,226 |
| $\quad$ Parr and Fry | 259 | 49 | 85 | 183 | 195 | 210 |
| $\quad$ Hatchery | 3,004 | 1,889 | 1,949 | 1,738 | 981 | 1,016 |
| Coho | 3,151 | 290 | 1,539 | 1,664 | 8,250 | 9,921 |
| $\quad$ Wild yearling |  |  |  |  |  |  |
| Wild subyearling | 0 | 0 | 1 | 1 | 3 | 4 |
| $\quad$ Hatchery yearling | 38 | 12 | 0 | 0 | 4 | 5 |
| Bull trout | 0 | 1 | 10 | 3 | 0 | 3 |
| $\quad$ Juvenile |  |  |  |  |  |  |
| $\quad$ Adult | 266 | 260 | 310 | 488 | 351 | 499 |
| Westslope cutthroat trout | 32 | 75 | 51 | 31 | 7 | 45 |
| Eastern brook trout | 72 | 59 | 86 | 60 | 38 | 54 |
| Mountain whitefish | 8 | 12 | 13 | 66 | 3 | 0 |
| Longnose dace | 5,544 | 2,970 | 2,108 | 3,291 | 990 | 778 |
| Northern pikeminnow | 2,663 | 2,633 | 2,257 | 1,762 | 1,526 | 1,393 |
| Sculpin spp. | 331 | 5 | 71 | 34 | 20 | 5 |
| Sucker spp. | 225 | 131 | 91 | 157 | 129 | 51 |
| Redside shiner | 30 | 4 | 6 | 0 | 0 | 0 |
| Yellow perch | 13 | 0 | 0 | 0 | 0 | 0 |

Appendix F. Monthly collection information for the Lower We natchee River smolt trap. 2015

| Species/Origin | Jan | Feb | Mar | Apr | May | Jun |  |  | Sep |  | Nov | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Chinook |  |  |  |  |  |  |  |  |  |  |  |  |
| Wild yearling | 9 | 154 | 405 | 751 | 220 | 20 | -- | -- | -- | -- | -- | 1,559 |
| Wild subyearling | 5 | 418 | 8,418 | 154,499 | 69,035 | 19,918 | -- | -- | -- | -- | -- | 252,293 |
| Hatchery yearling | 0 | 15 | 0 | 8,973 | 931 | 1 | -- | -- | -- | -- | -- | 9,920 |

Steelhead
Wild

| Smolt | 0 | 3 | 4 | 33 | 186 | 5 | -- | -- | -- | -- | -- | 231 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parr and fry | 1 | 18 | 16 | 15 | 19 | 31 | -- | -- | -- | -- | -- | 100 |
| Hatchery | 0 | 0 | 7 | 247 | 1,991 | 43 | -- | -- | -- | -- | -- | 2,288 |

Sockeye
Wild $\quad 0 \quad 0 \quad 35 \quad 3,997 \quad 146 \quad 0 \quad$-- - -- - -- - -- $--\quad 4,178$
Coho
Wild

| Smolt | 1 | 6 | 10 | 5 | 0 | 0 | -- | -- | -- | -- | -- | 22 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\quad$ Fry and parr | 2 | 280 | 313 | 968 | 2,153 | 1,256 | -- | -- | -- | -- | -- | 4,972 |
| Hatchery | 0 | 0 | 76 | 4,653 | 1,794 | 43 | -- | -- | -- | -- | -- | 6,566 |
| Unknown | 0 | 0 | 0 | 16 | 121 | 6 | -- | -- | -- | -- | -- | 143 |

Bull trout
Juvenile $\quad 0 \quad 0 \quad 0 \quad 0 \quad 0 \quad 0 \quad$-- $--\quad$-- - -- $--\quad 0$
Adult $\begin{array}{lllllllllllll}0 & 0 & 0 & 0 & 0 & 0 & -- & -- & -- & -- & -- & 0\end{array}$
Westslope cutthroat trout $0 \begin{array}{lllllllllllll} & 0 & 0 & 0 & 1 & 0 & -- & -- & -- & -- & -- & 1\end{array}$
$\begin{array}{lllllllllllll}\text { Mountain whitefish } & 0 & 0 & 0 & 1 & 3 & 5 & -- & -- & -- & -- & -- & 9\end{array}$
$\begin{array}{lllllllllllll}\text { Lamprey spp. } & 1 & 77 & 64 & 12 & 13 & 116 & -- & -- & -- & -- & -- & 283\end{array}$
$\begin{array}{lllllllllllll}\text { Longnose dace } & 1 & 29 & 6 & 5 & 49 & 152 & -- & -- & -- & -- & -- & 242\end{array}$
$\begin{array}{lllllllllllll}\text { Sculpin spp. } & 0 & 16 & 7 & 5 & 8 & 16 & -- & -- & -- & -- & -- & 52\end{array}$
$\begin{array}{lllllllllllll}\text { Sucker spp. } & 1 & 11 & 2 & 2 & 24 & 11 & -- & -- & -- & -- & -- & 51\end{array}$
$\begin{array}{lllllllllllll}\text { Redside shiner } & 0 & 0 & 0 & 4 & 2 & 13 & -- & -- & -- & -- & -- & 19\end{array}$
Stickleback (3-spined) $0 \begin{array}{lllllllllllll} & 0 & 0 & 0 & 2 & 11 & -- & -- & -- & -- & -- & 13\end{array}$
Northern pikeminnow $\begin{array}{lllllllllllll} & 0 & 2 & 0 & 2 & 5 & 3 & -- & -- & -- & -- & -- & 12\end{array}$
Chiselmouth $0 \begin{array}{llllllllllll} & 0 & 0 & 0 & 0 & 6 & -- & -- & -- & -- & -- & 6\end{array}$

| Peamouth | 0 | 0 | 0 | 0 | 0 | 3 | -- | -- | -- | -- | -- | 3 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

Appendix G. Annual collection information from the Lower Wenatchee River smolt trap.

| Species/Origin | 2015 | 2014 | 2013 |
| :--- | :---: | :---: | :---: |
| Chinook |  |  |  |
| $\quad$ Wild yearling | 1,559 | 1,700 | 1,854 |
| $\quad$ Wild subyearling | 252,293 | 81,445 | 52,652 |
| $\quad$ Hatchery yearling | 9,920 | 31,290 | 13,979 |
| Steelhead |  |  |  |
| $\quad$ Wild |  |  |  |
| $\quad$ Smolt | 331 | 182 | 710 |
| $\quad$ Parr | 231 | 80 | 173 |
| $\quad$ Hatchery | 100 | 102 | 537 |
| Sockeye | 2,288 | 494 | 819 |
| $\quad$ Wild |  |  |  |
| $\quad$ Hatchery | 4,178 | 7,678 | 4,520 |
| Coho | 0 | 0 | 72 |
| $\quad$ Wild yearling | 22 | 220 | 597 |
| Wild subyearling | 4,972 | 393 | 923 |
| $\quad$ Hatchery yearling | 6,566 | 16,908 | 12,960 |
| $\quad$ Unknown yearling | 143 | NA | NA |
| Bull trout |  |  |  |
| $\quad$ Juvenile | 0 | 3 | 6 |
| $\quad$ Adult | 0 | 0 | 0 |
| Westslope cutthroat trout | 1 | 3 | 0 |
| Mountain whitefish | 9 | 27 | 110 |
| Lamprey spp. | 283 | 292 | 762 |
| Longnose dace | 242 | 541 | 1,382 |
| Sculpin spp. | 52 | 128 | 242 |
| Sucker spp. | 51 | 134 | 240 |
| Redside shiner | 19 | 94 | 423 |
| Stickleback (3-spined) | 13 | 66 | 196 |
| Northern pikeminnow | 12 | 37 | 39 |
| Chiselmouth | 6 | 69 | 10 |
| Peamouth | 3 | 9 | 10 |
|  |  |  |  |

## Appendix C

Summary of PIT-Tagging Activities in the Wenatchee Basin, 2015

Appendix C. Numbers of fish captured, PIT tagged, lost, and released in the Wenatchee River basin during February through November, 2015.
$\left.\begin{array}{|c|c|c|c|c|c|c|c|c|}\hline \begin{array}{c}\text { Sampling } \\ \text { Location }\end{array} & \text { Species and Life Stage } & \begin{array}{c}\text { Number } \\ \text { collected }\end{array} & \begin{array}{c}\text { Number of } \\ \text { recaptures }\end{array} & \begin{array}{c}\text { Number } \\ \text { tagged }\end{array} & \begin{array}{c}\text { Number } \\ \text { died }\end{array} & \begin{array}{c}\text { Shed } \\ \text { tags }\end{array} & \begin{array}{c}\text { Total } \\ \text { releass } \\ \text { Percent }\end{array} \\ \text { mortality }\end{array}\right]$

| Sampling Location | Species and Life Stage | Number collected | Number of recaptures | Number tagged | Number died | Shed tags | Total tags released | Percent mortality |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Total: | Wild Subyearling Chinook | 286,401 | 263 | 12,983 | 771 | 1 | 12,982 | 0.27 |
|  | Wild Yearling Chinook | 8,095 | 219 | 7,681 | 66 | 0 | 7,681 | 0.82 |
|  | Wild Steelhead/Rainbow | 4,023 | 7 | 2,474 | 27 | 1 | 2,474 | 0.67 |
|  | Hatchery Steelhead/Rainbow | 5,888 | 2 | 2 | 1 | 0 | 2 | 0.02 |
|  | Wild Coho | 5,184 | 3 | 121 | 20 | 0 | 3 | 0.39 |
|  | Wild Sockeye | 4,178 | 3 | 3,922 | 64 | 0 | 3,922 | 1.53 |
| Grand Total: |  | 313,769 | 497 | 27,183 | 949 | 2 | 27,064 | 0.30 |

## Appendix D

Wenatchee Steelhead Spawning Escapement Estimates, 2015

# Wenatchee Steelhead Spawning Escapement Estimates in 2015 

Kevin See

March 15, 2016

## Introduction

Redd counts are an established method to provide an index of adult spawners (Gallagher et al. 2007). In the Wenatchee and Methow subbasins, index reaches are surveyed weekly during the steelhead spawning season (Mar 09, 2015 - May 28, 2015) and non-index reaches are surveyed once during the peak spawning period. The goal of this work is to:

- Predict observer net error, based on a model developed with data from steelhead redd surveys in the Methow, similar to that described in Murdoch et al. (2014).
- Use estimates of observer net error rates and the mean survey interval to estimate the number of redds in each index reach, using a Gaussian area under the curve (GAUC) technique described in Millar et al. (2012).
- Estimate the total number of redds in the non-index reaches by adjusting the observed counts with the estimated net error.
- Convert these estimates of redds in the mainstem areas (surveyed for redds) into estimates of spawners.
- Use PIT-tag based estimates of escapement for all tributaries in the Wenatchee, and combine those estimates with the redd-based estimates of spawners in the mainstem areas to estimate the total number of spawners in the Wenatchee.


## Methods

## Mainstem areas

The model for observer net error (observed redd counts / true number of redds) is a model averaging of the two best models that were fit to 43 data points in the Methow. Both models contained covariates of observed redd density (redds / m) and mean thalweg CV as a proxy for channel complexity. One model also contained discharge while the other also contained total redd survey experience as an additional covariate. Predictions were made using model averaged coefficients (based on AICc model weights) and the 2015 steelhead data. From these survey specific estimates of net error, a mean and standard error of net error was calculated for each reach. The standard deviation was calculated by taking the square root of the sum of the squared standard errors for all predictions within a reach.

Estimates of total redds were made for each index reach using the GAUC model described in Millar et al. (2012). The GAUC model was developed with spawner counts in mind. As it is usually infeasible to mark every individual spawner, only total spawner counts can be used, and an estimate of average stream life must be utilized to translate total spawner days to total unique spawners. However, in adapting this for redd surveys, two modifications could be used. The first would fit GAUC models to data showing all visible redds at each survey, and use an estimate of redd life as the equivalent of spawner stream life. However, because conditions led to many redds not disappearing before the end of the survey season, the estimates of redd life are biased low for this year. The second method relies on the fact that individual redds can be marked, and therefore the GAUC model can be fit to new redds only. The equivalent of stream life thus became the mean and standard deviation of the survey interval. We utilized the second method for this analysis.

For non-index reaches, which were surveyed only once during peak spawning, the estimate of total redds was calculated by dividing the observed redds by the estimate of net error associated with that survey. This assumes that no redds were washed out before the nonindex survey, and that no new redds appeared after that survey. As the number of redds observed in the non-index reaches ranged from 0 to 5 , any violation of this assumption should not affect the overall estimates very much. Based on the peak spawning time for the associated index reaches, the surveys in the non-index reaches were conducted either at peak spawning, or within 10 days after peak spawning (Figure 2\}).

To convert estimates of total redds into estimates of natural and hatchery spawners, total redds were multiplied by a fish per redd ( FpR ) estimate and then by the proportion of hatchery or wild fish. The fish per redd estimate was based on PIT tags from the branching patch-occupancy model (see below) observed to move into the lower or upper Wenatchee (below or above Tumwater dam). FpR was calculated as the ratio of male to female fish, plus 1. This was 1.78 above Tumwater dam, and 1.73 below Tumwater. Reaches W1 - W7 are below Tumwater, while reaches W8 - W10 are above Tumwater. Similarly, the proportion of hatchery and natural origin fish was calculated from the same group of PIT tags for areas above and below Tumwater. The proportion of hatchery origin fish was 0.6 above Tumwater dam, and 0.34 below Tumwater (Table 2).

## Tributary areas

Estimates of escapement to various tributaries in the Wenatchee were made using a branching patch-occupancy model based on PIT tag observations of fish tagged at Priest Rapids dam. All fish that escaped to the various tributaries were assumed to be spawners (i.e. pre-spawn mortality only occurs in the mainstem).

## Total spawners

When summing spawner estimates from index reaches to obtain estimates of total spawners in the Wenatchee, an attempt was made to incorporate the fact that the reaches within a stream are not independent. Estimates of correlation between the reaches within a stream were made based on weekly observed redds. Because correlations are often quite high between reaches, this is a better alternative than to naively assume the standard
errors between reaches are independent of one another. These estimates of correlation were combined with estimates of standard error for each index reach to calculate a covariance matrix for the Wenatchee index reaches (W2, W6, W8, W9, W10), which was used when summing estimates of spawners to estimate the total standard error. Failure to incorporate the correlations between reaches would result in an underestimate of standard error at the population scale. Non-index reaches were only surveyed once, so it is impossible to estimate a correlation coefficient between non-index reaches and index reaches. Therefore, they were assumed to be independent from the index reaches when summing the estimates of spawners. Because the estimates of tributary spawners were made separately (see above), they were also treated as independent when summing spawner estimates. The uncertainty in each step was carried through the entire analysis via the delta method (Casella and Berger 2002).

## Results

## Redd estimates

It should be noted that the GAUC parameters from index reaches were not used to estimate total redds in the associated non-index reaches. Figure 4 does illustrate that the non-index reach surveys were conducted close to the period of peak spawning (as determined by the associated index reaches), thus helping to validate the assumptions that go into estimating total redds in non-index reaches.

Table 1: Estimates of mean net error and total redds for each reach.

| Reach | Type | Index.Reach | Net.Error | Net.Error.CV | Redds.Counted | Redds.Est | Redds.CV |
| :--- | :--- | :--- | ---: | ---: | ---: | ---: | ---: |
| W1 | Non-Index | W2 | 0.55 | 0.24 | 0 | 0 | NA |
| W2 | Index | - | 0.59 | 1.40 | 2 | 3 | 1.50 |
| W3 | Non-Index | W2 | 0.44 | 0.30 | 1 | 2 | 0.30 |
| W4 | Non-Index | W6 | 0.46 | 0.23 | 0 | 0 | NA |
| W5 | Non-Index | W6 | 0.50 | 0.22 | 5 | 10 | 0.22 |
| W6 | Index | - | 0.99 | 0.85 | 54 | 53 | 0.88 |
| W6 | Non-Index | W6 | 0.46 | 0.15 | 0 | 0 | NA |
| W8 | Index | - | 0.92 | 0.90 | 9 | 10 | 0.95 |
| W9 | Index | - | 0.79 | 0.89 | 81 | 102 | 0.91 |
| W9 | Non-Index | W9 | 0.63 | 0.15 | 4 | 6 | 0.15 |
| W10 | Index | - | 0.83 | 0.61 | 99 | 120 | 0.65 |
| W10 | Non-Index | W10 | 0.59 | 0.13 | 3 | 5 | 0.13 |
| Total |  | NA | NA | NA | 258 | 311 | 0.63 |



Plots of observed redd counts (black dots) through time for each index reach, and the fitted curve from the GAUC model (blue line) with associated uncertainty (gray).


Observed redd counts for non-index reaches with non-zero peak redd counts. The blue curve shows the GAUC estimated spawning curve, demonstrating how close to peak spawning the non-index surveys were conducted.

## Spawner estimates

Table 2: Fish per redd and hatchery / natural origin proportion estimates.

| Area | Fish / redd | FpR Std. Error | Prop. Hatchery | Prop Std. Error |
| :--- | ---: | ---: | ---: | ---: |
| Above TUF | 1.777 | 0.059 | 0.599 | 0.026 |
| Below TUF | 1.728 | 0.089 | 0.343 | 0.040 |

Table 3: Estimates (CV) of spawners by area and origin.

| Area | Type | Hatchery | Natural |
| :--- | :--- | :--- | :--- |
| W1 | Non-Index | $0(--)$ | $0(--)$ |
| W2 | Index | $2(1.51)$ | $4(1.51)$ |
| W3 | Non-Index | $1(0.32)$ | $3(0.31)$ |
| W4 | Non-Index | $0(--)$ | $0(--)$ |
| W5 | Non-Index | $6(0.25)$ | $11(0.23)$ |
| W6 | Index | $32(0.89)$ | $60(0.88)$ |
| W6 | Non-Index | $0(--)$ | $0(--)$ |
| W8 | Index | $10(0.95)$ | $7(0.95)$ |
| W9 | Index | $108(0.92)$ | $73(0.92)$ |
| W9 | Non-Index | $7(0.16)$ | $5(0.16)$ |
| W10 | Index | $127(0.65)$ | $85(0.66)$ |
| W10 | Non-Index | $5(0.14)$ | $4(0.15)$ |
| Icicle | Trib | $52(0.32)$ | $83(0.25)$ |
| Peshastin | Trib | $40(0.37)$ | $206(0.16)$ |
| Mission | Trib | $23(0.49)$ | $71(0.28)$ |
| Chumstick | Trib | $0(--)$ | $38(0.39)$ |
| Chiwaukum | Trib | $12(0.72)$ | $48(0.34)$ |
| Chiwawa | Trib | $168(0.23)$ | $168(0.21)$ |
| Nason | Trib | $68(0.29)$ | $237(0.15)$ |
| Little Wenatchee | Trib | $0(--)$ | $0(--)$ |
| White River | Trib | $0(--)$ | $0(--)$ |
| Total |  | $661(0.45)$ | $1103(0.3)$ |

## Discussion

We have estimated the number of steelhead redds based on redd surveys, while incorporating potential observation error. After translating these to estimates of spawners by origin, we can then compare the spawner estimates to escapement estimates made using PIT tags, and estimate a pre-spawn mortality rate (Table 4). Taking the total PIT-tag based escapement estimate to the Wenatchee (after subtracting the number of hatchery fish removed at Tumwater), and subtracting the total estimate of spawners, including the tributaries, then dividing by the total escapement estimate provides an estimate of prespawn mortality across the entire Wenatchee population. We did this for natural and hatchery origin fish, and found that hatchery fish had a higher pre-spawn mortality rate, although the difference is not statistically significant.

Table 4: Wenatchee pre-spawn mortality rates.

| Origin | Pre-spawn_Mort | CV |
| :--- | ---: | ---: |
| Hatchery | 0.25 | 0.0016 |
| Natural | 0.16 | 0.0013 |

## Caveats

The predictions of surveyor net error were made using a model that had been fit to data in the Methow. Most covariates in the Wenatchee were within the range of values in the Methow study, but mean discharge was higher in all reaches in the Wenatchee than in the modeled reaches in the Methow (Figure 3). The mean discharge in the Methow study was 1069.2, while it was 2680 in the Wenatchee reaches in 2015 . That difference alone would change net error predictions by 0.29 , not an insignificant amount. However, the observed covariate values in the Wenatchee did not lead to unrealistic estimates of net error. The ranges of net error estimates for the Methow study and the Wenatchee in 2015 were very similar.


Source 官Methow Model 官 Wenatchee Prediction

Net error covariate values from the study in the Methow and the predicted reaches in the Wenatchee.

## References

Casella, G., and R. L. Berger. 2002. Statistical inference. Duxbury Pacific Grove, CA.
Gallagher, S., P. Hahn, and D. Johnson. 2007. Redd counts. Salmonid field protocols handbook: Techniques for assessing status and trends in salmon and trout populations. American Fisheries Society, Bethesda, Maryland:197-234.

Millar, R., S. McKechnie, and C. Jordan. 2012. Simple estimators of salmonid escapement and its variance using a new area-under-the-curve method. Canadian Journal of Fisheries and Aquatic Sciences 69:1002-1015.

Murdoch, A. R., C. J. Herring, K. E. See, and C. E. Jordan. 2014. Incorporating observer error in estimates of steelhead redd abundance in the Wenatchee River basin.

Truscott, B.l., A.R. Murdoch, J.M. Cram and K. See. 2015. Upper Columbia spring Chinook salmon and steelhead juvenile and adult abundance, productivity and spatial scale monitoring, Project \# 2010-034-00. Bonneville Power Administration, Portland OR. https://pisces.bpa.gov/release/documents/DocumentViewer.aspx?doc=P142786

## Appendix E

Genetic Diversity of Wenatchee Summer Steelhead

# Examining the Genetic Structure of Wenatchee Basin Steelhead and Evaluating the Effects of the Supplementation Program 

Developed for<br>Chelan County PUD and the Rock Island Habitat Conservation Plan Hatchery Committee<br>Developed by<br>Todd R. Seamons, Sewall Young, Cherril Bowman, and Kenneth I. Warheit WDFW Molecular Genetics Laboratory<br>Olympia, WA<br>and<br>Andrew R. Murdoch<br>Supplementation Research Team<br>Wenatchee, WA

17 January 2012

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## Executive Summary

In 1997, Wenatchee River summer steelhead, as part of the upper Columbia River evolutionarily significant unit (ESU), were listed as threatened under the Endangered Species Act (ESA). To address concerns about effects of hatchery supplementation, the hatchery program for hatchery produced (HOR) summer steelhead to be planted in the Wenatchee River changed from using mixed ancestry broodstock collected in the Columbia River to using Wenatchee River broodstock collected in the Wenatchee River. Three monitoring and evaluation (M\&E) indicators were developed to measure the genetic effects of hatchery production on wild fish populations. To address these indicators, temporal collections of tissue samples from Wenatchee River hatchery-produced (HOR) and natural origin (NOR) adults captured and sampled at Dryden and Tumwater dams and from NOR juveniles from three Wenatchee River tributaries and the Entiat River were surveyed for genetic variation with 132 genetic (SNPs) markers. Peshastin Creek (a Wenatchee River tributary) and the Entiat River served as no-hatchery-outplant controls, meaning they have stopped receiving HOR juvenile outplants. As per the M\&E plan, we interrogated these data for the presence or absence of spatial and temporal trends in allele frequencies, genetic distances, and effective population size.

Allele frequencies - Changes to the summer steelhead hatchery supplementation program had no detectable effect on genetic diversity of wild populations. On average, HOR adults had higher minor allele frequencies (MAF) than NOR adults, which may simply reflect the mixed ancestry of HOR adults. Both HOR and NOR adults had MAF similar to juveniles collected in spawning tributaries and in the Entiat River. There was no temporal trend in allele frequencies or observed heterozygosity in adult or juvenile collections and allele frequencies in control populations were no different than those still receiving hatchery outplants. This suggests that the hatchery program has had little effect on allele frequencies since broodstock sources changed in 1998.

Genetic distances - As intended, interbreeding of Wenatchee River HOR and NOR adults reduced the genetic differences between Wells Hatchery HOR adults and Wenatchee River NOR adults observed in the first few years after changing the broodstock collection protocol. Though there were detectable genetic differences between HOR and HOR adults, the magnitude of that
difference declined over time. HOR adults were genetically quite different from NOR adults and juveniles based on pair-wise $F_{\text {ST }}$ and principal components analysis (PCA), most likely because of the much smaller effective population size $\left(N_{\mathrm{b}}\right)$ in the hatchery population (see below). Pairwise $F_{\text {ST }}$ estimates and genetic distances between HOR and NOR adults collected the same year declined over time suggesting that the interbreeding of HOR and NOR adults in the hatchery (and presumably in the wild) is slowly homogenizing Wenatchee River summer steelhead. Analyses using brood year (the year fish were hatched, determined using scale-based age estimates) were inconclusive because of limitations of the data.

Effective population size $\left(N_{\mathrm{b}}\right)$ - Although the effective population size of the Wenatchee River hatchery summer steelhead program was consistently small, it does not appear to have caused a reduction in the effective population size of wild populations. On average, estimates of $N_{\mathrm{b}}$ were much lower and varied less for HOR adults than for NOR adults and juveniles. Estimates of $N_{\mathrm{b}}$ for HOR adults declined from the earliest brood years to a stable new low value after broodstock practices were changed in 1997. There was no indication that this had any effect on $N_{\mathrm{b}}$ in NOR adults and juveniles; $N_{\mathrm{b}}$ estimates for NOR adults and juveniles were, on average, higher and varied considerably over the time period covered by our dataset (1998-2010) and showed no temporal trend.

## Introduction

The National Marine Fisheries Service (NMFS) recognizes 15 Evolutionary Significant Units (ESU) for west coast steelhead (Oncorhynchus mykiss). The Upper Columbia ESU, which contains steelhead in the Wenatchee Basin, was listed as endangered under the Endangered Species Act (ESA) in 1997. Included in this listing were the Wells hatchery steelhead (program initiated in the late 1960s) that originated from a mixed group of native steelhead and are considered to be genetically similar to natural spawning populations above Wells Dam. Juvenile steelhead from Wells Fish Hatchery was the primary stock released into the Wenatchee River (Murdoch et al. 2003). The 1998 steelhead status review identified several areas of concern for this ESU including the risk of genetic homogenization due to hatchery practices and the high proportion ( $65 \%$ for the Wenatchee River) of hatchery fish present on the spawning grounds (Good et al. 2005). The Biological Review Team (BRT) further identified the relationship between the resident and anadromous forms of $O$. mykiss and possible changes in the population structure ('genetic heritage of the naturally spawning fish') in the basin as two areas requiring additional study. Furthermore, the West Coast Steelhead BRT (2003) recommended that stocks in the Wenatchee, Entiat, and Methow rivers, within the Upper Columbia ESU, be managed as separate populations.

A review of the presence of resident $O$. mykiss in the Upper Columbia ESU (Good et al. 2005) shows that rainbow trout are relatively abundant in upper Columbia River tributaries currently accessible to steelhead as well as in upriver tributaries unavailable to anadromous access by Chief Joseph and Grand Coulee dams (Kostow 2003). U.S. Fish and Wildlife Service (USFWS) biologists surveyed the abundance of trout and steelhead juveniles in the Wenatchee, Entiat, and Methow river drainages in the mid-1980s and found adult trout (defined as those with fork length $>20 \mathrm{~cm}$ ) in all basins (Mullan et al. 1992). The results also supported the hypothesis that resident $O$. mykiss are more abundant in tributary or mainstem areas upstream of the areas used by steelhead for rearing. No samples of rainbow trout from the Wenatchee were available for this study.

In addition to the mixed ancestry Wells Hatchery steelhead, Skamania Hatchery (Washougal River steelhead ancestry) steelhead were also released into the Wenatchee River basin for several years in the late 1980s (L. Brown, Washington Dept. of Fish and Wildlife [WDFW], personal communication). In 1996, broodstock for the Wenatchee River steelhead program were collected from Priest Rapids Dam and Dryden (rkm 24.9) and Tumwater (rkm 52.6) dams on the Wenatchee River. Because of the ESA listing, broodstock collection after 1996 was restricted to the Wenatchee River in an effort to develop a localized broodstock (Murdoch et al. 2003). Thus, starting in 1998, all juvenile steelhead released into the Wenatchee River and Wenatchee River tributaries were offspring of only Wenatchee River captured broodstock.

In response to the need for evaluation of the supplementation program, both a monitoring and evaluation plan (Murdoch and Peven 2005) and the associated analytical framework (Hays et al. 2006) were developed for the Habitat Conservation Plans Hatchery Committee through the joint effort of the fishery co-managers (Confederated Tribes of the Colville Reservation [CCT], NMFS, USFWS, WDFW, and Yakama Nation [YN]) and Chelan County, Douglas County, and Grant County Public Utility Districts (PUD). These reports outline 10 objectives to be applied to various species assessing the impacts of hatchery operations mitigating the operation of Rock Island and Rocky Reach Dams. This report pertains to Wenatchee River basin steelhead ( $O$. mykiss) and the steelhead supplementation program as addressed by objective 3 , specifically the first three evaluation indicators.

Objective 3: Determine if genetic diversity, population structure, and effective population size have changed in natural spawning populations as a result of the hatchery program. Additionally, determine if hatchery programs have caused changes in phenotypic characteristics of natural populations.

### 3.1 Allele Frequency

### 3.2 Genetic Distances Between Populations

### 3.3 Effective Spawning Population

To address these evaluation indicators the WDFW Molecular Genetics Lab (MGL) obtained pertinent tissue collections and samples, surveyed genetic variation with SNP markers using our standard laboratory protocols, and calculated the relevant genetic metrics and statistics. We used collections from both the Entiat River and Wenatchee River basins. Both have received hatchery plants from non-local stocks [i.e. Entiat was stocked with both Wenatchee and Wells program juveniles averaging 12 K and 18 K respectively during 1995-2001, and Wenatchee received on average 177 K juveniles from the Wells program during 1995-2001; (Good et al. 2005)], and both have all or some part of the basin designated as natural production "reference" drainage - no hatchery outplanting (i.e., the entire Entiat Basin, and Peshastin Creek in the Wenatchee River basin) (Good et al. 2005).

## Materials and methods

## Sample collections

To address objectives 3.1 through 3.3, we obtained samples from hatchery (HOR, adipose fin clipped) and natural origin (NOR, adipose fin intact) adult summer steelhead captured at Dryden or Tumwater diversion dams in the summer and fall of 1997 through 2009 (excepting 2004 and 2005; Table 1). All or some fraction of these fish was later used as hatchery broodstock the calendar year following the sampling year. In order to keep things simple we have reported years as the spawning year, i.e., the calendar year the fish were spawned, not the calendar year they were captured.

To address objective 3.2, it was necessary to have samples from natural origin fish from each of the spawning populations in the basin. It is difficult to obtain adult samples from known spawning populations due to the life history and behavior of steelhead, without tributary weirs or some other blocking method of collection. The NOR adult samples used as broodstock collected from Dryden and Tumwater Dams were a mixed collection representing all of the spawning populations located upstream. Therefore to determine population substructure within the basin we obtained collections of juvenile fish from smolt traps located within tributaries representing three major populations in the basin and from the Entiat River (Chiwawa River, Nason Creek, and Peshastin Creek; Table 2). We also obtained two collections of juvenile fish caught in a
smolt trap in the lower Wenatchee River. These, like the NOR adult collections, were a mixed collection presumably representing all populations located upstream. Fin tissue was taken from each fish and preserved in 95\% ethanol.

## Sample processing

Fin tissue samples were processed for 1468 HOR and NOR adult steelhead broodstock (Table 1) and for 1542 juvenile $O$. mykiss from the Wenatchee and Entiat Rivers (Table 2). Samples were genotyped at 152 single nucleotide polymorphism loci (SNPs, Tables 3, 4). We originally proposed to use microsatellites, but WDFW MGL and other regional genetic laboratories (Columbia River Inter-Tribal Fish Commission [CRITFC], Idaho Fish and Game [IDFG], USFWS) are moving toward using SNPs and they provide the same kinds of information with faster processing. Twenty SNP loci were developed to discriminate among trout species; 14 distinguish $O$. mykiss from coastal cutthroat trout ( $O$. clarkii clarkii) and westslope cutthroat ( $O$. clarkii lewisi), and 6 distinguish steelhead and coastal cutthroat from westslope cutthroat (Table 4). The remaining 132 SNP loci were developed to be used for population structure, parentage assignment, or other population genetic studies of $O$. mykiss (Table 3). These markers comprised the current standard set of SNP markers used for genetic studies of $O$. mykiss at WDFW MGL.

We used Qiagen DNEasy ${ }^{\circledR}$ kits (Qiagen Inc., Valencia, CA), following the recommended protocol for animal tissues, to extract and isolate DNA from fin tissue. SNP genotypes were obtained through PCR and visualization on Fluidigm EP1 integrated fluidic circuits (chips). Protocols followed Fluidigm's recommendations for TaqMan SNP assays as follows: Samples were pre-amplified by Specific Target Amplification (STA) following Fluidigm's recommended protocol with one modification. The 152 assays were pooled to a concentration of 0.2 X and mixed with 2X Qiagen Multiplexing Kit (Qiagen, Inc., Valencia CA), instead of TaqMan PreAmp Master Mix (Applied Biosystems), to a volume of $3.75 \mu$ l, to which $1.25 \mu \mathrm{l}$ of unquantified sample DNA was added for a total reaction volume of $5 \mu$ l. Pre-amp PCR was conducted on a MJ Research or Applied Biosystems thermal cycler using the following profile: $95^{\circ} \mathrm{C}$ for 15 min followed by 14 cycles of $95^{\circ} \mathrm{C}$ for 15 sec and $60^{\circ} \mathrm{C}$ for 4 minutes. Post-PCR reactions were diluted with $20 \mu \mathrm{l} \mathrm{d}_{2} \mathrm{O}$ to a final volume of $25 \mu \mathrm{l}$.

Specific SNP locus PCRs were conducted on the Fluidigm chips. Assay loading mixture contained 1X Assay Loading Reagent (Fluidigm), 2.5X ROX Reference Dye (Invetrogen) and 10X custom TaqMan Assay (Applied Biosystems); sample loading mixture contains 1X TaqMan Universal PCR Master Mix (Applied Biosystems), 0.05X AmpliTaq Gold DNA polymerase (Applied Biosystems), 1X GT sampling loading reagent (Fluidigm) and $2.1 \mu \mathrm{~L}$ template DNA. Four $\mu \mathrm{L}$ assay loading mix and $5 \mu \mathrm{~L}$ sample loading mix were pipetted onto the chip and loaded by the IFC loader (Fluidigm). PCR was conducted on a Fluidigm thermal cycler using a two step profile. Initial mix thermal profile was $70^{\circ} \mathrm{C}$ for $30 \mathrm{~min}, 25^{\circ} \mathrm{C}$ for $5 \mathrm{~min}, 52.3^{\circ}$ for $10 \mathrm{sec}, 50.1^{\circ} \mathrm{C}$ for $1 \mathrm{~min} 50 \mathrm{sec}, 98^{\circ} \mathrm{C}$ for $5 \mathrm{sec}, 96^{\circ} \mathrm{C}$ for $9 \mathrm{~min} 55 \mathrm{sec}, 96^{\circ} \mathrm{C}$ for $15 \mathrm{sec}, 58.6^{\circ} \mathrm{C}$ for 8 sec , and $60.1^{\circ} \mathrm{C}$ for 43 sec . Amplification thermal profile was 40 cycles of $58.6^{\circ} \mathrm{C}$ for $10 \mathrm{sec}, 96^{\circ} \mathrm{C}$ for 5 $\sec , 58.6^{\circ} \mathrm{C}$ for 8 sec and $60.1^{\circ} \mathrm{C}$ for 43 sec with a final hold at $20^{\circ} \mathrm{C}$.

The SNP assays were visualized on the Fluidigm EP1 machine using the BioMark data collection software and analyzed using Fluidigm SNP genotyping analysis software. To ensure all SNP markers were being scored accurately and consistently, all data were scored by two researchers and scores of each researcher were compared. Disputed scores were called missing data (i.e., no genotype).

## Evaluation of loci

A two-tailed exact test of Hardy-Weinberg equilibrium (HWE) was performed for each locus in each collection or population using the Markov Chain method implemented in GENEPOP v4.1 (dememorization number 1000, 100 batches, 1000 iterations per batch; Raymond and Rousset 1995; Rousset 2008). Significance of probability values was adjusted for multiple tests using false discovery rate (Verhoeven et al. 2005). $F_{\text {IS }}$, a measure of the fractional reduction in heterozygosity due to inbreeding in individuals within a subpopulation and an additional indicator of scoring issues, was calculated according to Weir and Cockerham (1984) using GENEPOP v4.1. Allele frequencies were calculated using CONVERT v1.0 (Glaubitz 2004). Expected and observed heterozygosities were calculated using GDA v1.1 (Lewis and Zaykin 2001).

## Allele frequencies, genetic distances and population differentiation

To evaluate Q1 of Objective 3.1 and 3.2, we evaluated trends and patterns in allele frequencies, genetic distances and population differentiation. To test for temporal patterns in allele frequencies, we compared sample or spawn year to two diversity metrics, allele frequency and observed heterozygosity, from each adult and juvenile collection. Each SNP locus had only one or two alleles, so we used the minor allele frequency (MAF) of each SNP locus for each adult collection and averaged across loci. We also calculated the average observed heterozygosity (Ho) for each SNP locus within each adult and juvenile collection. We examined the presence or absence of a temporal trend in average allele frequency and observed heterozygosity with logistic regression analysis in R ( R Development Core Team 2009).

To partition genetic variance into temporal, spatial (juvenile) and origin (adult) fractions, we performed hierarchical analysis of molecular variance (AMOVA) using ARLEQUIN v3.0 (Excoffier et al. 2005) with 1,000 permutations. We performed this analysis separately for juvenile and adult collections. Juveniles were grouped by sampling location (tributary) and adults were grouped by origin (HOR or NOR). To estimate the magnitude of genetic differences among temporal and spatial collections we calculated pairwise $F_{\text {ST }}$ estimates among collections using FSTAT (Goudet 1995) with 1000 permutations. Statistical significance was adjusted using false discovery rate (Verhoeven et al. 2005).

To evaluate the temporal changes in genetic relationships, we compared spawn year to within spawn year pairwise $F_{\text {ST }}$ estimates between NOR and NOR adults using beta regression (Simas and Rocha 2010). We used beta regression because the dependent variable was bound by zero and one but not binomial. Analysis was performed in R (package "betareg", Cribari-Neto and Zeileis 2010), with a loglog link.

We used principal component analyses (PCA) to explore the relationship between the covariation among the SNP loci within each collection and genetic differentiation between HOR and NOR collections, and to determine if the degree of differentiation has changed with time. Since each SNP is represented by only two alleles, only one allele per SNP is necessary to fully describe the covariation among all SNPs. We used matLab® scripts (2007a, The Mathworks, Natlick, MA)
to calculate the principal components from SNP allele frequencies using only the major allele (1MAF) for each SNP. We defined the major allele as the allele with the higher mean frequency across all collections, regardless of its status within any individual collection. We conducted three PCA analyses using: (1) all adult samples, aggregated based on origin (HOR versus NOR) and spawn year (i.e., the year the adult fish were used as broodstock) ( $\mathrm{N}=1437$, 22 collections), (2) same as \#1, but with the addition of all juvenile samples ( $\mathrm{N}=2938$, 37 collections), and (3) only those adults samples with available age information (Mike Hughes, WDFW, personal communication) aggregated based on origin, and spawn year or brood year (i.e., the year the fish were hatched) ( $\mathrm{N}=1313,20$ spawn-year or 25 brood-year collections).

Molecular differentiation between HOR and NOR adults within a year was calculated based on principal component scores using Euclidian distances. We calculated pair-wise Euclidian distances between HOR and NOR fish within a spawn year or brood year using the first three principal components, and standardized each distance by subtracting from it the mean Euclidian distance calculated across all pair-wise distances. We used Mahalanobis distances to calculate the variation among HOR and NOR collections (calculated separately), again using the first three principal components. Here, we calculated Mahalanobis distances as the Euclidian distances between each collection and the centroid of all collections (HOR and NOR combined), but the Euclidian distances are scaled based on the dispersion of collections around the centroid (i.e., the variance). Euclidian and Mahalanobis distances were calculated using MATLAB scripts.

## Effective spawning population

To evaluate Q 1 of Objective 3.3, we estimated $N_{e}$ using the single-sample linkage disequilibrium methods implemented in the program LDNE (Waples and Do 2008). This method requires that you input the $P_{\text {crit }}$ value, the minimum frequency at which alleles were included in the analysis, since results can be biased depending on this setting (Waples and Do 2010). SNP markers typically have only one or two alleles; if one of two alleles is excluded based on its frequency in the collection it essentially excludes the locus, reducing the overall dataset. Therefore, we used $P_{\text {crit }}$ values ranging from 0.1 to 0.001 to evaluate whether trends in $N_{\mathrm{e}}$ changed given which loci were used. Confidence intervals were calculated using a jackknife procedure.

We calculated an estimate of $N_{\mathrm{e}}$ for all adult and juvenile collections individually. However, the intention of an integrated hatchery program such as the Wenatchee River steelhead hatchery program is that HOR and NOR fish are integrated and progress as a single population through intentional interbreeding in the hatchery and presumed natural interbreeding in the wild. Thus, we also combined annual HOR and NOR collections to calculate an overall $N_{\mathrm{e}}$ estimate as has been done in other genetic monitoring and evaluation analyses (e.g., Small et al. 2007, [Chinook salmon, O. tshawytscha]).

Estimates of $N_{e}$ from linkage refer to the generations that produced the sample. To calculate the ratio of effective population size to census size $\left(N_{\mathrm{e}} / N\right)$, we obtained the number of fish spawned in the hatchery (1993 through 2006, i.e., those that produced the adipose fin clipped adults that returned to spawn in the Wenatchee River 1998 through 2010) and the estimated escapement of fish spawning naturally (HOR and NOR separately) for the same time period. Estimates of census population size in spawning tributaries was obtained by multiplying the fraction of redds counted within tributaries (Chad Herring, WDFW, unpublished data) by the total Wenatchee River census population estimate (Andrew Murdoch, WDFW, unpublished data). To calculate $N_{e} / N$, we performed two analyses. First, for adults, we assumed a five year generation time for natural origin adults and a four year generation time for hatchery origin adults and divided the $N_{\mathrm{e}}$ estimate by the census population estimate from four or five years earlier. For juveniles, we assumed an age at outmigration of two years and divided the $N_{\mathrm{e}}$ estimates by the estimate of census population size for the appropriate tributary. Second, we used available adult age data to parse individuals into cohorts originating in brood years (rather than spawn years) and then used LDNE to estimate $N_{\mathrm{e}}$ from cohort collections. We performed both analyses to make full use of all available data; age data were not available for many adults, and because of variable survival and sampling not all cohorts had sufficient numbers of HOR and NOR adults. According to Luikart et al. (2010), estimates produced using linkage disequilibrium should be interpreted as something between effective population size $\left(N_{e}\right)$ and the effective number of breeders $\left(N_{b}\right)$. Using cohorts, the estimate produced by LDNE is clearly an estimate of $N_{\mathrm{b}}$ rather than $N_{\mathrm{e}}$. In order to keep things simple, we have referred to all estimates as $N_{\mathrm{b}}$.

## Results and Discussion

## Collections and samples received

From 1468 samples from HOR and NOR adult steelhead broodstock, 1437 produced sufficient genetic data for further analysis (Table 1). From 1542 samples from NOR juvenile steelhead from Wenatchee River tributaries and the Entiat River, 1501 produced sufficient genetic data for further analysis and were genetically identified as $O$. mykiss (Table 2). Samples genetically identified as $O$. clarki ( 2 samples from the Chiwawa River, 1 from the Entiat River) or $O$. clarki/O. mykiss hybrids (4 - lower Wenatchee River, 4 - Nason Creek, 4 - Chiwawa River, and 1 - Entiat River) were omitted from further analysis.

## Evaluation of loci

Three loci showed deviations from HWE in 10 or more of 37 Wenatchee steelhead collections before correcting for multiple tests (AOmy016, AOmy051, AOmy252, Table A1) indicating possible scoring issues. These loci were omitted from further analysis. Nine of the remaining loci were monomorphic or nearly monomorphic in all collections (average MAF < 0.1, AOmy023, AOmy028, AOmy123, AOmy129, AOmy132, AOmy209, AOmy229, AOmy270, AOmy271, Table A1) contributing little or nothing to analytical power. These loci were also omitted from further analysis. No genetic data was available for collection 10FD due to poor PCR amplification at locus AOmy213 for the entire collection. AOmy213 had a relatively low MAF in most collections so rather than re-processing this collection at this locus or running different sets of loci for different tests, we omitted this locus from further analysis. Only six tests of deviation from HWE were significant after correcting for 4348 tests using false discovery rate. Two of these tests were in loci already omitted. The remaining four tests were spread among the remaining loci, indicating no more loci needed to be omitted from further analysis.

## Objective 3.1, 3.2 - Allele frequencies and Genetic distances

## Allele frequencies

Average MAF of SNP loci ranged from 0.00 to 0.60 in HOR adult collections and from 0.00 to 0.61 in NOR adult collections (Table A1). Observed heterozygosity ranged from 0.00 to 0.75 in HOR adult collections and from 0.01 to 0.67 in NOR adult collections. Juvenile collections produced similar ranges of MAF and Ho (Table A1). Average MAF and Ho of HOR adult collections appeared to be greater than those of natural origin collections. However, logistic regression analysis indicated there was no significant temporal trend in either diversity statistic (Figure 1). Similarly, there was no consistent temporal trend in MAF or Ho of juvenile collections (Figure 2). Both the Chiwawa River and Nason Creek, the two tributaries that currently still receive hatchery juvenile outplants, both appeared to have declining allele frequencies, but neither was statistically significant ( $P>0.90$ ). However, the power to detect significant trends was limited by the small sample sizes ( $\mathrm{n}=3$ sample years).

## Analysis of Molecular Variance

Analysis of molecular variance (AMOVA) of adult collections (i.e., temporal and origin structure) indicated most of the genetic variance was among individuals or among individuals within populations (99.04\%). Most of the remaining variance was temporal variation within hatchery and natural origin groups ( $0.61 \%$ ) with the remaining variation from origin ( $0.35 \%$ ). AMOVA of juvenile collections (i.e., spatial structure) indicated most of the genetic variance was among individuals $(98.44 \%$ ) or among individuals within populations $(0.94 \%)$. Most of the remaining variance existed among temporal collections within tributary collections ( $0.37 \%$ ) with the smallest fraction as among tributary variance ( $0.24 \%$ ). Thus, overall, there was more variability among years than among tributaries or origins, but no trend in the temporal variability.

## Pair-wise $\mathrm{F}_{\text {ST }}$ estimates

HOR adults were genetically different that NOR adults as estimated by $F_{\text {ST }}$ (full pair-wise table in Table A2, all pair-wise $F_{\text {ST }}$ estimates with $P$-values $\leq 0.05$ before correcting for multiple tests
were significantly different from zero after correcting for multiple tests using false discovery rate). On average, HOR adult collections were as different from one another (mean $F_{\mathrm{ST}}=0.011$ ) as they were from NOR adult collections among years (mean $F_{\mathrm{ST}}=0.009$ ) or from NOR adult collections within years (mean $F_{\mathrm{ST}}=0.010$ ). Among year comparisons of NOR adult collections were, on average, nearly an order of magnitude lower (mean $=0.002$ ). These patterns held whether spawn year or brood year (data not shown) was used to group individuals. Over time, within spawn year pair-wise $F_{\text {ST }}$ estimates between HOR and NOR adults declined over time ( $\beta$ $=-0.014, P=0.0185$; Figure 3), suggesting that the integration of hatchery and wild fish is slowly genetically homogenizing the groups. That relationship disappeared when adults were grouped by brood year (i.e., comparing fish produced the same year) and all brood years were used ( $\beta=-0.009, P=0.615$, data not shown). However, when the dataset was restricted to just those brood years when all typical (age at maturation frequency among all years $>0.10$ ) age classes were present in the dataset $(\mathrm{HOR}=$ age 3,$4 ; \mathrm{NOR}=$ age $4,5,6$; brood years 1996-1998, 2004-2005) a non-significant ( $P=0.278$ ) negative relationship ( $\beta=-0.12$ ) of $F_{\mathrm{ST}}$ and brood year was apparent. When the data were further restricted to just the years after the hatchery program changed to only collecting broodstock in the Wenatchee River (brood years 1998, 2004-2005), the slope was also negative $(\beta=-0.09)$, but the relationship was not statistically significant ( $P=$ 0.962 ).

Within tributary among sample year pair-wise comparisons of juvenile collections were, on average, only very slightly smaller than comparisons among tributaries ( 0.005 vs. 0.006 , respectively, Table 5, all pair-wise $F_{\text {ST }}$ estimates with $P$-values $\leq 0.05$ before correcting for multiple tests were significantly different from zero after correcting for multiple tests using false discovery rate). Nason Creek and Peshastin Creek on average showed higher among sample year $F_{\text {ST }}$ estimates ( 0.010 and 0.007 , respectively) than the Chiwawa or Entiat Rivers ( 0.004 and 0.002 , respectively). The pair-wise comparison of the two collections of lower Wenatchee River smolts, presumably a mix of Chiwawa, Nason, Peshastin smolts and smolts from other spawning tributaries, was an order of magnitude smaller $\left(F_{\mathrm{ST}}=0.0002\right)$, and not significantly different than zero (Table 5). There was no temporal trend in pair-wise comparisons of juvenile collections. However with, at most, four annual collections, detecting any temporal trend was unlikely. We also had no collections from years prior to 1998 (the first year of new hatchery program
broodstock collecting protocols) with which to compare contemporary data, nor could we find any reports or papers containing pre-hatchery-program-change genetic comparisons among Wenatchee River tributary populations, making it impossible to determine whether or not changing the hatchery program has had any effect at all on population structure. However, these data will be useful for future studies.

## Principal Components

Each principal component analysis (Figures 4,5) indicated that the genetic structure among HOR collections differed from that among NOR collections, and that this difference has decreased with time. When adult fish were aggregated based on origin and spawn-year, there was a clear differentiation between HOR and NOR adult collections along PC 1, and a separation among HOR collections, differentiating the early spawn-years (1998 - 2003) from the later spawn-years (2004 - 2010) along PC 2 and PC 3, respectively (Figure 4). The pair-wise genetic distances between HOR and NOR collections from the same spawn year (i.e., the HOR and NOR fish used as broodstock within the same year) decreased from the largest distance in 1998 to small distances in 2009 and 2010, although the smallest distance occurred in 2004 (Figure 4, top right). That is, within hatchery broodstock, the genetic difference between HOR and NOR fish decreased, on average, from 1998 to 2010, and the decrease appeared to be a mutual convergence of NOR fish shifting right along PC 1 and HOR fish shifting downward along PC 2 and PC 3. This increasing similarity in adult fish mirrored that seen in within year pair-wise $F_{\text {ST }}$ estimates between HOR and NOR adults which also declined over time (Figure 3).

Overall, there was considerably more genetic variation among the HOR collections than there was among the NOR collections with average Mahalanobis distances (distance between each collection and the overall centroid $[0,0,0]$ ) among the HOR and NOR collections being 4.2 and 1.5 , respectively. Since each NOR collection was generally composed of 3-4 brood-years, while HOR collections rarely were composed of more than two brood-years, we attributed the lower year-to-year genetic variability of the NOR broodstock to the greater homogenizing effect of including four or more brood-years compared with only two brood years for the HOR broodstock.

Including the 15 juvenile collections, along with the 22 adult collections, did not materially alter the principal component structure (Figure 6), although the total genetic variation accounted for by the three principal components decreased from $44 \%$ using only the adults to $33 \%$ when juveniles were included. For the most-part, the juvenile fish appeared intermediate between HOR and NOR fish, but there was greater overlap in principal component scores (and therefore greater genetic similarity) of the juvenile and NOR collections, than of the juvenile and HOR collections. The average Euclidian distance between the juvenile and HOR collections was 0.49 , compared to 0.23 between the juvenile and NOR collections, which was no different than 0.23 and 0.22 for the within juvenile and NOR collections, respectively.

By using the available adult age data, we were able to compare the genetic differentiation among the same set of fish when they are aggregated by origin (hatchery versus natural) and brood-year (year fish were hatched) with aggregates based on origin and spawn-year (year adult fish were spawned). A brood-year analysis compares within a year the genetic diversity generated from hatchery broodstock with that naturally produced in the spawning grounds. A spawn-year analysis compares the HOR and NOR genetic diversity that was mixed among cohorts of the parental generations. The same basic pattern of genetic structure that we have seen in spawnyear analyses (Figure 4, Figure 6, and the right side of Figure 5) also occurred in the brood-year analysis (left side of Figure 5). That is, from Figure 5 we saw (1) that HOR and NOR fish were differentiated from each other; (2) there was considerably more genetic variation (temporal variation) among the hatchery-origin collections than there was among the natural-origin collections (for brood-year, Mahalanobis distances $=5.18$ and 0.75 , respectively; for spawn-year, Mahalanobis distances $=4.25$ and 1.25 , respectively), and (3) that the genetic distances between HOR and NOR collections were lower in the more recent brood- and spawn-years, than in the earlier brood- and spawn-years (Figure $7 ; R^{2}=0.41$ or $41 \%, P<0.05$ ). This indicated that the HOR and NOR fish used as broodstock in 2010 were more similar to each other than they were at the inception of the new hatchery program.

The relationship between genetic distance and brood-year was not the same as the relationship between genetic distance and spawn-year. For brood-year, although the slope was negative (i.e.,
trending downward or decreased differentiation with time) and the two most-recent brood years (2005-2006) showed relatively small HOR and NOR adult differentiation, the negative slope was not significantly different from zero and the regression accounted for only $7 \%$ of the variation. This was likely the result of insufficient sampling of certain age classes from many brood years (especially from NOR adults) due to two un-processed sample years (2005 and 2006).

## Objective 3.3 - Effective spawning population

There was no difference in the temporal trends in estimates of $N_{b}$ with $P_{\text {crit }}$ set from 0.1 to 0.001 (Figure 8, data not shown for all collections), so we have reported only results with $P_{\text {crit }}=0.001$, i.e., the full genetic dataset. Using either spawn-year or brood year, estimates of NOR adult $N_{\mathrm{b}}$ were higher and varied more than those of HOR adults (Figures 9, 10), concordant with the PCA analysis. Estimates for HOR adults ranged from 17 to 174 (by spawn year, mean $=65$ ) or from 6 to 130 (by brood year, mean =39). Estimates for NOR adults ranged from 36 to 982 (by spawn year, mean $=405)$ or from 59 to 2966 (by brood year, mean $=645$ ). Many $N_{\mathrm{b}}$ estimates for NOR adults had confidence intervals extending to infinity on the upper bound. This reflected the difficulty in obtaining precise estimates of $N_{\mathrm{b}}$ for large populations (Waples and Do 2010).

Estimates of $N_{\mathrm{b}}$ for HOR steelhead dropped by approximately half from 1994, when broodstock were still collected at Wells Hatchery, to 1998, when the program used Wenatchee River trapped adults only, suggesting an effect of changing broodstock collection practices, which began in 1997 (Figures 8, 9). Since 1997, the hatchery population $N_{\mathrm{b}}$ remained at a relatively stable lower level (Figures 8, 9, and 10). There was no obvious change in $N_{\mathrm{b}}$ for NOR steelhead since 1993; the $N_{\mathrm{b}}$ estimate for 1993 was the largest, however the confidence interval overlapped estimates from many other years. The temporal trend in $N_{\mathrm{b}}$ estimates from combined collections mirrored those of the HOR collections alone, though estimates using combined collections were slightly larger (Figure 11).

As with $N_{\mathrm{b}}$ estimates, estimates of the ratio of $N_{\mathrm{b}} / N$ for NOR adults varied more than those of HOR adults (Figures 12, 13). However, using spawn year, i.e., mixtures of cohorts, the average $N_{\mathrm{b}} / N$ ratio for HOR adults was equal to that of NOR adults (mean $N_{\mathrm{b}} / N=0.26$ ), whereas when using brood year, the average $N_{\mathrm{b}} / N$ ratio for NOR adults was double that of HOR adults (NOR
average $=0.40$, HOR average $=0.20$ ). This is likely a consequence of the homogenizing effect of mixed cohorts. Estimates of $N_{\mathrm{b}}$ for HOR adults using spawn year were close to those estimated using brood year because of the lower diversity in age at maturation, whereas for NOR, grouping by brood year produces different estimates than when grouping by spawn year because of higher diversity in age at maturation. Regardless of which estimate was used, there was no temporal trend in $N_{\mathrm{b}} / N$ for either NOR or HOR adults.

## Summary

On average, HOR adults had higher minor allele frequencies (MAF) than NOR adults, and both had similar MAF as juveniles collected in spawning tributaries and in the Entiat River. There was no temporal trend in allele frequencies or observed heterozygosity in adult or juvenile collections and allele frequencies in control populations were no different than those still receiving hatchery outplants suggesting that the hatchery program has had little effect on allele frequencies since 1998.

HOR adults were genetically quite different from NOR adults and juveniles based on pair-wise $F_{\text {ST }}$ and principal components analysis (PCA), most likely because of the much smaller effective population size $\left(N_{\mathrm{b}}\right)$ in the hatchery population. Pair-wise $F_{\mathrm{ST}}$ estimates and genetic distances between HOR and NOR adults collected the same year declined over time suggesting that the interbreeding of HOR and NOR adults in the hatchery (and presumably in the wild) is slowly homogenizing Wenatchee River summer steelhead. Analyses using brood year (the year fish were hatched, determined using scale-based age estimates) were inconclusive because of limitations of the data.

On average, estimates of $N_{\mathrm{b}}$ were much lower and varied less for HOR adults than for NOR adults and juveniles. Estimates of $N_{\mathrm{b}}$ for HOR adults declined from the earliest brood years to a stable new low value after broodstock practices were changed in 1997. There was no indication that this had any effect on $N_{\mathrm{b}}$ in NOR adults and juveniles; $N_{\mathrm{b}}$ estimates for NOR adults and juveniles were, on average, higher and varied considerably over the time period covered by our dataset (1998-2010) and showed no temporal trend. Small $N_{\mathrm{b}}$ sizes increase the risk of loss of
genetic diversity due to inbreeding and random effects (genetic drift). The $N_{\mathrm{b}}$ of the hatchery component of the population may be increased by spawning more families, using specific mating designs, and minimizing variance in reproductive success. However, given the apparent lack of effects overall, changes to the hatchery protocol may not be necessary.

Overall, hatchery practices appear to have had little effect on natural origin Wenatchee summer steelhead neutral genetic diversity or $N_{\mathrm{b}}$. We cannot accurately assess their effects on population structure at this time. However, it is interesting to note that when juvenile collections are analyzed separately from adult collections, Peshastin Creek, which has received fewer hatchery outplants in the past and is currently a refuge from hatchery outplants, is genetically different than other tributaries and the Entiat River (data not shown). On the other hand, the Entiat River, which is also a refuge from hatchery outplants and is not a tributary of the Wenatchee River, is genetically very similar to Nason Creek and the Chiwawa River, both Wenatchee River tributaries. This suggests, though it does not conclude, that within basin population structure may have existed before summer steelhead hatchery production began in the upper Columbia River and that the population structure was eliminated by hatchery influence long before 1998.

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## Literature Cited

Aguilar, A., and J. C. Garza. 2008. Isolation of 15 single nucleotide polymorphisms from coastal steelhead, Oncorhynchus mykiss (Salmonidae). Molecular Ecology Resources 8(3):659662.

Campbell, N. R., and S. R. Narum. 2009. Identification and characterization of heat shock response-related single-nucleotide polymorphisms in $O$. mykiss and $O$. tshawytscha. Molecular Ecology Resources 9(6):1460-1466.

Campbell, N. R., K. E. N. Overturf, and S. R. Narum. 2009. Characterization of 22 novel single nucleotide polymorphism markers in steelhead and rainbow trout. Molecular Ecology Resources 9(1):318-322.

Cribari-Neto, F., and A. Zeileis. 2010. Beta regression in R. Journal of Statistical Software 34:124.

Excoffier, L., G. Laval, and S. Schneider. 2005. Arlequin (version 3.0): An integrated software package for population genetics data analysis. Evolutionary Bioinformatics 1:47-50.

Finger, A. J., M. R. Stephens, N. W. Clipperton, and B. May. 2009. Six diagnostic single nucleotide polymorphism markers for detecting introgression between cutthroat and rainbow trouts. Molecular Ecology Resources 9(3):759-763.

Glaubitz, J. C. 2004. CONVERT: A user-friendly program to reformat diploid genotypic data for commonly used population genetic software packages. Molecular Ecology Notes 4:309310.

Good, T. P., R. S. Waples, and P. Adams. 2005. Updated status of federally listed ESUs of West Coast salmon and steelhead. U.S. Dept. of Commerce, NMFS-NWFSC-66.

Goudet, J. 1995. FSTAT (Version 1.2): a computer program to calculate F-statistics. Journal of Heredity 86:485-486.

Hansen, M. H. H., and coauthors. 2011. Assembling a dual purpose TaqMan-based panel of single-nucleotide polymorphism markers in rainbow trout and steelhead (Oncorhynchus mykiss) for association mapping and population genetics analysis. Molecular Ecology Resources 11:67-70.

Hays, S., and coauthors. 2006. Analytical framework for monitoring and evaluating PUD hatchery programs.

Kostow, K. 2003. The biological implications of nonanadromous Oncorhynchus mykiss in Columbia Basin steelhead ESUs. Report to NOAA Fisheries and ODFW, 13 January 2003. (Available from Northwest Fisheries Science Center, 2725 Montlake Blvd. E., Seattle, WA 98112.).

Lewis, P. O., and D. Zaykin. 2001. Genetic Data Analysis: Computer program for the analysis of allelic data. Free program distributed by the authors over the internet from http://lewis.eeb.uconn.edu/lewishome/software.html.

Luikart, G., N. Ryman, D. Tallmon, M. Schwartz, and F. Allendorf. 2010. Estimation of census and effective population sizes: the increasing usefulness of DNA-based approaches. Conservation Genetics 11(2):355-373.

McGlauflin, M. T., and coauthors. 2010. High-resolution melting analysis for the discovery of novel single-nucleotide polymorphisms in rainbow and cutthroat trout for species identification. Transactions of the American Fisheries Society 139(3):676-684.

Mullan, J. W., K. R. Williams, G. Rhodus, T. W. Hillman, and J. D. McIntyre. 1992. Production and habitat of salmonids in mid-Columbia River tributary streams. U.S. Fish and Wildlife Service, Monograph I, Leavenworth, WA.

Murdoch, A., T. Miller, T. Maitland, M. Tonseth, and L. Prave. 2003. Annual progress report for Wenatchee summer steelhead, 2001 brood. Washington Dept. of Fish and Wildlife.

Murdoch, A., and C. Peven. 2005. Conceptual approach for monitoring and evaluating the Chelan County Public Utility District hatchery programs. Chelan County Public Utility District, Wenatchee, WA.

R Development Core Team. 2009. R: A language and environment for statistical computing
Raymond, M., and F. Rousset. 1995. An exact test for population differentiation. Evolution 49(6):1280-1283.

Rousset, F. 2008. GENEPOP'007: a complete re-implementation of the GENEPOP software for Windows and Linux. Molecular Ecology Resources 8(1):103-106.

Sánchez, C. C., and coauthors. 2009. Single nucleotide polymorphism discovery in rainbow trout by deep sequencing of a reduced representation library. BMC Genomics 10(1):559.

Simas, A. B., and A. V. Rocha. 2010. Improved estimators for a general class of beta regression models. Computational Statistics \& Data Analysis 54:348-366.

Small, M. P., K. I. Warheit, C. Dean, and A. Murdoch. 2007. Genetic monitoring of Methow Spring Chinook. Washington Department of Fish and Wildlife, Olympia, WA.

Sprowles, A. E., M. R. Stephens, N. W. Clipperton, and B. P. May. 2006. Fishing for SNPs: A targeted locus approach for single nucleotide polymorphism discovery in rainbow trout. Transactions of the American Fisheries Society 135(6):1698-1721.

Verhoeven, K. J. F., K. L. Simonsen, and L. M. McIntyre. 2005. Implementing false discovery rate control: increasing your power. Oikos 108(3):643-647.

Waples, R. S., and C. Do. 2008. LDNE: a program for estimating effective population size from data on linkage disequilibrium. Molecular Ecology Resources 8(4):753-756.

Waples, R. S., and C. Do. 2010. Linkage disequilibrium estimates of contemporary $N_{e}$ using highly variable genetic markers: a largely untapped resource for applied conservation and evolution. Evolutionary Applications 3(3):244-262.

Weir, B. S., and C. C. Cockerham. 1984. Estimating F-statistics for the analysis of population structure. Evolution 38(6):1358-1370.

## Figures

Figure 1. Observed average minor allele frequencies (MAF) and observed heterozygosities (Ho) of 119 SNP loci from 11 annual collections of hatchery-produced (HOR) and natural origin (NOR) adult steelhead from the Wenatchee River. Trend lines are from a logistic regression. Note the X axis does not cross the Y axis at the origin. Neither the slopes nor the intercepts were statistically significant.



Figure 2. Observed average minor allele frequencies (MAF) and observed heterozygosities (Ho) of 119 SNP loci from 15 collections of natural origin juvenile steelhead from Wenatchee River tributaries, the lower Wenatchee River and the Entiat River. There were no consistent temporal trends in MAF or Ho in these collections.



Figure 3. The relationship of time with pairwise $F_{\text {ST }}$ estimates between hatchery-produced (adipose fin clipped) and natural origin (unclipped) adults of the same sample year. The line is the prediction based on beta regression.


Figure 4. Principal component (PC) 1 versus 2 (top left), PC 1 versus 3 (bottom left), and PC 2 versus 3 (bottom right) based on an analysis using all adults aggregated into origin and spawn-year collections. Natural-origin spawn-years are shown in italicized typeface. The percentage within the label of each axis convey the percent of total genetic variance that is accounted for by that axis. Taken together, the three principal components account for $44 \%$ of the total SNP variation. Top right shows pairwise Euclidian distances versus spawn-year, with zero distance equal to average distance across all pairwise distances. Blue line is least-squares fit with $\mathrm{R}^{2}=0.45$.


Figure 5. Principal components (PC) 1 versus 2 (top) and 3 (bottom) for adults aggregated into brood-year (BY; left) and spawn-year (SY; right). Spawn-year analysis is the same as in Figure x1, except fewer individuals per collection were included (see methods). Note that for the SY analysis here PC 2 and 3 are similar to PC 3 and 2, respectively, in Figure x1. Only BY1995 (earliest year with paired hatchery-natural data), BY2000 (extreme PC 1 score), and BY2006 (latest year with paired hatchery-natural data) are labeled. Hatchery- and natural-origin individuals from BY1995, BY2000, and BY2006, returned to spawn (spawn-year) in 1999 (hatchery)/1999-2001 (natural), 2003-2004 (hatchery)/2004 and 2007 (natural), and 2009-2010 (hatchery)/2010 (natural), respectively. These years are labeled in the upper right figure. Only 4 year-old BY 2006 natural-origin fish are represented in the SY 2010 collection.


Figure 6. Principal component (PC) 1 versus 2 (top) and PC 1 versus 3 (bottom) based on an analysis using all adult and juvenile fish aggregated into age (juvenile versus adult), origin (hatchery versus adult) and spawn-year collections.


Figure 7. Pairwise Euclidian distances versus brood-year (top) and spawn-year (bottom), with zero distance equal to average distance across all pairwise distances. Blue lines are least-squares fits, which is not significant $($ slope $=0$ ) for brood-year, but significant (slope $>0$ ) for spawn-year.



Figure 8. Effective population size estimates $\left(N_{\mathrm{b}}\right)$ from Wenatchee River adult hatcheryproduced steelhead annual collections calculated using single sample methods implemented in the program LDNE (Waples and Do 2008). Each line connects annual estimates of $N_{\mathrm{b}}$ estimated with a different value of $P_{\text {crit }}$, the smallest allelic proportion allowed during analysis. With SNP data, omitting an allele omits the locus. Estimates of $N_{\mathrm{b}}$ changed very little when $P_{\text {crit }}$ varied from 0.1 to 0.001 . Setting $P_{\text {crit }}=0.001$ forced the use of all available loci.


Figure 9. Estimates of Wenatchee River steelhead effective number of breeders ( $N_{\mathrm{b}}$ ) estimated using the single sample methods incorporated in the program LDNE (Waples and Do 2008). Estimates of $N_{\mathrm{b}}$ refer to parental (and even grantparental) generations. $N_{\mathrm{b}}$ data were plotted against their estimated parental brood year. We assumed a 5 year generation time for natural origin adults (NOR), a 4 year generation time for hatchery-produced adults (HOR) and an age of smolt outmigration of age 2 for smolt collections from Wenatchee River tributaries (Chiwawa River, Nason Creek, Peshastin Creek), the lower Wenatchee River, and the Entiat River. Bars represent the $95 \%$ confidence interval estimated by jackknife procedure. Bars that exceed the upper limit of the Y axis are labeled with the upper bound (Inf. = infinity).


Figure 10. Estimates of $N_{\mathrm{b}}$ for collections of hatchery-produced (HOR) and natural origin (NOR) Wenatchee River summer steelhead grouped by brood year rather than spawn year. Brood year was estimated using scale-based age data. Error bars that extend past the top of the chart are all bounded by infinity.


Figure 11. Estimates of $N_{\mathrm{b}}$ for combined annual adult hatchery-produced (HOR) and natural origin (NOR) steelhead and for HOR adults alone. The temporal patterns are similar, though estimates from combined collections are larger than those from HOR collections alone.


Figure 12. $N_{\mathrm{b}} / N$ ratios for hatchery-produced (HOR) and natural origin (NOR) adult Wenatchee River summer steelhead grouped by spawn year. The average $N_{\mathrm{b}} / N$ ratios are not different, though in later years NOR adults appear to have lower $N_{\mathrm{b}} / N$ ratios.


Figure 13. $N_{\mathrm{b}} / N$ ratios for hatchery-produced (HOR) and natural origin (NOR) adult Wenatchee River summer steelhead collections with individuals grouped in brood years rather than spawn years. Individual brood year was estimated using scale-based age data.


## Tables

Table 1. Samples of adult steelhead collected for Wenatchee Program broodstock and used for genetic monitoring and evaluation.

| Origin | Sampling Location | Year <br> spawned | WDFW <br> Collection <br> code | Samples (N) | Unused <br> Samples $^{\text {a }}$ |
| :--- | :--- | :---: | :---: | :---: | :---: |
| Hatchery | Dryden/Tumwater Dams | 1998 | 98 AE | 32 | 4 |
|  |  | 1999 | 98 LJ | 62 | 2 |
|  |  | 2000 | 99 NE | 60 | 5 |
|  |  | 2001 | 00 DQ | 99 | 1 |
| Natural | 2002 | 01 MS | 64 |  |  |
|  |  | 2003 | 02 NP | 89 |  |
|  |  | 2004 | 03 KW | 61 |  |
|  |  | 2007 | 06 CW | 64 | 1 |
|  |  | 2008 | 08 AG | 56 |  |
|  |  | 2009 | 09 AV | 74 |  |
|  |  | 2010 | 10 FE | 76 | 1 |
|  |  | Total | 737 | 14 |  |
|  |  | 1998 | 98 AF | 30 | 5 |
|  |  | 1999 | 99 AA | 51 | 1 |
|  |  | 2000 | 99 ND | 33 | 3 |
|  |  | 2001 | 00 DP | 50 |  |
|  |  | 2002 | 01 MR | 95 |  |
|  |  | 2003 | 02 NO | 50 |  |
|  | 2004 | 03 KV | 71 | 3 |  |
|  |  | 2007 | 06 CX | 74 |  |
|  |  | 2008 | 08 AF | 74 | 1 |
|  | 2009 | 09 AU | 82 | 2 |  |
|  | 2010 | 10 FD | 90 | 2 |  |
|  |  | Total | 700 | 17 |  |

${ }^{2}$ Samples were not used if they had incomplete ( $\leq 80 \%$ or 95 of 119 loci) or duplicate genotypes.

Table 2. Samples of natural origin juvenile steelhead and rainbow trout collected from four Wenatchee basin rivers or creeks and the Entiat River.

| Sampling Location | WDFW |  |  | Unused samples $^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: |
|  | Collection Year | Collection Code | Samples (N) |  |
| Chiwawa River | 2007 | 07AO | 127 | 5 |
|  | 2008 | 08CG | 143 | 1 |
|  | 2009 | 09NF | 35 | 2 |
| Entiat River | 2007 | 07AL | 134 | 4 |
|  | 2008 | 08CI | 82 | 4 |
|  | 2009 | 09NC | 74 | 1 |
|  | 2010 | 100X | 82 | 1 |
| Lower Wenatchee River | 2007 | 07AM | 139 | 5 |
|  | 2008 | 08CE | 98 | 2 |
| Nason Creek | 2007 | 07AN | 81 | 4 |
|  | 2008 | 08CF | 133 | 6 |
|  | 2009 | 09NG | 103 | 2 |
| Peshastin Creek | 2008 | 08 CH | 142 | 2 |
|  | 2009 | 09NE | 34 | 1 |
|  | 2010 | 100Y | 94 | 1 |
|  |  | Total | 1501 | 41 |

${ }^{\text {a }}$ Samples were not used if they were genetically identified as cutthroat trout or cutthroat/rainbow trout hybrids, or if they had incomplete ( $\leq 80 \%$ or 95 of 119 loci) or duplicate genotypes.

Table 3. List of 132 general use, diploid single nucleotide polymorphic (SNP) loci genotyped in Wenatchee River basin and Entiat River steelhead.

| WDFW Name | Locus Name | Allele 1 | Allele 2 | Reference |
| :---: | :---: | :---: | :---: | :--- |
| AOmy005 | Omy_aspAT-123 | T | C | (Campbell et al. 2009) |
| AOmy014 | Omy_e1-147 | G | T | (Sprowles et al. 2006) |
| AOmy015 | Omy_gdh-271 | C | T | (Campbell et al. 2009) |
| AOmy016 | Omy_GH1P1_2 | C | T | (Aguilar and Garza 2008) |
| AOmy021 | Omy_LDHB-2_e5 | T | C | (Aguilar and Garza 2008) |
| AOmy023 | Omy_MYC_2 | T | C | (Aguilar and Garza 2008) |
| AOmy027 | Omy_nkef-241 | C | A | (Campbell et al. 2009) |
| AOmy028 | Omy_nramp-146 | G | A | (Campbell et al. 2009) |
| AOmy047 | Omy_u07-79-166 | G | T | WDFW - S. Young unpubl. |
| AOmy051 | Omy_121713-115 | T | A | (Abadía-Cardoso et al. 2011) |
| AOmy056 | Omy_128693-455 | T | C | (Abadía-Cardoso et al. 2011) |
| AOmy059 | Omy_187760-385 | A | T | (Abadía-Cardoso et al. 2011) |
| AOmy061 | Omy_96222-125 | T | C | (Abadía-Cardoso et al. 2011) |
| AOmy062 | Omy_97077-73 | T | A | (Abadía-Cardoso et al. 2011) |
| AOmy063 | Omy_97660-230 | C | G | (Abadía-Cardoso et al. 2011) |
| AOmy065 | Omy_97954-618 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy067 | Omy_aromat-280 | A | T | WSU - J. DeKoning unpubl. |
| AOmy068 | Omy_arp-630 | G | A | (Campbell et al. 2009) |
| AOmy071 | Omy_cd59-206 | C | T | WSU - J. DeKoning unpubl. |
| AOmy073 | Omy_colla1-525 | C | T | WSU - J. DeKoning unpubl. |
| AOmy079 | Omy_g12-82 | T | C | WSU - J. DeKoning unpubl. |
| AOmy081 | Omy_gh-475 | C | T | (Campbell et al. 2009) |
| AOmy082 | Omy_gsdf-291 | T | C | WSU - J. DeKoning unpubl. |
| AOmy089 | Omy_hsp90BA-193 | C | T | (Campbell and Narum 2009) |
| AOmy094 | Omy_inos-97 | C | A | WSU - J. DeKoning unpubl. |
| AOmy095 | Omy_mapK3-103 | A | T | CRITFC - N. Campbell unpubl. |
| AOmy096 | Omy_mcsf-268 | T | C | WSU - J. DeKoning unpubl. |
| AOmy100 | Omy_nach-200 | A | T | WSU - J. DeKoning unpubl. |
|  |  |  |  |  |


| AOmy107 | Omy_Ots249-227 | C | T | (Campbell et al. 2009) |
| :---: | :---: | :---: | :---: | :---: |
| AOmy108 | Omy_oxct-85 | A | T | WSU - J. DeKoning unpubl. |
| AOmy110 | Omy_star-206 | A | G | WSU - J. DeKoning unpubl. |
| AOmy111 | Omy_stat3-273 | G | Deletion | WSU - J. DeKoning unpubl. |
| AOmy113 | Omy_tlr3-377 | C | T | WSU - J. DeKoning unpubl. |
| AOmy117 | Omy_u09-52-284 | T | G | WDFW - S. Young unpubl. |
| AOmy118 | Omy_u09-53-469 | T | C | WDFW - S. Young unpubl. |
| AOmy120 | Omy_u09-54.311 | C | T | WDFW - S. Young unpubl. |
| AOmy123 | Omy_u09-55-233 | A | G | WDFW - S. Young unpubl. |
| AOmy125 | Omy_u09-56-119 | T | C | WDFW - S. Young unpubl. |
| AOmy129 | Omy_BAMBI4.238 | T | C | WDFW - S. Young unpubl. |
| AOmy132 | Omy_G3PD_2.246 | C | T | WDFW - S. Young unpubl. |
| AOmy134 | Omy_Il-1b-028 | T | C | WDFW - S. Young unpubl. |
| AOmy137 | Omy_u09-61.043 | A | T | WDFW - S. Young unpubl. |
| AOmy151 | Omy_p53-262 | T | A | CRITFC - N. Campbell unpubl. |
| AOmy 173 | BH2VHSVip10 | C | T | Pascal \& Hansen unpubl. |
| AOmy174 | OMS00003 | T | G | (Sánchez et al. 2009) |
| AOmy176 | OMS00013 | A | G | (Sánchez et al. 2009) |
| AOmy177 | OMS00018 | T | G | (Sánchez et al. 2009) |
| AOmy179 | OMS00041 | G | C | (Sánchez et al. 2009) |
| AOmy181 | OMS00052 | T | G | (Sánchez et al. 2009) |
| AOmy182 | OMS00053 | T | C | (Sánchez et al. 2009) |
| AOmy183 | OMS00056 | T | C | (Sánchez et al. 2009) |
| AOmy184 | OMS00057 | T | G | (Sánchez et al. 2009) |
| AOmy185 | OMS00061 | T | C | (Sánchez et al. 2009) |
| AOmy186 | OMS00062 | T | C | (Sánchez et al. 2009) |
| AOmy187 | OMS00064 | T | G | (Sánchez et al. 2009) |
| AOmy189 | OMS00071 | A | G | (Sánchez et al. 2009) |
| AOmy190 | OMS00072 | A | G | (Sánchez et al. 2009) |
| AOmy191 | OMS00078 | T | C | (Sánchez et al. 2009) |
| AOmy192 | OMS00087 | A | G | (Sánchez et al. 2009) |


| AOmy193 | OMS00089 | A | G | (Sánchez et al. 2009) |
| :--- | :--- | :---: | :---: | :--- |
| AOmy194 | OMS00090 | T | C | (Sánchez et al. 2009) |
| AOmy195 | OMS00092 | A | C | (Sánchez et al. 2009) |
| AOmy196 | OMS00094 | T | G | (Sánchez et al. 2009) |
| AOmy197 | OMS00103 | A | T | (Sánchez et al. 2009) |
| AOmy198 | OMS00105 | T | G | (Sánchez et al. 2009) |
| AOmy199 | OMS00112 | A | T | (Sánchez et al. 2009) |
| AOmy200 | OMS00116 | T | A | (Sánchez et al. 2009) |
| AOmy201 | OMS00118 | T | G | (Sánchez et al. 2009) |
| AOmy202 | OMS00119 | A | T | (Sánchez et al. 2009) |
| AOmy203 | OMS00120 | A | G | (Sánchez et al. 2009) |
| AOmy204 | OMS00121 | T | C | (Sánchez et al. 2009) |
| AOmy205 | OMS00127 | T | G | (Sánchez et al. 2009) |
| AOmy206 | OMS00128 | T | G | (Sánchez et al. 2009) |
| AOmy207 | OMS00132 | A | T | (Sánchez et al. 2009) |
| AOmy208 | OMS00133 | A | G | (Sánchez et al. 2009) |
| AOmy209 | OMS00134 | A | G | (Sánchez et al. 2009) |
| AOmy210 | OMS00153 | T | G | (Sánchez et al. 2009) |
| AOmy211 | OMS00154 | A | T | (Sánchez et al. 2009) |
| AOmy212 | OMS00156 | A | T | (Sánchez et al. 2009) |
| AOmy213 | OMS00164 | T | G | (Sánchez et al. 2009) |
| AOmy215 | OMS00175 | T | C | (Sánchez et al. 2009) |
| AOmy216 | OMS00176 | T | G | (Sánchez et al. 2009) |
| AOmy218 | OMS00180 | T | G | (Sánchez et al. 2009) |
| AOmy220 | Omy_1004 | A | T | (Hansen et al. 2011) |
| AOmy221 | Omy_101554-306 | T | C | (Abadía-Cardoso et al. 2011) |
| AOmy222 | Omy_101832-195 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy223 | Omy_101993-189 | A | T | (Abadía-Cardoso et al. 2011) |
| AOmy225 | Omy_102505-102 | A | G | (Abadí-Cardoso et al. 2011) |
| AOmy226 | Omy_102867-443 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy227 | Omy_103705-558 | T | C | (Abadía-Cardoso et al. 2011) |


| AOmy228 | Omy_104519-624 | T | C | (Abadía-Cardoso et al. 2011) |
| :--- | :---: | :---: | :---: | :--- |
| AOmy229 | Omy_104569-114 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy230 | Omy_105075-162 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy231 | Omy_105385-406 | T | C | (Abadía-Cardoso et al. 2011) |
| AOmy232 | Omy_105714-265 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy233 | Omy_107031-704 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy234 | Omy_107285-69 | C | G | (Abadía-Cardoso et al. 2011) |
| AOmy235 | Omy_107336-170 | C | G | (Abadía-Cardoso et al. 2011) |
| AOmy238 | Omy_108007-193 | A | G | (Abadía-Cardoso et al. 2011) |
| AOmy239 | Omy_109243-222 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy240 | Omy_109525-403 | A | G | (Abadía-Cardoso et al. 2011) |
| AOmy241 | Omy_110064-419 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy242 | Omy_110078-294 | A | G | (Abadía-Cardoso et al. 2011) |
| AOmy243 | Omy_110362-585 | G | A | (Abadía-Cardoso et al. 2011) |
| AOmy244 | Omy_110689-148 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy245 | Omy_111005-159 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy246 | Omy_111084-526 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy247 | Omy_111383-51 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy248 | Omy_111666-301 | T | A | (Abadía-Cardoso et al. 2011) |
| AOmy249 | Omy_112301-202 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy250 | Omy_112820-82 | G | A | (Abadía-Cardoso et al. 2011) |
| AOmy252 | Omy_114976-223 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy253 | Omy_116733-349 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy254 | Omy_116938-264 | A | G | (Abadía-Cardoso et al. 2011) |
| AOmy255 | Omy_117259-96 | T | C | (Abadía-Cardoso et al. 2011) |
| AOmy256 | Omy_117286-374 | A | T | (Abadía-Cardoso et al. 2011) |
| AOmy257 | Omy_117370-400 | A | G | (Abadía-Cardoso et al. 2011) |
| AOmy258 | Omy_117540-259 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy260 | Omy_117815-81 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy261 | Omy_118175-396 | T | A | (Abadía-Cardoso et al. 2011) |
| AOmy262 | Omy_118205-116 | A | G | (Abadía-Cardoso et al. 2011) |


| AOmy263 | Omy_118654-91 | A | G | (Abadía-Cardoso et al. 2011) |
| :---: | :---: | :---: | :---: | :--- |
| AOmy265 | Omy_120255-332 | A | T | (Abadía-Cardoso et al. 2011) |
| AOmy266 | Omy_128996-481 | T | G | (Abadía-Cardoso et al. 2011) |
| AOmy267 | Omy_129870-756 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy268 | Omy_131460-646 | C | T | (Abadía-Cardoso et al. 2011) |
| AOmy269 | Omy_98683-165 | A | C | (Abadía-Cardoso et al. 2011) |
| AOmy270 | Omy_cyp17-153 | C | T | WSU - J. DeKoning unpubl. |
| AOmy271 | Omy_ftzf1-217 | A | T | WSU - J. DeKoning unpubl. |
| AOmy272 | Omy_GHSR-121 | T | C | CRITFC - N. Campbell unpubl. |
| AOmy273 | Omy_metA-161 | T | G | CRITFC - N. Campbell unpubl. |
| AOmy274 | Omy_UBA3b | A | T | (Hansen et al. 2011) |

Primer and probe sequences for unpublished loci available by request.

Table 4. List of 20 species identification single nucleotide polymorphic (SNP) loci genotyped in Wenatchee River basin and Entiat River steelhead.

|  |  | Expected genotype |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :--- |
| WDFW Name | Locus Name | O. mykiss | O. clarkii clarkii | O. clarkii lewisi | Reference |
| ASpI001 | Ocl_Okerca | T | C | C | (McGlauflin et al. 2010) |
| ASpI002 | Ocl_Oku202 | A | C | C | (McGlauflin et al. 2010) |
| ASpI003 | Ocl_Oku211 | G | T | T | (McGlauflin et al. 2010) |
| ASpI004 | Ocl_Oku216 | C | C | A | (McGlauflin et al. 2010) |
| ASpI005 | Ocl_Oku217 | C | C | A | (McGlauflin et al. 2010) |
| ASpI006 | Ocl_SsaHM5 | A | A | G | (McGlauflin et al. 2010) |
| ASpI007 | Ocl_u800 | T | C | C | (McGlauflin et al. 2010) |
| ASpI008 | Ocl_u801 | A | T | T | (McGlauflin et al. 2010) |
| ASpI009 | Ocl_u802 | C | C | T | (McGlauflin et al. 2010) |
| ASpI010 | Ocl_u803 | C | T | T | (McGlauflin et al. 2010) |
| ASpI011 | Ocl_u804 | G | G | C | (McGlauflin et al. 2010) |
| ASpI012 | Omy_B9_228 | A | A | C | (Finger et al. 2009) |
| ASpI013 | Omy_CTDL1_243 | C | A | A | (Finger et al. 2009) |
| ASpI014 | Omy_F5_136 | C | G | G | (Finger et al. 2009) |
| ASpI016 | Omy_myclarp404-111 | T | G | G | CRITFC - S. Narum - unpubl. |
| ASpI017 | Omy_myclgh1043-156 | C | T | T | CRITFC - S. Narum - unpubl. |
| ASpI018 | Omy_Omyclmk436-96 | A | C | C | CRITFC - S. Narum - unpubl. |
| ASpI019 | Omy_RAG11_280 | T | A | A | (Sprowles et al. 2006) |
| ASpI020 | Omy_URO_302 | T | C | C | (Finger et al. 2009) |
| ASpI021 | Omy_BAC-F5.238 | C | G | G | WDFW - S. Young unpubl. |

Primer and probe sequences for unpublished loci available by request.

Table 5. Pairwise $F_{\text {ST }}$ estimates for collections from Wenatchee River tributaries and the Entiat River (below diagonal) and associated bootstrap estimated $P$-values (above diagonal).

| Population | Year | Chiwawa River |  |  | Nason Creek |  |  |  Lower <br>  Wenatchee <br> Peshastin Creek River |  |  |  |  | Entiat River |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 2007 | 2008 | 2009 | 2007 | 2008 | 2009 | 2008 | 2009 | 2010 | 2007 | 2008 | 2007 | 2008 | 2009 | 2010 |
| Chiwawa | 2007 |  | 0.000 | 0.003 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.000 | 0.001 | 0.001 | 0.000 | 0.001 | 0.000 | 0.000 |
| River | 2008 | 0.004 |  | 0.004 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
|  | 2009 | 0.004 | 0.003 |  | 0.000 | 0.001 | 0.061 | 0.000 | 0.001 | 0.000 | 0.086 | 0.050 | 0.022 | 0.108 | 0.005 | 0.045 |
| Nason | 2007 | 0.011 | 0.010 | 0.007 |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| Creek | 2008 | 0.007 | 0.007 | 0.005 | 0.009 |  | 0.003 | 0.000 | 0.002 | 0.000 | 0.079 | 0.000 | 0.001 | 0.000 | 0.000 | 0.000 |
|  | 2009 | 0.007 | 0.007 | 0.003 | 0.014 | 0.006 |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| Peshastin | 2008 | 0.010 | 0.011 | 0.008 | 0.013 | 0.010 | 0.013 |  | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| Creek | 2009 | 0.005 | 0.005 | 0.006 | 0.010 | 0.007 | 0.008 | 0.003 |  | 0.002 | 0.002 | 0.047 | 0.028 | 0.004 | 0.005 | 0.001 |
|  | 2010 | 0.010 | 0.011 | 0.008 | 0.015 | 0.008 | 0.011 | 0.003 | 0.003 |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| Lower |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Wenatchee | 2007 | 0.003 | 0.003 | 0.000 | 0.005 | 0.008 | 0.007 | 0.009 | 0.010 | 0.008 |  | 0.112 | 0.020 | 0.012 | 0.002 | 0.017 |
| River | 2008 | 0.002 | 0.005 | 0.002 | 0.003 | 0.004 | 0.005 | 0.007 | 0.009 | 0.006 | 0.000 |  | 0.049 | 0.459 | 0.047 | 0.002 |
| Entiat | 2007 | 0.005 | 0.006 | 0.002 | 0.005 | 0.006 | 0.005 | 0.005 | 0.007 | 0.006 | 0.001 | 0.002 |  | 0.451 | 0.173 | 0.000 |
| River | 2008 | 0.004 | 0.004 | 0.000 | 0.007 | 0.005 | 0.007 | 0.008 | 0.009 | 0.011 | 0.002 | 0.001 | 0.000 |  | 0.644 | 0.002 |
|  | 2009 | 0.005 | 0.006 | 0.002 | 0.003 | -0.001 | 0.003 | 0.002 | 0.003 | 0.004 | 0.003 | 0.002 | 0.002 | 0.000 |  | 0.028 |
|  | 2010 | 0.005 | 0.006 | 0.003 | 0.006 | 0.004 | 0.006 | 0.006 | 0.008 | 0.009 | 0.002 | 0.003 | 0.003 | 0.003 | 0.002 |  |

$P$-values in bold were significant at $\alpha=0.05$ after correcting for multiple tests using false discovery rate.

## Appendix F

NPDES Hatchery Effluent Monitoring, 2015

## NPDES MONITORING FOR WDFW FACILITIES

The WDFW facilities requiring discharge reports are: Chelan Hatchery, Chelan Falls Hatchery, Eastbank Hatchery, Wells Hatchery, Chiwawa Ponds, Methow Hatchery, Similkameen Hatchery, Dryden Acclimation Pond, and Priest Rapids Hatchery. Carlton Acclimation Pond permit became inactive January 2014. An inactive permit is exempt from sampling and submitting discharge reports because production is below the permit requirements for monitoring discharges. NPDES permits are not required for the Twisp and Chewuch acclimation facilities, because they are below the levels that require a discharge permit.

The Wells Hatchery Pollution Abatement (PA) pond has no effluent data January through December. Priest Rapids Hatchery Pollution Abatement (PA) pond has no effluent data January through March, and September through December. The PA ponds for these facilities had no discharge throughout these months.

There were no violations reported at these NPDES permitted facilities during the period 1 January 2015 through 31 December 2015.

All WDFW hatcheries monitor their discharge in accordance with the National Pollutant Discharge Elimination System (NPDES) permit. This permit is administered in Washington by the Washington Department of Ecology under agreement with the United States Environmental Protection Agency. The current permit was extended until 31 March 2016. The permit was renewed effective 1 April 2016 and will expire 31 March 2021.

Facilities are exempted from sampling during any month that pounds of fish on hand fall below $20,000 \mathrm{lbs}$ and pounds of feed used fall below $5,000 \mathrm{lbs}$, with the exception of offline settling basin discharges, which are to be monitored once per month when ponds are in use and discharging to receiving waters. Inactive permitted facilities retain a permit, but are not required to monitor discharges because the pounds of fish and pounds of feed remain below monitoring guidelines set by the permit.

Sampling at permitted facilities includes the following parameters:
$<$ FLOW Measured in millions of gallons per day (MGD) discharge.
$<$ SS EFF Average net settleable solids in the hatchery effluent, measured in $\mathrm{ml} / \mathrm{L}$.
$<$ TSS COMP Average net total suspended solids, composite sample ( $6 \mathrm{x} / \mathrm{day}$ ) of the hatchery effluent, measured in mg/L.
<TSS MAX Maximum daily net total suspended solids, composite sample ( $6 \mathrm{x} /$ day) of the hatchery effluent, measured in $\mathrm{mg} / \mathrm{L}$.
$<$ SS PA Maximum settleable solids discharge from the pollution abatement pond, measured in $\mathrm{ml} / \mathrm{L}$.
$<$ SS \% Removal of settleable solids within the pollution abatement pond from inlet to outlet, measured as a percent. No longer required under permit effective 1 June 2000.
$<$ TSS PA Maximum total suspended solids effluent grab from the pollution abatement pond discharge, measured in $\mathrm{mg} / \mathrm{L}$.
$<$ TSS \% Removal of suspended solids within the pollution abatement pond from inlet to outlet, measured as a percent. No longer required under permit effective June 1, 2000.
<SS DD Settleable solids discharged during drawdown for fish release. One sample per pond drawdown, measured in $\mathrm{ml} / \mathrm{L}$.
$<$ TRC Total residual chlorine discharge after rearing vessel disinfection and after neutralization with sodium thiosulfate. One sample per disinfection, measured in ug/L.

In addition, at Similkameen Hatchery only, the following sampling was conducted at the request of WA Dept of Ecology, but is not required under NPDES permit:
$<$ SS IW Settleable solids influent grab taken as wastes are pumped into the pollution abatement pond, measured in mg/L. No longer monitored as of January 2008.
$<$ TSS IW $\quad$ Total suspended solids influent grab as wastes are pumped into the pollution abatement pond, measured in mg/L. No longer monitored as of January 2008.

| Eastbank Hatchery <br> NPDES Permit Number WAG13- $5011$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | FLOW | SS EFF | $\begin{gathered} \text { TSS } \\ \text { COMP } \end{gathered}$ | $\begin{gathered} \hline \text { TSS } \\ \text { MAX } \end{gathered}$ | FLOW PA | SS PA | $\begin{aligned} & \text { SS } \\ & \% \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \text { PA } \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \% \end{aligned}$ | lbs of Fish | Ibs of Feed |  |  |
| 2015 | JAN | 28.43 | 0 | 0.4 | 0.4 | 7000 | 0.01 |  | 26.6 |  | 25412 | 6743 |  |  |
|  | FEB | 28.43 | 0 | 0.4 | 0.4 | 8500 | 0.01 |  | 24.8 |  | 33757 | 4618 |  |  |
|  | MAR | 20.68 | 0 | 0 | 0 | 10000 | 0.01 |  | 10 |  | 26814 | 5033 |  |  |
|  | APR | 22.29 | 0 | 0 | 0 | 3000 | 0.01 |  | 21.4 |  | 19553 | 5573 |  |  |
|  | MAY | 22.96 | 0 | 0 | 0 | 5000 | 0.01 |  | 14.2 |  | 27705 | 8855 |  |  |
|  | JUN | 29.73 | 0 | 0.2 | 0.2 | 7500 | 0.01 |  | 15 |  | 37051 | 9782 |  |  |
|  | JUL | 25.85 | 0 | 0.4 | 0.4 | 5000 | 0.01 |  | 10.6 |  | 35599 | 5821 |  |  |
|  | AUG | 27.14 | 0 | 1 | 1.4 | 7500 | 0.01 |  | 20.8 |  | 17833 | 6587 |  |  |
|  | SEP | 27.78 | 0 | 0.4 | 0.4 | 15000 | 0.01 |  | 39.8 |  | 24733 | 10184 |  |  |
|  | OCT | 31.03 | 0 | 0.2 | 0.2 | 10000 | 0.01 |  | 2.6 |  | 35072 | 9143 |  |  |
|  | NOV | 23.59 | 0 | 0 | 0 | 7500 | 0.01 |  | 17.6 |  | 24480 | 3504 |  |  |
|  | DEC | 23.59 | 0 | 0.6 | 0.6 | 5000 | 0.01 |  | 15.6 |  | 19478 | 4759 |  |  |
| Wells <br> Hatchery <br> NPDES Permit Number WAG13- <br> 5009 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | FLOW | SS EFF | $\begin{gathered} \text { TSS } \\ \text { COMP } \end{gathered}$ | $\begin{gathered} \text { TSS } \\ \text { MAX } \\ \hline \end{gathered}$ | FLOW PA | SS PA | $\begin{aligned} & \text { SS } \\ & \% \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \text { PA } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \% \\ & \hline \end{aligned}$ | lbs of Fish | Ibs of Feed | $\begin{aligned} & \text { SS } \\ & \text { DD } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \text { DD } \\ & \hline \end{aligned}$ |
| 2015 | JAN | 16.85 | 0 | 0.2 | 0.2 | ** | ** |  | ** |  | 69543 | 14511 |  |  |
|  | FEB | 19.41 | 0 | 0.2 | 0.2 | ** | ** |  | ** |  | 79660 | 17750 |  |  |
|  | MAR | 18.96 | 0 | 0.2 | 0.2 | ** | ** |  | ** |  | 101677 | 15519 |  |  |
|  | APR | 16.13 | 0 | 0.2 | 0.4 | ** | ** |  | ** |  | 85708 | 9827 | 0.1 | 1.4 |
|  | MAY | 11.54 | 0 | 0.6 | 0.6 | ** | ** |  | ** |  | 30900 | 5296 | 0.17 | 3 |
|  | JUN | 5.54 | 0 | 0.8 | 0.8 | ** | ** |  | ** |  | 9177 | 1887 |  |  |
|  | JUL | 5.38 | 0 | 0.4 | 0.4 | ** | ** |  | ** |  | 7459 | 2459 |  |  |
|  | AUG | 5.69 | 0.01 | 0.2 | 0.2 | ** | ** |  | ** |  | 11132 | 6628 |  |  |
|  | SEP | 7.06 | 0.01 | 1 | 1 | ** | ** |  | ** |  | 21400 | 7904 |  |  |
|  | OCT | 8.49 | 0.01 | 0.9 | 1 | ** | ** |  | ** |  | 30343 | 8420 |  |  |
|  | NOV | 9.95 | 0 | 1.2 | 1.2 | ** | ** |  | ** |  | 39509 | 13790 |  |  |
|  | DEC | 10.53 | 0.01 | 1.4 | 1.4 | ** | ** |  | ** |  | 53633 | 14376 |  |  |

[^13]Chiwawa Ponds - Chiwawa River
NPDES Permit Number WAG13-
5015

|  |  | FLOW | SS EFF | TSS |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COMP | TSS |  |  |  |  |  |  |  |  |
| MAX | lbs of Fish | lbs of <br> Feed | SS <br> DD | TSS <br> DD |  |  |  |  |  |
| 2015 | JAN | 4.25 | 0 | 0.8 | 0.8 | 10040 | 300 |  |  |
|  | FEB | 3.62 | 0 | 1.4 | 1.4 | 15765 | 390 |  |  |
| MAR | 4.52 | 0 | -0.4 | -0.4 | 9775 | 260 |  |  |  |
| APR | 3.85 | 0 | -0.2 | -0.2 | 8194 | 132 | 0.04 | 4 |  |
| MAY | No Monitoring |  |  | 0 | 0 |  |  |  |  |
| JUN | No Monitoring |  |  | 0 | 0 |  |  |  |  |
| JUL | No Monitoring |  |  | 0 | 0 |  |  |  |  |
| AUG | No Monitoring |  |  | 0 | 0 |  |  |  |  |
| SEP | No Monitoring |  |  | 0 | 0 |  |  |  |  |
| OCT | 4.22 | 0 | 1 | 1 | 6042 | 1012 |  |  |  |
| NOV | 3.65 | 0 | -0.2 | -0.2 | 11234 | 348 |  |  |  |
| DEC | 3.49 | 0 | 2 | 2 | 10026 | 341 |  |  |  |

Chiwawa Ponds - Wenatchee
River
NPDES Permit Number WAG13-
$\left.\begin{array}{ccccccccc}\mathbf{5 0 1 5} & & & & & & \\ \hline & & \text { FLOW } & \text { SS EFF } & \begin{array}{c}\text { TSS } \\ \text { COMP }\end{array} & \begin{array}{c}\text { TSS } \\ \text { MAX }\end{array} & \text { lbs of Fish } & \text { Feed } & \begin{array}{c}\text { SS } \\ \text { DD }\end{array} \\ \hline \text { TSS } \\ \text { DD }\end{array}\right]$

Methow Hatchery
NPDES Permit Number WAG13-
5000

|  |  | FLOW | SS EFF | $\begin{gathered} \text { TSS } \\ \text { COMP } \\ \hline \end{gathered}$ | $\begin{aligned} & \text { TSS } \\ & \text { MAX } \\ & \hline \end{aligned}$ | FLOW PA | SS PA | $\begin{aligned} & \text { SS } \\ & \% \\ & \hline \end{aligned}$ | TSS PA | TSS \% | lbs of Fish | lbs of Feed | $\begin{aligned} & \text { SS } \\ & \text { DD } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \text { DD } \\ & \hline \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | JAN | 11.52 | 0 | 0.1 | 0.2 | 14400 | 0.1 |  | 0 |  | 9700 | 1300 |  |  |
|  | FEB | 11.52 | 0 | 0 | 0 | 14400 | 0.1 |  | 5.4 |  | 10500 | 1420 |  |  |
|  | MAR | 10.08 | 0 | 1.3 | 1.8 | 14400 | 0.1 |  | 0.2 |  | 9600 | 828 |  |  |
|  | APR | 10.08 | 0 | -0.4 | -0.4 | 14400 | 0.1 |  | 6.4 |  | 9700 | 900 | 0 | 0 |
|  | MAY | 2.6 | 0 | 1.4 | 1.4 | 14400 | 0.1 |  | 5.2 |  | 1223 | 455 | 0.1 | 3.8 |
|  | JUN | 3.77 | 0 | 0 | 0 | 14400 | 0.1 |  | 0 |  | 2036 | 757 |  |  |
|  | JUL | 4.32 | 0 | 0.6 | 0.6 | 14400 | 0.1 |  | 17.2 |  | 2600 | 600 |  |  |
|  | AUG | 4.32 | 0 | 0.2 | 0.2 | 14400 | 0.1 |  | 1 |  | 4000 | 1100 |  |  |
|  | SEP | 5.33 | 0 | 0 | 0 | 14400 | 0.1 |  | 0.8 |  | 6200 | 852 |  |  |
|  | OCT | 5.33 | 0 | 0 | 0 | 14400 | 0.1 |  | 0.8 |  | 10000 | 800 |  |  |
|  | NOV | 5.62 | 0 | 0 | 0 | 14400 | 0.1 |  | 3.2 |  | 10600 | 875 |  |  |
|  | DEC | 7.98 | 0 | 0.2 | 0.2 | 14400 | 0.1 |  | 0.2 |  | 11200 | 930 |  |  |

Similkameen Hatchery
NPDES Permit Number WAG13-

| $\mathbf{5 0 0 7}$ |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |



Chelan Falls Hatchery
NPDES Permit Number WAG13-

|  |  | FLOW | SS EFF | $\begin{gathered} \hline \text { TSS } \\ \text { COMP } \end{gathered}$ | $\begin{gathered} \hline \text { TSS } \\ \text { MAX } \end{gathered}$ | FLOW PA | SS PA | $\begin{aligned} & \hline \text { SS } \\ & \% \end{aligned}$ | $\begin{array}{r} \hline \text { TSS } \\ \text { PA } \end{array}$ | $\begin{aligned} & \hline \text { TSS } \\ & \% \end{aligned}$ | Ibs of Fish | Ibs of Feed |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | JAN | 12.9 | 0.05 | -8.8 | -8.8 | 857 | 0.05 |  | 0.8 |  | 31994 | 4568 |
|  | FEB | 12.9 | 0.05 | -1.9 | -1.8 | 857 | 0.05 |  | 1.2 |  | 33820 | 1650 |
|  | MAR | 12.8 | 0.05 | -2.2 | -2.2 | 857 | 0.05 |  | 1.4 |  | 34262 | 3766 |
|  | APR | 12.6 | 0.05 | 0.2 | 0.2 | 857 | 0.05 |  | 1.6 |  | 35344 | 17751 |
|  | MAY | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | JUN | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | JUL | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | AUG | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | SEP | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | OCT | No Moni | ring |  |  |  |  |  |  |  | 0 | 0 |
|  | NOV | 7 | 0.04 | -9.4 | -9.4 | 3000 | 0.05 |  | 0.2 |  | 17614 | 2227 |
|  | DEC | 7 | 0.04 | -1 | -1 | 3000 | 0.05 |  | 0.8 |  | 19753 | 2481 |


| Dryden Acclimation <br> Pond <br> NPDES Permit Number WAG13- <br> 5014 |  |  |  |  |  | lbs of Fish | $\begin{gathered} \text { lbs of } \\ \text { Feed } \\ \hline \end{gathered}$ | $\begin{aligned} & \text { SS } \\ & \text { DD } \end{aligned}$ | $\begin{aligned} & \hline \text { TSS } \\ & \text { DD } \\ & \hline \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | FLOW | SS EFF | $\begin{gathered} \text { TSS } \\ \text { COMP } \end{gathered}$ | $\begin{gathered} \text { TSS } \\ \text { MAX } \end{gathered}$ |  |  |  |  |
| $2015$ | JAN | No Moni | ring |  |  | 0 | 0 |  |  |
|  | FEB | No Moni | ring |  |  | 0 | 0 |  |  |
|  | MAR | 14.21 | 0 | -0.8 | -0.8 | 33366 | 2948 |  |  |
|  | APR | 15.26 | -0.1 | 0.3 | 0.4 | 46973 | 5236 | 0.01 | 3.8 |
|  | MAY | No Moni | ring |  |  | 0 | 0 |  |  |
|  | JUN | No Moni | ring |  |  | 0 | 0 |  |  |
|  | JUL | No Moni | ring |  |  | 0 | 0 |  |  |
|  | AUG | No Moni | ring |  |  | 0 | 0 |  |  |
|  | SEP | No Moni | ring |  |  | 0 | 0 |  |  |
|  | OCT | No Moni | ring |  |  | 0 | 0 |  |  |
|  | NOV | No Moni | ring |  |  | 0 | 0 |  |  |
|  | DEC | No Moni | ring |  |  | 0 | 0 |  |  |

Priest Rapids
NPDES Permit Number WAG13
7013

|  |  | FLOW | SS EFF | $\begin{gathered} \text { TSS } \\ \text { COMP } \end{gathered}$ | $\begin{gathered} \text { TSS } \\ \text { MAX } \end{gathered}$ | FLOW PA | SS PA | $\begin{aligned} & \text { TSS } \\ & \text { PA } \end{aligned}$ | Ibs of Fish | Ibs of Feed | $\begin{aligned} & \text { SS } \\ & \text { DD } \end{aligned}$ | $\begin{aligned} & \text { TSS } \\ & \text { DD } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | JAN | 23.93 | 0 | 0.4 | 0.4 | ** | ** | ** | 9211 | 202 |  |  |
|  | FEB | 26.98 | 0 | 0 | 0 | ** | ** | ** | 10229 | 1180 |  |  |
|  | MAR | 28.24 | 0 | 0.4 | 0.4 | ** | ** | ** | 14796 | 4440 |  |  |
|  | APR | 30.55 | 0 | -0.4 | -0.4 |  | 0 |  | 26695 | 13102 |  |  |
|  | MAY | 46.88 | 0 | 0 | 0 |  | 0 |  | 78430 | 30166 |  |  |
|  | JUN | 44.43 | 0 | 0.6 | 0.6 |  | 0 |  | 135899 | 34962 | 0 | 1.32 |
|  | JUL | No Moni | oring |  |  |  |  |  | 0 | 0 |  |  |
|  | AUG | No Moni | ring |  |  |  |  |  | 0 | 0 |  |  |
|  | SEP | 60.35 |  |  |  | ** | ** | ** | 0 | 0 |  |  |
|  | OCT | 65.95 | 0 |  |  | ** | ** | ** | 0 | 0 |  |  |
|  | NOV | 65.95 | 0 |  |  | ** | ** | ** | 0 | 0 |  |  |
|  | DEC | 24.25 | 0 | -0.4 | -0.4 | ** | ** | ** | 8632 | 0 |  |  |

**PA pond - No discharge this month
Appendix G

Steelhead Stock Assessment at Priest Rapids Dam, 2013-2014

## Priest Rapids Dam 2013-2014 Adult Upper Columbia River Steelhead Run-Cycle Stock Assessment Report

## Introduction

Upper Columbia River (UCR) steelhead stock assessment sampling at Priest Rapids Dam (PRD) is authorized through the Endangered Species Act (ESA) Section 10 Permit 1395 (NMFS 2003). Permit authorizations include interception and biological sampling of up to 10 percent of the UCR steelhead passing PRD to determine upriver population size, estimate hatchery to wild ratios, determine age class contribution and evaluate the need for managing hatchery steelhead consistent with ESA recovery objectives, which include fully seeding spawning habitat with naturally produced UCR steelhead supplemented with artificially propagated enhancement steelhead (NMFS 2003).

## Stock Assessment

The 2013 steelhead sampling at Priest Rapids Dam began 8 July and concluded 14 November. Sampling consisted of operating the Priest Rapids Off Ladder Trap (OLAFT), located on the left bank Priest Rapids Dam, eight hours per day, up to three days per week, for a total of 57 sampling days. Steelhead were trapped, handled, and released in accordance with Section 2.1 and 2.2.1 of the National Marine Fisheries Service (NMFS) Biological Opinion for ESA Permit 1395 (NMFS 2003). The cumulative sample rate attained during 2013 totaled 13.5\%.

The Washington Department of Fish and Wildlife (WDFW) sampled 2,318 steelhead of the 2013/2014 run-cycle passing PRD, totaling 15,072 steelhead, for an overall sampling rate of $14.6 \%$. Of the 2,196 steelhead sampled, $1,426(64.9 \%)$ were hatchery origin and 770 ( $35.1 \%$ ) were wild origin. The estimated 2013-2014 run- cycle total wild steelhead return was 4,657 , representing $166.6 \%$ of the 1986-2012 average and about $88.6 \%$ of the most recent five-year average (Table 1).

Based on external marks and external and internal tags, 1,426 hatchery-origin steelhead were sampled at Priest Rapids Dam during the 2013 return cycle and included 19.5\% Wenatchee hatchery-origin steelhead and $49.6 \%$ "above Wells Dam" hatchery-origin steelhead ${ }^{1}$ (Table 2), while $12.0 \%$ of the hatchery-origin steelhead sampled could not be assigned to a specific hatchery program. Ringold FH origin steelhead represented about $12.5 \%$ of the hatchery sample (Table 2).

[^14]Table 1. Priest Rapids Dam adult steelhead returns and stock composition, 1974-2013.

| Run-cycle ${ }^{1 /}$ | Hatchery | Wild | Wild percent | Total run |
| :---: | :---: | :---: | :---: | :---: |
| 1974 |  |  |  | 2,950 |
| 1975 |  |  |  | 2,560 |
| 1976 |  |  |  | 9,490 |
| 1977 |  |  |  | 9,630 |
| 1978 |  |  |  | 4,510 |
| 1979 |  |  |  | 8,710 |
| 1980 |  |  |  | 8,290 |
| 1981 |  |  |  | 9,110 |
| 1982 |  |  |  | 10,770 |
| 1983 |  |  |  | 32,000 |
| 1984 |  |  |  | 26,200 |
| 1985 |  |  |  | 34,010 |
| 1986 | 20,022 | 2,342 | 10.5 | 22,364 |
| 1987 | 9,955 | 4,058 | 29.0 | 14,013 |
| 1988 | 7,530 | 2,670 | 26.2 | 10,200 |
| 1989 | 8,033 | 2,685 | 25.1 | 10,718 |
| 1990 | 6,252 | 1,585 | 20.2 | 7,837 |
| 1991 | 11,169 | 2,799 | 20.0 | 13,968 |
| 1992 | 12,102 | 1,618 | 11.8 | 13,720 |
| 1993 | 4,538 | 890 | 16.4 | 5,428 |
| 1994 | 5,880 | 855 | 12.7 | 6,735 |
| 1995 | 3,377 | 993 | 22.7 | 4,370 |
| 1996 | 7,757 | 843 | 9.8 | 8,600 |
| 1997 | 8,157 | 785 | 8.8 | 8,942 |
| 1998 | 4,919 | 928 | 15.9 | 5,847 |
| 1999 | 6,903 | 1,374 | 16.6 | 8,277 |
| 2000 | 9,023 | 2,341 | 20.6 | 11,364 |
| 2001 | 24,362 | 5,715 | 19.0 | 30,077 |
| 2002 | 12,884 | 2,983 | 18.8 | 15,867 |
| 2003 | 14,890 | 2,837 | 16.0 | 17,729 |
| 2004 | 15,670 | 2,985 | 16.0 | 18,655 |
| 2005 | 10,352 | 3,127 | 23.2 | 13,479 |
| 2006 | 8,738 | 1,677 | 16.1 | 10,415 |
| 2007 | 12,160 | 3,097 | 20.3 | 15,257 |
| 2008 | 13,528 | 3,030 | 18.3 | 16,558 |
| 2009 | 32,557 | 7,439 | 18.6 | 39,996 |
| 2010 | 18,784 | 7,647 | 28.9 | 26,431 |
| 2011 | 15,910 | 4,896 | 23.5 | 20,806 |
| 2012 | 13,908 | 3,284 | 19.1 | 17,192 |
| 2013 | 10,415 | 4,657 | 30.9 | 15,072 |
| 1986-2012 average | 11,828 | 2,796 | 18.7 | 14,181 |
| 2008-2012 average | 18,939 | 5,257 | 21.7 | 24,197 |

${ }^{1 /}$ A return cycle is the combined total of steelhead passing PRD from 1 June - 30 November during year ( x ), plus steelhead passing PRD between 15 April and 31 May on year ( $\mathrm{x}+1$ ).

Table 2. Origin classification of steelhead sampled at Priest Rapids Dam, 8 July - 14 November 2013.


Reconciliation of salt water age of wild and hatchery steelhead sampled at Priest Rapids Dam during 2013 was accomplished through scale sample analysis. Salt-age analysis of the 2013 UCR steelhead run-cycle provides an estimated hatchery-origin return dominated by 1 - salt and 2 -salt age composition of $60.1 \%$ and $39.7 \%$, respectively (Table 3). Natural-origin steelhead salt ages were $68.6 \%$ and $31.2 \%$ for salt ages 1 and 2, respectively. Three-salt age fish represented only $0.2 \%$ of the combined hatchery/wild sample (Table 3).

Table 3. Salt-water age composition of 2013 - 2014 return cycle Upper Columbia River steelhead sampled at Priest Rapids Dam, corrected by scale age/origin determination.

| Salt-age | Origin |  |  |  | Combined |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Hatchery |  | Wild |  |  |  |
|  | N | \% | $N$ | \% | $N$ | \% |
| 1-salt | 845 | 60.1 | 521 | 68.6 | 1,366 | 63.1 |
| 2-salt | 559 | 39.7 | 237 | 31.2 | 796 | 36.7 |
| 3-salt | 3 | 0.2 | 1 | 0.1 | 4 | 0.2 |
| 4-salt | 0 | 0.0 | 0 | 0.0 | 0 | 0.0 |
| Total | 1,407 | 65.0 | 759 | 35.0 | 2,166 |  |

Freshwater residency of naturally produced Upper Columbia River steelhead present in the 2013-2014 run cycle were dominated by age-2 freshwater fish (70.8\%), and was only slightly lower than the 1986-2012 average of $74.7 \%$ (Table 4).

Table 4. 2013 return year freshwater age of wild Upper Columbia River steelhead sampled at Priest Rapids Dam during steelhead stock assessment activities, compared to July - November 1986-2012 average.

| Freshwater age | 2013-2014 run cycle |  | 1986-2012 proportion |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $N$ | \% | $N$ | \% |
| 1.x | 31 | 4.4 | 458 | 8.4 |
| 2.x | 495 | 70.8 | 4,086 | 74.7 |
| 3.x | 161 | 23.0 | 885 | 16.2 |
| 4.x | 12 | 1.7 | 39 | 0.7 |
| 5.x | 0 | 0.0 | 3 | $>0.1$ |
| Total | 699 |  | 5,471 |  |

Wild and hatchery origin steelhead exhibited similar saltwater growth in the 2013 runcycle. Wild 1- and 2-salt adults were slightly larger than their hatchery cohorts (Table 5). Age 1-salt hatchery and age 1- and 2-salt wild steelhead observed in the 2013-2014 adult run-cycle return past PRD were comparable in size to the 1986-2012 run-cycle average (Table 5).

Table 5. Average fork length of 1-salt and 2-salt, Upper Columbia River steelhead sampled at Priest Rapids Dam during July - November 2013 and the period between 1986-2012.

|  | Average fork length (cm) |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  | 2013-2014 run cycle |  |  |  |
| Salt age | Wild | Hatchery | 1986-2012 run cycle <br> Wild | Hatchery |
| X.1 | 57.7 | 57.2 | 60.3 | 59.0 |
| x. 2 | 70.5 | 69.7 | 72.7 | 71.8 |

## Appendix H

Wenatchee Sockeye Salmon Spawning Escapement, 2015

# PUBLIC UTILITY DISTRICT NUMBER 1 OF CHELAN COUNTY Natural Resource Division <br> Fish and Wildlife Department <br> 327 N. Wenatchee Ave., Wenatchee WA 98801 (509) 663-8121 

March 20, 2015
To: HCP Hatchery Committee
From: Catherine Willard

## Subject: 2015 Wenatchee Sockeye Mark/Recapture-Based Sockeye Escapement Estimates to Tributaries

## Introduction

In 2015, the Chelan County Public Utility District (District) estimated sockeye escapement to tributaries based on mark-recapture methodology. The purpose of this document is to report the spawning escapement estimates for the Little Wenatchee and White River subbasins. This information is used to track and/or estimate viable salmonid population parameters (VSP): abundance, productivity, spatial structure, and diversity (McElhaney et al. 2000).

## Methods

## Mark-Recapture Method:

Detection efficiencies of the in-stream arrays were calculated for the Little Wenatchee River and White River in 2015. The in-stream arrays include a series of upstream and downstream coils (Figure 1). Combined, these coils represented the upstream and downstream detection arrays, respectively. Overall detection efficiency $P_{\text {all }}$ of the arrays was calculated based on observed detection probabilities of individual arrays:

$$
P_{\text {all }}=1-\left(1-P_{\text {array } 1}\right)\left(1-P_{\text {array } 2}\right)
$$

where the probability of missing a fish on both the upstream $P_{\text {array1 }}$ and downstream $P_{\text {array } 2}$ arrays were combined for an overall efficiency $P_{\text {all }}$ (Connolly et al. 2008).

Adult sockeye salmon were tagged at adult fishways within the Columbia River and at Tumwater Dam. Additionally, adult returns that were PIT tagged as juveniles were used in the analyses. Total passage of adult sockeye salmon through Tumwater Dam was obtained from Columbia River Data Access in Real Time (DART 2015). Resulting tag files were queried in PTAGIS (2015), providing detection histories for each study fish.


Figure 1. Schematic of a PIT array configuration.

Resulting data from passage at Tumwater Dam, mark and recapture using PIT tags, and detection efficiency estimates can provide estimation of escapement to spawning tributaries. Assumptions include: (1) the study population is "closed," i.e., no individuals die or emigrate between the initial mark and subsequent recaptures; (2) tags are not lost and detections are correctly identified; (3) all individuals have the same probability of being detected, and (4) the number of recapture events are proportional to the total population. Lastly, it was assumed that PIT-tagging efforts at Tumwater have negligible influence on fish behavior and tagged individuals behave similarly to untagged individuals. The resulting escapement rate, adjusted for detection efficiency, was then applied to the total population as such:

$$
\text { Escapement }=\left(\frac{\left(\frac{O b s_{L W N}}{E f f_{L W N}}+\frac{O b s_{W T L}}{E f f_{W T L}}\right)}{P I T s_{T U M}}\right) \times \text { Counts }_{T U M}
$$

where the PIT tag detections ( $O b s$ ) at the Little Wenatchee ( $L W N$ ) and White River (WTL) were adjusted for detection efficiency (Eff), compared to the number released (PITs) at Tumwater Dam (TUM), and the resulting proportion was applied to the population observed (Counts) passing Tumwater Dam.

## Results

## Sockeye Salmon Mark-Recapture Method

Fishway enumeration at Tumwater Dam indicated that 51,410 adult sockeye salmon passed the facility during the 2015 migration, which was a sufficient return to open a recreational fishery in Lake Wenatchee for 2015. PIT tags were implanted in 943 of these fish at Tumwater and seven of these fish were PIT-tagged before passing Tumwater; 76 fish were subsequently detected at the Little Wenatchee PIT tag array and 371 fish were subsequently detected at the White River PIT tag array (Table 1). Based on the recapture of PIT-tagged adult sockeye and assigned detection efficiency, total estimated escapement from Tumwater Dam to the Little Wenatchee River was 4,113 adult sockeye and 20,087 adult sockeye to the White River (Table 2).

Table 1. Number of adult sockeye salmon PIT-tagged, released, and detected upstream of Tumwater Dam in 2009 through 2015, and mark/recapture based tributary escapement estimates. Obs. $=$ observed, D.E. $=$ detection efficiency, Est $=$ estimated (Obs./D.E.), and NA $=$ not available.

| Year | Number of PIT-tagged adults detected or tagged at Tumwater ${ }^{1}$ | White River |  |  | Little Wenatchee River |  |  | Chiwawa River Obs. | Nason Creek Obs. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Obs. | $\begin{aligned} & \text { D.E. } \\ & \left(\boldsymbol{p}_{\text {all }}\right) \end{aligned}$ | Est | Obs. | $\begin{aligned} & \text { D.E. } \\ & \left(\boldsymbol{p}_{\text {all }}\right) \end{aligned}$ | Est |  |  |
| 2009 | 1,085 | 381 | 0.406 | 939 | 38 | 0.971 | 39 | 37 | 7 |
| 2010 | 1,164 | 571 | $0.900^{2}$ | 635 | 67 | 1.000 | 67 | 3 | 1 |
| 2011 | 484 | 40 | NA ${ }^{3}$ | $N A$ | 84 | -- | 0 | 0 | 0 |
| 2012 | 1,154 | 410 | 0.943 | 435 | 74 | 0.987 | 75 | 0 | 0 |
| 2013 | 719 | 152 | NA ${ }^{3}$ | $N A$ | 55 | 0.818 | 67 | 0 | 0 |
| 2014 | 1,729 | 848 | 0.999 | 848 | 76 | 1.000 | 76 | 0 | 3 |
| $2015{ }^{4}$ | 950 | 371 | 0.999 | 371 | 76 | 1.000 | 76 | 76 | 4 |

[^15]Table 2. Estimated escapement of adult sockeye salmon to Little Wenatchee and White rivers based on mark-recapture events, in-stream detection efficiency, and adult enumeration at Tumwater Dam, 2009-2015.

| Year | Tumwater <br> count | Recreational <br> harvest | Little <br> Wenatchee | White <br> River | Combined | Escapement |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | 16,034 | 2,229 | 576 | 13,876 | 14,452 | 0.901 |
| 2010 | 35,821 | 4,129 | 2,062 | 19,542 | 21,604 | 0.603 |
| $2011^{1}$ | 18,634 | 0 | 2,431 | 14,582 | 17,013 | 0.913 |
| 2012 | 66,520 | 12,107 | 4,607 | 23,866 | 28,473 | 0.428 |
| $2013^{1}$ | 29,015 | 6,262 | 2,426 | 14,294 | 16,720 | 0.576 |
| 2014 | 99,898 | 16,281 | 4,319 | 49,021 | 53,340 | 0.534 |
| 2015 | 51,410 | 7,916 | 4,113 | 20,087 | 24,200 | 0.470 |
| Average | $\mathbf{4 5 , 3 3 3}$ | $\mathbf{6 , 9 8 9}$ | $\mathbf{2 , 9 3 3}$ | $\mathbf{2 2 , 1 8 1}$ | $\mathbf{2 5 , 1 1 5}$ | $\mathbf{0 . 6 8 4}$ |

${ }^{1}$ Escapement was calculated using AUC counts for the Little Wenatchee River and a linear regression relationship to the Little Wenatchee River for the White River.

## References

Hillborn, R. B.G. Bue, and S. Sharr. 1999. Estimating spawning escapements from periodic counts: a comparison of methods. Can. J. Fish. Aquat. Sci. 56: 888-896.

Hyatt, K.D., M.M. Stockwell, H. Wright, K. Long, J. Tamblyn, and M. Walsh. 2006. Fish and Water Management Tool Project Assessments: Okanogan Adult Sockeye Salmon (Oncorhynchus nerka) Abundance and Biological Traits in 2005. DRAFT Report to file: JSID-SRe 3-05. Salmon and Freshwater Ecolsystems Division, Fisheries and Oceans Canada, Nanaimo, B.C. V9T 6N7.

McElhany, P., M. H. Ruckelshaus, M. J. Ford, T. C. Wainwright, and E. P. Bjorkstedt. 2000. Viable salmonid populations and the recovery of Evolutionarily Significant Units. NOAA Technical Memorandum.

Mullan, J. W. 1987. Status and propagation of Chinook salmon in the mid-Columbia River through 1985. U.S. Fish and Wildlife Serv. Biol. Rep. 87(3) 111 pp.

Murdoch, A. and C. Peven. 2005. Conceptual Approach to Monitoring and Evaluating the Chelan County Public Utility District Hatchery Programs. Prepared for: Chelan PUD Habitat Conservation Plan's Hatchery Committee. Chelan PUD, Wenatchee, WA.

Perrin, C.J. and J.R. Irvine. 1990. A review of survey life estimates as they apply to the area-under-the-curve method for estimating the spawning escapement of Pacific salmon. Canadian Tech. Rep. of Fisheries and Aquatic Sciences No. 1733. Department of Fisheries and Oceans, Nanaimo, B.C. V9R 5K6.

Peven, C. M. 1990. The life history of naturally produced steelhead trout from the MidColumbia River Basin. MS Thesis, University of Washington, Seattle.

## Appendix I

Genetic Diversity of Wenatchee Sockeye Salmon

# Assessing the Genetic Diversity of Lake Wenatchee Sockeye Salmon And Evaluating The Effectiveness Of Its Supportive Hatchery Supplementation Program 

Developed for<br>Chelan County PUD<br>and the<br>Habitat Conservation Plan's Hatchery Committee

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## Executive Summary

Nine spawning populations of sockeye (Oncorhynchus nerka) salmon have been identified in Washington, including stocks in the Lake Wenatchee basin (SaSI 5800) (Washington Department of Fisheries et al. 1993). Lake Wenatchee sockeye are classified as an Evolutionary Significant Unit (ESU), and consists of sockeye salmon that spawn primarily in tributaries above Lake Wenatchee (the White River, Napeequa River, and Little Wenatchee Rivers). Since 1990, the Wenatchee Sockeye Program has released juveniles into Lake Wenatchee to supplement natural production of sockeye salmon in the basin. The program's broodstock are predominantly natural-origin sockeye adults returning to the Wenatchee River captured at Tumwater Dam (Rkm 52.0), where a netpen system is used to house both maturing adults and juveniles prior to release into Lake Wenatchee to over-winter.

Previous genetic studies have generally found a lack of concordance between population genetic relationships and their geographic distributions. These studies indicate that the nearest geographic neighbors of sockeye salmon populations are not necessarily the most genetically similar. Specifically for the Columbia River Basin, sockeye from Lake Wenatchee, Okanogan River, and Redfish Lake may be more closely related to a population from outside the Columbia River (depending on marker used) then to each other.

In this study we investigated the temporal and spatial genetic structure of Lake Wenatchee sockeye collections, without regard to sockeye populations outside of the Lake Wenatchee area. Our primary objective here was to determine if the Wenatchee Sockeye Program affected the natural Lake Wenatchee sockeye population. More specifically, we were tasked to determine if the genetic composition of Lake Wenatchee sockeye population had been altered by a supplementation program that was based on the artificial propagation of a small subset of that population. Using microsatellite DNA allele frequencies, we investigated population differentiation between temporally replicated collections of natural-origin Lake Wenatchee sockeye and program broodstock. We analyzed thirteen collections of Lake Wenatchee sockeye (Table 1), eight temporally replicated collections of natural-origin Lake Wenatchee sockeye ( $\mathrm{N}=786$ ) and five temporally replicated collections of Wenatchee Sockeye Program broodstock ( $\mathrm{N}=248$ ). Paired natural - broodstock collections were available from years 2000, 2001, 2004, 2006, and 2007.

## Conclusions

We observed that allele frequency distributions were consistent over time, irrespective of collection origin, resulting in small and statistically insignificant measures of genetic differentiation among collections. We interpreted these results to indicate no year-to-year differences in allele frequencies among natural-origin or broodstock collections.
Furthermore, there were no observed difference between pre- and post-supplementation collections. Therefore, we accepted our null hypothesis that the allele frequencies of the broodstock collections equaled the allele frequencies of the natural collections, which
equaled the allele frequency of the donor population. Given the small differences in genetic composition among collections, the genetic model for estimating $\mathrm{N}_{\mathrm{e}}$ produced estimates with extremely large variances, preventing the observation of any trend in $\mathrm{N}_{\mathrm{e}}$.

## Introduction

A report titled "Conceptual Approach to Monitoring and Evaluating the Chelan County Public Utility District Hatchery Programs" was prepared July 2005 by Andrew Murdoch and Chuck Peven for the Chelan PUD Habitat Conservation Plan's Hatchery Committee. This report outlined 10 objectives to be applied to various species assessing the impact (positive or negative) of hatchery operations mitigating the operation of Rock Island Dam. This current study pertains only to Lake Wenatchee sockeye and objective 3:

> Determine if genetic diversity, population structure, and effective population size have changed in natural spawning populations as a result of the hatchery program. Additionally, determine if hatchery programs have caused changes in phenotypic characteristics of natural populations.

In order to evaluate cause and effect of hatchery supplementation, WDFW Molecular Genetics Lab surveyed genetic variation of Lake Wenatchee sockeye. The conceptual approach for this project follows that of a parallel study regarding the Wenatchee River spring Chinook supplementation program (Blankenship et al. 2007). We determined the genetic diversity present in the Lake Wenatchee sockeye population by analyzing temporally replicated collections spanning 1989-2007, which included collections from before and following the inception of the Wenatchee Sockeye Program. Documenting the genetic composition of the Lake Wenatchee sockeye population is necessary to assess the effect of the hatchery program on the Lake Wenatchee population. In addition, this work provides a genetic baseline for future projects requiring genetic data. See study objectives below for specific details about how this project addresses Murdoch and Peven (2005) objective 3.

## Lake Wenatchee Sockeye Salmon

Nine spawning populations of sockeye (Oncorhynchus nerka) salmon have been identified in Washington (Washington Department of Fisheries et al. 1993): 1) Baker

River, 2) Ozette Lake, 3) Lake Pleasant, 4) Quinault Lake, and 5) Okanogan River (classified as native stock); 6) Cedar River (classified as non-native stock); 7) Lake Wenatchee, classified as mixed stock); 8) Lake Washington/Lake Sammamish tributaries; and 9) Lake Washington beach spawners (classified as unknown origin). Chapman et al. (1995) listed four additional spawning aggregations of sockeye salmon that appear consistently in Columbia River tributaries: the Methow, Entiat, and Similkameen Rivers; and Icicle Creek in the Wenatchee River drainage.

Located in north central Washington, the Wenatchee River basin drains a portion of the eastern slope of the Cascade Mountains, including high mountainous regions of the Cascade crest. The headwater area of the Wenatchee River is Lake Wenatchee, a typical low productivity oligotrophic or ultra-oligotrophic sockeye salmon nursery lake (Allen and Meekin 1980, Mullan 1986, Chapman et al. 1995). Sockeye salmon bound for Lake Wenatchee enter the Columbia River in April and May and arrive at Lake Wenatchee in late July to early August (Chapman et al. 1995; Washington Department of Fisheries et al. 1993). The run timing of Lake Wenatchee sockeye salmon, classified as an Evolutionary Significant Unit (ESU), appears to have become earlier by 6-30 days during the past 70 years (Chapman et al. 1995; Quinn and Adams 1996). Additionally, scale pattern analysis suggests Wenatchee sockeye migrate past Bonneville Dam earlier than the sockeye bound for the Okanogan River (Fryer and Schwartzberg 1994). The Wenatchee population spawns from mid-September through October in the Little Wenatchee, White, and Napeequa Rivers above Lake Wenatchee (Washington Department of Fisheries et al. 1993), peaking in late September (Chapman et al. 1995). Limited beach spawning is believed to occur in Lake Wenatchee (L. Lavoy pers. com.; Mullan 1986), although Gangmark and Fulton (1952) reported two lakeshore seepage areas in Lake Wenatchee that were used by spawning sockeye salmon. Sockeye salmon fry enter Lake Wenatchee between March and May (Dawson et al. 1973), and typically rear in the lake for one year before leaving as smolts (Gustafson et al. 1997; Peven 1987).

Both the physical properties of the habitat and ecological/biological factors of the sockeye populations differ between the Lake Wenatchee ESU and the geographically
proximate Okanogan ESU. For example: 1) Different limnology is encountered by sockeye salmon in Lakes Wenatchee and Osoyoos; 2) Lake Wenatchee sockeye predominantly return at ages four and five (a near absence of 3-year-olds), where a large percentage of 3-year-olds return to the Okanogan population; and 3) the apparent one month separation in juvenile outmigration-timing between Okanogan- and Wenatcheeorigin fish (Gustafson et al. 1997 and references therein).

## Sockeye Artificial Propagation In Lake Wenatchee

The construction of Grand Coulee Dam completely blocked fish passage to the upper Columbia River, and $85 \%$ of sockeye salmon passing Rock Island Dam between 1935 and 1936 were estimated to be from natural stocks bound for areas up-river to Grand Coulee Dam (Mullan 1986; Washington Department of Fisheries et al. 1938). To compensate for loss of habitat resulting from Grand Coulee Dam, the federal government initiated the Grand Coulee Fish-Maintenance Project (GCFMP) in 1939 to maintain fish runs in the Columbia River above Rock Island Dam. Between 1939 and 1943, all sockeye salmon entering the mid-Columbia River were trapped at Rock Island Dam, and over 32,000 mixed Lake Wenatchee, Okanogan River, and Arrow Lake adult sockeye salmon were released into Lake Wenatchee (Gustafson et al. 1997 Appendix Table D-2). In addition to adult relocation, between 1941 and 1969 over 52.8 million fry descended from original spawners collected at Rock Island and Bonneville Dams, were released into Lake Wenatchee (Gustafson et al. 1997 Appendix Table D-2).

No releases of artificially-reared sockeye salmon occurred in the Wenatchee watershed during the years 1970 to 1989 (Gustafson et al. 1997 Appendix Table D-2). Since 1990, the Wenatchee Sockeye Program has released juveniles into Lake Wenatchee to supplement natural production of sockeye salmon in the basin. Sockeye adults returning to the Wenatchee River are captured at Tumwater Dam (Rkm 52.0) and transferred to Lake Wenatchee net pens until mature. The Wenatchee Sockeye Program goals are 260 adults with an equal sex ratio, $<10 \%$ hatchery-origin returns (identified by coded wire tags), and the adults removed for broodstock account for $<10 \%$ of the run size. Fish are spawned at Lake Wenatchee and their gametes are taken to Rock Island Fish Hatchery

Complex (i.e., Eastbank) for fertilization and incubation. Fry are returned to the Lake Wenatchee net -pens after they are large enough to be coded wire tagged, and are housed in the pens until fall (one year after spawning), when they are liberated into the lake to over-winter. For brood years 1991 - 2004 an average of 218,683 (std. dev. $=71,090$ ) pen-reared Lake Wenatchee-origin juvenile sockeye salmon have been released yearly into Lake Wenatchee.

## Previous Genetic Studies

Protein (allozyme) variation - Surveying genetic variation at 12 allozyme loci, Utter et al. (1984) reported moderate population structure among 16 sockeye collections from southeast Alaska through the Columbia River Basin, including Okanogan and Wenatchee stocks, with an apparent genetic association between upper Fraser River and Columbia River sockeye salmon. Winans et al. (1996) surveyed variation at 55 allozyme loci for 25 sockeye salmon and two kokanee collections from 21 sites in Washington, Idaho, and British Columbia, and reported the lowest level of allozyme variability of any species of Pacific salmon and a highest level of inter-population differentiation. Furthermore, these authors reported that there was no clear relationship between geographic and genetic differentiation among the populations within there study. Other studies corroborate the results of Winans et al. (1996), finding a lack of discernible geographic patterning for sockeye salmon populations in British Columbia, Alaska, and Kamchatka (Varnavskaya et al. 1994, Wood et al. 1994, Wood 1995). These studies indicate that the nearest geographic neighbors of sockeye salmon populations are not necessarily the most genetically similar, which contrasts with the other Pacific salmon species that exhibit concordance between geographic and genetic differentiation (Utter et al. 1989, Winans et al. 1994, Shaklee et al. 1991). As part of the comprehensive status review of west coast sockeye salmon (Gustafson et al. 1997), NMFS biologists collected new allozyme genetic information for 17 sockeye salmon populations and one kokanee population in Washington and combined these data for analysis with the existing Pacific Northwest sockeye salmon and kokanee data from Winans et al. (1996). Results of the updated study were consistent with Winans et al. (1996), with no clear concordance between geographic and genetic distances. Sockeye salmon from Lake Wenatchee, Redfish Lake,

Ozette Lake, and Lake Pleasant are very distinct from other collections in the study, and Columbia River populations were not necessarily most closely related to each other. Gustafson et al. (1997) also examined between-year variability within a collection location and found low levels of statistical significance among the five Lake Wenatchee collections included in the study (For 10 pair-wise comparisons using sum-G test, five were statistically significant). Lake Wenatchee brood year 1987 accounted for three of the significant comparisons, which were driven by unusually high frequencies of two allozyme alleles (ALAT*95 and ALAT*108) (Winans et al. 1996). Nevertheless, Gustafson et al. (1997) conclude that, in general, temporal variation at a locale was considerably less than between-locale variation.

Nucleic acid variation - Beacham et al. (1995) reported levels of variation in nuclear DNA of $O$. nerka using minisatellite probes. They analyzed 10 collections, including a sample from Lake Wenatchee. Cluster analysis showed the Lake Wenatchee sample was different from all the other collections, including those from the Columbia River. Using a similar molecular technique, Thorgaard et al. (1995) examined the use of multi-locus DNA fingerprinting (i.e., banding patterns) to discriminate among 14 sockeye salmon and kokanee populations. Dendrograms based on analysis of banding patterns produced different genetic affinity groups depending on the probes used. While none of the five DNA probes showed a close relationship between Lake Wenatchee and Okanogan River sockeye salmon, if information from all probes were combined, $O$. nerka from Redfish Lake, Wenatchee, and Okanogan were separate from kokanee of Oregon and Idaho and a sockeye salmon sample from the mid-Fraser River.

## Study Objective

We documented temporal variation in genetic diversity (i.e., heterozygosity and allelic diversity), and investigated population differentiation between temporally replicated collections of natural-origin Lake Wenatchee sockeye and program broodstock, using microsatellite DNA allele frequencies. Temporally replicated collections from the same location can also be used to estimate effective population size $\left(\mathrm{N}_{\mathrm{e}}\right)$. If populations are "ideal", the census size of a population is equal to the "genetic size" of the population.

Yet, numerous factors lower the "genetic size" below census, such as, non-equal sex ratios, changes in population size, and variance in the numbers of offspring produced from parent pairs. $\mathrm{N}_{\mathrm{e}}$ is thought to be between 0.10 and 0.33 of the estimated census size (Bartley et al. 1992; RS Waples pers. comm.), although numerous observations differ from this general rule. $\mathrm{N}_{\mathrm{e}}$ can be calculated directly from demographic data, or inferred from observed differences in genetic variance over time. Essentially, when calculated from genetic data, $\mathrm{N}_{\mathrm{e}}$ is the estimated size of an "ideal" population that accounts for the genetic diversity changes observed, irrespective of abundance.

We will address the hypotheses associated with Objective 3 in Murdock and Peven (2005) using the following four specific tasks:

Task 1 - Document the observed genetic diversity.
Task 2 - Test for population differentiation among Lake Wenatchee collections and the associated supplementation program.

Task 2 was designed to address two hypotheses listed as part of Objective 3 in Murdoch and Peven (2005):

- Ho: Allele frequency Hatchery $=$ Allele frequency ${ }_{\text {Naturally produced }}=$ Allele frequency $_{\text {Donor pop }}$.
- Ho: Genetic distance between subpopulations Year $x={\text { Genetic distance between subpopulations }{ }_{\text {Year }} \mathrm{y}}$ Murdoch and Peven (2005) proposed these two hypotheses to help evaluate supplementation programs through a "Conceptual Process" (Figure 5 in Murdoch and Peven 2005). There are two components to the first hypothesis, which must be considered separately for Lake Wenatchee sockeye. The first component involves comparisons between natural-origin populations from Lake Wenatchee to determine if there have been changes in allele frequencies through time starting with the donor population. Documenting a change does not necessarily indicate that the supplementation program has directly affected the natural-origin fish, as additional tests would be necessary to support that hypothesis. The intent of the second component is to determine if the hatchery produced populations have the same genetic composition as the naturally produced populations.

Task 3 - Calculate $\mathrm{N}_{\mathrm{e}}$ using the temporal method for multiple samples from the same location to document trend.

Task 4 - Compare $\mathrm{N}_{\mathrm{e}}$ estimates with trend in census size for Lake Wenatchee sockeye.

## Methods and Materials

## Sampling

Thirteen collections of Lake Wenatchee sockeye were analyzed, eight temporally replicated collections of natural Lake Wenatchee sockeye ( $\mathrm{N}=786$ ) and five temporally replicated collections of Wenatchee Sockeye Program broodstock (N=248) (Table 1). Paired natural - broodstock collections were available from years 2000, 2001, 2004, 2006, and 2007 (Table 1). All collections were made at Tumwater Dam on the Wenatchee River. Note that collections classified as broodstock were predominantly natural-origin sockeye. A majority of the genetic samples were from dried scales. The tissue collections from 2006 and 2007 were fin clips stored immediately in ethanol after collection. DNA was extracted from stored tissue using Nucleospin 96 Tissue following the manufacturer's standard protocol (Macherey-Nagel, Easton, PA, U.S.A.).

## Laboratory Analysis

Polymerase chain reaction (PCR) amplification was performed using 17 fluorescently end-labeled microsatellite marker loci, One 2 (Scribner et al 1996) One 100, 101, 102, 105, 108, 110, 114, and 115 (Olsen et al. 2000), Omm 1130, 1135, 1139, 1142, 1070, and 1085 (Rexroad et al. 2001), Ots 3M (Banks et al. 1999) and Ots 103 (Small et al. 1998). PCR reaction volumes were $10 \mu \mathrm{~L}$, with the reaction variables being $2 \mu \mathrm{~L} 5 \mathrm{x}$ PCR buffer (Promega), $0.6 \mu \mathrm{~L} \mathrm{MgCl}_{2}(1.5 \mathrm{mM})$ (Promega), $0.2 \mu \mathrm{~L} 10 \mathrm{mM}$ dNTP mix (Promega), and $0.1 \mu \mathrm{~L}$ Go Taq DNA polymerase (Promega). Loci were amplified as part of multiplexed sets, so primer molarities and annealing temperatures varied. Multiplex one had an annealing temperature of $55^{\circ} \mathrm{C}$, and used 0.09 Molar (M) One 108, 0.06 M One 110, and 0.11 M One 100. Multiplex two had an annealing temperature of $53^{\circ} \mathrm{C}$, and used 0.08 M One 102, 0.1 M One 114, and 0.05 M One 115. Multiplex three had an annealing temperature of $55^{\circ} \mathrm{C}$, and used 0.08 M One 105 and 0.07 M Ots 103. Multiplex four had
an annealing temperature of $53^{\circ} \mathrm{C}$, and used 0.09 M Omm 1135 and 0.08 M Omm 1139. Multiplex five had an annealing temperature of $60^{\circ} \mathrm{C}$, and used $0.2 \mathrm{M} \mathrm{Omm} 1085,0.09 \mathrm{M}$ Omm 1070, and 0.05 M Ots 3M. Multiplex six had an annealing temperature of $48^{\circ} \mathrm{C}$, and used 0.06 M One 2, $0.08 \mathrm{M} \mathrm{Omm} \mathrm{1142}$,and 0.08 M Omm 1130. One 101 was run in isolation with a primer molarity of 0.06 . Thermal cycling was conducted on either PTC200 (MJ Research) or GeneAmp 9700 thermal cyclers as follows: $94^{\circ} \mathrm{C}$ ( 2 min ); 30 cycles of $94^{\circ} \mathrm{C}$ for 15 sec ., 30 sec . annealing, and $72^{\circ} \mathrm{C}$ for 1 min .; a final $72^{\circ} \mathrm{C}$ extension and then a $10^{\circ} \mathrm{C}$ hold. PCR products were visualized by denaturing polyacrylamide gel electrophoresis on an ABI 3730 automated capillary analyzer (Applied Biosystems). Fragment analysis was completed using GeneMapper 3.7 (Applied Biosystems).

## Genetic data analysis

Assessing within collection genetic diversity - Heterozygosity measurements were reported using Nei's (1987) unbiased gene diversity formula (i.e., expected heterozygosity) and Hedrick's (1983) formula for observed heterozygosity. Both tests were implemented using the microsatellite toolkit (Park 2001). For each locus and collection FSTAT version 2.9.3.2 (Goudet 1995) was used to assess Hardy-Weinberg equilibrium, where deviations from the neutral expectation of random associations among alleles were calculated using a randomization procedure. Alleles were randomized among individuals within collections (4160 randomizations for this dataset) and the $\mathrm{F}_{\text {IS }}$ (Weir and Cockerham 1984) calculated for the randomized datasets were compared to the observed $\mathrm{F}_{\text {IS }}$ to obtain an unbiased estimation of the probability that the null hypothesis was true. The 5\% nominal level of statistical significance was adjusted for multiple tests (Rice 1989). Genotypic linkage disequilibrium was calculated following Weir (1979) using GENETIX version 4.05 (Belkhir et al. 1996). Statistical significance of linkage disequilibrium results was assessed using a permutation procedure implemented in GENETIX for each locus by locus combination within each collection.

Assessing among collection genetic differentiation - The temporal stability of allele frequencies was assessed by the randomization chi-square test implemented in FSTAT version 2.9.3.2 (Goudet 1995). Multi-locus genotypes were randomized between
collections. The G-statistic for observed data was compared to G-statistic distributions from randomized datasets (i.e., null distribution of no differentiation between collections). Population differentiation was also investigated using pairwise estimates of $\mathrm{F}_{\text {ST }}$. Multi-locus estimates of pairwise $\mathrm{F}_{\mathrm{ST}}$, estimated by a "weighted" analysis of variance (Weir and Cockerham, 1984), were calculated using GENETIX version 4.05 (Belkhir et al.1996). $\mathrm{F}_{\text {ST }}$ was used to quantify population structure, the deviation from statistical expectations (i.e., excess homozygosity) due to non-random mating between populations. To determine if the observed $\mathrm{F}_{\mathrm{ST}}$ estimate was consistent with statistically expectations of no population structure, a permutation test was implemented in GENETIX (1000 permutations).

Effective population size $\left(\mathbf{N}_{\mathbf{e}}\right)$ - Estimates of the effective population size were obtained using a multi-collection temporal method (Waples 1990a). The temporal method assumes that cohorts are used, but we did not decompose the collection year samples into their respective cohorts using age data. Therefore, $\mathrm{N}_{\mathrm{e}}$ estimates that pertain to individual year classes of breeders are not valid; however the harmonic mean over all samples will estimate an $\mathrm{N}_{\mathrm{e}}$ that pertains to the time period from which the collections are derived. Comparing samples from years $i$ and $j$, Waples’ (1990a) temporal method estimates the effective number of breeders ( $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ ) according to:

$$
\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}=\frac{\mathrm{b}}{2\left(\hat{\mathrm{~F}}-1 / \widetilde{\mathrm{S}}_{\mathrm{i}, \mathrm{j}}\right)}
$$

The standardized variance in allele frequency ( $\hat{\mathrm{F}}$ ) is calculated according to Pollack (1983). The parameter b is calculated analytically from age structure information and the number of years between samples (Tajima 1992). The age-at-maturity information required to calculate b was obtained from ecological data (Hillman et al. 2007). The harmonic mean of sample sizes from years $i$ and $j$ is $\tilde{S}_{\mathrm{i}, \mathrm{j}}$. The harmonic mean over all pairwise estimates of $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ is $\tilde{\mathrm{N}}_{\mathrm{b}}$. SALMONNb (Waples et al. 2007) was used to calculate $\tilde{\mathrm{N}}_{\mathrm{b}}$.

## Results and Discussion

In this section we combine our presentation and interpretations of the genetic analyses. Additionally, this section is organized based on the task list presented in the study plan.

Task 1 - Document the observed genetic diversity.

Substantial genetic diversity was observed over all Lake Wenatchee sockeye collections analyzed (Table 1), with heterozygosity estimates over all loci having a mean of 0.79 . Genetic diversity was consistent with expected Hardy-Weinberg random mating genotypic proportions for all collections. The $\mathrm{F}_{\text {IS }}$ observed for each collection was not statistically significant given the distribution of $\mathrm{F}_{\text {IS }}$ generated using a randomization procedure. Additionally, there were no statistically significant associations observed between alleles across loci (i.e., linkage equilibrium) (data not shown). We concluded from these results that the genetic data from each collection was consistent with statistical expectations for random association of alleles within and between loci. In other words, each collection represents samples from a single gene pool (i.e., populations), and the genetic diversity observed has no detectable technical artifacts or evidence of natural selection.

Task 2 - Test for differentiation among Lake Wenatchee collections and the associated supplementation program.

We explicitly tested the hypothesis of no significant differentiation within natural-origin or broodstock collections from Lake Wenatchee using a randomization chi-square test. The null hypothesis for these tests was that the allele frequencies from two different populations were drawn from the same underlying distribution. We show the results for the pairwise comparisons among eight temporally replicated natural-origin collections from Lake Wenatchee ( 28 pairwise tests), and report all tests were non-significant (Table 2A). Similarly, for five temporally replicated broodstock collections, 10 of 10 pairwise tests were non-significant (Table 2B). We also tested if natural-origin and broodstock
collections were differentiated from each other over time, and report that 40 of 40 tests were non-significant (Table 2C). The nominal level of statistical significance ( $\alpha=0.05$ ) was adjusted for multiple comparisons using strict Bonferroni correction (Rice 1989). Yet, there are perhaps slight differences between paired natural-broodstock collections. Note that the p-values for comparisons regarding 2006 and 2007 paired collections are lower than for comparisons regarding 2000, 2001, and 2004. The small sample sizes for broodstock collections in 2006 and 2007 may not have been random samples from the Lake Wenatchee sockeye population.

Given the consistencies observed for allele frequency distributions over time, metrics of population structure were expected to be small. This was the case, as the estimated $\mathrm{F}_{\text {ST }}$ over all thirteen collections was 0.0003 . This observed value fell within the distribution of $\mathrm{F}_{\mathrm{ST}}$ values expected if there were no population structure present (permutation test pvalue 0.12). Analysis of the paired natural-broodstock collections corroborated this result. Pairwise estimates of $\mathrm{F}_{\text {ST }}$ were 0.000 for years 2000, 2001, 2004, and 2007, and 0.002 for 2006. All five estimates were non-significant. Essentially, all 13 sockeye collections could be considered samples from the same population. Given these results, it is valid to combine all collections for statistical analysis. Therefore, we did not calculate genetic distances among any collections, as it is inappropriate to estimate distances that are effectively zero.

## Conclusions

We interpret these data to indicate that there appears to be no significant year-to-year differences in allele frequencies among natural-origin or broodstock collections, nor are there observed differences between collections pre- and post-supplementation. As a result, we accept the null hypothesis that the allele frequencies of the broodstock collections equal the allele frequencies of the natural collections, which equals the allele frequency of the donor population. Furthermore, the observed genetic variance that can be attributed to among collection differences was negligible.

Task 3 - Calculate $\mathrm{N}_{\mathrm{e}}$ using the temporal method for multiple samples from the same location to document trend.

The fundamental parameter for inferring $\mathrm{N}_{\mathrm{e}}$ using genetic data is the standardized variance in allele frequency ( $\hat{\mathrm{F}}$ ) (Pollack 1983). Methods estimate $\mathrm{N}_{\mathrm{e}}$ from observed changes in $\hat{F}$ over temporally replicated collections from the same location. Yet, as previously shown, there were no statistically significant differences detected in allele frequencies. The underlying model for estimating $\mathrm{N}_{\mathrm{e}}$ produced estimates with extremely large variances, given small temporal differences in $\hat{F}$, which rendered any trend in $\mathrm{N}_{\mathrm{e}}$ unobservable. Table 3 shows $\mathrm{N}_{\mathrm{e}}$ estimates calculated using temporally replicated natural collections.

Task 4-Compare $\mathrm{N}_{\mathrm{e}}$ estimates with trend in census size for Lake Wenatchee sockeye.

See Task 3

## Acknowledgements

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## Literature Cited

Allen RL and Meekin TK (1980) Columbia River sockeye salmon study, 1971-1974. Wash. Dep. Fish. Prog. Rep. 120, 75 p.

Banks MA, Blouin MS, Baldwin BA, Rashbrook VK, Fitzgerald HA, Blankenship SM, Hedgecock D (1999) Isolation and inheritance of novel microsatellites in chinook salmon (Oncorhynchus tschawytscha). Journal of Heredity, 90:281-288.

Bartley D, Bentley B, Brodziak J, Gomulkiewicz R, Mangel M, and Gall GAE (1992) Geographic variation in population genetic structure of chinook salmon from California and Oregon. Fish. Bull., U.S. 90:77-100.

Blankenship SM, Von Bargen J, Warheit KI, and Murdoch AR (2007) Assessing the Genetic Diversity of Natural Chiwawa River Spring Chinook Salmon and Evaluating the Effectiveness of its Supportive Hatchery Supplementation Program. WDFW report to Chelan County PUD, March 2007.

Belkhir K, Borsa P, Chikhi Let al (1996) GENETIX, logiciel sous Windows TM pour la Génétique des populations. Laboratoire Génome, Populations, Interactions, CNRS UMR 5000, Université de Montpellier II, Montpellier (France).

Chapman D, Peven C, Giorgi A, Hillman T, and Utter F (1995) Status of spring chinook salmon in the mid-Columbia River. Don Chapman Consultants, Inc., 477 p. (Available from Don Chapman Consultants, 3653 Rickenbacker, Ste. 200, Boise, ID 83705.)

Crawford BA (1979) The origin and history of trout brood stocks of the Washington Department of Game. Wash. State Game Dep., Fish. Res. Rep. 76 p.

Dawson JJ, Thorne RE, and Traynor JJ (1973) Acoustic surveys of Lake Wenatchee and Lake Osoyoos in 1973. Final Report, Service Contr. 526, to Wash. Dep. Fish., by Fish. Res. Inst., Coll. of Fish., Univ. Washington, Seattle, WA, 18 p.

Fryer JK and Schwartzberg M (1994) Identification of Columbia Basin sockeye salmon stocks using scale pattern analyses in 1993. Columbia River Inter-Tribal Fish Commission, Tech. Rep. 94-2, 39 p.

Gangmark HA and Fulton LA (1952) Status of Columbia blueback salmon runs, 1951. U. S. Fish Wildl. Serv. Spec. Sci. Rep. 74, 29 p.

Goudet J (1995) FSTAT (Version 1.2): A computer program to calculate F-statistics. Journal of Heredity 86: 485-486.

Gustafson RG, Wainwright TC, Winans GA, Waknitz FW, Parker LT, and Waples RS (1997) Status review of sockeye salmon from Washington and Oregon. U.S. Dept. Commer., NOAA Tech. Memo. NMFS-NWFSC-33, 282 p.

Hedrick PW (1983) Genetics of Populations, Science Books International, Boston
Hillman T, Miller M, Peven C, Tonseth M, Miller T, Truscott K, and Murdoch A (2007) Monitoring and Evaluation of the Chelan County PUD Hatchery Programs: 2007 Annual Report.

Knutzen D (1995) Letter to R. Gustafson, NMFS, from D. Knutzen, WDFW, re. Historical kokanee planting records for Lake Wenatchee, Lake Pleasant, Lake Ozette, Lake Shannon, and Baker Lake from 1981-1994, dated 17 July 1995. 1 p. plus attachment. (Available from West Coast Sockeye Salmon Administrative Record, Environmental and Technical Services Division, Natl. Mar. Fish. Serv., 525 N. E. Oregon Street, Portland, OR 97232.)

Mullan JW (1986) Determinants of sockeye salmon abundance in the Columbia River, 1880's-1982: A review and synthesis. U.S. Fish Wildl. Serv. Biol. Rep. 86(12), 135 p.

Murdoch AR and Peven C (2005) Conceptual Approach to Monitoring and Evaluating the Chelan County Public Utility District Hatchery Programs, Final Report.

Nei M (1978) Estimation of average heterozygosity and genetic distance from a small number of individuals. Genetics 89:583-590.

Olsen JB, Wilson SL, Kretschmer EJ, Jones KC, Seeb JE (2000) Characterization of 14 tetranucleotide microsatellite loci derived from sockeye salmon. Molecular Ecology 9, 2185-2187.

Park SDE (2001) Trypanotolerance in West African Cattle and the Population Genetic Effects of Selection [ Ph.D. thesis], University of Dublin

Peven CM (1987) Downstream migration timing of two stocks of sockeye salmon on the Mid-Columbia River. Northwest Sci. 61(3):186-190.

Pollak E (1983) A new method for estimating the effective population size from allele frequency changes. Genetics 104, 531-548.

Quinn TP and Adams DJ (1996) Environmental changes affecting the migratory timing of American shad and sockeye salmon. Ecology 77:1151-1162.

Rexroad CE, Coleman RL, Martin AM, Hershberger WK, Killefer J (2001) Thirty-five polymorphic microsatellite markers for rainbow trout (Oncorhynchus mykiss). Animal Genetics, 32:317-319.

Rice WR, (1989) Analyzing tables of statistical tests. Evolution. 43:223-225.
Scribner KT, Gust JR, Fields RL (1996) Isolation and characterization of novel salmon microsatellite loci: cross-species amplification and population genetic applications. Candian Journal of Fisheries and Aquatic Sciences. 53, 833-841.

Shaklee JB, Klaybor DC, Young S, and White BA (1991) Genetic stock structure of oddyear pink salmon, Oncorhynchus gorbuscha (Walbaum), from Washington and British Columbia and potential mixed-stock fisheries applications. J. Fish. Biol. 39(A):21-34.

Small MP, Beacham TD, Withler RE, and Nelson RJ (1998) Discriminating coho salmon (Oncorhynchus kisutch) populations within the Fraser River, British Columbia. Molecular Ecology 7: 141-155.

Tajima F (1992) Statistical Method for Estimating the Effective Population Size in Pacific Salmon. J Hered 83, 309-311.

Utter F, Aebersold P, Helle J, and Winans G (1984) Genetic characterization of populations in the southeastern range of sockeye salmon. In J. M. Walton and D. B. Houston (editors), Proceedings of the Olympic wild fish conference, p. 17-31. Fisheries Technology Program, Peninsula College, Port Angeles, WA.

Utter F, Milner G, Stahl G, and Teel D (1989) Genetic population structure of chinook salmon, Oncorhynchus tshawytscha, in the Pacific Northwest. Fish. Bull., U.S. 87:239-264.

Varnavskaya NV, Wood CC, and Everett RJ (1994) Genetic variation in sockeye salmon (Oncorhynchus nerka) populations of Asia and North America. Can. J. Fish. Aquat. Sci. 51(Suppl. 1):132-146.

Waples RS (1990a) Conservation genetics of Pacific salmon. III. Estimating effective population size. Journal of Heredity 81:277-289

Waples RS, Masuda M, Pella J (2007) SALMONNb: a program for computing cohortspecific effective population sizes $\left(\mathrm{N}_{\mathrm{b}}\right)$ in Pacific salmon and other semelparous species using the temporal method. Molecular Ecology Notes 7, 21-24.

Washington Department of Fisheries (WDF), Washington Department of Game (WDG), and United States Bureau of Fisheries (USBF) (1938) A report on the preliminary investigations into the possible methods of preserving the Columbia River salmon and steelhead at the Grand Coulee Dam. Wash. Dep. Fish, Olympia, WA, 120 p.

Washington Department of Fisheries (WDF), Washington Department of Wildlife (WDW), and Western Washington Treaty Indian Tribes (WWTIT) (1993) 1992 Washington State salmon and steelhead stock inventory (SASSI). Wash. Dep. Fish Wildl., Olympia, WA, 212 p. plus 5 regional volumes.

Washington Department of Fish and Wildlife (WDFW) (1996) Letter to M. Schiewe, NMFS, from R. Lincoln, Assistant Director, Fish Management Program, Washington Department of Fish and Wildlife, dated 12 July 1996. 3 p. plus appendix. (Available from West Coast Sockeye Salmon Administrative Record, Environmental and Technical Services Division, Natl. Mar. Fish. Serv., 525 N. E. Oregon Street, Portland, OR 97232. )

Weir BS (1979) Inferences about linkage disequilibrium. Biometrics 35:235-254.
Weir BS, Cockerham CC (1984) Estimating F-Statistics for the Analysis of PopulationStructure. Evolution 38:1358-1370.

Winans GA, Aebersold PB, Urawa S, and Varnavskaya NV (1994) Determining continent of origin of chum salmon (Oncorhynchus keta) using genetic stock identification techniques: status of allozyme baseline in Asia. Can. J. Fish. Aquat. Sci. 51 (Suppl. 1):95-113.

Winans GA, Aebersold PB, and Waples RS (1996) Allozyme variability of Oncorhynchus nerka in the Pacific Northwest, with special consideration to populations of Redfish Lake, Idaho. Trans. Am. Fish. Soc. 205:645-663.

Wood CC, Riddell BE, Rutherford DT, and Withler RE (1994) Biochemical genetic survey of sockeye salmon (Oncorhynchus nerka) in Canada. Can. J. Fish. Aquat. Sci. 51(Suppl. 1):114-131.

Wood CC (1995) Life history variation and population structure in sockeye salmon. In J. L. Nielsen (editor), Evolution and the aquatic ecosystem: defining unique units in population conservation. Am. Fish. Soc. Symp. 17:195-216.

Table 1 Lake Wenatchee sockeye collections analyzed. MNA is the mean number of alleles per locus, Hz is unbiased heterozygosity, Obs Hz is observed heterozygosity, and HW is the p-value of the null hypothesis of random association of alleles (i.e., Hardy - Weinberg equilibrium). For reference, the nominal level of statistical significance at $\alpha=0.05$ is 0.0002 after correction for multiple tests.

|  | Collection <br> Code | Tissue <br> Type | Source | N | MNA | Hz | Obs Hz | HW |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1989 | $89^{1}$ | Scales | Natural | 96 | 14.35 | 0.792 | 0.791 | 0.424 |
| 1990 | $90^{1}$ | Scales | Natural | 96 | 13.19 | 0.793 | 0.779 | 0.131 |
| 2000 | $00 A A E$ | Scales | Broodstock | 96 | 12.31 | 0.787 | 0.776 | 0.213 |
| 2000 | $00^{1}$ | Scales | Natural | 96 | 11.76 | 0.801 | 0.826 | 0.868 |
| 2001 | 01 AAS | Scales | Broodstock | 53 | 9.47 | 0.788 | 0.793 | 0.392 |
| 2001 | $01^{1}$ | Scales | Natural | 96 | 14.35 | 0.786 | 0.794 | 0.456 |
| 2002 | $02^{1}$ | Scales | Natural | 96 | 14.53 | 0.794 | 0.777 | 0.780 |
| 2004 | $04^{1}$ | Scales | Natural | 96 | 14.65 | 0.798 | 0.803 | 0.704 |
| 2004 | 04 AAV | Scales | Broodstock | 43 | 14.35 | 0.796 | 0.795 | 0.051 |
| 2006 | 06 CN | Tissue | Broodstock | 38 | 14.59 | 0.793 | 0.785 | 0.688 |
| 2006 | 06 CO | Tissue | Natural | 96 | 14.53 | 0.806 | 0.803 | 0.408 |
| 2007 | 07 EE | Tissue | Broodstock | 18 | 14.00 | 0.790 | 0.790 | 0.221 |
| 2007 | $07 E F$ | Tissue | Natural | 96 | 14.35 | 0.789 | 0.800 | 0.347 |

[^16]Table 2 Allelic differentiation for Lake Wenatchee sockeye collections. A single analysis tested (pairwise) the allelic differentiation between all thirteen collections; however p -values for G -statistics are partitioned in the table by A) natural-origin, B) broodstock, and C) natural versus broodstock. Underlined values are for paired naturalbroodstock collections from the same year. For reference, the nominal level of statistical significance at $\alpha=0.05$ is 0.0006 after correction for multiple tests. No significant values were observed.
A) Natural-Origin Collections

|  | 89 | 90 | 00 | 01 | 02 | 04 | 06 CO | 07 EF |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 89 |  | 0.257 | 0.359 | 0.531 | 0.331 | 0.127 | 0.031 | 0.263 |
| 90 |  | 0.953 | 0.148 | 0.753 | 0.903 | 0.077 | 0.283 |  |
| 00 |  |  |  | 0.328 | 0.527 | 0.607 | 0.604 | 0.400 |
| 01 |  |  |  |  | 0.209 | 0.081 | 0.127 | 0.093 |
| 02 |  |  |  |  |  | 0.085 | 0.707 | 0.235 |
| 04 |  |  |  |  |  | 0.312 | 0.577 |  |
| 06 CO |  |  |  |  |  | 0.435 |  |  |
| 07 EF |  |  |  |  |  |  |  |  |

B) Broodstock Collections

|  | 00AAE | 01AAS | 04AAV | 06 CN | 07EE |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 00AAE |  | 0.189 | 0.090 | 0.008 | 0.058 |
| 01AAS |  |  | 0.122 | 0.020 | 0.116 |
| 04AAV |  |  |  | 0.008 | 0.031 |
| 06CN |  |  |  |  | 0.326 |
| 07EE |  |  |  |  |  |

C) Natural vs. Broodstock

|  | 89 | 90 | 00 | 01 | 02 | 04 | 06 CO | 07 EF |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 00AAE | 0.027 | 0.309 | $\underline{0.572}$ | 0.018 | 0.041 | 0.012 | 0.093 | 0.040 |
| 01AAS | 0.115 | 0.471 | 0.160 | $\underline{0.219}$ | 0.519 | 0.049 | 0.654 | 0.133 |
| 04AAV | 0.136 | 0.219 | 0.210 | 0.423 | 0.208 | $\underline{0.328}$ | 0.037 | 0.153 |
| 06CN | 0.029 | 0.004 | 0.053 | 0.007 | 0.022 | 0.004 | $\underline{0.019}$ | 0.001 |
| 07EE | 0.099 | 0.229 | 0.053 | 0.015 | 0.093 | 0.178 | 0.090 | $\underline{0.037}$ |

Table 3 Estimation of $\mathrm{N}_{\mathrm{e}}$ for temporally replicated natural-original sockeye collections. Above the diagonal are pairwise estimates of $\mathrm{N}_{\mathrm{e}}$, where negative values mean sampling variance can account for genetic variance observed (i.e., genetic drift unnecessary).
Below the diagonal are variances for pairwise estimates of $\mathrm{N}_{\mathrm{e}}$. Absent variance values (denoted by - ) were too large for SalmonNb to display.

|  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |
| Collection | 89 | 90 | 00 | 01 | 02 | 04 | 06 CO | 07 EF |
| 89 |  | -3936.6 | -1414 | -2636.3 | 671.4 | 1871.1 | 1066.1 | 1951.2 |
| 90 | $2.59 \mathrm{E}+09$ |  | -1490.3 | 3649.1 | -31144 | -6808.4 | 817.6 | 93190.2 |
| 00 | $1.40 \mathrm{E}+09$ | $4.45 \mathrm{E}+09$ |  | -592.2 | -6842.2 | -667.1 | -1736.9 | -1350.1 |
| 01 | $1.21 \mathrm{E}+09$ | $1.47 \mathrm{E}+09$ | $2.33 \mathrm{E}+09$ |  | 977.1 | 6160.4 | 387.8 | 2531.5 |
| 02 | $1.91 \mathrm{E}+09$ | $1.33 \mathrm{E}+09$ | $1.16 \mathrm{E}+09$ | $2.29 \mathrm{E}+09$ |  | 1495.6 | -848.5 | 3213.6 |
| 04 | $2.21 \mathrm{E}+09$ | $3.62 \mathrm{E}+09$ | $4.08 \mathrm{E}+09$ | $1.27 \mathrm{E}+09$ | $1.14 \mathrm{E}+09$ |  | 896.6 | 2155.3 |
| 06 CO | $1.34 \mathrm{E}+09$ | $1.39 \mathrm{E}+09$ | $1.73 \mathrm{E}+09$ | - | $4.51 \mathrm{E}+09$ | $1.2 \mathrm{E}+09$ |  | 3278.6 |
| 07 EF | $2.15 \mathrm{E}+09$ | $1.51 \mathrm{E}+09$ | $1.18 \mathrm{E}+09$ | $1.68 \mathrm{E}+09$ | - | $1.36 \mathrm{E}+09$ | $2.65 \mathrm{E}+09$ |  |

## Appendix J

Genetic Diversity of Chiwawa River Spring Chinook Salmon

# Assessing the Genetic Diversity of Natural Chiwawa River Spring Chinook Salmon and Evaluating the Effectiveness of its Supportive Hatchery Supplementation Program 

Developed for<br>Chelan County PUD<br>and the<br>Habitat Conservation Plan's Hatchery Committee

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## Executive Summary

The main objective of this study was to determine the potential impacts of the Chiwawa River Supplementation Program on natural spring Chinook in the upper Wenatchee system. We did this by investigating population differentiation between temporally replicated Chiwawa River natural and hatchery samples from the Wenatchee River watershed using microsatellite DNA allele frequencies and the statistical assignment of individual fish to specific populations. Additionally, to assess the genetic effect of the hatchery program, we investigated the relationship between census and effective population sizes using collections obtained before and after the supplementation program. In this summary, we briefly describe the salient results contained within this report; however, each "Task" within the Results/Discussion section below contains extended coverage for each topic along with an expanded interpretation of each result.

Overall, we observed substantial genetic diversity within collections, with heterozygosities equal to roughly $80 \%$, over thirteen microsatellite markers. Microsatellite allele frequencies among temporally replicated collections from the same population (i.e., location) were variable, resulting in significant genetic differentiation among these collections. However, these difference are likely the result of salmon life history in this area, as four-year-old Chinook comprise a majority of returns each year. That is, the genetic tests are detecting the differences of contributing parents from each cohort, rather than a hatchery effect.

## Analysis of Chiwawa River Collections

To assess the multiple competing hypotheses regarding population differentiation within and among Chiwawa River collections, we found it necessary to organized the Chiwawa genetic data into three data sets: (1) fish origin (hatchery versus natural), (2) spawning location (hatchery broodstock versus in-river (natural) spawners), and (3) four "treatment" groups (1. hatchery-origin hatchery broodstock, 2. hatchery-origin natural spawner, 3. natural-origin natural spawner, and 4. natural-origin hatchery broodstock). We conducted separate analyses using each of the three data sets, with each analysis
touching on some aspect of the components necessary to move through the Conceptual Process outlined by Murdoch and Peven (2005).

Origin Dataset - We report that allele frequencies within and between natural- and hatchery-origin collections are significantly different, but there does not appear to be a robust signal indicating that the recent natural-origin collections have diverged greatly from the pre- or early post-supplementation collections. Genetic drift will occur in all populations, but does not appear to be a major factor affecting allele frequencies within the Chiwawa collections.

Spawning Location Dataset - There are significant allele frequency differences within and between hatchery broodstock and natural spawner collections. However, in recent years the allele frequency differences between the hatchery broodstock and natural spawner collections have declined. Furthermore, based on linkage disequilibrium, there is a genetic signal that is consistent with increasing homogenization of allele frequencies within hatchery broodstock collections, but a similar homogenization within the natural spawner collection is not apparent. These data suggest that there exists consistent year-to-year variation in allele frequencies among hatchery and natural spawning collections, but there is a trend toward homogenization of the allele frequencies of the natural- and hatchery-origin fish that compose the hatchery broodstock.

Four Treatment dataset - Although there are signals of allelic differentiation among Chiwawa River collections, there are no robust signs that these collections are substantially different from each other. We used two different analyses to measure the degree of genetic variation that exists among individuals and collections within the Chiwawa River. First, we conducted a principal component analysis using all Chiwawa samples with complete genotypes (i.e., no missing alleles from any locus). Although the first two principal component axes account for only $10.5 \%$ of the total molecular variance, a substantially greater portion of that variance is among individual fish, regardless of their identity, rather than among hatchery and natural collections. The
variances in principal component scores among individuals are 11 and 13 times greater than the variance in scores among collections.

Secondly, using an Analysis of Molecular Variance (AMOVA), we were able to determine how best to group populations, with "best" being defined as that grouping that accounts for the greatest proportion of among group (i.e., population) variance. Furthermore, by partitioning molecular variance into different hierarchical components, we are able to determine what level accounts for the majority of the molecular variance. The AMOVA results clearly show that nearly all molecular variation, no matter how the data are organized, resides within a collection. The percentage of total molecular variance occurring within collections ranged from $99.68 \%$ to $99.74 \%$. These results indicate that the significant differences among collections of Chiwawa fish account for less than one percent of the total molecular variance, and these differences cannot be attributed to fish origin or spawning location.

## Effective Population Size ( $N_{e}$ )

The contemporary estimate of $\mathrm{N}_{\mathrm{e}}$ calculated using genetic data combined for Chiwawa natural-origin spawners (NOS) and hatchery-origin spawners (HOS) Chinook is $\mathrm{N}_{\mathrm{e}}=386.8$, which is slightly larger than the pre-hatchery $\mathrm{N}_{\mathrm{e}}$ we estimated using demographic data from 1989 - 1992. Additionally, the $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratio calculated using 386.8 for $\mathrm{N}_{\mathrm{e}}$ and the arithmetic mean yearly census of NOS and HOS Chinook from 1989 2005 for N is 0.40 . These results suggest the $\mathrm{N}_{\mathrm{e}}$ has not declined during the period of Chiwawa Hatchery Supplementation Program operation.

## Analysis Of Upper Wenatchee Tributary Collections

We compared genetic data for spring Chinook collected from the major spawning aggregates of the Wenatchee River. We observed significant differences in allele frequencies among temporally replicated collections within populations, and among populations within the upper Wenatchee. However, these differences account for a very small portion of the overall molecular variance, and these populations overall are very similar to each other. Of all the populations within the Wenatchee River, the White River
appears to be the most distinct. Yet, this distinction is more a matter of detail than of large significance, as the median $\mathrm{F}_{\text {ST }}$ between White River collections and all other collections (except the Little Wenatchee collection; see Results/Discussion) is less than $1.5 \%$ among population variance. We consider the implications of these results in the Conclusion section that follows the Results/Discussion section. Additionally, there is no evidence that the Chiwawa River Supplementation Program has changed the allele frequencies in the Nason Creek and White River populations, despite the presence of hatchery-origin fish in both these systems.

## Introduction

Murdoch and Peven (2005) outlined 10 objectives to assess the impact (positive or negative) of hatchery operations mitigating the operation of Rock Island Dam. Two objectives relate to monitoring the genetic integrity of populations:

Objective 3: Determine if genetic diversity, population structure, and effective population size have changed in natural spawning populations as a result of the hatchery program. Additionally, determine if hatchery programs have caused changes in phenotypic characteristics of natural populations.

Objective 5: Determine if the stray rate of hatchery fish is below the acceptable levels to maintain genetic variation between stocks.

This study addresses Objective 3 (above), and documents analyses and results WDFW completed for populations of spring Chinook (Oncorhynchus tshawytscha) in the Wenatchee River watershed. This study was not intended to specifically address Objective 5 (above); however, genetic data provide results relevant to Objective 5. The critical component of Objective 3 is to determine if hatchery supplementation has effected change. Furthermore, change in this context means altering census size and/or genetic marker allele frequencies; we did not attempt to measure changes in fitness. Perhaps a more meaningful rewording of Objective 3 is, "Did the hatchery supplementation program succeed at increasing the census size of a target population while leaving genetic integrity intact?" In order to evaluate cause and effect of hatchery supplementation, we surveyed and compared genetic variation in samples collected before and after potential effects from the Chiwawa Hatchery Supplementation Program. Samples were acquired from the primary spawning aggregates in the upper Wenatchee River watershed: Nason Creek, Little Wenatchee River, White River, and Chiwawa River. Hatchery samples were acquired from programs that could potentially affect genetic composition of Wenatchee stocks, the integrated Chiwawa River stock (local stock), Leavenworth National Fish Hatchery spring Chinook (Carson Stock - non local), and Entiat NFH (Carson Stock - non local). Additionally, the genetic markers used were the Genetic Analysis of Pacific Salmonids (GAPS) (Seeb et al. in review) standardized
microsatellites, so all data from the Wenatchee study will be available for inclusion in the GAPS Chinook coastwide microsatellite baseline.

## History of Artificial Propagation

Artificial propagation in the upper Columbia River began in 1899 when hatcheries were constructed on the Wenatchee and Methow rivers (Mullan 1987). These initial operations were small, with the Tumwater Hatchery on the Wenatchee River releasing several hundred thousand fry, and the Methow River hatchery producing few Chinook salmon before it was closed in 1913 (Craig and Suomela 1941, Nelson and Bodle 1990). The Leavenworth State Hatchery operated in the Wenatchee River Basin between 1913 and 1931 using eggs from non-native stocks (Willamette River spring-run and lower Columbia Chinook hatchery fall-run). These early attempts at hatchery production were largely unsuccessful for spring-run Chinook (WDF 1934). Between 1931 and 1939, no Chinook salmon hatcheries were in operation above Rock Island Dam (Rkm 730).

In 1938, the last salmon was allowed to pass upstream through the uncompleted Grand Coulee Dam (Rkm 959). To mitigate the loss of habitat, adult Chinook salmon were trapped, under the auspices of the Grand Coulee Fish Maintenance Project (GCFMP), at Rock Island Dam beginning in May 1939, and relocated into three of the remaining accessible tributaries to the upper Columbia River: the Wenatchee, Entiat, and Methow Rivers. GCFMP transfers continued through the autumn of 1943. Spring- and summer/fall-run fish were differentiated at Rock Island Dam based on a 9 July cutoff date for Chinook arrivals at Rock Island Dam (Fish and Hanavan 1948). Spring-run adults collected at Rock Island Dam (pre 9 July fish) were either transported to Nason Creek on the Wenatchee River to spawn naturally (1939-43), or to the newly constructed Leavenworth NFH (1940) for holding and subsequent spawning (1940-43). Eggs were incubated on site or transferred to the Entiat NFH (1941) and Winthrop NFH (1941). In 1944 spring-run adults were allowed to freely pass Rock Island Dam. The GCFMP did not differentiate among late-run stocks (post 9 July fish) passing Rock Island Dam. Laterun offspring reared at the Leavenworth NFH, Entiat NFH, and Winthrop NFHs were an
amalgamation of summer and fall upper Columbia River populations (Fish and Hanavan 1948). Late-run fish were transplanted into the upper and lower Wenatchee, Methow, and Entiat Rivers.

After 1943, the Winthrop NFH continued to use local spring-run Chinook for hatchery production, while the other NFHs largely focused on summer-run Chinook salmon. Renewed emphasis on spring run production in the mid-1970s saw the inclusion of local and non-local eggs (Carson NFH stock, Klickitat River stock, and Cowlitz River stock) to the NFHs. In the early 1980s, imports of non-native eggs were reduced significantly, and thereafter the Leavenworth, Entiat, and Winthrop NFHs have relied on adults returning to their facilities for their egg needs (Chapman et al. 1995). Regarding late-run Chinook, due to the variety of methods employed to collect broodstock at dams, hatcheries, or the result of juvenile introductions into various areas, Chinook populations and runs (i.e., summer and fall) have been mixed considerably in the upper Columbia system over the past five decades (reviewed in Chapman et al. 1994).

Washington Department of Fish and Wildlife (WDFW) operates two facilities producing spring-run Chinook, the Methow Fish Hatchery (MFH) owned by Douglas County PUD that began operation in 1992 and Eastbank Fish Hatchery (EFH) owned by Chelan County PUD that began operation in 1989. Both programs were designed to implement supplementation (supportive breeding) programs for naturally spawning populations on the Methow and Wenatchee Rivers, respectively (Chapman et al. 1995). As part of the Rock Island Mitigation Agreement between Chelan County Public Utility District and the fishery management parties (RISPA 1989), a supplementation (supportive breeding) program was initiated in 1989 on the Chiwawa River to mitigate smolt mortality resulting from the operation of Rock Island Hydroelectric Project. EFH uses broodstock collected at a weir on the Chiwawa River, although in recent years hatchery fish have been collected at Tumwater Dam. Similarly, the MFHC uses returning adults collected at weirs on the Methow River and its tributaries, the Twisp and Chewuch Rivers (Chapman et al. 1995; Bugert 1998). Although low run size and trap efficiency has resulted in most broodstock being collected from the hatchery outfall or in some years Wells Dam,
progeny produced from these programs are reared at and released from satellite sites on the tributaries where the adults were collected. Numerous other facilities have reared spring-run Chinook salmon on an intermittent basis.

## Previous Genetic Studies - Population differentiation

Waples et al. (1991a) examined 21 polymorphic allozyme loci in samples from 44 populations of Chinook salmon in the Columbia River Basin. These authors reported three major clusters of Columbia River Basin Chinook salmon: 1) Snake River springand summer-run Chinook salmon, and mid and upper Columbia River spring-run Chinook salmon, 2) Willamette River spring-run Chinook salmon, 3) mid and upper Columbia River fall- and summer-run Chinook salmon, Snake River fall-run Chinook salmon, and lower Columbia River fall- and spring-run Chinook salmon. Utter et al. (1995) examined allele frequency variability at 36 allozyme loci in samples of 16 upper Columbia River Chinook populations. Utter et al. (1995) indicated that spring-run populations were distinct from summer- and fall-run populations, where the average genetic distance between spring-run and late-run Chinook were about eight times the average of genetic distances between samples within each group. Additionally, allele frequency differences among spring-run populations were considerably greater than that among summer- and fall-run populations in the upper Columbia River. Utter et al. (1995) also reported hatchery populations of spring-run Chinook salmon were genetically distinct from natural spring-run populations, but hatchery populations of fall-run Chinook salmon were not genetically distinct from natural fall-run populations.

As part of an evaluation of the relative reproductive success for the Chiwawa River supplementation program, Murdoch et al. (2006), used eleven microsatellite loci to assess population differentiation among spring Chinook salmon population samples in the upper Wenatchee River. Murdoch et al. (2006) reported a $>99 \%$ accuracy of correctly identifying spring-run and fall-run Chinook from the Wenatchee River. They also reported slight, but significantly different genetic variation among wild spring populations and between wild and hatchery stocks. Yet, since the spring-run populations
are genetically similar, identifying individuals genetically from the upper tributaries of the Wenatchee River was difficult. This result is exemplified in their individual assignment results, where $<8 \%$ of spring-run individuals, hatchery or wild, were correctly assigned using their criterion of an LOD ( $\log$ of odds) score greater than 2. Murdoch et al. (2006) also reported contemporary natural spring Chinook show heterozygote deficit and low linkage disequilibrium (LD), while contemporary hatchery spring Chinook show heterozygote excess and high LD.

Williamson et al. (submitted) have continued the work of Murdoch et al. (2006) by analyzing Chiwawa River demographic data from 1989 - 2005 to estimate the proportions of recruits that were produced by Chinook with hatchery or wild origin. In an "ideal" population, the genetic size (i.e., effective size or $\mathrm{N}_{\mathrm{e}}$ ) and the census size are equal; however various demographic factors such as unequal sex ratios and variance in reproductive success among individuals reduces the genetic size below the census size. It is generally thought that the genetic size is approximately $10-33 \%$ the census size (Bartley et al. 1992; RS Waples pers. comm.), although values have been reported outside this range (Araki et al. 2007; Arden and Kapuscinski 2003; Heath et al. 2002). Despite being difficult to estimate, the effective population size in many respects is a more important parameter to know than census size, because $\mathrm{N}_{\mathrm{e}}$ determines how genetic diversity is distributed within populations and how the forces of evolution (i.e., forces that change genetic diversity over time) will affect the genetic variation present.

Williamson et al. (submitted) used demographic data to 1 ) investigate the effect of unequal sex ratio on genetic diversity, 2) investigate the effect of variation in reproductive success on genetic diversity, 3) investigate the effect of fluctuations in population size on genetic diversity, and 4) estimate the effective population size, using the inbreeding method (Ryman and Laikre 1991). Most importantly, they use demographic data from 1989 - 2000 to assess the impact of the Chiwawa Hatchery Supplementation Program on the effective population size of natural-origin Chiwawa River spring Chinook. They estimate that the $\mathrm{N}_{\mathrm{e}}$ of naturally spawning Chiwawa Chinook (i.e., both hatchery- and wild-origin fish on the spawning grounds) from 1989 -

1992 was $\mathrm{N}_{\mathrm{e}}=2683$ and in $1997-2000$ was $\mathrm{N}_{\mathrm{e}}=989$. They compare spawning ground $\mathrm{N}_{\mathrm{e}}$ to estimates calculated from combined broodstock and naturally spawning Chinook demographic data. The combined inbreeding $\mathrm{N}_{\mathrm{e}}$ estimate from $1989-1992$ was $\mathrm{N}_{\mathrm{e}}=$ 147 and in $1997-2000$ was $\mathrm{N}_{\mathrm{e}}=490$. Williamson et al. (submitted) argue that since the combined $\mathrm{N}_{\mathrm{e}}$ estimate is lower than the naturally spawning estimate, the supplementation program has had a negative impact on the Chiwawa River $\mathrm{N}_{\mathrm{e}}$.

Williamson et al. (submitted) also present genetic data for Chinook recovered on spawning grounds in upper Wenatchee River tributaries in 2004 and 2005. These genetic data are derived from the Murdoch et al. (2006) study. They compare samples collected from Chiwawa River (i.e., hatchery and wild), White River, Nason Creek, and Leavenworth Hatchery. Additionally, they include a 1994 Chiwawa River wild smolt sample for comparison with the 2004 brood year. Williamson et al. (submitted) report statistically significant genetic differentiation among Chiwawa River, White River and Nason Creek. Additionally, they report that the 1994 and 2004 Chiwawa River wild samples are not statistically different, but the 2004 Chiwawa wild and hatchery collections are statistically different.

## Study Objectives

This study investigated within and among population genetic diversity to assess the effect of the Chiwawa Hatchery's supplemental program on the natural Chiwawa River spring Chinook population. Differences among temporal population samples, the census size, heterozygosity, and allelic diversity were documented. We investigated population differentiation between the Chiwawa River natural and hatchery samples, and among all temporally replicated samples from the Wenatchee River watershed using microsatellite DNA allele frequencies and the statistical assignment of individual fish to specific populations. To assess the genetic effect of the hatchery program, correlation between census and effective population sizes were investigated using temporally replicated samples obtained before and after the supplementation program operation. To address the hypotheses associated with Objective 3 in Murdock and Peven (2005) we developed
eleven specific "Tasks" (Blankenship and Murdoch 2006), to which we analyzed specific genetic data. We present the results from these analyses specific to each individual Task.

## Methods and Materials

## Tissue collection and DNA extraction

We analyzed thirty-two population collections of adult spring Chinook salmon (Oncorhynchus tshawytscha) obtained from the Wenatchee River between 1989 and 2006 (Table 1). Nine collections of natural Chinook adults from the Chiwawa River ( $\mathrm{n}=501$ ), and nine collections of Chiwawa Hatchery Chinook ( $\mathrm{n}=595$ ) were collected at a weir located in the lower Chiwawa River. The 1993 and 1994 Chiwawa Hatchery samples are smolt samples from the 1991 and 1992 hatchery brood years, respectively. Additional samples were collected from upper Wenatchee River tributaries, White River, Little Wenatchee River, and Nason Creek. Six collections of natural White River Chinook ( $\mathrm{n}=179$ ), one collection from the Little Wenatchee ( $\mathrm{n}=19$ ), and six collections from Nason Creek ( $\mathrm{n}=268$ ) were obtained. Single collections were obtained for Chinook spawning in the mainstem Wenatchee River and Leavenworth National Fish Hatchery. An additional out-of-basin collection from Entiat River was also included in the analysis. Samples collected in 1992 or earlier are scale samples. All other samples were either fin clips or operculum punches, stored immediately in ethanol after collection. DNA was extracted from stored tissue using Nucleospin 96 Tissue following the manufacturer's standard protocol (Macherey-Nagel, Easton, PA, U.S.A.).

## Laboratory analysis

We performed polymerase chain reaction (PCR) amplification on each fish sample using the 13 fluorescently end-labeled microsatellite marker loci standardized as part of the GAPS project (Seeb et al. in review). GAPS genetic loci are: $\mathrm{Ogo2}$, $\mathrm{Ogo4}$ (Olsen et al. 1998); Oki100 (unpublished); Omm 1080 (Rexroad et al. 2001); Ots201b (unpublished); Ots208b, Ots211, Ots212, and Ots213 (Grieg et al. 2003); Ots 3 M , Ots 9 (Banks et al.
1999); OtsG474 (Williamson et al. 2002); Ssa408 (Cairney et al. 2000). PCR reaction volumes were $10 \mu \mathrm{~L}$, and contained $1 \mu \mathrm{~L} 10 \mathrm{x}$ PCR buffer (Promega), $1.0 \mu \mathrm{~L} \mathrm{MgCl2} \mathrm{(1.5}$ mM final) (Promega), $0.2 \mu \mathrm{~L} 10 \mathrm{mM}$ dNTP mix (Promega), and 0.1 units/mL Taq DNA polymerase (Promega). Loci were amplified as part of multiplexed sets, so primer molarities and annealing temperatures varied. Multiplex one had an annealing temperature of $50^{\circ} \mathrm{C}$, and used 0.37 Molar (M) Oki100, 0.35 M Ots 201 b , and 0.20 M Ots208b, and 0.20 M Ssa 408 . Multiplex two had an annealing temperature of $63^{\circ} \mathrm{C}$, and used $0.10 \mathrm{M} \mathrm{Ogo2}$, and 0.25 M of a non-GAPS locus (Ssa 197). Multiplex three had an annealing temperature of $56^{\circ} \mathrm{C}$, and used $0.18 \mathrm{M} \mathrm{Ogo4}, 0.18 \mathrm{M}$ Ots 213 , and 0.16 M OtsG474. Multiplex four had an annealing temperature of $53^{\circ} \mathrm{C}$, and used 0.26 M Omm1080, and 0.12 M Ots 3 M . Multiplex five had an annealing temperature of $60^{\circ} \mathrm{C}$, and used 0.30 M Ots $212,0.20 \mathrm{M}$ Ots 211 , and 0.10 M Ots 9 . Thermal cycling was conducted on either a PTC200 thermal cycler (MJ Research) or GeneAmp 9700 (Applied Biosystems) as follows: $95^{\circ} \mathrm{C}(2 \mathrm{~min}) ; 30$ cycles of $95^{\circ} \mathrm{C}$ for 30 sec ., 30 sec . annealing, and $72^{\circ} \mathrm{C}$ for 30 sec .; a final $72^{\circ} \mathrm{C}$ extension and then a $10^{\circ} \mathrm{C}$ hold. PCR products were visualized by electrophoresis on an ABI 3730 automated capillary analyzer (Applied Biosystems). Fragment analysis was completed using GeneMapper 3.7 (Applied Biosystems). Standardization of genetic data to GAPS allele standards was conducted following Seeb et al. (in review).

## Genetic data analysis

Assessing within population genetic diversity - Heterozygosity measurements are reported using Nei's (1987) unbiased gene diversity formula (i.e., expected heterozygosity) and Hedrick's (1983) formula for observed heterozygosity. Both tests are implemented using the microsatellite toolkit (Park 2001). We used GENEPOP version 3.4 (Raymond and Rousset 1995) to assess Hardy-Weinberg equilibrium (HWE), where deviations from the neutral expectation of random associations among alleles are calculated using a Markov chain method (5000 iterations in this study) to obtain unbiased estimates of Fisher's exact test. Global estimates of $\mathrm{F}_{\text {IS }}$ according to Weir and Cockerham (1984) were calculated using GENEPOP version 3.4. Genotypic linkage disequilibrium was calculated following Weir (1979) using GENEPOP version 3.4.

Linkage results for population collections are reported as the proportion of pairwise (locus by locus) tests that are significant (alpha $=0.01$ ). Linkage disequilibrium is considered statistically significant if more than $5 \%$ of the pairwise tests based on permutation are significant for a collection.

Within- and among-population genetic differentiation - The temporal stability of allele frequencies within populations, and pairwise differences in allele frequencies among populations were assessed using several different procedures. First, we tested for differences in allele frequencies among populations defined in Table 1 using a randomization chi-square test implemented in GENEPOP version 3.4 (Raymond and Rousset 1995). This procedure tests for differences between pairs of populations where alleles are randomized between the populations (i.e., genic test). The null hypothesis for this test is that the allele frequency distributions between two populations are the same. A low p-value should be interpreted as the allele frequency distributions being compared are unlikely to be samples drawn from the same underlying distribution.

Second, to graphically describe allele frequency differences among populations we conducted a nonmetric multidimensional scaling analysis using allele-sharing distance matrices from two different data sets. Pairwise allele-sharing distances are calculated as 1 - (mean over all loci of the sums of the minima of the relative frequencies of each allele common to a pair of populations). To calculate the allele-sharing distances for each pair of populations we used PowerMarker v3.25 (Liu and Muse 2005). Nonmetric multidimensional scaling is a technique designed to construct an n-dimensional "map" of populations, given a set of pairwise distances between populations (Manly 1986). The output from this analysis is a set of coordinates along n-axes, with the coordinates specific to the number of n-dimensions selected. To simplify our analysis we selected a 2-dimensional analysis to represent the relative positions of each population in a typical bivariate plot. The goodness of fit between the original allele-sharing distances and the pairwise distances between all populations along the 2-dimensional plot is measured by a "stress" statistic. Kruskal (in Rohlf 2002) developed a five-tier guide for evaluating stress levels, ranging from a perfect fit (stress=0) to a poor fit (stress=0.40). We
conducted the nonmetric multidimensional scaling analysis for one data set containing Chiwawa natural- and hatchery-origin collections, and another data set containing Chiwawa broodstock and in-river spawner collections. We used the mdscale module in MATLAB R2006b (The Mathworks 2006) to generate the nonmetric multidimensional scaling coordinates.

We examined the geographic and temporal structure of populations in the upper Wenatchee (Chiwawa River, Nason Creek, and White River, only) using a series of analyses of molecular variance (AMOVAs). Here, we defined an AMOVA as an analysis of variance of allele frequencies, as originally designed by Cockerham (1969), but implemented in Arlequin v2.1 (Schneider et al. 2000). These analyses permit populations to be aggregated into groups, and molecular variance is then partitioned into within collections, among collections, but within groups, and among group components. With this approach, we were able to determine how best to group populations, with "best" being defined as that grouping that accounts for the greatest proportion of among group variance. Furthermore, by partitioning molecular variance into three different hierarchical components, we are able to determine what level accounts for the majority of the molecular variance.

Finally, we explored the partitioning of molecular variance between among-individuals and among-populations using a principal component analysis and multi-locus estimates of pairwise $\mathrm{FST}_{\text {, }}$, estimated by a "weighted" analysis of variance (Weir and Cockerham, 1984). Principal component analysis is a data-reduction technique whereby the correlation structure among variables can be used to combine variables into a series of multivariate components, with each original variable receiving a weighted value for each component based on its correlation with that component. Here, we used a program written by Warheit in MATLAB R2006b (The Mathworks 2006) that treats each allele for each locus as a single variable ( 13 loci = 26 alleles or variables), and these 26 "variables" were arranged into 26 components, with each component accounting for a decreasing amount of molecular variance. Estimates of $\mathrm{F}_{\text {ST }}$ were calculated using GENETIX version 4.05 (Belkhir et al.1996). To determine if the FST estimates were
statistically different from random (i.e., no structure), 1000 permutations were implemented in GENETIX version 4.05 (Belkhir et al.1996).

Effective population size ( $\mathbf{N}_{\mathbf{e}}$ ) - Estimates of the effective population size were obtained using two methods, a multi-collection temporal method (Waples 1990), and a singlecollection method (Waples 2006) using linkage disequilibrium data. The temporal method assumes that cohorts are used, but we did not decompose the collection year samples into their respective cohorts using age data. Therefore, $\mathrm{N}_{\mathrm{e}}$ estimates that pertain to individual year classes of breeders are not valid; however the harmonic mean over all samples will estimate the contemporary $\mathrm{N}_{\mathrm{e}}$. Comparing samples from years $i$ and $j$, Waples' (1990) temporal method estimates the effective number of breeders ( $\left.\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j},}\right)$ according to:

$$
\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}=\frac{\mathrm{b}}{2\left(\hat{\mathrm{~F}}-1 / \hat{\mathrm{S}}_{\mathrm{i}, \mathrm{j}}\right)}
$$

The standardized variance in allele frequency ( $\hat{F}$ ) is calculated according to Pollack (1983). The parameter $b$ is calculated analytically from age structure information and the number of years between samples (Tajima 1992). The age-at-maturity information required to calculate b was obtained from Murdoch et al. (2006) for this analysis. They observed for Chiwawa Hatchery Chinook that $8.6 \%$ matured at age 2, $4 \%$ at age 3, $87 \%$ at age 4, and $0.4 \%$ at age 5. For Chiwawa natural Chinook, Murdoch et al. (2006) observed that $1.8 \%$ matured at age 3, 81.6\% at age 4, and $16.7 \%$ at age 5. The harmonic mean of sample sizes from years $i$ and $j$ is $\tilde{\mathrm{S}}_{\mathrm{i}, \mathrm{j}}$. Over all pairwise comparisons the harmonic mean of all $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ is $\tilde{\mathrm{N}}_{\mathrm{b}}$, the contemporary estimate of the effective population size $\left(\mathrm{N}_{\mathrm{e}}\right)$. SALMONNb (Waples et al. 2007) was used to calculate $\widetilde{\mathrm{N}}_{\mathrm{b}}$. As suggested by authors, alleles with a frequency below 0.05 were excluded from the analysis to reduce potential bias.

The method of Waples (2006) uses linkage disequilibrium (i.e., mean squared correlation of allele frequencies at different gene loci) as a means of estimating effective population size $\left(N_{e}\right)$ from a single sample. While this method is biased in some cases where $N_{e} / \mathrm{N}$
ratio is less the 0.1 and the sample size is less than the true $\mathrm{N}_{\mathrm{e}}$, it has been shown to produce comparable results to the temporal method. Burrows' delta method is used to estimate LD, and a bias corrected estimate of $\mathrm{N}_{\mathrm{e}}$ is calculated after eliminating alleles with frequency less than 0.05 . This test was implemented using $\operatorname{LDN}_{\mathrm{e}}$ (Do and Waples unpublished). In age-structured species, $\mathrm{N}_{\mathrm{e}}$ estimates based on LD are best interpreted as the effective number of breeders $\left(\mathrm{N}_{\mathrm{b}}\right)$ that produced the sample (Waples 2006). $\mathrm{N}_{\mathrm{b}}$ should be multiplied by the mean generation length (i.e., 4 in this case) to obtain an overall estimate of $\mathrm{N}_{\mathrm{e}}$ based on an $\mathrm{N}_{\mathrm{b}}$ estimate. We analyzed collections categorized by spawning location (i.e., hatchery broodstock or in-river) and did not analyze collections categorized by origin (i.e., hatchery or natural). Waples' (2006) method estimates $\mathrm{N}_{\mathrm{e}}$ from observed LD, therefore the corresponding $\mathrm{N}_{\mathrm{e}}$ estimates for the hatchery collections would be low and the estimates for the natural collections would be high. Yet, since the supplementation program is integrated, and hatchery fish can spawn naturally, we feel it inappropriate to analyze the hatchery and natural samples as if they were separate, which would essentially partition all the LD into the hatchery samples.

Each collection has an $\mathrm{N}_{\mathrm{b}}$ estimate and an associated confidence interval. If the confidence interval includes infinity, it means that sampling error accounts for all the LD observed (i.e., empirical LD is less than expected LD). The usual interpretation is that there is no evidence for any disequilibrium caused by genetic drift in a finite number of parents. Since the LD method estimates the number of breeders that contributed to the sample being analyzed, in order to calculate an $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratio, the appropriate census size must be used. The census size used to derive a ratio was the estimate four years prior to the collection analyzed using LD, which assumed a strict four-year-old lifecycle, although the observed proportion of four-year-olds was approximately $85 \%$ each year. The census numbers (Table 2) used to calculate the ratios for Chiwawa broodstock and in-river spawners were combined NOS (natural-origin spawners) and HOS (hatcheryorigin spawners) census estimates.

Individual assignment - A population baseline file was constructed containing all 1704 individual Chinook from 34 population collections (Table 1; Chiwawa origin data set
plus all samples from other populations). All individuals in the baseline had geneotypes that included nine or more loci. Individual Chinook were assigned to their most likely population of origin based on the partial Bayesian criteria of Rannala and Mountain (1997), using a "jack-knife" procedure, where each individual to be assigned was removed from the baseline prior to the calculation of population likelihoods. This procedure was implemented in a program written by Warheit in MATLAB R2006b (The Mathworks 2006). Two assignment criteria were used, 1) the population with the largest posterior probability for an individual was the "most-likely" population of origin (i.e., all individuals assigned to a collection), and 2) an assignment was consider valid only if the posterior probability was greater than or equal to 0.9 . Please note that while the analysis used 34 population collections to assign Rannala and Mountain likelihoods for each individual, these likelihoods were aggregated based on "population" (i.e., Chiwawa, Nason, White, and so on) and posterior probabilities were calculated for population location, rather than individual collections.

## Results and Discussion

In this section we combine our presentation and interpretations of the genetic analyses. Additionally, this section will be organized based on the task list presented in the study plan. Overall conclusions are provided following this section.

## Task 1: Determine trend in census size for Chiwawa River spring Chinook.

Census data from 1989-2005 are provided in Table 2 for the Chiwawa Hatchery broodstock and spring Chinook present in the Chiwawa River. The demographic data for naturally spawning Chinook are based on redd sampling and carcass surveys, while broodstock data are based on Chiwawa hatchery records. As the supplementation program is integrated by design, we also present the proportion of natural-origin broodstock ( pNOB ) incorporated into the hatchery, in addition to the number of naturalorigin (NOS) and hatchery-origin (HOS) spawners present in Chiwawa River. The
census size fluctuated yearly, and a general reduction in census size was observed in the mid to late 1990's. This trend was apparent in both the broodstock and in the river. The arithmetic mean census size from 1989 - 2005 for the Chiwawa Hatchery (i.e., broodstock) was $\mathrm{N}=87.5$ per year. The arithmetic mean census size from 1989 - 2005 for the Chiwawa River (i.e., NOS and HOS combined) was N=961.9 per year. For collection years when adult Chiwawa hatchery-origin fish would have been absent in the Chiwawa River (1989 - 1992), the arithmetic mean of natural Chiwawa Chinook census size is $\mathrm{N}=962.7$. We will use this number as the baseline census size to assess if census size has changed. We used two different values for the contemporary census size in the Chiwawa River, NOS only and NOS + HOS. Additionally, we used collection years 2002-2005 for the contemporary NOS and HOS estimates, as these are the most recent data and the number of years included for estimation is the same as the pre-hatchery estimate above (i.e., four years). For NOS only, the arithmetic mean census size from 2002-2005 was $\mathrm{N}=536.0$. For total census size (i.e., NOS and HOS combined), the arithmetic mean census size from 2002 - 2005 was $\mathrm{N}=1324.0$. For the demographic data presented here, the contemporary census size is larger than the census estimate derived from the years prior to hatchery operation.

## Task 2: Document the observed genetic diversity.

## Genetic Diversity Categorized By Origin

For Chiwawa River collections categorized by origin (Table 1A), substantial genetic diversity was observed, with heterozygosity estimates over all loci, having a mean of 0.80. Genetic diversity was consistent with expected Hardy-Weinberg random mating genotypic proportions for ten of the eighteen collections. Eight of the nine Chiwawa natural collections were consistent with HWE, and two of nine Chiwawa Hatchery collections were consistent with HWE. $\mathrm{F}_{\text {IS }}$ is observed to be slight for all Chiwawa population collections, suggesting individuals within collections do not show excessive homozygosity.

The deviations from HWE observed were generally associated with hatchery collections. The two smolt collections (i.e., 1993 and 1994) showed significant deviations from HWE, which may be a function of non-random hatchery practices involving the contributing natural-origin parental broodstocks (i.e., 1991 and 1992 cohort). Deviations from HWE in the remaining hatchery collections may be the result of few individuals being represented in the broodstock (see below).

Additionally, linkage disequilibrium (LD) was also common for Chiwawa hatcheryorigin collections and minimal for Chiwawa natural-origin collections. The random association of alleles between loci (i.e., linkage equilibrium) is expected under ideal conditions. LD is observed when particular genotypes are encountered more than expected by chance. Laboratory artifacts (e.g. null alleles) or physical linkage of loci on the same chromosome can cause LD, but the LD we observed was not associated with certain locus combinations, which you would expect if either artifacts or physical linkage were the cause of LD. LD was observed for seven of the nine hatchery-origin collections. As with the deviations from HWE, the high LD in the 1993 and 1994 hatchery-origin collections may be a result of non-random hatchery practices. The substantial LD observed in the hatchery-origin adult collections (collection years 2000, 2001, 2004, and 2006) might be the result of small parental broodstock sizes contributing to those returning adults. During the mid 1990's, the Chiwawa broodstock size was low, with zero individuals collected in 1995 and 1999; so fewer individuals would be contributing to the hatchery adult returns than the natural. This idea is corroborated by the lower LD observed for the 2005 hatchery-origin collection, which had a contributing parental broodstock size in 2001 (i.e., the major contributing parental generation) approximately eight times as large as the previous few collection years (Table 2). LD reappears in the 2006 Chiwawa hatchery-origin collection, which had a contributing parental broodstock size (i.e., for the most-part, the 2002 hatchery brood year) five times lower (Table 2) than that of the 2005 collection.

While seven of nine hatchery-origin collections showed significant LD, only one natural origin collection showed LD, and for this collection, only $10 \%$ of the loci-pairs were in
disequilibrium (Table 1). The fact that LD predominated in the hatchery samples, suggests that variance in reproductive success (i.e., overrepresentation of particular parents) is higher in the hatchery-origin than in natural-origin collections.

## Genetic Diversity Categorized By Spawning Location

For upper Wenatchee River collections categorized by spawning location (Table 1B), substantial genetic diversity was observed, with heterozygosity estimates over all loci, having a mean of 0.79 and ranging from a low of 0.69 (1993 White River) to 0.85 (1993 Little Wenatchee). Genetic diversity was consistent with HWE for nineteen of twentynine population collections. For the collections that departed from HWE, seven were from the Chiwawa River, one was from Leavenworth Hatchery, one was the Wenatchee mainstem collection of hatchery-origin - naturally spawning fish, and one was from the White River. FIS is observed to be slight for all population collections except the 1993 White River collection ( $10 \%$ heterozygote deficit) (Table 1B). Collections deviating with HWE generally correlated with collections having high LD. Twelve population collections showed a proportion of pairwise linkage disequilibrium tests (across all loci) greater than 5\% (Table 1B), eight of which were Chiwawa collections.

Starting in 1996, spawning location collections are composed of both natural- and hatchery-origin samples. The LD seen in the later spawning location collections may be caused by an admixing effect (i.e., mixing two populations), where random mating has not had the chance to freely associate alleles into genotypes. Interestingly, there appears to be a trend of reducing LD through time within the broodstock collections (Table 1B), which suggests that a "homogenizing" effect is taking place within the Chiwawa River. This observation is discussed more fully in Task 3 below.

## Task 3: Test for population differentiation among collections within the Chiwawa River and associated supplementation program.

## Introduction

Task 3 was designed to address two hypotheses listed as part of Objective 3 in Murdoch and Peven (2005):

- Ho: Allele frequency Hatchery $=$ Allele frequency Naturally produced $=$ Allele frequency ${ }_{\text {Donor pop }}$.
- Ho: Genetic distance between subpopulations Year $x=$ Genetic distance between subpopulations year y

Murdoch and Peven (2005) proposed these two hypotheses to help evaluate the Chiwawa supplementation program through the "Conceptual Process" (Figure 5 in Murdoch and Peven 2005; repeated here as Figure 1). There are two components to the first hypothesis, which must be considered separately. The first component involves comparisons between natural-origin populations in the Chiwawa to determine if there have been changes in allele frequencies or genetic distances, through time starting with the donor population. Documenting a change does not necessarily indicate that the supplementation program has directly affected the natural origin fish, as additional tests would be necessary to support that hypothesis. The intent of the second component is to determine if the hatchery produced populations have the same genetic composition as the naturally produced populations.

Although on the surface these two components and their associated comparisons may appear simple, from a hypothesis-testing perspective the analyses are complicated by the fact that natural-origin fish may have had hatchery-origin parents, and hatchery-origin fish may have had natural-origin parents. As such, we organized the Chiwawa genetic data into three data sets: (1) fish origin (hatchery versus natural), (2) spawning location (hatchery broodstock versus in-river (natural) spawners), and (3) four "treatment" groups (1. hatchery-origin hatchery broodstock, 2 . hatchery-origin natural spawner, 3. naturalorigin natural spawner, and 4. natural-origin hatchery broodstock). We conducted separate analyses using each of the three data sets, with each analysis touching on some aspect of the components necessary to move through the Conceptual Process (Figure 1).

## Hatchery- Versus Natural-Origin

We address the following questions with the origin data set:

1. Are there changes in allele frequencies and allele sharing distances in the naturalorigin collections from pre-supplementation to today?
2. Are there changes in allele frequencies and allele sharing distances in the hatchery-origin collections from early supplementation to today?
3. Are there significant differences in allele frequencies and large allele sharing distances between hatchery- and natural-origin adults from a collection year, and has this pattern changed through time?

Genic Differentiation Tests - We explicitly tested the hypothesis of no significant differentiation within natural- or hatchery-origin collections from the Chiwawa River using a randomization chi-square test. We show the results for the pairwise comparisons among natural-origin collections from the Chiwawa River populations in the first block of the second page of Table 3. Ten of the 36 ( $28 \%$ ) pairwise comparisons have highly significant allele frequency differences, while only 12 of the 36 comparisons ( $33 \%$ ) showed no significant differences. Eight of these 12 comparisons involved the 1996 collection, which included only eight samples and therefore provided little power to differentiate allele frequencies. If we exclude the 1996 collection, only $14 \%$ of the pairwise comparisons showed no significant differences, and here all but one of these comparisons involved the 1989 collection. The 1989 collection appeared to be the least differentiated collection in the natural-origin data set in that all pairwise comparisons were either not significant, or only mildly significant at the nominal critical value. No comparisons involving the 1989 collection were significant using a Bonferroni-corrected critical value, and 1989 is the only natural-origin collection in our data set that can be classified as "pre-supplementation."

We can interpret these results to indicate that although there appears to be significant year-to-year differences in allele frequencies among post-supplementation collections, the allele frequencies between each post-supplementation collection and the 1989 presupplementation collection are not greatly different. However, the level of differentiation
does increase from the early post-supplementation years to the more recent years (2001, 2004-2006), although the statistical level of this significance never exceeds the Bonferroni-corrected critical value. Finally, sample sizes were also small for the 1989 collection ( $\mathrm{n}=36$ ) and we cannot eliminate a reduction in power as a contributing factor for the lack of significance for these tests.

As with the hatchery-origin collections, most pairwise comparisons of allele frequencies between hatchery-origin samples were significant (Table 3, first page, upper block). Out of the 36 pairwise comparisons, all but three are significant at some level, and most comparisons are highly significant. Similar to the natural-origin analysis, the nonsignificant results were limited to comparisons involving the 1996, which included only eight samples.

As a result of this analysis we reject the hypothesis that there was no significant differentiation among natural- or hatchery-origin collections from the Chiwawa River. Furthermore, the allele frequencies of the hatchery-origin collections are significantly different from those of natural-origin collections (Table 3, first page, second block). For those fish collected in the same year, allele frequencies are significantly different between hatchery- and natural-origin collections, although in 2005 the level of significance was below the Bonferroni critical value (Table 3). The next step is to examine the pattern of allelic differentiation to discover first if there is a trend among the data, and second, if this trend suggests that the allele frequency differences among Chiwawa River natural-origin fish collections has been affected by the hatchery-origin fish.

Allele-sharing and Nonmetric Multidimensional Scaling - We constructed a pairwise allele-sharing distance matrix for all hatchery- and natural-origin collections from the Chiwawa River and subjected this matrix to a nonmetric multidimensional scaling analysis, restricting the analysis to two dimensions (Figure 2). The stress statistic for this analysis is 0.09 , a value Kruskal (in Rohlf 2002) listed as a good to excellent fit between the actual allele-sharing distances and the Euclidean (straight-line) distances in the plot.

In other words, Figure 2 is a good visual representation of the allele sharing distance matrix; collections with a high percentage of alleles shared will be closer to each other than collections with a lower percentage of alleles shared.

With the exception of the two outlier years (1996 and 1998) the Chiwawa natural-origin collections form a tight cluster indicating an overall common set of shared alleles among these collections. Even if we ignore the 1996 and 1998 hatchery-origin collections, there appears to be a greater variance in shared alleles among the Chiwawa hatchery-origin collections than the natural-origin collections (Figure 2). In fact, the median percentage of alleles shared among the Chiwawa natural-origin collections is $76 \%$ compared with $69 \%$ alleles shared among the Chiwawa hatchery-origin collections.

Also, there appears to be a convergence in allele sharing distances (i.e., a decrease in allele frequency differences) between the hatchery- and natural-origin fish from the late 1980s/early 1990s to 2006. The series of red arrows in Figure 2 represent the progression of change in hatchery-origin allele sharing distances from 1996 (first adult hatchery origin fish in our analysis) to 2006 and this progression is decidedly in the direction of the natural-origin cluster. However, the most recent natural-origin collections (2001, 2004-2006) appear to have pulled closer to the hatchery-origin collections, compared with the 1989 natural-origin collection (note the close proximity of the 2000 and 1989 natural-origin collections). Nevertheless, the cluster of natural-origin collections adjacent to the hatchery-origin collections in Figure 2 also includes the 1993 natural-origin collection. Qualitatively, it appears that the initial hatchery-origin and natural-origin collections were more different from each other in terms of the percentage of shared alleles than are the most recent hatchery- and natural-origin collections. This may have been a result of a non-random sample of natural-origin fish that was used as broodstock in the initial years of the supplementation program (see discussion in Task 2 concerning deviations from HWE and linkage disequilibrium).

That being said, we do need to emphasize that Figure 2 is dominated by five outlier collections (two each from the 1996 and 1998 collections, and the 1994 smolt collection).

The 1996 and 1998 collections are characterized by small samples sizes, and the 1994 smolt collection has nearly all pairs of loci in linkage disequilibrium (Table 1). If we eliminate these five outlier groups, both the hatchery- and natural-origin collections form a relatively tight cluster. Excluding the five outliers, the median percentage of shared alleles among all pairwise combinations of Chiwawa hatchery versus Chiwawa natural collections is $76 \%$. This compares with a median pairwise percentage of $79 \%$ among only Chiwawa natural-origin collections. That is, there are nearly as many alleles shared between the hatchery-origin and natural-origin collections as there are among the naturalorigin collections themselves. There is also a narrowing of differences between naturaland hatchery-origin fish from the same collection years from 1993 ( $76 \%$ shared alleles) through 2006 (83\% shared alleles).

If allelic differentiation among collections is a function of genetic drift, we would expect a positive correlation between the number of years between two collections and the allele sharing distance. That is, if genetic drift is the primary cause of allele frequency differences between two collections, the greater the number of years between the two collections the larger the allele-sharing distance. For both the natural- and hatcheryorigin collections we examined the relationship between the number of years between a pair of collections and the collections' allele-sharing distance (Figure 3). Although the relationship between time interval and allele distance appears to be a positive function in the natural collections, the slope of the regression line is 0.0017 , and is not significantly different from zero. Furthermore, the correlation coefficient $\left(\mathrm{r}^{2}\right)$ equals 0.1068 , which means that the time interval between collections accounts for only $10 \%$ of the pairwise differences in allelic distance. The hatchery-origin collections do show a significantly positive slope ( $0.0037 ; \mathrm{p}=0.0254$ ) and a regression coefficient nearly three times greater than that for the natural-origin collections. However, the correlation coefficient is still relatively small ( $r^{2}=0.3290$ ), indicating that the time interval between collections accounts for one-third of the pairwise differences in allelic distance. The results suggest that if genetic drift is a factor in allelic differentiation between collections, it is only a minor factor, and appears to have affected the hatchery-origin collections more than the natural-origin collections.

If four-year-old fish dominate each collection year, we would expect a closer relationship among collections that are spaced at intervals of four years. The average percentage of alleles shared between two natural-origin collections that are separated by four years or a multiple of four years is $81 \%$, compared with $78 \%$ for natural-origin collections separated by years that are not divisible by four. Likewise, for hatchery-origin collections the average percentage of alleles shared is $80 \%$ and $75 \%$ for collections separated by years divisible and not divisible by four, respectively. Although the percent differences described above are relatively small, they are consistent with the idea that allelic differences between collections are a function of year-to-year variability among different cohorts of four year-old fish.

Summary - The allele frequencies within and between natural- and hatchery-origin collections are significantly different, but there does not appear to be a robust signal indicating that the recent natural-origin collections have diverged greatly from the pre- or early post-supplementation collections. Genetic drift will occur in all populations, but does not appear to be a major factor with the Chiwawa collections. We propose that the differences among collections are a function of differences in allele frequencies among cohorts of the four year-old fish that dominate each collection.

## Hatchery Broodstock Versus Natural (In-River) Spawners

We address the following questions with the spawner data set:

1. Are there changes in allele frequencies and allele sharing distances in the natural spawning collections from pre-supplementation to today?
2. Are there changes in allele frequencies and allele sharing distances in the hatchery broodstock collections from early supplementation to today?
3. Are there significant differences in allele frequencies and large allele sharing distances between hatchery and natural spawning adults from a collection year, and has this pattern changed through time?

Genic Differentiation Tests - For the most part there are significant differences in allele frequencies among collections for both the hatchery broodstock and natural spawners (Table 4), and these differences are consistent with the origin data set (Table 3). There are four collection years with paired samples (2001, 2004-2006) where we can compare allele frequency differences between the hatchery broodstock and natural spawners, within the same year. The 2001 hatchery broodstock and natural spawner collections have significantly different allele frequencies, but the level of significance decreased from 2001 to 2004, and become non-significant in 2005 and 2006 (Table 4). This indicates that by 2005, the hatchery broodstock and natural spawners collections were effectively sampling from the same population of fish. Additionally, the percentage of alleles shared between the hatchery broodstock and the natural spawners increased from $76 \%$ in 2001 to $86 \%$ in 2006 (allele sharing distance matrix, not shown). From this analysis, we conclude that although there are year-to-year differences in allele frequencies within the natural and hatchery spawner collections, there appears to be a convergence of allele frequencies within collection-year, between the natural and hatchery spawner populations.

Linkage Disequilibrium - Linkage disequilibrium is the correlation of alleles between two loci, and can occur for several reasons. If two loci are physically linked on the same chromosome, than alleles from each of these loci should be correlated. However, linkage between two loci can occur as a result of population bottlenecks, small population sizes, and natural selection. If any of these conditions had occurred or were occurring within the Chiwawa River system, we would expect to find substantial linkage disequilibrium in many or perhaps all Chiwawa collections. However, many Chiwawa collections, especially the natural-origin collections, do not show linkage disequilibrium (Table 1), and it would appear that the linkage disequilibrium within certain Chiwawa collections is not a function of the processes listed above. Linkage disequilibrium can also result if the collection is composed of an admixture. That is, if two or more reproductively isolated populations are combined into a single collection, the collection will show linkage disequilibrium. Each broodstock and natural spawning collection is composed of naturaland hatchery-origin fish. If these hatchery- and natural-origin fish are drawn from the
same population, the spawning collections should not show substantial linkage disequilibrium. However, if the hatchery- and natural-origin fish are from different populations (i.e., full hatchery - natural integration has not been achieved), the spawning collections should show substantial linkage disequilibrium.

There are only three Chiwawa spawning collections that are not composed of both hatchery- and natural-origin samples: 1989 (natural-origin, natural spawner), 1993 (natural-origin, hatchery broodstock), and 2001 (natural-origin, natural spawner). Of the 10 spawning collections with both hatchery- and natural-origin fish, seven show significant linkage disequilibrium. Two of the three collections that did not show linkage disequilibrium are the 1996 and 1998 hatchery broodstock collections, which are composed of only seven natural- and six hatchery-origin fish, and two natural- and 19 hatchery-origin fish, respectively. Within the hatchery broodstock collections with linkage disequilibrium, the percent of loci pairs showing linkage decreased from $32 \%$ in 2000 to $13 \%$ in 2001 and 2004, to only $1 \%$ and $5 \%$ in 2005 and 2006, respectively (Table 1). If the homogenization of allele frequencies of natural- and hatchery-origin fish was increasing from 2000 to 2006, we would expect a decrease in linkage disequilibrium among the broodstock collections. This is what occurred within the hatchery broodstock collections, but did not occur within the natural spawner collections, where the percent of loci pairs showing linkage was $18 \%$ in $2004,6 \%$ in 2005, and $10 \%$ in 2006 (Table 1). Furthermore, the 2001 natural spawner collection, with no hatchery-origin component showed linkage disequilibrium with $9 \%$ of loci pairs.

There is no correlation between percent of loci pairs showing linkage disequilibrium and percent of broodstock composed of hatchery-origin fish $\left(r^{2}=0.0045\right)$. Furthermore, the natural spawner and hatchery broodstock collections were each composed of roughly the same average percentage of hatchery-origin fish ( $57 \%$ and $53 \%$, respectively). If the decrease in linkage disequilibrium among the hatchery broodstock collections from 2000 to 2006 was a result of a homogenization of allele frequencies of natural- and hatcheryorigin fish in the broodstock, the same degree of homogenization did not occur within the
natural spawner collections. This would occur if natural- and hatchery-origin fish spawning within the river remain segregated, either by habitat or by fish behavior.

Summary - As with the origin data set, there are significant allele frequency differences within and between hatchery broodstock and natural spawner collections. However, in recent years the allele frequency differences between the hatchery broodstock and natural spawner collections has declined. Furthermore, based on linkage disequilibrium, there is a genetic signal that is consistent with increasing homogenization of allele frequencies within hatchery broodstock collections, but a similar homogenization within the natural spawner collection is not apparent. These data suggest that there exists consistent year-to-year variation in allele frequencies among hatchery and natural spawning collections, but there is a trend toward homogenization of the allele frequencies of the natural- and hatchery-origin fish that compose the hatchery broodstock.

## Four Treatment Groups

Analyses of genetic differences between hatchery (broodstock) and natural spawner collections is confounded by the fact that each these two groups are composed of fish of natural- and hatchery-origin. To understand the effects of hatchery supplementation on natural-origin fish that spawn naturally, we needed to divide the Chiwawa data set into four mutually exclusive groups: (1) hatchery-origin hatchery broodstock, (2) hatcheryorigin natural spawner, (3) natural-origin hatchery broodstock, and (4) natural-origin natural spawner, with each group consisting of multiple collection years, for a total of 25 different groups.

Allele-sharing and Nonmetric Multidimensional Scaling -As with previous analyses discussed above, we constructed a pairwise allele-sharing distance matrix for all collections from each of these treatment groups and subjected this matrix to a nonmetric multidimensional scaling analysis, restricting the analysis to two dimensions. Figure 4 shows that five outlier groups dominate the allele-sharing distances within this data set. These outlier groups are also present in Figure 2, as discussed above, and Figure 2 and 4 resemble each other because the same fish are included in each analysis. The difference
between Figures 2 and 4 is that in Figure 4 the fish are grouped into collection year and the four treatment groups, rather than collection year and two treatment groups (hatcheryversus natural-origin).

Figure 4 does not provide useful resolution of the groups within the polygon, because the outlier groups dominate the allele sharing distances. We removed the five outlier groups from Figure 4, recalculated the allele sharing distances and subjected this new matrix to a multidimensional scaling analysis (Figure 5). Figure 5 shows separation among the 2001, 2004-2006 collections, but this separation does not necessarily indicate that within-year collections are more similar to each other than any collection is to a collection from another year. For example, the 2006 natural-origin natural spawner and the 2005 naturalorigin hatchery broodstock collections share $81 \%$ alleles, while the 2006 natural-origin natural spawner and 2006 hatchery-origin hatchery broodstock collections share $75 \%$ alleles. There does not appear to be any discernable pattern of change in allele-sharing distance among the collections relevant to pre- or post-supplementation. Although the 1989 pre-supplementation natural-origin collection appears distinct (Figure 5), the 1993 natural-origin hatchery broodstock collection appears quite similar to the 2005 and 2006 natural-origin collections (Figure 5). The 1993 natural-origin hatchery broodstock collection, although not technically pre-supplementation, is composed of fish whose ancestry cannot be traced to any Chiwawa hatchery fish. Therefore, there is no clear pattern of allele sharing change from pre-supplementation to recent collections.

There does appear to be some change in the average percentage of alleles shared within the 2001 to 2006 collections, with an increase from $74 \%$ in 2001 and 2004 to $78 \%$ and $79 \%$ in 2005 and 2006, respectively. The results provided by this analysis are consistent with the results presented in the origin and spawner data sets. That is, there are allele frequency and allele sharing differences among the collections, but analyses do not strongly suggest that these differences are a function of the supplementation program. Furthermore, there is also a weak signal that the hatchery and natural collections within the most recent years are more similar to each other than in the previous years.

Overall Genetic Variance - Although there are signals of allelic differentiation among Chiwawa River collections, there are no robust signs that these collections are substantially different from each other. We used two different analyses to measure the degree of genetic variation that exists among individuals and collections within the Chiwawa River. First, we conducted a principal component analysis using all Chiwawa samples with complete genotypes (i.e., no missing alleles from any locus). Although the first two principal component axes account for only $10.5 \%$ of the total molecular variance, a substantially greater portion of that variance is among individual fish, regardless of their identity, rather than among hatchery and natural collections (Figure 6). The variances in principal component scores among individuals are 11 and 13 times greater than the variance in scores among collections, along the first and second axes, respectively.

Second, we conducted a series of analyses of molecular variance (AMOVA) to ascertain the percentage of molecular variance that could be attributed to differences among collections. We organized these analyses to test also for differences in the hierarchical structure of the data. That is, we tested for differences among collections using the following framework:

- No organizational structure - all 25 origin-spawner collections considered separately
- Origin-spawner collections organized into 10 collection year groups
- Origin-spawner collections organized into 2 breeding location groups (hatchery versus natural)
- Origin-spawner collections organized into 2 origin groups (hatchery versus natural)
- Origin-spawner collections organized into the 4 origin-spawner groups

It is clear from this analysis that nearly all molecular variation, no matter how the data are organized, resides within a collection (Table 5). The percentage of total molecular variance occurring within collections ranged from $99.68 \%$ to $99.74 \%$. The among group variance component was limited to less than $0.26 \%$ and in all organizational structures,
except "no structure," the among group percentage was not significantly greater than zero. Furthermore, none of the organizational structures provided better resolution than "no structure" in terms of accounting for molecular variance within the data set. These results indicate that if there are significant differences among collections of Chiwawa fish, these differences account for less than one percent of the total molecular variance, and these differences cannot be attributed to fish origin or spawning location.

## Summary and Conclusions

We reject the null hypothesis that the allele frequencies of the hatchery collections equal the allele frequencies of the natural collections, which equals the allele frequency of the donor population. Furthermore, because the allele-sharing distances are not consistent within and among collections years, we also reject the second stated hypothesis discussed above. However, there is an extremely small amount of genetic variance that can be attributed to among collection differences. The allelic differentiation that does exist among collections does not appear to be a function of fish origin, spawning location, genetic drift, or collection year. Figure 5 and related statistics does suggest that hatchery and natural collections in 2005 and 2006 are more similar to each other than previous years' collections, and this would be expected in a successful integrated hatchery supplementation program.

Since each of these collection years are generally composed of four-year-old fish, the differentiation among these collections for the most part is differentiation among specific cohorts. The slightly greater percentage of alleles shared among collections that are separated in time by multiples of four years, compared with collections that are not separated in time as such, suggests that cohort differences may be the most important factor accounting for differences in allele frequencies among collections.

## Task 4: Develop a model of genetic drift.

See Task 3

# Task 5: Analyze spring Chinook population samples from the Chiwawa River and Chiwawa Hatchery from multiple generations. 

See Task 3

## Task 6: Analyze among population differences for upper Wenatchee spring Chinook.

Supplementation of the Chiwawa River spring Chinook population may affect populations within the Wenatchee River watershed other than the Chiwawa River stock. If the stray rate for Chiwawa hatchery-origin fish is greater than that for natural-origin fish, an increase in gene flow from the Chiwawa population into other populations may result. If this gene flow is high enough, Chiwawa River fish may alter the genetic structure of these other populations. Records from field observations indicate that hatchery-origin fish are present in all major spawning aggregates (A.R Murdoch, unpublished data), and these fish are successfully reproducing (Blankenship et al 2006). The intent of this task is to investigate if there have been changes to the genetic structure of the spring Chinook stocks within upper Wenatchee tributaries during the past 15-20 years, and if changes have occurred, are they a function of the Chiwawa River Supplementation Program? Therefore, we ask the following two questions:

1. Are allele frequencies within populations in the upper Wenatchee stable through time? That is, is there significant allelic differentiation among collections within upper Wenatchee populations?
2. Are the recent collections from the upper Wenatchee populations more similar to the Chiwawa population than earlier collections from the same populations?

For this task we analyzed natural spawning collections from the White River (naturalorigin), Little Wenatchee River (natural-origin), Nason Creek (natural-origin), and

Wenatchee mainstem (hatchery-origin), and hatchery collections from Leavenworth NFH and Entiat River NFH (Table 1). We also included in the analysis the natural- and hatchery-origin collections from the Chiwawa River. There are no repeated collections from Leavenworth, Entiat, Little Wenatchee, and Wenatchee mainstem (Table 1), so for many of the analyses we have limited our discussion to the Chiwawa River, White River, and Nason Creek collections. Furthermore, genetic structure of the Little Wenatchee collection, which consisted of only 19 samples, was unexpectedly quite different from the other collections. For example, the $\mathrm{F}_{\text {ST }}$ statistic measures the percent of total molecular variation that can be attributed to differences between populations. The median $\mathrm{F}_{\text {ST }}$ for all pairwise combinations of collections from all populations, except Little Wenatchee (33 populations, 528 individual $\mathrm{F}_{\text {ST }}$ statistics) equals $0.010(1 \%)$, with a range of 0.000 to 0.037 (Table 6). The median $\mathrm{F}_{\text {ST }}$ for the Little Wenatchee paired with all other collections ( 33 individual $\mathrm{F}_{\text {ST }}$ statistics) equals 0.106 ( $10.6 \%$ ), with a range of 0.074 to 0.121 . The ten-fold increase in the $\mathrm{F}_{\text {ST }}$ statistic indicates that either the Little Wenatchee spring Chinook is unique among the upper Wenatchee River stocks, or this 1993 collection is somehow aberrant. Therefore, we exclude the Little Wenatchee collection from many other analyses.

Population Differentiation - Table 3 provides the levels of significance for all pairwise genic differentiation tests. Most between-collection comparisons are highly significant, with no pattern of increasing or decreasing differentiation with time, and no differences when comparisons are made with Chiwawa hatchery- versus Chiwawa natural-origin fish. For example, excluding the outlier 1996 and 1998 Chiwawa hatchery- and naturalorigin collections, Nason Creek showed highly significant allele frequency differences between the Chiwawa hatchery- and natural-origin collections at $100 \%$ and $86 \%$ of the comparisons, respectively. The same comparisons with the White River produced 100\% and $93 \%$ highly significant allele frequency comparisons, respectively. Allele frequencies between Nason Creek and White River were likewise differentiated from each other.

The collection allele frequencies within the upper Wenatchee system are significantly different, and these differences do not appear to change as a function of time (Table 3). Nason Creek shows greater within-population year-to-year variation in allele frequencies than does the White River, with 47\% of the pairwise comparisons showing highly significant differences, compared with only $13 \%$ for the White River. However, the 2005 and 2006 collections from the White River appear to be somewhat more differentiated from not only each other, but from the earlier collections from the White River.

Despite the high degree of temporal and spatial structure suggested by the genic differentiation tests, as described above for within-Chiwawa analysis (Task 3), most of the genetic variation within this data set occurs within populations, rather than between populations (Table 6). The $\mathrm{F}_{\text {ST }}$ values for most population comparisons are between 0.01 and 0.02 , indicating $1 \%$ to $2 \%$ among-population variance, with the remaining $98 \%$ to 99\% variance occurring within populations. The White River shows the highest median $\mathrm{F}_{\text {ST }}$ among the natural-origin collections, equal to 0.014 , compared with 0.009 for both the Nason Creek and Chiwawa natural-origin collections. The median FST for the Chiwawa hatchery-origin collections (0.012) was higher than that for the Chiwawa natural-origin collections.

Table 7 summarizes the information from the $\mathrm{F}_{\text {ST }}$ analyses, under five different temporal and spatial scenarios. Under all scenarios, over $99 \%$ of the molecular variance is within populations. There is significantly greater spatial structure among populations ("Origin") in 2005 and 2006 than from 1989 to 1996. That is, there appears to be more spatial structure among the Chiwawa hatchery-origin, Chiwawa natural-origin, White River, and Nason Creek now, than in 1989 to 1996, despite the potential homogenizing and cumulative effect of hatchery strays. However, we stress that the amount of molecular variance associated with the among population differences, despite being significantly greater than $0.00 \%$, is limited to only $0.43 \%$.

Allele-sharing and Nonmetric Multidimensional Scaling - As in the Chiwawa River data discussed above, we constructed an allele-sharing distance matrix and then subjected
that matrix to a multidimensional scaling analysis (Figure 7). Consistent with all previously discussed multidimensional scaling analyses, the 1996 and 1998 adult, and the 1994 smolt collections are outliers. There is clear separation between the White River collections and all other natural-origin and Chiwawa hatchery-origin collections, indicating that there are more alleles shared among the Nason Creek and Chiwawa collections, than with the White River collections. Furthermore, there is a slight separation between the Chiwawa natural-origin natural spawner collections and Nason Creek collections, suggesting different groups of shared alleles between these populations. There is more variation in the allele-sharing distances among collections involved with the Chiwawa hatchery (origin or broodstock) than any of the natural-origin collections, even if we exclude the 1994, 1996, and 1998 collections. This suggests that there is more year-to-year variation in the composition of hatchery-origin and hatchery broodstock than within natural-origin populations throughout the upper Wenatchee. All Wenatchee mainstem fish are hatchery-origin, and if these fish are from the Chiwawa Supplementation Program (rather than from Leavenworth), it is not unexpected that this collection would be plotted within the Chiwawa polygon (Figure 7).

Assignment of Individual to Populations - Finally, we conducted individual assignment tests whereby we assigned each individual fish to a population, based on a procedure developed by Rannala and Mountain (1997) (Table 8 and 9). Individual fish may be correctly assigned to the population from which they were collected, or incorrectly assigned to a different population. Incorrect assignments may occur if the fish is an actual migrant (i.e., source population different from population where collected), or because the genotype for that fish matches more closely with a population different from its source. If there are many individuals from a population incorrectly assigned to populations other than its source population, that original population is either unreal (i.e., an admixture), or there is considerable gene flow between that population and other populations. Furthermore, in assigning individuals to populations, we can either accept the assignment with the highest probability, regardless of how low that probability may be, or we can establish a more stringent criterion, such as to not accept an assignment unless the posterior probability is equal to or greater than 0.90 . This value is roughly
equal to having the likelihood of the most-likely population equal to 10 times that of the second most-likely population.

We provide a summary of the assignments in Tables 8 and 9. On average, nearly $50 \%$ of the fish are assigned incorrectly if we accept all assignments (Table 8), but the incorrect assignment rate drops to roughly $10 \%$ when we accept only those assignments with probabilities greater than 0.90 . However, with this more stringent criterion, nearly $64 \%$ of the fish go unassigned. These results indicate that the allele frequency distributions for these populations are very similar, and it would be very difficult to assign an individual fish of unknown origin to the correct population. If all fish are assigned, there is a $50 \%$ chance, overall, of a correct assignment. If you accept only those assignment with the 0.90 criterion, nearly two-thirds of the fish would be unassigned, but there is a $90 \%$ chance of correctly assigning those fish that are indeed assigned.

Of all the populations in the data set, there are fewer errors associated with assigning fish to the White River. If all fish are assigned (Table 8), $72 \%$ of those fish assigned to the White River, are actually from the White River (115 fish out of a total of 159 fish assigned to the White River). This compares to a rate of only $52 \%$ and $53 \%$ for Nason Creek and Chiwawa natural-origin, respectively, and $60 \%$ for the Chiwawa hatcheryorigin collections. With the 0.90 criterion (Table 9), $89 \%$ of the fish assigned to the White River, are actually from the White River, compared with $70 \%$ and $65 \%$ for Nason Creek and Chiwawa natural origin, respectively, and $81 \%$ for the Chiwawa hatchery origin.

When all fish are assigned, most of the incorrectly assigned fish from Nason Creek and White River are assigned to Chiwawa River, at roughly equal frequencies to the hatcheryand natural-origin populations. Incorrectly assigned fish to other populations occur at a slightly higher rate in Nason Creek than in the White River. However, when only those fish meeting the 0.90 criterion are assigned (Table 9), incorrectly assigned fish from Nason Creek are distributed among White and Chiwawa Rivers, as well as Leavenworth NFH, and the Entiat NFH. Mis-assignment to the Chiwawa hatchery-origin was the
highest among the Nason Creek collections, equal to nearly $14 \%$. This contrasts with the White River where mis-assignments do not exceed $7 \%$ anywhere, and there is a roughly even distribution of mis-assignments among Nason Creek and Chiwawa River collections.

Summary and Conclusions - There is little geographic or temporal structure among populations within the upper Wenatchee systems. Among population molecular variance is limited to $1 \%$ or less. The little variance that can be attributed to among populations indicates that the White River is more differentiated from the Chiwawa and Nason populations than these populations are from each other. Furthermore, although we cannot rule out a hatchery effect on the Nason Creek and White River populations, there is no indication there has been any temporal changes in allele frequencies within these populations that can be attributed directly to the Chiwawa River Supplementation Program. In fact, Table 7 weakly suggests that there is more differentiation among these populations now, than there was before or at the early stages of Chiwawa supplementation.

Therefore, returning to our two original questions, there are significant differences in allele frequencies among collections within populations, and among populations within the upper Wenatchee spring Chinook stocks. However, these differences account for a very small portion of the overall molecular variance, and these populations overall are very similar to each other. There is no evidence that the Chiwawa River Supplementation Program has changed the allele frequencies in the Nason Creek and White River populations, despite the presence of hatchery-origin fish in both these systems. Finally, of all the populations within the Wenatchee River, the White River appears to be the most distinct. Yet, this distinction is more a matter of detail than of large significance, as the median $\mathrm{F}_{\text {ST }}$ between White River collections and all other collections (except the Little Wenatchee) is less than $1.5 \%$ among population variance.

## Task 7: Calculate the inbreeding effective population size using demographic data for each sample year, and document the ratio of census to effective size.

This analysis was completed by Williamson et al. (submitted).

## Task 8: Calculate LD $\mathrm{N}_{\mathrm{b}}$ using genetic data for each sample year, and document the ratio of census to effective size.

We report $\mathrm{N}_{\mathrm{e}}$ estimated for the Chiwawa River collections based on the bias correction method of Waples (2006) implemented in LDNe (Do and Waples unpublished). $\mathrm{N}_{\mathrm{e}}$ estimates based on LD are best interpreted as the effective number of breeders $\left(\mathrm{N}_{\mathrm{b}}\right)$ that produced the sample (Waples 2006).

For collections categorized by spawning location (i.e., hatchery broodstock or natural), estimates of $\mathrm{N}_{\mathrm{b}}$ are shown in Table 10. Considering the hatchery broodstock, $\mathrm{N}_{\mathrm{b}}$ estimates range from 30.4 (1996) to 274.3 (2005). To obtain $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratios, the $\mathrm{N}_{\mathrm{b}}$ estimate is multiplied by four (i.e., mean generation length) and divided by the total in river (i.e., NOS [natural-origin spawners] plus HOS [hatchery-origin spawners]) census data from four years prior (i.e., major cohort; see Table 2). The observed $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratios for the broodstock collections range from $11 \%$ to $54 \%$ of the census estimate, excluding the 2000 collection which is $106 \%$. A ratio greater than one is possible under special circumstances, and certain artificial mating schemes within hatcheries can inflate $\mathrm{N}_{\mathrm{e}}$ above N ; yet, it is unknown if this is the case for this collection. While no direct comparisons are possible, the $\mathrm{N}_{\mathrm{b}}$ estimates reported by Williamson et al. (submitted) for Chiwawa broodstock collections from 2000-2003 are similar in magnitude to our estimates. For Chiwawa natural spawner collections, the $\mathrm{N}_{\mathrm{b}}$ estimates range from 5.2 (1989) to 231.5 (2005), with observed $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratios of $22 \%-48 \%$ of the census estimate.

## Task 9: Calculate $\mathbf{N}_{\mathrm{b}}$ using the temporal method for multiple samples from the same location.

Estimates of effective number of breeders $\left(\mathrm{N}_{\mathrm{b}}\right)$ derived from Waples' (1990) temporal method are shown in Tables 11-13. Eight collection years were used for the Chiwawa broodstock collections (Table 11). The harmonic mean of all pairwise estimates of $\mathrm{N}_{\mathrm{b}}($ $\tilde{\mathrm{N}}_{\mathrm{b}}$ ) was 269.4. This estimate is the contemporary $\mathrm{N}_{\mathrm{e}}$ for Chiwawa broodstock collections. For the five collection years of Chiwawa in-river spawners (Table 12), the estimated $\tilde{\mathrm{N}}_{\mathrm{b}}=224.2$. This estimate is the contemporary $\mathrm{N}_{\mathrm{e}}$ for Chiwawa River natural spawner collections. Since the Chiwawa Supplementation Program is integrated by design, we also performed another estimation of $\mathrm{N}_{\mathrm{e}}$ using composite hatchery and natural samples. There are paired samples from 2004-2006. We combined genetic data for hatchery (HOS) and natural (NOS) origin fish from 2004-2006 to create a single Chiwawa River natural spawner sample for each year. The three composite samples from 2004 - 2006 were then analyzed using the temporal method (Table 13), resulting in a $\widetilde{\mathrm{N}}_{\mathrm{b}}$ $=386.8$. This estimate is the contemporary $\mathrm{N}_{\mathrm{e}}$ for Chiwawa River.

Williamson et al. (submitted) estimated $\mathrm{N}_{\mathrm{e}}$ using Waples' (1990) temporal method for Chinook captured in 2004 and 2005, and used age data to decompose brood years into consecutive cohorts from 2000-2003. They report for Chiwawa broodstock a $\tilde{\mathrm{N}}_{\mathrm{b}}=$ 50.4. This estimate is not similar to our Chiwawa broodstock estimate. However, if we analyze the hatchery-origin Chinook only, our estimate is $\widetilde{\mathrm{N}}_{\mathrm{b}}=80.1$ for collection years 1989 - 2006 (data not shown). Williamson et al. (submitted) report for Chiwawa naturally spawning Chinook a $\tilde{\mathrm{N}}_{\mathrm{b}}=242.7$, which is slightly higher than our estimate for in-river spawners from 1989 - 2006, but lower than our estimate from combined NOS and HOS Chinook from 2004 - 2006 collection years.

## Task 10: Use available data and the Ryman-Laikre and Wang-Ryman models to determine the expected change of $\mathrm{N}_{\mathrm{e}}$ for natural spring Chinook salmon in the Wenatchee River due to hatchery operation.

$\mathrm{N}_{\mathrm{e}}$ is generally thought to be between 0.10 and 0.33 of the estimated census size (Bartley et al. 1992; RS Waples pers. comm.). We used this range to generate an estimate of $\mathrm{N}_{\mathrm{e}}$ for Chiwawa natural spawners prior to hatchery operation. For brood years 1989 - 1992, the arithmetic mean census size was $\mathrm{N}=962.7$ (Table 2), resulting in an estimated $\mathrm{N}_{\mathrm{e}}$ ranging from 96.3 - 317.7. The contemporary estimate of $\mathrm{N}_{\mathrm{e}}$ calculated using genetic data for the Chiwawa in-river spawners is $\mathrm{N}_{\mathrm{e}}=224.2$ (Table 12), falling in the middle of the pre-hatchery range. The $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratio calculated using 224.2 and the arithmetic census of NOS Chinook from 1989 - 2005 is 0.42 . A more appropriate contemporary $\mathrm{N}_{\mathrm{e}}$ to compare with the pre-hatchery estimate (i.e., $96.3-317.7$ ) is the combined NOS and HOS estimate from natural spawners, since the supplementation program is integrated. As discussed above, the contemporary estimate of $\mathrm{N}_{\mathrm{e}}$ calculated using genetic data for Chiwawa NOS and HOS Chinook is $\mathrm{N}_{\mathrm{e}}=386.8$ (Table 13), which is slightly larger than the pre-hatchery range, suggesting the $\mathrm{N}_{\mathrm{e}}$ has not declined during the period of hatchery operation. The $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ ratio calculated using 386.8 and the arithmetic census of NOS and HOS Chinook from 1989 - 2005 is 0.40 . These results suggest the Chiwawa Hatchery Supplementation Program has not resulted in a smaller $\mathrm{N}_{\mathrm{e}}$ for the natural spawners from the Chiwawa River.

Williamson et al. (submitted) argued that since their combined (i.e., broodstock and natural) $\mathrm{N}_{\mathrm{e}}$ estimate was lower than the naturally spawning estimate, the supplementation program likely had a negative impact on the Chiwawa River $\mathrm{N}_{\mathrm{e}}$. We disagree with this interpretation of these data. Since the natural spawning component is mixed hatchery and natural ancestry, the $\mathrm{N}_{\mathrm{e}}$ estimates from natural spawning data are the results that bear on possible hatchery impacts. The census data show the population declined in the mid 1990's and rebounded by 2000 (Table 2). This trend is reflected in the $\mathrm{N}_{\mathrm{e}}$ results, as shown above, and Williamson et al. (submitted) clearly show in their Table 4 the $\mathrm{N}_{\mathrm{e}}$ was lower in $2000\left(\mathrm{~N}_{\mathrm{e}}=989\right)$ than it was in $1992\left(\mathrm{~N}_{\mathrm{e}}=2683\right)$. Yet, the important comparison
they make in our view was the natural spawning $\mathrm{N}_{\mathrm{e}}$ versus the natural only component $\mathrm{N}_{\mathrm{e}}$ (i.e., hypothetically excluding hatchery program). Williamson et al. (submitted) report the 1989 - $1992 \mathrm{~N}_{\mathrm{e}}$ estimated from naturally spawning Chinook (i.e., NOS and HOS integrated) was essentially the same as the natural only component estimate, 2683 and 2776, respectively. This result is not surprising since no HOS fish were present between 1989 - 1992. They also report that the $1997-2000 \mathrm{~N}_{\mathrm{e}}$ estimated from naturally spawning Chinook (i.e., NOS and HOS integrated) was $\mathrm{N}_{\mathrm{e}}=989$, while the natural-origin estimate of $\mathrm{N}_{\mathrm{e}}$ in $1997-2000$ was $\mathrm{N}_{\mathrm{e}}=629$. Since the natural-origin estimate of 629 is lower than 989 , the $\mathrm{N}_{\mathrm{e}}$ estimate from all in-river spawners, we argue that their analysis of demographic data show the $\mathrm{N}_{\mathrm{e}}$ estimated from naturally spawning Chinook (i.e., NOS and HOS integrated) is larger only if the hatchery Chinook in the river are ignored.

## Task 11: Use individual assignment methods to determine the power of self-assignment for upper Wenatchee River tributaries.

See "Assignment of Individual to Populations" in Task 6

## Conclusions

Has the Chiwawa Hatchery Supplementation Program succeeded at increasing the census size of the target population while leaving genetic integrity intact? This is an important question, as hatcheries can impact natural populations by reducing overall genetic diversity (Ryman and Laikre 1991), reducing the fitness of the natural populations through relaxation of selection or inadvertent positive selection of traits advantageous in the hatchery (Ford 2002; Lynch and O’Hely 2001), and by reducing the reproductive success of natural populations (McLean et al. 2003). The census data presented here show that the current natural spawning census size is similar to the pre-supplementation census size. Despite large numbers of hatchery-origin fish on the Chiwawa River spawning grounds, the genetic diversity of the natural-origin collections appear unaffected by the supplementation program; heterozygosities are high, and contemporary $\mathrm{N}_{\mathrm{e}}$ is similar (perhaps slightly higher) than pre-supplementation $\mathrm{N}_{\mathrm{e}}$. We did find
significant year-to-year differences in allele frequencies in both the origin and spawner datasets, but these differences do not appear to be related to fish origin, spawning area, or genetic drift. However, we do suggest that cohort differences may be the most important factor accounting for differences in allele frequencies among collections.

The main objective of this study was to determine the potential impacts of the hatchery program on natural spring Chinook in the upper Wenatchee system. We did this by analyzing temporally replicated collections from the Chiwawa River, and by comparing genetic diversity prior to the presumed effect of the Chiwawa Hatchery Supplementation Program, with contemporary collections. We report that the genetic diversity present in the Chiwawa River is unchanged (allowing for differences among cohorts) from 1989 2006, and the contemporary estimate of the effective population size $\left(\mathrm{N}_{\mathrm{e}}\right)$ using genetic data is approximately the same as the $\mathrm{N}_{\mathrm{e}}$ estimate extrapolated from 1989 - 1992 census data (i.e., pre-hatchery collection years). We observed substantial genetic diversity, with heterozygosities $\sim 80 \%$ over thirteen microsatellite markers. Yet, temporal variation in allele frequencies was the norm among temporal collections from the same populations (i.e., location). The genetic differentiation of replicated collections from the same population is likely the result of salmon life history in this area, as four-year-old Chinook comprise a majority of returns each year. The genetic tests are detecting the differences of contributing parents for each cohort. An important point related to the temporal variation, is that the hatchery broodstock is composed in part of the natural origin Chinook from the Chiwawa River. When we compared the genetic data (within a collection year) for Chinook brought into the hatchery as broodstock with the Chinook that remained in the river (years 2001, 2004 - 2006), there was a trend of decreasing statistical differences in allele frequencies from 2001 to 2004, and no differences were detected for 2005 and 2006. While the replicated collections may have detectable differences in allele frequencies, those differences reflect actual differences in cohorts, not the result of hatchery operations, and the hatchery broodstock collection method captures the differences in returning Chiwawa River spring adults each year. We conclude from these results that the genetic diversity of natural spring Chiwawa Chinook has been maintained during the Chiwawa Hatchery Supplementation Program.

We observe slight, but statistically significant population differentiation between Chiwawa River, White River, and Nason Creek collections. Murdoch et al (2006) and Williamson et al. (submitted) also observed population differentiation between Chiwawa River, White River, and Nason Creek collections. Yet, $99.3 \%$ of the genetic variation observed was within samples, very little variance could be attributed to population differences (i.e., population structure). The AMOVA analysis and poor individual assignment results suggest the occurrence of gene flow among Wenatchee River locations or a very recent divergence of these groups. While Murdoch et al. 2006 did not perform an AMOVA analysis, their $\mathrm{F}_{S T}$ results provide comparable data to our amongpopulation results. Murdoch et al. 2006 report $\mathrm{F}_{\text {ST }}$ ranging from $2 \%-3 \%$ for pairwise comparisons between of Chiwawa, White, and Nason River collections. Since FST is an estimate of among-sample variance, these results also imply a majority of the genetic variance (i.e., $97 \%-98 \%$ ) resides within collections. To provide further context for the magnitude of these variance estimates, we present the among-group data from Murdoch et al. 2006 comparing summer-run and spring-run Chinook from the Wenatchee River. They report that approximately $91 \%$ of observed genetic variance is within-collection for comparisons between collections of summer- and spring-run Chinook. Ultimately, the information provided by this and other reports will be incorporated into the management process for Wenatchee River Chinook. However, we would like to emphasize that the application of these genetic data to management is more about the goals related to the distribution of genetic diversity in the future than specific data values reported. If Chinook are collected at Tumwater Dam instead of within the upper Wenatchee River tributaries, a vast majority of the genetic variation present in the basin would be captured, although any differences among tributaries would be mixed. Alternatively, management policies could be crafted to promote and maintain the among-group genetic diversity that genetic studies consistently observe to be non-zero within the Wenatchee River.

We agree with Murdoch et al. (2006) that it appears hatchery Chinook are not contributing to reproduction in proportion to their abundance. Additionally, if the total census size (i.e., NOS and HOS combined) within the Chiwawa River does not continue
to increase, genetic diversity may decline within this system, given the smaller $\mathrm{N}_{\mathrm{e}}$ within the hatchery-origin collections compared with the natural-origin collections.

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## Literature Cited

Araki H, Waples RS, Ardren WR, Cooper B, Blouin MS (2007) Effective population size of steelhead trout: influence of variance in reproductive success, hatchery programs, and genetic compensation between life-history forms. Molecular Ecology, 16:953966.

Arden WR and Kapuscinski AR (2003) Demographic and genetic estimates of effective population size $\left(\mathrm{N}_{\mathrm{e}}\right)$ reveals genetic compensation in steelhead trout. Molecular Ecology 12: 35-49

Banks MA, Blouin MS, Baldwin BA, Rashbrook VK, Fitzgerald HA, Blankenship SM, Hedgecock D (1999) Isolation and inheritance of novel microsatellites in chinook salmon (Oncorhynchus tschawytscha). Journal of Heredity, 90:281-288.

Banks MA, Rashbrook VK, Calavetta MJ et al (2000) Analysis of microsatellite DNA resolves genetic structure and diversity of chinook salmon (Oncorhynchus tshawytscha) in California's Central Valley. Canadian Journal of Fisheries and Aquatic Sciences 57:915-927.

Bartley D, Bentley B, Brodziak J, Gomulkiewicz R, Mangel M, and Gall GAE (1992) Geographic variation in population genetic structure of chinook salmon from California and Oregon. Fish. Bull., U.S. 90:77-100.

Blankenship SM, Von Bargen J, and Truscott KD (2006) Genetic analysis of White River juveniles retained for captive brood at AquaSeed to assess the hatchery status of contributing parents. Developed for Grant County PUD.

Blankenship SM and Murdoch AR (2006) Study Plan For Assessing the Genetic Diversity of Natural Chiwawa River Spring Chinook Salmon And Evaluating The Effectiveness Of Its Supportive Hatchery Supplementation Program. Developed for Chelan County PUD and the Habitat Conservation Plan's Hatchery Committee.

Bugert R (1998) Mechanics of supplementation in the Columbia River. Fisheries 23:1120.

Cairney M, Taggart JB, Hoyheim B (2000) Characterization of microsatellite and minisatellite loci in Atlantic salmon (Salmo salar L.) and cross-species amplification in other salmonids. Mol Ecol, 9:2175-2178.

Campton DE (1987) Natural hybridisation and introgression in fishes: methods of detection and genetic interpretations. In: Population genetics and fisheries management. (Eds. Ryman, N. and Utter, F.), pp. 161-192. Washington Sea Grant Program, University of Washington Press, Seattle, USA.

Chapman D, Giorgi A, Hillman T, Deppert D, Erho M, Hays S, Peven C, Suzumoto B, and Klinge R (1994) Status of summer/fall chinook salmon in the mid-Columbia region. Report for Chelan, Douglas, and Grant County PUDs. 412 p. + app. (Available from Don Chapman Consultants, 3653 Rickenbacker, Ste. 200, Boise, ID 83705.)

Chapman D, Peven C, Giorgi A, Hillman T, and Utter F (1995) Status of spring chinook salmon in the mid-Columbia River. Don Chapman Consultants, Inc., 477 p. (Available from Don Chapman Consultants, 3653 Rickenbacker, Ste. 200, Boise, ID 83705.)

Cockerham CC (1969) Variance of gene frequencies. Evolution 23:72-83.
Craig JA, and Suomela (1941) Time of appearance of the runs of salmon and steelhead trout native to the Wenatchee, Entiat, Methow, and Okanogan rivers. U.S. Fish Wildl. Serv.

Fish FF, and Hanavan MG (1948) A report on the Grand Coulee Fish Maintenance. Project 1939-1947. U.S. Fish Wildl. Serv. Spec. Sci. Rep 55.

Ford MJ (2002) Selection in captivity during supportive breeding may reduce fitness in the wild. Conservation Biology 16(3):815-825

Frankham R, Ballou JD, Briscoe DA (2002). Introduction to Conservation Genetics, Cambridge University Press, Cambridge, UK.

Greig C, Jacobson DP, Banks MA (2003) New tetranucleotide microsatellites for finescale discrimination among endangered chinook salmon (Oncorhynchus tshawytscha). Mol Ecol Notes, 3:376-379.

Heath DD, Busch C, Kelly J, and Atagi DY (2002) Temporal change in genetic structure and effective population size in steelhead trout (Oncorhynchus mykiss). Molecular Ecology 11:197-214

Hedrick P, Hedgecock D (1994) Effective population size in winter-run chinook salmon. Conservation Biology, 8:890-892.

Hill WG (1981) Estimation of effective size from data on linkage disequilibrium. Genetical Research 38: 209-216.

Jensen LF, Hansen MM, Carlsson J et al (2005) Spatial and temporal genetic differentiation and effective population size of brown trout (Salmo trutta, L.) in small Danish rivers. Conservation Genetics 6:615-621.

Liu, K and Muse SV (2005) PowerMarker: Integrated analysis environment for genetic marker data. Bioinformatics 21:2128-2129.

Lynch M and O'Hely M (2001) Captive breeding and the genetic fitness of natural populations. Conservation Genetics 2:363-378

Manly, BFJ. (1986) Multivariate Statistical Methods. A Primer. Chapman and Hall. London.. 159 + x pp.

McLean JE, Bentzen P, Quinn TP (2003) Differential reproductive success of sympatric, naturally spawning hatchery and wild steelhead trout (Oncorhynchus mykiss) through the adult stage. Can. J. Fish. Aquat. Sci., 60:433-440.

Mullan, JW (1987) Status and propagation of chinook salmon in the mid-Columbia River through 1985. U.S. Fish Wildl. Serv. Biol. Rep. 87:111.

Murdoch AR and Peven C (2005) Conceptual Approach to Monitoring and Evaluating the Chelan County Public Utility District Hatchery Programs, Final Report.

Murdoch AR, Pearsons TN, Maitland TW, Ford M, and Williamson K (2006) Monitoring the reproductive success of naturally spawning hatchery and natural spring Chinook salmon in the Wenatchee River. BPA Project No. 2003-039-00, Contract No. 00021391 . pp. 96

Nelson WR, and Bodle J (1990) Ninety years of salmon culture at the Little White Salmon National Fish Hatchery. U.S. Fish Wildl. Serv. Biol. Rep. 90:22.

Palm S, Laikre L, Jorde PE, et al (2003) Effective population size and temporal genetic change in stream resident brown trout (Salmo trutta, L.). Conservation Genetics 4:249-264.

Olsen JB, Bentzen P, Seeb JS (1998) Characterization of seven microsatellite loci derived from pink salmon. Molecular Ecology, 7:1087-1089

Rannala B, Mountain JL (1997) Detecting immigration by using multilocus genotypes. Proceedings of the National Academy of Sciences 94:9197-9201.

Rexroad CE, Coleman RL, Martin AM, Hershberger WK, Killefer J (2001) Thirty-five polymorphic microsatellite markers for rainbow trout (Oncorhynchus mykiss). Animal Genetics, 32:317-319.

Rohlf, F. J. (2002) NTSYSpc: Numerical Taxonomy System, ver. 2.1. Exeter Publishing, Ltd.

Ryman N, Laikre L (1991) Effects of supportive breeding on the genetically effective population size. Conservation Biology, 5:325-329.

Seeb L, et al. (in review) Development of a Standardized DNA Database for Chinook Salmon. Fisheries

Schneider S, Roessli D, Excoffier L (2000) Arlequin ver 2.000: A software for population genetic data analysis. Genetics and Biometry Laboratory. University of Geneva, Switzerland.

The Mathworks (2006) MatLab Release R2006b. Massachusetts.
Utter FM, Chapman DW, and Marshall AR (1995) Genetic population structure and history of chinook salmon of the Upper Columbia River. Am. Fish. Soc. Symp. 17:149-165.

Wang J (2005) Estimation of effective size from data on genetic markers. Trans. Royal. Phil. Soc. B 360: 1395-1409.

Wang J, and Ryman N (2001) Genetic effects of multiple generations of supportive breeding. Conservation Biology 15: 1615-1631.

Wang J, Whitlock MC (2003) Estimating Effective Population Size and Migration Rates From Genetic Samples Over Space and Time. Genetics 163:429-446

Waples RS (1989) A generalized approach for estimating effective population size from temporal changes in allele frequency. Genetics, 121:379-391.

Waples RS (1990) Conservation genetics of Pacific salmon. III. Estimating effective population size. J. Hered. 81: 277-289.

Waples RS (1991) Genetic interactions between hatchery and wild salmonids: Lessons from the Pacific Northwest. Can. J. Fish. Aquat. Sci. 48(Suppl. 1):124-133.

Waples RS (2005) Genetic estimates of contemporary effective population size: to what time periods do the estimates apply? Molecular Ecology, 14:3335-3352

Waples RS (2006) A bias correction for estimates of effective population size based on linkage disequilibrium at unlinked gene loci. Conservation Genetics 7:167-184

Washington Department of Fisheries (WDF). 1934. Forty-second and forty-fifth inclusive annual reports of the State Department of Fisheries for the period from April 1, 1931-March 31, 1935, fiscal years of 1931 to 1934 inclusive. Wash. Dep. Fish., pp. 78

Williamson K, Cordes J, May B (2002) Characterization of microsatellite loci in chinook salmon (Oncorhynchus tshawytscha) and cross-species amplification in other salmonids. Molecular Ecology Notes, 2:17-19.

Williamson KS, Murdoch AR, and Ford MJ (submitted) Influence of supportive breeding on genetic diversity of hatchery and natural Wenatchee River spring Chinook salmon.


Figure 1. Conceptual process for evaluating potential changes in genetic variation in the Chiwawa naturally produced populations as a result of the supplementation hatchery programs (From Murdoch and Peven 2005).


Figure 2. Multidimensional scaling plot from an allele-sharing distance matrix calculated from the Chiwawa data set organized by fish origin (i.e., hatchery versus natural). The red arrows connect consecutive hatchery-origin collections starting with the first adult collection (1996) and ending with the 2006 collection (see Table 1 for collection years).


Figure 3. Relationships between the time interval in years and allele sharing distances, with each circle representing the pairwise relationship between two Chiwawa collections. Separate regression lines for the natural- and hatchery-origin collections. The slope for the natural-origin collection is not significantly different from zero ( $\mathrm{p}=0.1483$ ), while the slope for hatchery-origin collection is significantly greater than zero ( $\mathrm{p}=0.0254$ ) indicating a positive relationship between time interval and allele sharing distance.


Figure 4. Multidimensional scaling plot from an allele-sharing distance matrix calculated from the Chiwawa data set organized by four treatment groups, as discussed in the text. Each circle represents a single collection within each of the four treatment groups, and the polygon encloses all groups that are not outliers. Each outlier group is specifically labeled.


Figure 5. As in Figure 4, but allele-sharing distance matrix recalculated without the five outlier groups shown in Figure 4. Polygons group together treatment groups from the same collection year. Dates associated with symbols also refer to collection year. Collection years 2004-2006 included all four treatment groups, while collection year 2001 did not include a hatchery-origin natural spawner group. Legend is read as follows: Open circles refer to hatchery-origin hatchery spawner group, while filled box refers to natural-origin hatchery spawner group, and so on.


Figure 6. Principal component (PC) analysis of individual fish from the Chiwawa River. Only fish with complete microsatellite genotypes were included in the analysis ( $\mathrm{n}=757$ ). Open circles are the PC scores for individual fish, and the filled circles are the centroids (bivariate means) for each of the 25 groups discussed in the text. PC axes 1 and 2 account for only $10.5 \%$ of the total molecular variance.


Figure 7. Multidimensional scaling plot from an allele-sharing distance matrix calculated from the Chiwawa origin data set and all other non-Chiwawa collections, except Little Wenatchee River. Legend is read with abbreviations beginning with origin and then spawning location. $\mathrm{H}=$ hatchery, $\mathrm{N}=$ natural, and $\mathrm{S}=$ smolts. Polygons with solid lines enclose the naturalorigin natural spawner collections from each population (i.e., river). The polygon with the dotted lines enclose all Chiwawa collections, except for the five outlier collections, as discussed in text.

Table 1 Summary of within population genetic data. Chiwawa collection data are summarized in A) by origin of the sample (i.e., clipped vs. non-clipped). All collection data are summarized in B) by spawning location (i.e., hatchery broodstock or on spawning grounds). Hz is heterozygosity, HWE is the statistical significance of deviations from Hardy-Weinberg expectations $(*=0.05, * *=0.01$, and $* * *=0.001$ ), LD is the proportion of pairwise locus tests (across all populations) exhibiting linkage disequilibrium (bolded values are statistically significant), and the last column is mean number of alleles per locus.

|  | Sample <br> size | Gene <br> Diversity | Observed <br> Hz | HWE | F IS | LD | Mean \# <br> Alleles |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |
| Collection |  |  |  |  |  |  |  |
| A) Origin |  |  |  |  |  |  |  |
| 1993 Chiwawa Hatchery | 95 | 0.77 | 0.79 | $* * *$ | -0.02 | $\mathbf{0 . 8 6}$ | 14.00 |
| 1994 Chiwawa Hatchery | 95 | 0.76 | 0.77 | $* * *$ | -0.01 | $\mathbf{0 . 9 1}$ | 11.38 |
| 1996 Chiwawa Hatchery | 8 | 0.75 | 0.81 | - | -0.01 | 0.00 | 8.23 |
| 1998 Chiwawa Hatchery | 27 | 0.81 | 0.82 | - | 0.00 | 0.04 | 12.62 |
| 2000 Chiwawa Hatchery | 43 | 0.75 | 0.78 | $* * *$ | -0.01 | $\mathbf{0 . 1 9}$ | 12.46 |
| 2001 Chiwawa Hatchery | 69 | 0.77 | 0.80 | $* * *$ | -0.02 | $\mathbf{0 . 1 4}$ | 15.31 |
| 2004 Chiwawa Hatchery | 72 | 0.77 | 0.77 | $* * *$ | 0.01 | $\mathbf{0 . 4 5}$ | 15.92 |
| 2005 Chiwawa Hatchery | 91 | 0.79 | 0.82 | $*$ | -0.03 | $\mathbf{0 . 0 5}$ | 16.15 |
| 2006 Chiwawa Hatchery | 95 | 0.80 | 0.84 | $* * *$ | -0.05 | $\mathbf{0 . 4 9}$ | 15.85 |
|  |  |  |  |  |  |  |  |
| 1989 Chiwawa Natural | 36 | 0.76 | 0.78 | - | 0.01 | 0.00 | 12.77 |
| 1993 Chiwawa Natural | 62 | 0.78 | 0.81 | - | -0.02 | 0.04 | 15.85 |
| 1996 Chiwawa Natural | 8 | 0.72 | 0.78 | - | -0.02 | 0.00 | 7.54 |
| 1998 Chiwawa Natural | 10 | 0.78 | 0.84 | - | 0.00 | 0.00 | 8.23 |
| 2000 Chiwawa Natural | 39 | 0.78 | 0.79 | $* * *$ | 0.00 | $\mathbf{0 . 1 0}$ | 14.00 |
| 2001 Chiwawa Natural | 75 | 0.78 | 0.80 | - | -0.03 | 0.03 | 15.31 |
| 2004 Chiwawa Natural | 85 | 0.78 | 0.77 | - | 0.02 | 0.01 | 15.77 |
| 2005 Chiwawa Natural | 90 | 0.79 | 0.79 | - | 0.01 | 0.01 | 16.15 |
| 2006 Chiwawa Natural | 96 | 0.80 | 0.81 | - | -0.01 | 0.01 | 16.46 |
|  |  |  |  |  |  |  |  |

Table 1 Within population genetic data analysis summary continued.

|  | Sample <br> size | Gene <br> Diversity | Observed <br> Hz | HW | FIS | LD | Mean \# <br> Alleles |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |

B) Spawning Location

| 1993 Chiwawa Broodstock | 62 | 0.78 | 0.81 | - | -0.02 | 0.00 | 15.85 |
| :--- | :--- | :--- | :--- | :--- | ---: | :--- | :--- |
| 1996 Chiwawa Broodstock | 16 | 0.75 | 0.79 | - | -0.02 | 0.00 | 10.92 |
| 1998 Chiwawa Broodstock | 37 | 0.82 | 0.83 | - | 0.00 | 0.01 | 14.38 |
| 2000 Chiwawa Broodstock | 82 | 0.78 | 0.78 | $* * *$ | 0.00 | $\mathbf{0 . 3 2}$ | 15.62 |
| 2001 Chiwawa Broodstock | 89 | 0.78 | 0.80 | $*$ | -0.02 | $\mathbf{0 . 1 3}$ | 15.77 |
| 2004 Chiwawa Broodstock | 61 | 0.77 | 0.76 | $*$ | 0.02 | $\mathbf{0 . 1 3}$ | 14.92 |
| 2005 Chiwawa Broodstock | 75 | 0.79 | 0.78 | $*$ | 0.02 | 0.01 | 15.85 |
| 2006 Chiwawa Broodstock | 89 | 0.80 | 0.83 | - | -0.03 | $\mathbf{0 . 0 5}$ | 16.46 |
|  |  |  |  |  |  |  |  |
| 1989 Chiwawa River | 36 | 0.76 | 0.78 | - | 0.01 | 0.00 | 12.77 |
| 2001 Chiwawa River | 55 | 0.78 | 0.80 | - | -0.02 | $\mathbf{0 . 0 9}$ | 14.00 |
| 2004 Chiwawa River | 96 | 0.78 | 0.78 | $*$ | 0.01 | $\mathbf{0 . 1 8}$ | 17.23 |
| 2005 Chiwawa River | 106 | 0.79 | 0.82 | $*$ | -0.02 | $\mathbf{0 . 0 6}$ | 16.69 |
| 2006 Chiwawa River | 102 | 0.80 | 0.83 | $* * *$ | -0.03 | $\mathbf{0 . 1 0}$ | 16.77 |
|  |  |  |  |  |  |  |  |
| 1989 White River | 48 | 0.75 | 0.75 | - | 0.01 | 0.01 | 12.85 |
| 1991 White River | 19 | 0.76 | 0.76 | - | 0.03 | 0.00 | 10.92 |
| 1992 White River | 22 | 0.75 | 0.79 | - | -0.02 | 0.01 | 11.00 |
| 1993 White River | 21 | 0.75 | 0.69 | $*$ | 0.10 | 0.00 | 10.15 |
| 2005 White River | 29 | 0.75 | 0.77 | - | -0.01 | 0.03 | 12.23 |
| 2006 White River | 40 | 0.76 | 0.76 | - | 0.01 | 0.04 | 13.38 |
|  |  |  |  |  |  |  |  |

Table 1 Within population genetic data analysis summary continued.

| Collection | Sample <br> size | Gene <br> Diversity | Observed <br> Hz | HW | FIS | LD | Mean \# <br> Alleles |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 Little Wenatchee R. | 19 | 0.84 | 0.85 | - | 0.02 | 0.00 | 11.23 |
| 1993 Nason Creek | 45 | 0.78 | 0.80 | - | -0.01 | 0.01 | 13.77 |
| 2000 Nason Creek | 51 | 0.76 | 0.78 | - | -0.02 | $\mathbf{0 . 1 3}$ | 13.92 |
| 2001 Nason Creek | 41 | 0.79 | 0.81 | - | -0.01 | $\mathbf{0 . 0 8}$ | 14.23 |
| 2004 Nason Creek | 38 | 0.76 | 0.76 | - | 0.02 | 0.03 | 13.23 |
| 2005 Nason Creek | 45 | 0.78 | 0.82 | - | -0.04 | 0.03 | 14.92 |
| 2006 Nason Creek | 48 | 0.80 | 0.82 | - | -0.01 | 0.00 | 15.77 |
| 2001 Wenatchee River | 32 | 0.79 | 0.80 | $*$ | 0.00 | 0.04 | 12.85 |
| 2000 Leavenworth NFH | 73 | 0.80 | 0.82 | $*$ | -0.02 | $\mathbf{0 . 1 5}$ | 16.23 |
| 1997 Entiat NFH | 37 | 0.81 | 0.83 | - | -0.01 | $\mathbf{0 . 0 6}$ | 14.38 |

Table 2 Demographic data for Chiwawa Hatchery and Chiwawa natural spring Chinook salmon. BS is census size of hatchery broodstock, pNOB is the proportion of hatchery broodstock of natural origin, NOS is the census size of natural-origin spawners present in Chiwawa River, HOS is the census size of hatchery-origin spawners present in Chiwawa River, Total is NOS and HOS combined, and pNOS is the proportion of spawners present in Chiwawa River of natural origin.

| Brood Year | Hatchery |  | In River |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | BS | pNOB | NOS | HOS | Total | pNOS |
| 1989 | 28 | 1 | 1392 | 0 | 1392 | 1.00 |
| 1990 | 18 | 1 | 775 | 0 | 775 | 1.00 |
| 1991 | 32 | 1 | 585 | 0 | 585 | 1.00 |
| 1992 | 78 | 1 | 1099 | 0 | 1099 | 1.00 |
| 1993 | 94 | 1 | 677 | 491 | 1168 | 0.58 |
| 1994 | 11 | 0.64 | 190 | 90 | 280 | 0.68 |
| 1995 | 0 | 0 | 8 | 50 | 58 | 0.14 |
| 1996 | 18 | 0.44 | 131 | 51 | 182 | 0.72 |
| 1997 | 111 | 0.29 | 210 | 179 | 389 | 0.54 |
| 1998 | 47 | 0.28 | 134 | 45 | 178 | 0.75 |
| 1999 | 0 | 0 | 119 | 13 | 132 | 0.90 |
| 2000 | 30 | 0.3 | 378 | 310 | 688 | 0.55 |
| 2001 | 371 | 0.3 | 1280 | 2850 | 4130 | 0.31 |
| 2002 | 71 | 0.28 | 694 | 919 | 1613 | 0.43 |
| 2003 | 94 | 0.44 | 380 | 223 | 603 | 0.63 |
| 2004 | 215 | 0.39 | 820 | 788 | 1608 | 0.51 |
| 2005 | 270 | 0.33 | 250 | 1222 | 1472 | 0.17 |

Table 3 Levels of significance for pairwise tests of genic differentiation among all hatchery- and natural-origin collections used in this analysis. HS = highly significant ( $\mathrm{P}<0.000095$; the Bonferroni corrected p-value for an alpha $=0.05$ ); * $=\mathrm{P}<0.05$ (nominal critical value for most statistical test); - = $>0.05$ (not significant). A significant result between pairs of populations indicates that the allele frequencies between the pair are significantly different. Results are read by comparing the collections along the rows to collections along columns. The top block for each section is a symmetric matrix, as it compares collections within the same group.

|  |  | Chiwawa - Hatchery Origin |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1993 | 1994 | 1996 | 1998 | 2000 | 2001 | 2004 | 2005 | 2006 |
|  | 1993 |  | HS | * | HS | HS | HS | HS | HS | HS |
|  | 1994 | HS |  | HS | HS | HS | HS | HS | HS | HS |
|  | 1996 | * | HS |  | * | - | * | - | - | * |
|  | 1998 | HS | HS | * |  | HS | HS | HS | HS | HS |
|  | 2000 | HS | HS | - | HS |  | HS | * | HS | HS |
|  | 2001 | HS | HS | * | HS | HS |  | HS | * | HS |
|  | 2004 | HS | HS | - | HS | * | HS |  | HS | HS |
|  | 2005 | HS | HS | - | HS | HS | * | HS |  | HS |
|  | 2006 | HS | HS | * | HS | HS | HS | HS | HS |  |
|  | 1989 | HS | HS | - | HS | HS | * | HS | HS | HS |
|  | 1993 | HS | HS | - | HS | HS | - | HS | * | HS |
|  | 1996 | * | HS | - | * | - | - | - | - | - |
|  | 1998 | HS | HS | - | - | HS | * | * | * | - |
|  | 2000 | HS | HS | - | HS | HS | HS | * | HS | HS |
|  | 2001 | HS | HS | - | HS | HS | HS | HS | * | HS |
|  | 2004 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2005 | HS | HS | - | HS | HS | * | HS | * | HS |
|  | 2006 | HS | HS | - | * | HS | HS | HS | HS | HS |
| $\begin{aligned} & \text { ¢ } \\ & \text { H/ } \\ & \text { Z } \end{aligned}$ | 1996 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2000 | HS | HS | * | HS | HS | HS | HS | HS | HS |
|  | 2001 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2004 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2005 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2006 | HS | HS | - | * | HS | HS | HS | HS | HS |
| $\stackrel{y}{4}$ | 1989 | HS | HS | HS | HS | HS | HS | HS | HS | HS |
|  | 1991 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 1992 | HS | HS | * | HS | HS | HS | HS | HS | HS |
|  | 1993 | HS | HS | * | HS | HS | HS | HS | HS | HS |
|  | 2005 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2006 | HS | HS | HS | HS | HS | HS | HS | HS | HS |
| $\begin{aligned} & \text { む̀ } \\ & \stackrel{ \pm}{0} \end{aligned}$ | Wen-M | HS | HS | * | HS | HS | * | * | - | HS |
|  | Leaven | HS | HS | * | HS | HS | HS | HS | HS | HS |
|  | Entiat | HS | HS | * | HS | HS | HS | HS | HS | HS |

Table 3 (con't)

|  |  | Chiwawa - Natural Origin |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1989 | 1993 | 1996 | 1998 | 2000 | 2001 | 2004 | 2005 | 2006 |
|  | 1989 |  | - | - | - | - | * | * | * | * |
|  | 1993 | - |  | - | * | * | * | HS | * | HS |
|  | 1996 | - | - |  | - | - | - | - | - | - |
|  | 1998 | - | * | - |  | * | * | HS | * | * |
|  | 2000 | - | * | - | * |  | HS | - | HS | HS |
|  | 2001 | * | * | - | * | HS |  | HS | * | HS |
|  | 2004 | * | HS | - | HS | - | HS |  | HS | HS |
|  | 2005 | * | * | - | * | HS | * | HS |  | * |
|  | 2006 | * | HS | - | * | HS | HS | HS | * |  |
|  | 1996 | * | * | - | * | * | HS | HS | HS | HS |
|  | 2000 | HS | HS | HS | HS | HS | HS | HS | HS | HS |
|  | 2001 | HS | * | - | * | HS | HS | HS | HS | HS |
|  | 2004 | HS | HS | - | HS | HS | HS | HS | HS | HS |
|  | 2005 | * | * | - | * | HS | HS | HS | HS | HS |
|  | 2006 | HS | HS | - | - | HS | HS | HS | HS | HS |
| $\begin{aligned} & \text { \#2 } \\ & \frac{1}{3} \end{aligned}$ | 1989 | HS | HS | * | HS | HS | HS | HS | HS | HS |
|  | 1991 | HS | HS | * | - | HS | HS | HS | HS | HS |
|  | 1992 | HS | HS | - | * | HS | HS | HS | HS | HS |
|  | 1993 | HS | * | - | * | HS | HS | HS | HS | HS |
|  | 2005 | HS | * | * | * | HS | HS | HS | * | HS |
|  | 2006 | HS | HS | * | HS | HS | HS | HS | HS | HS |
| $\begin{aligned} & \text { む } \\ & \text { む } \end{aligned}$ | Wen-M | * | - | - | - | * | * | HS | * | * |
|  | Leaven | HS | HS | * | * | HS | HS | HS | HS | HS |
|  | Entiat | HS | HS | * | HS | HS | HS | HS | HS | HS |

Table 3 (con't)

|  |  | Nason |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1996 | 2000 | 2001 | 2004 | 2005 | 2006 |
|  | 1996 |  | HS | - | HS | - | * |
|  | 2000 | HS |  | HS | HS | HS | HS |
|  | 2001 | - | HS |  | * | - | * |
|  | 2004 | HS | HS | * |  | * | HS |
|  | 2005 | - | HS | - | * |  | - |
|  | 2006 | * | HS | * | HS | - |  |
| $\begin{aligned} & \text { \#ٍ } \\ & \frac{1}{3} \end{aligned}$ | 1989 | HS | HS | HS | HS | HS | HS |
|  | 1991 | * | HS | HS | HS | * | * |
|  | 1992 | HS | HS | HS | HS | HS | HS |
|  | 1993 | * | HS | HS | HS | HS | HS |
|  | 2005 | * | HS | HS | HS | HS | HS |
|  | 2006 | HS | HS | HS | HS | HS | HS |
| $\begin{aligned} & \text { む } \\ & \stackrel{ \pm}{ \pm} \end{aligned}$ | Wen-M | HS | HS | HS | HS | * | HS |
|  | Leaven | HS | HS | HS | HS | HS | HS |
|  | Entiat | HS | HS | HS | HS | HS | HS |

Table 3 (con't)

|  |  | White |  |  |  |  |  | Other |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1989 | 1991 | 1992 | 1993 | 2005 | 2006 | $\begin{array}{\|c\|} \text { Wen-M } \\ 2001 \end{array}$ | $\begin{gathered} \text { Leaven } \\ 2000 \end{gathered}$ | Entiat 1997 |
|  | 1989 |  | - | * | - | HS | HS | HS | HS | HS |
|  | 1991 | - |  | - | - | * | * | * | HS | HS |
|  | 1992 | * | - |  | - | * | * | HS | HS | HS |
|  | 1993 | - | - | - |  | * | * | HS | HS | HS |
|  | 2005 | HS | * | * | * |  | * | HS | HS | HS |
|  | 2006 | HS | * | * | * | * |  | HS | HS | HS |
| $\begin{aligned} & \text { む } \\ & \pm \end{aligned}$ | Wen-M | HS | * | HS | HS | HS | HS |  | HS | HS |
|  | Leaven | HS | HS | HS | HS | HS | HS | HS |  | HS |
|  | Entiat | HS | HS | HS | HS | HS | HS | HS | HS |  |

Table 4 Probabilities (above diagonal) and levels of significance (below diagonal) for pairwise tests of genic differentiation among all Chiwawa hatchery broodstock and Chiwawa natural spawner collections used in this analysis. HS = highly significant ( $\mathrm{P}<0.000476$; the Bonferroni corrected pvalue for an alpha $=0.05$ ); * $=\mathrm{P}<0.05$ (nominal critical value for most statistical test); $-=\mathrm{P}>0.05$ (considered not significant). A significant result between pairs of populations indicates that the allele frequencies between the pair are significantly different. Pairwise comparisons between the hatchery broodstock and natural spawner collections from 2001, 2004, 2005, and 2006, respectively, are highlighted.

|  |  | Smolt |  | Hatchery Broodstock |  |  |  |  |  |  |  | Natural Spawners |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1993 | 1994 | 1993 | 1996 | 1998 | 2000 | 2001 | 2004 | 2005 | 2006 | 1989 | 2001 | 2004 | 2005 | 2006 |
| $\begin{aligned} & \stackrel{\#}{\mathbf{O}} \\ & \underset{\sim}{6} \end{aligned}$ | 1993 | HS 0.0000 |  | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 1994 |  |  | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 1993 | HS | HS |  | 0.9155 | 0.0000 | 0.0073 | 0.3647 | 0.0003 | 0.0694 | 0.0000 | 0.2220 | 0.0039 | 0.0008 | 0.0095 | 0.0000 |
|  | 1996 | HS | HS | - |  | 0.0151 | 0.8388 | 0.0452 | 0.4916 | 0.3189 | 0.0716 | 0.5591 | 0.0759 | 0.8101 | 0.2364 | 0.0786 |
|  | 1998 | HS | HS | HS | * |  | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0043 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0005 |
|  | 2000 | HS | HS | * | - | HS |  | 0.0000 | 0.4720 | 0.0000 | 0.0000 | 0.0036 | 0.0000 | 0.0712 | 0.0000 | 0.0000 |
|  | 2001 | HS | HS | - | * | HS | HS |  | 0.0000 | 0.0059 | 0.0000 | 0.0003 | 0.0000 | 0.0000 | 0.0126 | 0.0000 |
|  | 2004 | HS | HS | * | - | HS | - | HS |  | 0.0000 | 0.0000 | 0.0001 | 0.0000 | 0.0012 | 0.0000 | 0.0000 |
|  | 2005 | HS | HS | - | - | HS | HS | * | HS |  | 0.0005 | 0.0024 | 0.0137 | 0.0025 | 0.7782 | 0.0018 |
|  | 2006 | HS | HS | HS | - | * | HS | HS | HS | * |  | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.5770 |
|  | 1989 | HS | HS | - | - | HS | * | * | HS | * | HS |  | 0.0023 | 0.0317 | 0.0000 | 0.0003 |
|  | 2001 | HS | HS | * | - | HS | HS | HS | HS | * | HS | * |  | 0.0000 | 0.2641 | 0.0000 |
|  | 2004 | HS | HS | * | - | HS | - | HS | * | * | HS | * | HS |  | 0.0000 | 0.0000 |
|  | 2005 | HS | HS | * | - | HS | HS | * | HS | - | HS | HS | - | HS |  | 0.0000 |
|  | 2006 | HS | HS | HS | - | * | HS | HS | HS | * | - | * | HS | HS | HS |  |

Table 5 Analysis of molecular variance (AMOVA) for the Chiwawa collections, showing the partition of molecular variance into (1) within collections, (2) among collections but within group, and (3) among group components. Each column in the table represents a separate analysis testing for differences under a different spatial or temporal hypothesis. The different analyses are grouped together in a single table for comparisons. The values within the table are percentages and the parenthetical values are P -values, or probabilities, associated with that percentage. P values greater than 0.05 indicate that the percentage is not significantly different from zero. For example, when collections are organized by hatchery- versus natural-origin ("Origin" - fourth column), $0.11 \%$ of the molecular variance is attributed to among group (i.e., hatchery- versus natural-origin), which is not significantly different from zero. No collections (first column) indicates no organization or grouping among all collections, and the among-group percentage is equal to the FST for the entire data set.

|  | No Structure | Collection <br> Year | Spawning <br> Location | Origin | Origin- <br> Spawning <br> Location |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Among Groups | 0.26 | 0.20 | 0.05 | 0.11 | 0.11 |
|  | $(0.00)$ | $(0.43)$ | $(0.48)$ | $(0.15)$ | $(0.06)$ |
| Among collections - | - | 0.08 | 0.24 | 0.21 | 0.18 |
| Within groups |  | $(0.003)$ | $(0.00)$ | $(0.00)$ | $(0.06)$ |
|  |  | 99.72 | 99.71 | 99.68 | 99.71 |
| Within collections | 99.74 | $(0.00)$ | $(0.00)$ | $(0.00)$ | $(0.00)$ |

Table 6 Fst $_{\text {ST }}$ values for all pairwise combinations of populations. Each $\mathrm{F}_{\text {ST }}$ is the median value for all pairwise combinations of collections within each population (the number of collections within each population is shown parenthetically next to each population name on each row). For example, the FST for the Chiwawa hatchery versus the White River (0.019) is the median value of 54 pairwise comparisons. The bold values along the center diagonal are the median $\mathrm{F}_{\text {ST }}$ values within each collection. For those populations with only one collection, the diagonal value was set at 0.000 .

|  | ChiwawaHatchery | ChiwawaNatural | Entiat | Leavenworth | Nason | Wenatcheemain | White | Little Wenatchee |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Chiwawa-Hatchery (9) | 0.013 | 0.008 | 0.016 | 0.012 | 0.011 | 0.005 | 0.019 | 0.111 |
| Chiwawa-Natural (9) |  | 0.003 | 0.012 | 0.011 | 0.007 | 0.003 | 0.014 | 0.105 |
| Entiat (1) |  |  | 0.000 | 0.005 | 0.010 | 0.008 | 0.019 | 0.078 |
| Leavenworth (1) |  |  |  | 0.000 | 0.007 | 0.008 | 0.014 | 0.092 |
| Nason (6) |  |  |  |  | 0.006 | 0.008 | 0.015 | 0.099 |
| Wenatchee-main (1) |  |  |  |  |  | 0.000 | 0.012 | 0.098 |
| White (6) |  |  |  |  |  |  | 0.005 | 0.113 |
| Little Wenatchee (1) |  |  |  |  |  |  |  | 0.000 |

Table 7 As in Table 5, except data includes Chiwawa hatchery- and natural-origin, Nason Creek, and White River collections

|  | All Years | All Years | 1989-1996 | $2005-2006$ | $2005-2006$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | No Structure | Origin | Origin | Origin | Collection Year |
| Among Groups | 0.28 | 0.33 | -0.07 | 0.43 | -0.06 |
|  | $(0.00)$ | $(0.00)$ | $(0.67)$ | $(0.01)$ | $(0.57)$ |
| Among Collections - |  | 0.04 | 0.22 | 0.25 | 0.64 |
| Within groups |  | $(0.00)$ | $(0.00)$ | $(0.00)$ | $(0.00)$ |
| Within Collections | 99.72 | 99.63 | 99.85 | 99.32 | 99.41 |

Table 8 Individual assignment results reported are the numbers of individuals assigned to each population using the partial Bayesian criteria of Rannala and Mountain (1997) and a "jack-knife" procedure (see Methods). The population with the highest posterior probability is considered the stock of origin (i.e., no unassigned individuals). Individuals from each population are assigned to specific populations (along rows). Bold values indicate correct assignment back to population of origin. Individuals assigned to a population are read down columns. For example, of the 595 individuals from Chiwawa hatchery origin, 134 individuals were assigned to Chiwawa natural origin (reading across). Of the 511 individuals assigned to Chiwawa natural origin (reading down), 60 were from Nason Creek.

| Population | Total | Unassigned | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ | $\mathbf{6}$ | $\mathbf{7}$ | $\mathbf{8}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1) Chiwawa Hatchery | 595 | 0 | $\mathbf{3 7 1}$ | 134 | 2 | 16 | 0 | 45 | 15 | 12 |
| 2) Chiwawa Natural | 501 | 0 | 156 | $\mathbf{2 6 9}$ | 4 | 5 | 0 | 42 | 9 | 16 |
| 3) Entiat | 37 | 0 | 4 | 5 | $\mathbf{1 3}$ | 8 | 0 | 6 | 1 | 0 |
| 4) Leavenworth | 73 | 0 | 9 | 8 | 3 | 33 | 0 | 17 | 0 | 3 |
| 5) Little Wenatchee | 19 | 0 | 0 | 0 | 0 | 0 | $\mathbf{1 9}$ | 0 | 0 | 0 |
| 6) Nason | 268 | 0 | 49 | 60 | 5 | 11 | 0 | $\mathbf{1 3 1}$ | 1 | 11 |
| 7) Wenatchee Mainstem | 32 | 0 | 12 | 9 | 0 | 1 | 0 | 2 | $\mathbf{6}$ | 2 |
| 8) White | 179 | 0 | 22 | 26 | 0 | 2 | 0 | 13 | 1 | $\mathbf{1 1 5}$ |
| TOTAL | 1704 | 0 | 623 | 511 | 27 | 76 | 19 | 256 | 33 | 159 |

Table 9 As in Table 8, except the posterior probability from the partial Bayesian criteria of Rannala and Mountain (1997) must be 0.90 or greater, to be assigned to a population. Those individuals with posterior probabilities less than 0.90 are unassigned.

| Aggregate | Total | Unassigned | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ | $\mathbf{6}$ | $\mathbf{7}$ | $\mathbf{8}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1) Chiwawa Hatchery | 595 | 332 | $\mathbf{2 1 4}$ | 31 | 1 | 4 | 0 | 10 | 3 | 0 |
| 2) Chiwawa Natural | 501 | 375 | 30 | $\mathbf{8 2}$ | 0 | 1 | 0 | 5 | 2 | 6 |
| 3) Entiat | 37 | 24 | 1 | 1 | $\mathbf{5}$ | 4 | 0 | 2 | 0 | 0 |
| 4) Leavenworth | 73 | 51 | 0 | 1 | 1 | $\mathbf{1 9}$ | 0 | 1 | 0 | 0 |
| 5) Little Wenatchee | 19 | 2 | 0 | 0 | 0 | 0 | $\mathbf{1 7}$ | 0 | 0 | 0 |
| 6) Nason | 268 | 188 | 11 | 6 | 2 | 5 | 0 | 53 | 0 | 3 |
| 7) Wenatchee Mainstem | 32 | 23 | 4 | 3 | 0 | 0 | 0 | 0 | $\mathbf{2}$ | 0 |
| 8) White | 179 | 92 | 4 | 3 | 0 | 1 | 0 | 5 | 1 | $\mathbf{7 3}$ |
| TOTAL | 1704 | 1087 | 264 | 127 | 9 | 34 | 17 | 76 | 8 | 82 |

Table 10 Estimates of $\mathrm{N}_{\mathrm{e}}$ based on bias correction method of Waples (2006) implemented in LDNe (Do and Waples unpublished). Collections are categorized by spawning location. Sample size is the harmonic mean of the sample size, $95 \%$ CI is the confidence interval calculated using Waples' (2006) equation 12, and Major Cohort assumes that each collection is $100 \%$ four-year-olds.

|  | Sample <br> size | Estimated <br> $\mathrm{N}_{\mathrm{b}}$ | $95 \% \mathrm{CI}$ | Major <br> Cohort | Census | $\mathrm{N}_{\mathrm{e}} / \mathrm{N}$ |
| :--- | :---: | :--- | :--- | :--- | :--- | :--- |
| 1993 Chiwawa Broodstock | 58.4 | 103.1 | $77.0-149.7$ | 1989 | 1392 | 0.30 |
| 1996 Chiwawa Broodstock | 15.5 | 30.4 | $19.6-58.1$ | 1992 | 1099 | 0.11 |
| 1998 Chiwawa Broodstock | 33.4 | 37.7 | $29.8-49.7$ | 1994 | 280 | 0.54 |
| 2000 Chiwawa Broodstock | 77.8 | 48.4 | $41.4-57.2$ | 1996 | 182 | 1.06 |
| 2001 Chiwawa Broodstock | 80.4 | 49.6 | $42.2-59.2$ | 1997 | 389 | 0.51 |
| 2004 Chiwawa Broodstock | 56.6 | 48.1 | $39.0-60.9$ | 2000 | 688 | 0.28 |
| 2005 Chiwawa Broodstock | 73 | 274.3 | $148.9-1131.8$ | 2001 | 4130 | 0.27 |
| 2006 Chiwawa Broodstock | 88.4 | 198.3 | $136.1-340.5$ | 2002 | 1613 | 0.49 |
|  |  |  |  |  |  |  |
| 1989 Chiwawa River | 26.6 | 5.2 | $3.9-6.3$ | 1985 |  |  |
| 2001 Chiwawa River | 46.7 | 38.6 | $31.0-49.3$ | 1997 | 389 | 0.40 |
| 2004 Chiwawa River | 88.5 | 82.6 | $67.3-104.4$ | 2000 | 688 | 0.48 |
| 2005 Chiwawa River | 104.2 | 231.5 | $161.8-382.7$ | 2001 | 4130 | 0.22 |
| 2006 Chiwawa River | 101.1 | 107.3 | $87.2-136$ | 2002 | 1613 | 0.27 |
|  |  |  |  |  |  |  |

Table 11 Summary of output from program SALMONNb and data for eight Chiwawa broodstock collections from Wenatchee River. For each pairwise comparison of samples $i$ and $j, \widetilde{\mathrm{~S}}$ is the harmonic mean sample size, $n$ is the number of independent alleles used in the comparison, $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ are the pairwise estimates of $\mathrm{N}_{\mathrm{b}}$, and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j}}\right]$ is the variance of $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. $\widetilde{\mathrm{N}}_{\mathrm{b}}$ is the harmonic mean of the $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. Alleles with a frequency below 0.05 were excluded from the analysis to reduce potential bias.

| Year | 1993 | 1996 | 1998 | 2000 | 2001 | 2004 | 2005 | 2006 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

Pairwise $\widetilde{\mathrm{S}}$ (above diagonal) and $n$ (below diagonal):

| 1993 | - | 24.5 | 42.5 | 66.4 | 67.2 | 57.2 | 64.6 | 70.3 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1996 | 82 | - | 21.2 | 25.8 | 26.0 | 24.4 | 25.6 | 26.4 |
| 1998 | 80 | 81 | - | 46.7 | 47.2 | 42.0 | 45.8 | 48.4 |
| 2000 | 80 | 82 | 84 | - | 78.6 | 65.2 | 75.1 | 82.7 |
| 2001 | 73 | 77 | 81 | 76 | - | 66.0 | 76.2 | 84.2 |
| 2004 | 77 | 81 | 75 | 76 | 78 | - | 63.5 | 69.0 |
| 2005 | 71 | 75 | 82 | 73 | 73 | 69 | - | 80.0 |
| 2006 | 81 | 80 | 84 | 75 | 74 | 75 | 72 | - |

Pairwise $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ (above diagonal) and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}\right]$ (below diagonal):

| 1993 | - | -742.7 | 406.9 | 1240.8 | -5432.0 | 829.8 | 808.9 | 729.0 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1996 | 22491.2 | - | 110.4 | -1786.5 | 765.9 | 162.8 | 824.7 | 382.7 |
| 1998 | 10910.4 | 67299.1 | - | 101.8 | 237.1 | 69.6 | 307.0 | 140.0 |
| 2000 | 6910.0 | 742895.8 | 19122.7 | - | 490.6 | 1498.2 | 706.9 | 201.6 |
| 2001 | 49318.3 | 21402.8 | 9754.2 | 6126.6 | - | 307.8 | 82.0 | 362.5 |
| 2004 | 8338.4 | 257267.7 | 24283.0 | 145043.4 | 7095.7 | - | 269.7 | 140.1 |
| 2005 | 31511.8 | 22242.5 | 10015.8 | 6596.6 | 114931.1 | 8240.4 | - | 599.6 |
| 2006 | 6223.8 | 43935.2 | 73518.7 | 10152.5 | 5885.3 | 12827.0 | 6370.8 | - |

$\tilde{\mathrm{N}}_{\mathrm{b}}=269.4$

Table 12 Summary of output from program SALMONNb and data for five Chiwawa in-river spawner collections from Wenatchee River. For each pairwise comparison of samples $i$ and $j, \tilde{\mathrm{~S}}$ is the harmonic mean sample size, $n$ is the number of independent alleles used in the comparison, $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ are the pairwise estimates of $\mathrm{N}_{\mathrm{b}}$, and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}\right.$ ] is the variance of $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. $\tilde{\mathrm{N}}_{\mathrm{b}}$ is the harmonic mean of the $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. Alleles with a frequency below 0.05 were excluded from the analysis to reduce potential bias.

| Year | 1989 | 2001 | 2004 | 2005 | 2006 |
| :--- | :--- | :--- | :--- | :--- | :--- |

Pairwise $\widetilde{\mathrm{S}}$ (above diagonal) and $n$ (below diagonal):

| 1989 | - | 33.3 | 40.2 | 41.7 | 42.2 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2001 | 72 | - | 60.5 | 63.9 | 63.3 |
| 2004 | 72 | 77 | - | 95.3 | 94.0 |
| 2005 | 69 | 72 | 75 | - | 102.5 |
| 2006 | 76 | 76 | 77 | 78 | - |

Pairwise $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ (above diagonal) and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}\right]$ (below diagonal):

| 1989 | - | 118.4 | 299.0 | 143.3 | 165.3 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2001 | 40378.8 | - | 181.7 | -1537.3 | 153.5 |
| 2004 | 10455.2 | 7265.5 | - | 387.1 | 329.4 |
| 2005 | 20923.6 | 68660.6 | 5040.7 | - | 356.8 |
| 2006 | 16227.2 | 8886.9 | 3802.0 | 4522.8 | - |

$\tilde{\mathrm{N}}_{\mathrm{b}}=224.2$

Table 13 Summary of output from program SALMONNb and data for three brood years that combined Chiwawa natural- and hatchery-origin samples from Wenatchee River. For each pairwise comparison of samples $i$ and $j, \tilde{\mathrm{~S}}$ is the harmonic mean sample size, $n$ is the number of independent alleles used in the comparison, $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ are the pairwise estimates of $\mathrm{N}_{\mathrm{b}}$, and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j}}\right]$ is the variance of $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. $\tilde{\mathrm{N}}_{\mathrm{b}}$ is the harmonic mean of the $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$. Alleles with a frequency below 0.05 were excluded from the analysis to reduce potential bias.

| Year 2004 | 2005 | 2006 |  |
| :--- | :--- | :--- | :--- |

Pairwise $\widetilde{\mathrm{S}}$ (above diagonal) and $n$ (below diagonal):

| 2004 | - | 162 | 164.3 |
| :--- | :--- | :--- | :--- |
| 2005 | 77 | - | 188.2 |
| 2006 | 76 | 75 | - |

Pairwise $\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}$ (above diagonal) and $\operatorname{Var}\left[\hat{\mathrm{N}}_{\mathrm{b}(\mathrm{i}, \mathrm{j})}\right]$ (below diagonal):

| 2004 | - | 611.3 | 210.8 |
| :--- | :--- | :--- | :--- |
| 2005 | 9351.5 | - | 727.5 |
| 2006 | 14965.5 | 8673.9 | - |

$\tilde{\mathrm{N}}_{\mathrm{b}}=386.8$

## Appendix K

Fish Trapping at the Nason Creek Smolt Trap 2015

# Population Estimates for Juvenile Salmonids in Nason Creek, WA 

## 2015 Annual Final Report

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#### Abstract

In 2015, Yakama Nation Fisheries Resource Management (YNFRM) monitored emigration of Endangered Species Act (ESA) listed Upper Columbia River (UCR) spring Chinook salmon and summer steelhead as well as naturally spawned juvenile coho salmon in Nason Creek. This report summarizes juvenile abundance and freshwater survival estimates for each of these species. Fish were captured using a 1.5 m rotary smolt trap between March 1 and November 30, 2015. We collected 745 spring Chinook salmon, 430 summer steelhead, 1 bull trout, and 5 coho; all of natural origin and varying age classes. Daily fish abundances for spring Chinook, steelhead, and coho were expanded by stream discharge-to-trap efficiency regression or pooled estimates. All estimates were made with a $95 \%$ confidence interval (CI) with total emigration estimates for BY2013 spring Chinook juveniles and coho juveniles of $57,525( \pm 39,889)$ and 161 $( \pm 714)$, respectively. We estimated the total BY2012 summer steelhead emigration at the trap to be 25,566 ( $\pm 6,020$ ). Egg-to-emigrant survival rates for BY2013 spring Chinook and BY2012 summer steelhead were $5.8 \%$ and $3.0 \%$, respectively. The egg-to-emigrant survival rate for BY2011 summer steelhead was $0.9 \%$. Productivity, as measured by emigrants-per-redd, for spring Chinook and summer steelhead, was 271 and 162, respectively. With no coho redds on Nason Creek in 2013, egg-to-emigrant survival and productivity could not be estimated for the 2013 brood.


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### 1.0 INTRODUCTION

Beginning in the fall of 2004, Yakama Nation Fisheries Resource Management (YNFRM) began operating a rotary smolt trap in Nason Creek for nine months per year. Prior to 2004, the smolt trap was operated on a limited basis solely for hatchery coho predation studies. This project is a cost share between the YNFRM's Mid-Columbia Coho Reintroduction Program (MCCRP) and Grant County PUD's Hatchery Monitoring Plan. Trap operations were conducted in compliance with ESA consultation specifically to address abundance and productivity of spring Chinook, steelhead trout, and coho salmon in Nason Creek.

Within this document we will report:

1) Juvenile abundance and productivity of spring Chinook salmon (tkwínat)

Oncorhynchus tshawytscha, steelhead trout (shúshaynsh) Oncorhynchus mykiss and coho salmon (súnx) Oncorhynchus kisutch in Nason Creek.
2) Emigration timing of spring Chinook salmon, steelhead trout and coho salmon emigrating from Nason Creek.

The data presented will be directly used to address Objective 2 in the Monitoring and Evaluation Plan for PUD Hatchery Programs (Hillman et al. 2015) on a 5-year analytic cycle:

## Objective 2: Determine if the proportion of hatchery fish on the spawning grounds affects the freshwater productivity of supplemented stocks (Hillman et al. 2013).

### 1.1 Watershed Description

The Nason Creek watershed drains 65,600 acres of alpine glaciated landscape where high precipitation and moderate rain on snow recurrence controls the hydrology and aquatic communities. Nason Creek originates near the Cascade crest at Stevens Pass and flows east for approximately 37 river kilometers (rkm) until joining the Wenatchee River at rkm 86.3 just below Lake Wenatchee. Both smolt trap locations employed in 2014 (see section 2.1 Trapping Equipment and Operations) were downstream from the majority of spring Chinook and steelhead spawning grounds (Figure 1). There are 26.4 rkm along the mainstem accessible to anadromous fish in Nason Creek. Private land ownership comprises 52,300 acres (79.7\%) of the watershed while 12,800 acres ( $19.5 \%$ ) are federal and 480 acres ( $0.1 \%$ ) are state owned (USFS et al. 1996).

The channel morphology of the lower 25 kilometers of Nason Creek has been impacted by development of highways, railroads, power lines, and residential development resulting in channel confinement and reduced side-channel habitat. The present condition is a low gradient ( < $1.1 \%$ ), low sinuosity ( $1: 2$ to 2:0 channel-to-valley length ratio) and depositional channel (USFS et al. 1996). Peak runoff typically occurs in May and June with occasional high water produced by rain on snow events in October and November.

In 2015, mean daily discharge for Nason Creek was 285 cfs with mean daily stream temperatures ranging from $0.0^{\circ} \mathrm{C}$ to $21.3^{\circ} \mathrm{C}$ (Figure $2 \& 3$ ). Spring discharge was extremely limited due to deminished snowpack by the onset of the trapping season. Maximum daily mean discharge in the spring of 2015 was 733 cfs ; normal maximum mean (12-year) daily flows during spring freshets on Nason Creek are approximately 2,000cfs. The lack of snowpack prompted the earlyonset of base-flow conditions (<100cfs) by the end of June. Base-flow conditions persisted into late October, at which time multiple rain-on-snow events pushed Nason Creek into flood conditions.


Figure 1. Map of Wenatchee River Subbasin with the Nason Creek rotary trap location.


Figure 2. Mean daily stream discharge at the Nason Creek WDOE stream monitoring station in 2015.


Figure 3. Mean daily water temperature at the Nason Creek DOE stream monitoring station in 2015.

### 2.0 METHODS

### 2.1 Trapping Equipment and Operation

The smolt trap was operated continually 24 hours per day, 7 days per week when conditions permitted. During spring snowmelt, operations occurred only during hours of darkness in order to minimize trap damage and capture mortality, while retaining the ability to sample during periods of peak fish movement. Without the threat of vandalism posed during periods of peak use at the previously-used campground location, summer operations at the Bolser location were not modified (daytime suspension).

On a daily basis, fish were removed from the primary collection box and retained in separate shore-anchored holding boxes until removed for efficiencies trials (up to 72 hours; Section 7 permit 2011/05645). A rotating drum-screen constantly removed small debris from the live box to avoid fish injury. All changes/modifications to the trap as well as periods of stoppage were noted. During periods when the trap was not operating (e.g. high discharge, high debris or mechanical malfunction), the number of target species captured was estimated. The estimated number of fish captured was calculated using the average number of fish captured three days prior and three days after the break in operation. This estimate of daily capture was incorporated into the overall emigration estimate.

### 2.2 Biological Sampling

Trap operating procedures and techniques followed a standardized basin-wide monitoring plan developed by the Upper Columbia Regional Technical Team (RTT) for the Upper Columbia Salmon Recovery Board (UCSRB; Hillman 2004), which was adapted from Murdoch and Petersen (2000).

All fish were enumerated by species and size class. Fish to be sampled were anesthetized in a solution of MS-222, weighed with an electronic scale and measured in a wetted trough-type measuring board. Anesthetized fish received oxygen through aquarium bubblers and were allowed to fully recover before being either released downstream of the trap or used in efficiency trials. Fork length (FL) and weight were recorded for all fish except when large numbers of fry or non-target species were collected; a sub-sample of 25 fish were measured and weighed while the remaining fish were tallied. Weight was measured to the nearest 0.1 gram and FL to the nearest millimeter. We used these data to calculate a Fulton-type condition factor (Kfactor) using the formula:
$K=\left(W / L^{3}\right) \times 100,000$
Where $\mathrm{K}=$ Fulton-type condition metric, $\mathrm{W}=$ weight in grams, $\mathrm{L}=$ fork length in millimeters and 100,000 is a scaling constant.

Scale samples were collected from steelhead measuring $\geq 60 \mathrm{~mm}$ FL so that age and brood year could be assigned. Samples were collected according to the needs and protocols set by Washington Department of Fish and Wildlife (WDFW), who conducted the analysis and provided YNFRM with results. Tissue samples were collected from spring Chinook and
steelhead for DNA analysis. Samples from spring Chinook and steelhead were retained for reproductive success analyses conducted by WDFW and National Marine Fisheries Service (NMFS). All target salmonids were classified as either natural or hatchery origin by physical appearance, presence/absence of coded wire tags (CWTs), or post-orbital elastomer tags. Developmental stages were visually classified as fry, parr, transitional, or smolt. Fry were defined as newly emerged fish with or without a visible yolk sac and a FL measuring < 50 mm . Age- 0 coho and spring Chinook salmon captured before July 1 were considered 'fry' and were excluded from subyearling population estimates because of the uncertainity that these fish were actively migrating (UCRTT, 2001).

### 2.3 PIT Tagging

All natural origin Chinook, steelhead and coho measuring $\geq 60 \mathrm{~mm}$ were PIT tagged. Once anesthetized, each fish was examined for external wounds or descaling, then scanned for the presence of a previously implanted PIT tag. If a tag was not detected, a pre-loaded 12 mm Digital Angel 134.2 kHz type TX 1411ST PIT tag was inserted into the body cavity using a Biomark MK-25 Rapid Implant Gun. Each unique tag code was electronically recorded along with date of tag implantation, date of fish release, tagging personnel, FL, weight, and anesthetic bath temperature. Data were entered using P3 software and submitted to the PIT Tag Information System (PTAGIS). PIT tagging methods were consistent with methodologies described in the PIT Tag Marking Procedures Manual (CBFWA 1999) as well as in 2008 ISEMP protocols (Tussing 2008).

After marking and sampling, fish were held for a minimum of 24-hours in holding boxes at the trap to; a) ensure complete recovery, b) assess tagging mortality, and c) determine a PIT tag shed rate. Mark groups were released by hand 0.8 rkm above the trap at nautical twilight. At each release, fish were distributed evenly along apposing banks in pools and other protected areas. Fish that were not used in mark-recapture trials were released downstream from the trap.

### 2.4 Mark-Recapture Trials

Groups of marked juveniles were released during a range of stream discharges in order to determine the trapping efficiency. PIT tags were the only method of marking used in 2015. These releases followed the protocols described in Hillman (2004), in which the author suggests a minimum sample size of 100 fish for each mark-recapture trial. Although 100 fish/trial represented the ideal mark group, low abundance of fish often required mark-recapture trials be completed with smaller sample sizes. To achieve the largest marked group possible, we combined catch over a maximum of 72 hours. Fish being held for mark-recapture trials were kept in auxiliary live boxes attached to the end of each pontoon or floating holding boxed anchored to the stream bank. A pre-season, minimum mark group size for each species/life stage was initially determined based on past regression models. In light of high abundance, minimum trial sizes could be raised to a more robust mark group with the intention of strengthening existing regression models.

Each mark-recapture trial was conducted over a three-day (72 hour) period to allow time for passage or capture. Completed trials were only considered invalid if an interruption to trapping occurred or proper pre-release procedures were not followed. Trials resulting in zero recaptures
were included in the efficiency regression (if determined valid once vetted through release/recapture protocols) as allowed by the new method of observed trap efficiency calculation. The model used (Bailey) employs use of recaptures +1 in the calculation of efficiency as a mode of bias correction. As a result, even trials yeilding no recaptures can be included in regression modeling (See equation 3 in 2.5.1 Estimate of Abundance).

In the event that low juvenile abundaces could not provide any opportunities for efficiency trials, releases were performed to allow for a pooled estimate. These releases did not have a minimum size and were released at equal intervals across the migratory period. Pooled estimates at the Nason Creek trap were utilized as an alternative method of estimation prior to the development of a viable regression model.

### 2.5 Data Analysis

### 2.5.1 Estimate of Abundance During Smolt Trapping

Seasonal juvenile migration, N , was estimated as the sum of daily migrations, $N_{i}$, i.e., $N=\sum_{i} N_{i}$, and daily migration was calculated from catch and efficiency:

$$
\begin{equation*}
\hat{N}_{i}=\frac{C_{i}}{\hat{e}_{i}} \tag{1}
\end{equation*}
$$

where $C_{i}=$ number of fish caught in period $I$;
$\hat{e}_{i}=$ trap efficiency estimated from the flow-efficiency relationship, $\sin ^{2}\left(b_{0}+b_{1}\right.$ flow $\left.w_{i}\right)$,
where $b_{0}$ is estimated intercept and $b_{1}$ is the estimated slope of the regression.
The regression parameters $b_{0}$ and $b_{1}$ are estimated using linear regression for the model:

$$
\begin{equation*}
\arcsin \left(\sqrt{e_{k}^{\text {obs }}}\right)=\beta_{0}+\beta_{1} \text { flow }_{k}+\varepsilon \tag{2}
\end{equation*}
$$

where $e_{k}^{o b s}=$ observed trap efficiency of Eq. 2 for trapping period $k$;
$\beta_{0}=$ intercept of the regression model;
$\beta_{1}=$ slope parameter;
$\varepsilon=$ error with mean 0 and variance $\sigma^{2}$.
In Equation 2, the observed trap efficiency, $e_{k}^{o b s}$, is calculated as follows,

$$
\begin{equation*}
e_{k}^{o b s}=\frac{r_{k}+1}{m} . \tag{3}
\end{equation*}
$$

The estimated variance of seasonal migration is calculated from daily estimates as:

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)=\underbrace{\sum_{i} \operatorname{Var}\left(N_{i}\right)}_{\text {Part A }}+\underbrace{\sum_{i} \sum_{j} \operatorname{Cov}\left(N_{i}, N_{j}\right)}_{\text {Part B }},
$$

or,

$$
\begin{equation*}
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)=\underbrace{\sum_{i} \operatorname{Var}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}\right)}_{\text {Part } A}+\underbrace{\sum_{i} \sum_{j} \operatorname{Cov}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}, \frac{\left(C_{j}+1\right)}{\hat{e}_{j}}\right)}_{\text {Part B}} . \tag{4}
\end{equation*}
$$

Part A of equation 4 is the variance of daily estimates. Part B is the between-day covariance. Note that the between-day covariance exists only for days that use the same trap efficiency model. If, for example, day 1 is estimated with one trap efficiency model, and day 2 estimated from a different model, then there is no covariance between day 1 and day 2 . The full expression for the estimated variance:

$$
\begin{aligned}
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)= & \underbrace{\sum_{i} \hat{N}_{i}^{2}\left(\frac{N_{i} \hat{e}_{i}\left(1-\hat{e}_{i}\right)}{\left(C_{i}+1\right)^{2}}+\frac{4\left(1-\hat{e}_{i}\right)}{\hat{e}_{i}} \operatorname{Var}\left(b_{0}+b_{1} \text { flow }_{i}\right)\right)}_{\text {PartA }}+ \\
& \underbrace{\sum_{i} \sum_{j} 4\left(\hat{N}_{i}\left(1-\hat{e}_{i}\right)\right)\left(\hat{N}_{j}\left(1-\hat{e}_{j}\right)\right) \cdot\left[\hat{\operatorname{Var}\left(b_{0}\right)+\text { flow }_{i} \text { flow }_{j} \hat{\left.\operatorname{Var}\left(b_{1}\right)\right]}}\right.}_{\text {PartB }} .
\end{aligned}
$$

where $\operatorname{Var}\left(b_{0}+b_{1}\right.$ flow $\left._{i}\right)=\operatorname{MS} E\left(1+\frac{1}{n}+\frac{\left(\text { flow }_{i}-\overline{\text { flow }}\right)^{2}}{(n-1) s_{\text {flow }}^{2}}\right)$, and $\hat{\operatorname{Var}}\left(b_{0}\right)$ and $\hat{\operatorname{Var}}\left(b_{1}\right)$ are
obtained from regression results. In Excel, the standard error (SE) of the coefficients is provided. The variance is calculated as the square of the standard error, $S E^{2}$.

In cases when there was no significant flow-efficiency relationship (i.e., low correlation), then a pooled, or average trap efficiency will suffice for the stratum. The estimator is calculated as follows:

$$
\hat{\bar{e}}=\frac{\sum_{j=1}^{k} r_{j}}{\sum_{j=1}^{k} m_{j}}
$$

where $\hat{\bar{e}}=$ the average or pooled trap efficiency for the stratum;
$m_{j}=$ the number of smolts marked and released in efficiency trial $j$ for the stratum;
$r_{j}=$ the number of smolts recaptured out of $m_{j}$ marked fish in efficiency trial $j$.
Abundance for a trapping period is estimated as:

$$
\hat{N}_{i}^{\text {pooled }}=\frac{C_{i}}{\hat{\bar{e}}},
$$

and total stratum abundance is:

$$
N^{\text {pooled }}=\sum_{i} \hat{N}_{i}^{\text {pooled }} .
$$

The variance of seasonal abundance takes into account the variability in catch numbers that are a result of binomial sampling (Part A), the pooled variance of trap efficiency, $\hat{\bar{e}}$ (Part B), and the covariance in daily estimates that arises from using a common estimate of efficiency across all trapping days (Part C):

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}^{\text {pooled }}\right)=\underbrace{\left(\sum_{i} \frac{\hat{N}_{i}(1-\hat{\bar{e}})}{\hat{\bar{e}}}\right)}_{\text {PartA }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}} \sum_{i} \hat{N}_{i}^{2}}_{\text {PartB }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}} \sum_{i} \sum_{j} \hat{N}_{i} \hat{N}_{j}}_{\text {Part }} .
$$

The Part B and Part C terms are combined in the calculation as a new Part B:

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}^{\text {pooled }}\right)=\underbrace{\left(\sum_{i} \frac{\hat{N}_{i}(1-\hat{\bar{e}})}{\hat{\bar{e}}}\right)}_{\text {Part }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}}\left[\sum_{i} \hat{N}_{i}^{2}+\sum_{i} \sum_{j} \hat{N}_{i} \hat{N}_{j}\right]}_{\text {PartB }} .
$$

The variance of $\hat{\bar{e}}$ is calculated as:

$$
\operatorname{Var}\left((\hat{\bar{e}})=\operatorname{Var}\left(\frac{\sum_{k=1}^{n} r_{k}}{\sum_{k=1}^{n} m_{k}}\right)=\frac{\sum_{k=1}^{n}\left(r_{k}-\hat{\bar{e}}_{k} m_{k}\right)^{2}}{\bar{m}^{2} n(n-1)}\right.
$$

where $\bar{m}$ is the average release size across all efficiency trial, $\frac{\sum_{k=1}^{n} m_{k}}{n}$. Confidence intervals were calculated using the following formulas:

$$
95 \% \text { confidence interval }=1.96 \times \sqrt{\sum \operatorname{var}}\left[\hat{N}_{i}\right]
$$

The single M-R estimator of abundance carries a set of well documented assumptions (Everhart and Youngs 1981; Seber 1982),

1. The population is closed to mortality.
2. The probability of capturing a marked or unmarked fish is equal.
3. Marked fish were randomly dispersed in the population prior to recapture.
4. Marking does not affect probabilities of capture.
5. Marks were not lost between the time of release and recapture.
6. All marks are reported upon recapture.
7. The number of fish in the trap, C , is fully enumerated and known without error.

### 2.5.2 Estimate of Abundance During The Non-Trapping Period

An estimate of spring chinook emmigration during the non-trapping period (December 1 through February 28) was calculated using remote-tagged spring chinook parr and the lower Nason Creek PIT tag array (NAL). A flow-detection efficiency regression was developed using mark-groups previously released to test the efficiency of the smolt trap. Daily spring Chinook detections at the NAL array and the developed regression were then applied to the Bailey estimator, as was peformed with daily trap abundance data (See equation 2.5.1 Estimate of Abundance). Tag rate determined at the Nason Creek smolt trap was used to account for unmarked emmigrants passing the NAL array.

Tag rate, $t_{i}$, was calculated as:

$$
t_{i}=\frac{t}{p}
$$

where $t=$ total smolt trap recaptures subsequent to the tagging effort;
$p=$ total catch at the smolt trap.

Daily abundace during the non-trapping period is calculated as:

$$
\hat{N}_{i}=\left(\frac{C_{i}}{\hat{e}_{i}}\right) / t_{i},
$$

where $C_{i}=$ number of fish caught in period $I$;
$\hat{e}_{i}=$ trap efficiency estimated from the flow-efficiency relationship, $\sin ^{2}\left(b_{0}+b_{1} f l o w_{i}\right) ;$
$t_{i}=$ tag rate.

### 2.5.3 Production and Survival

Production estimates by age class were summed to produce a total emigration estimate. For spring Chinook and coho, estimates of fall migrant parr were added to subsequent spring smolt estimates to generate a single brood year estimate. For steelhead, a single brood year may require up to three years for emigration from Nason Creek to occur. Pending scale analysis, steelhead captured in 2015 were aged via an age-length histogram built upon previously analyzed scale samples. For all three species, egg-to-emigrant estimates were calculated by dividing estimated emigrants by approximated egg deposition during a spawning brood (average fecundity used to determine egg deposition derived from WDFW Chiwawa broodstock spawning). The number of emigrants-per-redd for each brood year was calculated by dividing the total emigrant estimate by the number of redds counted during spawning ground surveys.

### 3.0 RESULTS

### 3.1 Dates of Operation

The Nason Creek smolt trap was installed on February 25, and operated in its fixed position for the entirety of the trapping season (March 1 to November 30). Removal of the trap occurred on December 2. We attempted to run the trap continuously 24 hours a day, 7 days per week. Intentional suspension of trapping activities occurred for a prolonged period in the summer-early fall due to extreme base flows (July 18 - October 20; Table 1). Pulling of the trap also occurred in the fall as a precaution during two major flood events. Trap stoppages were most frequent from July through November, as heavy debris loads and ice formation prevented continuous operation.

Table 1. Summary of Nason Creek rotary trap operation.

| Date of Trap <br> Operations | Trap <br> Status | Description | Days |
| :--- | :--- | :--- | :---: |
|  | Operating | Continuous data collection | 119 |
| March 1 to | Interrupted | Interrupted by debris | 3 |
| June 30 | Pulled | Intentionally pulled during periods of high flow, low flow, or significant <br> ice formation | 0 |
|  | Operating | Continuous data collection | 34 |
| July 1 to | Interrupted | Interrupted by debris, ice and/or low flows |  |
| November 30 | Pulled | Intentionally pulled during periods of high flow, low flow, or significant <br> ice formation | 105 |

### 3.2 Daily Captures and Biological Sampling

### 3.2.1 Spring Chinook Yearlings (BY2013)

Between March 1 and June 30, a total of 152 wild Chinook yearlings were captured at the trap (Figure 4). The majority of these fish were collected following an intial spike in flow immediately following operation commencement. A peak catch of 10 yearling smolts coincided with a secondary spike in discharge occurring on March 27. Following the final freshets of March, catch dropped substantially with the last emigrating Chinook yearling captured on May 21. Although three trap stoppages occurred during this period, they likely did not adversely affect total Chinook smolts captured and therefore, estimates were forgone. Mean FL and weight for Chinook yearlings was $93 \mathrm{~mm}(n=152 ; S D=7.0)$ and $8.4 \mathrm{~g}(n=152 ; S D=2.2$; Table 2$)$, respectively. Tissue sample were collected from 138 fish for an ongoing, parental-based DNA analysis by WDFW. Five wild spring Chinook mortalities were incurred.


Figure 4. Daily catch of BY2013 spring Chinook yearlings with mean daily stream discharge at the Nason Creek rotary trap, March 1 to June 30, 2015.

Table 2. Summary of length and weight sampling of juvenile spring Chinook captured at the Nason Creek rotary trap in 2015.

| Brood <br> Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | K- <br> Factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | $n$ | SD | Mean | $n$ | SD |  |
| 2013 | Wild Spring Chinook Yearling Smolt | 93 | 152 | 7.0 | 8.4 | 152 | 2.2 | 1.03 |
| 2014 | Wild Spring Chinook Subyearling Fry | 45 | 338 | 9.9 | 1.0 | 338 | 0.9 | 0.87 |
| 2014 | Wild Spring Chinook Subyearling Parr | 84 | 210 | 8.0 | 6.5 | 209 | 1.7 | 1.08 |
| 2013 | Hatchery Spring Chinook Yearling Smolt | 136 | 284 | 12.3 | 29.5 | 284 | 8.8 | 1.13 |

### 3.2.2 Spring Chinook Subyearlings (BY2014)

A total of 210 wild spring Chinook subyearling parr were captured between July 1 and November 30, with an additional 338 subyearling fry captured prior to July 1 (Figure 5). A peak daily capture of 89 subyearling Chinook parr occurred on November 3, following the first fall high-water event of the year. Mean FL and weight among fall subyearling parr was 84 mm ( $n=$ 210; $S D=8.0$ ) and $6.5 \mathrm{~g}(n=209 ; S D=1.7)$, respectively. We estimate that an additional 16 Chinook subyearling parr would have been captured during short stoppages ( $\leq 3$ days) had the trap run without interruption. Estimates of daily abundance during the prolonged period of suspended trapping (July 14 - October 10) were not made due to a lack of documented pre- and post-suspension movement, as well as the duration of the suspenstion. Tissue samples were collected from 213 fish for an ongoing, parental-based DNA analysis by WDFW. A total of 10
subyearling Chinook ( 9 fry and 1 parr) mortalities occurred in 2015. Causes of death included trapping mortality, tagging/handing mortality, and pre-existing fungal infection/poor condition.


Figure 5. Daily catch of BY2014 spring Chinook subyearlings with mean daily stream discharge at the Nason Creek rotary trap, July 1 to November 30, 2015.

### 3.2.3 Hatchery Spring Chinook Smolts (BY2013)

During the months of April and May, a total of 43,082 hatchery spring Chinook smolts were released into Nason Creek (M. Babiar, personal communication, January 14, 2016). All hatchery spring Chinook were released directly from the Grant County Public Utility District (GCPUD) Nason Creek Acclimation Facility located at rkm17.3. Subsequently, a total of 714 smolts were captured with a mean FL and weight of $136 \mathrm{~mm}(n=284 ; S D=12.3)$ and $29.5 \mathrm{~g}(n=284 ; S D=$ 8.8), respectively (Figure 6). Hatchery spring Chinook were not captured at the smolt trap beyond May 10. There were no mortalities incurred.


Figure 6. Daily catch of BY2013 hatchery spring Chinook smolts with mean daily stream discharge at the Nason Creek rotary trap, March 1 to June 30, 2015.

### 3.2.4 Summer Steelhead

A total of 430 wild summer steelhead juveniles were captured throughout the season from March 1 to November 30 with a peak catch of 89 juveniles on November 2 (Figure 6). We estimated that an additional 2 age- 1 juveniles would have been captured had there been no interruptions to trapping during the migratory period (Mar 1 to July 31). Histogram analysis of known steelhead ages sampled from 2005 to 2014 allowed us to estimate ages of fish captured in 2015 using FL. We estimate that of the total steelhead captured, 182 were young-of-the-year, 233 were age-1, 14 were age-2, and 1 was age- 3 . Subyearling steelhead caught had a mean FL of 70 mm ( $n=182$; $S D=15.5$ ), and a mean weight of $4.3(n=176 ; S D=2.0)$. The majority of steelhead juveniles captured were age-1 parr emigrating past the trap in spring. Mean FL and weight of age-1 fish was 88 mm ( $n=233$; $S D=20.2$; Table 3 ) and $8.3 \mathrm{~g}(n=233 ; S D=6.7$ ), respectively. Age-2 steelhead were caught primarily in the spring, with only one fish being captured after July 31. Mean FL and weight of age-2 fish was $149 \mathrm{~mm}(n=14 ; S D=13.5)$ and $33.7 \mathrm{~g}(n=14 ; S D=8.2)$, respectively. A single age- 3 fish with a FL of 175 mm and weight of 51.3 g was also captured. Scales were taken from a sub-sample $(n=188)$ to be used for future age analyses. Two trapping mortalities were incurred (See 3.6 ESA Compliance).


Figure 7. Daily catch of wild summer steelhead with mean daily stream discharge at the Nason Creek rotary trap, March 1 to November 30, 2015. Estimates of fish passage during trap interruptions are not depicted.

Table 3. Summary of length, weight and condition factor by age class of wild summer steelhead emigrants and hatchery steelhead captured at the Nason Creek rotary trap.

| Brood <br> Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | K- <br> Factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | $n$ | SD | Mean | $n$ | SD |  |
| 2015 | Wild Summer Steelhead (Age-0) | 70 | 182 | 15.5 | 4.3 | 176 | 2.0 | 1.06 |
| 2014 | Wild Summer Steelhead (Age-1) | 88 | 233 | 20.2 | 8.3 | 233 | 6.7 | 1.04 |
| 2013 | Wild Summer Steelhead (Age-2) | 149 | 14 | 13.5 | 33.7 | 14 | 8.2 | 1.00 |
| 2012 | Wild Summer Steelhead (Age-3) | 191 | 1 | - | 73.8 | 1 | - | 1.06 |
| 2014 | Hatch. Summer Steelhead Smolt | 175 | 273 | 15.2 | 51.3 | 273 | 12.5 | 0.94 |

### 3.2.5 Hatchery Steelhead Smolts (BY2014)

During April and May, WDFW directly planted a total of 86,613 hatchery summer steelhead smolts into Nason Creek (M. Babiar, personal communication, January 14, 2016). Subsequently, a total of 448 hatchery steelhead were captured at the smolt trap with a mean FL and weight of $175 \mathrm{~mm}(n=273 ; S D=15.2)$ and $51.3 \mathrm{~g}(n=273 ; S D=12.5)$, respectively (Figure 7). The presence of hatchery-origin steelhead at the trap was limited to 45 days after initial release, and did not continue into the summer. Hatchery origin was determined by the presence of coded wire tags (CWT). One mortality was incurred.


Figure 8. Daily catch of BY2014 hatchery steelhead smolt with mean daily stream discharge at the Nason Creek rotary trap, March 1 to June 30, 2015.

### 3.2.6 Bull Trout

Bull trout presence at the trap in 2015 was limited to a single fish with a FL of 180 mm and weight of 50.1 g . The bull trout was released immediately after morphometric measurements were taken. No other sampling/tagging activities were performed.

### 3.2.7 Coho Yearlings (BY2013)

Two naturally produced coho yearlings were captured during spring emigration between March 1 and June 30 (Figure 8). Mean FL and weight were $109 \mathrm{~mm}(n=2 ; S D=4.9)$ and $12.0 \mathrm{~g}(n=2$;
$S D=0.1$ ), respectively (Table 5). Scale and tissue samples were not taken from naturallyproduced coho smolts in 2015. There were no coho yearling mortalities.


Figure 9. Daily catch of BY2013 naturally-produced coho yearlings with mean daily stream discharge at the Nason Creek rotary trap, March 1 to June 30, 2015.

Table 4. Summary of length and weight sampling of juvenile coho salmon captured at the Nason Creek rotary trap in 2015.

| Brood <br> Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | K- <br> Factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | $n$ | SD | Mean | $n$ | SD |  |
| 2013 | Naturally Produced Coho Yearling Smolts | 109 | 2 | 4.9 | 12.0 | 2 | 0.1 | 0.95 |
| 2014 | Naturally Produced Coho Subyearling Fry | 47 | 7 | 13.7 | 1.4 | 7 | 1.5 | 0.86 |
| 2014 | Naturally Produced Coho Subyearling Parr | 69 | 3 | 7.0 | 4.0 | 3 | 1.3 | 1.20 |
| 2013 | Hatchery Coho Yearling Smolts | 131 | 952 | 9.9 | 23.3 | 952 | 4.8 | 1.03 |

### 3.2.8 Coho Subyearlings (BY2014)

A total of three naturally produced coho subyearling parr were captured during between July 1 and November 30 (Figure 9). Mean FL and weight were $69 \mathrm{~mm}(n=3 ; S D=7.0)$ and $4.0 \mathrm{~g}(n=$ 3; $S D=1.3$ ), respectively. An additional seven subyearling coho fry were also captured with a mean FL of 47 mm . There were no coho subyearling mortalities.


Figure 10. Daily catch of BY2014 naturally-produced coho subyearlings with mean daily stream discharge at the Nason Creek rotary trap, July 1 to November 30, 2015.

### 3.2.9 Hatchery Coho Smolts (BY2013)

A total of 253,242 hatchery coho were released into Nason Creek above the trap in spring of 2015. All hatchery coho released were acclimated in natural ponds adjacent to Nason Creek and reared to smolt stage prior to volitional release. Between March 1 and June 30, a total of 1,798 hatchery coho were captured at the trap (Figure 10). Mean FL was $131 \mathrm{~mm}(n=952 ; S D=9.9)$ and mean weight was $23.3 \mathrm{~g}(n=952 ; S D=4.8$; Table 2 ). A peak daily catch of 215 hatchery coho smolts occurred on May 5 following volitional release into Nason Creek. One trapping mortality was incurred. Hatchery coho emigration data at the Nason Creek trap assists the MCCRP by providing size-at-emigration, emigration timing and duration of residence in Nason Creek.


Figure 11. Daily catch of BY2013 hatchery coho smolt with mean daily stream discharge at the Nason Creek rotary trap, March 1 to June 30, 2015.

### 3.3 Remote Parr Tagging (BY2013 Spring Chinook)

YNFRM and WDFW personnel PIT tagged and released a total of 1,821 BY2013 spring Chinook parr between September 22 and October 24, 2014. The total surveyed area included Nason Creek from rkm 0.8 to 26.1. All collections were performed via backpack electrofisher. Equal capture effort (measured in electrofisher seconds used) was applied across all reaches.

Between October 1 and March 30, a total of 311 re-sights of the remote tagged Chinook were documented at the NAL array (Figure 12). Of these detections, only 13 were during the winter non-trapping period. PTAGIS event logs for the NAL array indicated that it operated continuously for the duration of this time with no alterations (PTAGIS 2015).

Subsequent to the remote tagging effort, 30 remote-tagged BY2013 spring Chinook were recaptured at the Nason Creek smolt trap. Total spring Chinook catch at the smolt trap was 798 emigrants during the same period. The pooled tag rate for remote-tagged spring Chinook captured at the Nason smolt trap was $3.8 \%$.


Figure 12. Daily detections of remote-tagged BY2013 spring Chinook at the lower Nason Creek PIT tag antenna array (NAL) between October 2014 and March 2015.

### 3.4 Trap Efficiency Calibration and Population Estimates

### 3.4.1 Spring Chinook Yearlings (BY2013)

Infrequent releases, low abundance, and a lack of recaptures did not allow a species-specific model to be used on BY2013 yearling emigrants. In order to produce an estimate, a pooled efficiency ( $2.07 \%$ ) composed of spring Chinook yearling and hatchery-origin coho yearling surrogate trials was used (Table 5). We recognize the sub-optimal nature of this estimation methodology, and will recalculate the estimates using linear regression analysis as soon as feasible. We estimated a total of $6,992( \pm 32,823 ; 95 \%$ CI) BY2013 Chinook yearlings emigrated in spring of 2015 (Table 7). Parr emmigration during the non-trapping period was estimated using a flow-efficiency regression ( $r^{2}=0.61 ; p=0.0002$ ) based on detections at the NAL pit tag array. We estimated that $6,822( \pm 9,035 ; 95 \% \mathrm{CI})$ BY2013 spring Chinook emigrated out of Nason Creek during the non-trapping period. Combined with a recalculated BY2013 subyearling estimate of $43,711( \pm 20,788 ; 95 \% \mathrm{CI})$, we estimated that a total of 57,526 ( $\pm 39,889$; 95\% CI) BY2013 spring Chinook juveniles emigrated from Nason Creek.

Table 5. Trap efficiency trials conducted with BY2013 wild spring Chinook yearlings and hatchery-origin coho yearling surrogates.

| Origin/Species/Stage | Age | Date | Marked | Recaptured | Discharge <br> $(\mathrm{cfs})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Yearlings | $1+$ | $4 / 23 / 2015$ | 7 | 0 | 337 |
| Wild Chinook Yearlings | $1+$ | $4 / 27 / 2015$ | 2 | 0 | 269 |


| Wild Chinook Yearlings | $1+$ | $5 / 6 / 2015$ | 5 | 0 | 330 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Yearlings | $1+$ | $5 / 10 / 2015$ | 1 | 0 | 334 |
| Wild Chinook Yearlings | $1+$ | $5 / 14 / 2015$ | 22 | 0 | 418 |
| Wild Chinook Yearlings | $1+$ | $5 / 22 / 2015$ | 1 | 0 | 421 |
| Hatchery-Origin Coho Yearlings | $1+$ | $5 / 5 / 2015$ | 98 | 2 | 370 |
| Hatchery-Origin Coho Yearlings | $1+$ | $5 / 12 / 2015$ | 224 | 8 | 408 |
| Hatchery-Origin Coho Yearlings | $1+$ | $5 / 14 / 2015$ | 101 | 3 | 418 |
| Hatchery-Origin Coho Yearlings | $1+$ | $5 / 19 / 2015$ | 102 | 0 | 421 |
| Hatchery-Origin Coho Yearlings | $1+$ | $5 / 23 / 2015$ | 66 | 0 | 416 |
| Total |  |  | $\mathbf{6 2 9}$ | $\mathbf{1 3}$ |  |

Table 6. Estimated egg-to-emigrant survival and smolts-per-redd production for Nason Creek spring Chinook salmon.

| Brood Year | $\begin{gathered} \text { No. } \\ \text { of } \\ \text { Redds } \end{gathered}$ | Fecundity ${ }^{\text {a }}$ | Est. Egg Deposition | No. of Emigrants |  |  |  | Egg-toEmigrant | Emigrants per Redd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | $\begin{gathered} \text { Age- } \\ 0^{\mathrm{b}} \end{gathered}$ | $\begin{gathered} \text { Non } \\ \text { Trap }^{\mathrm{d}} \\ \hline \end{gathered}$ | Age1 | Total $\pm 95 \%$ CI |  |  |
| 2002 | 294 | 4,654 | 1,368,276 | DNOT |  | 4,683 | - | - | - |
| 2003 | 83 | 5,844 | 485,052 | 8,829 |  | 6,358 | $15,187 \pm 1,605$ | 3.1\% | 183 |
| 2004 | 169 | 4,799 | 811,031 | 11,822 |  | 2,597 | $14,419 \pm 2,766$ | 1.8\% | 85 |
| 2005 | 193 | 4,327 | 835,111 | 11,814 |  | 8,696 | 20,510 $\pm 5,018$ | 2.5\% | 106 |
| 2006 | 152 | 4,324 | 657,248 | 4,144 |  | 7,798 | $11,942 \pm 1,744$ | 1.8\% | 79 |
| 2007 | 101 | 4,441 | 448,541 | 15,556 |  | 5,679 | $21,235 \pm 2,864$ | 4.7\% | 210 |
| 2008 | 336 | 4,592 | 1,542,912 | 23,182 |  | 3,611 | $26,793 \pm 6,756$ | 1.7\% | 80 |
| 2009 | 167 | 4,573 | 763,691 | 27,720 |  | 1,705 | $29,425 \pm 12,777$ | 3.9\% | 176 |
| 2010 | 188 | 4,314 | 811,032 | 8,491 |  | 3,535 | $12,026 \pm 1,954$ | 1.5\% | 64 |
| 2011 | 170 | 4,385 | 745,450 | 17,991 |  | 2,422 | $20,413 \pm 3,889$ | 2.7\% | 120 |
| 2012 | 413 | 4,223 | 1,744,099 | 28,110 |  | 4,561 | $32,671 \pm 4,863$ | 1.9\% | 79 |
| 2013 | 212 | 4,716 | 999,792 | 43,711 | 6,822 | 6,992 | $57,525 \pm 39,889$ | 5.8\% | 271 |
| 2014 | 115 | 4,467 | 513,705 | 13,903 | - | - | - | - | - |
| Avg.c | 199 | 4,594 | 894,905 | 18,306 | - | 4,905 | 23,831 | 2.9\% | 132 |

a Data provided by Hillman et al. 2015.
${ }^{\mathrm{b}}$ Does not include subyearling fry prior to July 1.
${ }^{\text {c }}$ 11-year average of complete brood data, BY2003-2013.
${ }^{\mathrm{d}}$ Estimated emigration during the winter non-trapping period (December 1 - February 28).



Figure 13. Relationships between estimated egg deposition and total emigrants produced, egg-to-emigrant survival, and emigrants per redd for Nason Creek spring Chinook, BY 2003 to 2013. *2013 brood (denoted by red border) does not include non-trapping estimate.

### 3.4.2 Spring Chinook Subyearlings (BY2014)

A linear regression model was developed using subyearling mark groups released in the fall of 2014 and 2015. This weighted regression was not significant $\left(r^{2}=0.36 ; p=0.09\right)$ at our accepted limit $(\alpha=0.05)$. However, previous comparisons to pooled estimates suggest that linear regression analysis would be a more viable means of estimation despite less than optimal significance. Also, extreme high flows, low yearling Chinook abundance, and sporadic trap operation in the month of November would have greatly hindered the development of a pooled estimate. As a multi-year regression, this initial flow-efficiency relationship represents the starting point from which we will build further estimates. Using this model, we estimated that a total of 13,903 ( $\pm 11,963 ; 95 \%$ CI) BY2014 spring Chinook emigrated past the trap in the Fall of 2013 (Table 6).

Table 7. Trap efficiency trials conducted with BY2014 wild spring Chinook subyearlings.

| Origin/Species/Stage | Age | Date | Marked | Recaptured | Discharge <br> (cfs) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Subyearlings | 0 | $11 / 3 / 2015$ | 138 | 0 | 460 |
| Wild Chinook Subyearlings | 0 | $11 / 23 / 2015$ | 9 | 0 | 520 |

### 3.4.3 Summer Steelhead

Low abundance of summer steelhead emigrants in the spring of 2015 required a pooled estimate be used in light of the inability to meet minimum mark-group sizes $(n=50)$ for regression analysis (Table 8). Releases of PIT-tagged steelhead were subsequently released every four days upstream at the established release location (Table 9). In a total of 13 separate trials, 116 wild summer steelhead were released upstream with only 1 recapture ( $0.86 \%$ ). Estimates of age- 0 fry and parr were not made due to insufficient evidence that active migration is occurring at this young age. Previous attempts at the old location to build a model based on young-of-the-year steelhead parr in the fall have yielded weak flow-efficiency relationships; further suggesting that age-0 parr catch is the result of displacement rather than active migration. We estimated that 22,504 ( $\pm 3,175 ; 95 \%$ CI) BY2014 age-1, 1,508 ( $\pm 897 ; 95 \%$ CI) BY2013 age-2, and 116 ( $\pm 436$; $95 \%$ CI) BY2012 age-3 steelhead emigrated past the trap in 2015 (Table 10). We estimate that total (age 1-3) BY2012 emigration to be $25,566( \pm 6,020 ; 95 \% \mathrm{CI})$.

Table 8. Efficiency trials conducted with wild summer steelhead juveniles.

| Origin/Species/Stage | Date | Marked | Recaptured | Discharge (cfs) |
| :---: | :---: | :---: | :---: | :---: |
| Wild Steelhead Parr/Smolt | $4 / 23 / 2015$ | 17 | 1 | 337 |
| Wild Steelhead Parr/Smolt | $4 / 27 / 2015$ | 3 | 0 | 269 |
| Wild Steelhead Parr/Smolt | $5 / 2 / 2015$ | 8 | 0 | 338 |
| Wild Steelhead Parr/Smolt | $5 / 6 / 2015$ | 13 | 0 | 330 |
| Wild Steelhead Parr/Smolt | $5 / 10 / 2015$ | 3 | 0 | 334 |
| Wild Steelhead Parr/Smolt | $5 / 14 / 2015$ | 1 | 0 | 418 |
| Wild Steelhead Parr/Smolt | $5 / 18 / 2015$ | 6 | 0 | 392 |
| Wild Steelhead Parr/Smolt | $5 / 22 / 2015$ | 10 | 0 | 421 |
| Wild Steelhead Parr/Smolt | $5 / 26 / 2015$ | 9 | 0 | 337 |
| Wild Steelhead Parr/Smolt | $5 / 30 / 2015$ | 26 | 0 | 365 |
| Wild Steelhead Parr/Smolt | $6 / 4 / 2015$ | 9 | 0 | 218 |
| Wild Steelhead Parr/Smolt | $6 / 8 / 2015$ | 4 | 0 | 192 |
| Wild Steelhead Parr/Smolt | $6 / 16 / 2015$ | 7 | 0 | 109 |
| Total |  | $\mathbf{1 1 6}$ | $\mathbf{1}$ |  |

Table 9. Estimated egg-to-emigrant survival and emigrants-per-redd production for Nason Creek summer steelhead.

| Brood Year | No. of Redds | Fecundity ${ }^{\text {a }}$ | Est. Egg Deposition | No. of Emigrants |  |  |  | Egg-to- <br> Emigra <br> nt | Emigrant <br> s per <br> Redd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 1+ | 2+ | 3+ | Total $\pm 95 \%$ CI |  |  |
| 2001 | 27 | 5,951 | 160,677 | DNOT | DNOT | 846 | - | - | - |
| 2002 | 80 | 5,776 | 462,080 | DNOT | 2,475 | 0 | - | - | - |
| 2003 | 121 | 6,561 | 793,881 | 4,906 | 1,054 | 27 | $5,987 \pm 1,193$ | 0.8\% | 49 |
| 2004 | 127 | 5,118 | 649,986 | 5,107 | 906 | 22 | $6,035 \pm 885$ | 0.9\% | 48 |
| 2005 | 412 | 5,545 | 2,284,540 | 7,416 | 2,502 | 298 | $10,216 \pm 2,147$ | 0.4\% | 25 |
| 2006 | 77 | 5,688 | 437,976 | 19,609 | 2,673 | 37 | $22,319 \pm 5,722$ | 5.1\% | 290 |
| 2007 | 78 | 5,840 | 455,520 | 26,518 | 2,325 | 117 | $28,960 \pm 7,739$ | 6.4\% | 371 |


| 2008 | 88 | 5,693 | 500,984 | 8,782 | 1,164 | 0 | $9,946 \pm 2,382$ | $2.0 \%$ | 113 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | 126 | 6,199 | 781,074 | 13,606 | 608 | 312 | $14,526 \pm 2,868$ | $1.9 \%$ | 115 |
| 2010 | 270 | 5,458 | $1,473,660$ | 12,767 | 3,999 | 0 | $16,776 \pm 3,885$ | $1.1 \%$ | 62 |
| 2011 | 235 | 6,276 | $1,474,860$ | 13,109 | 482 | 0 | $13,591 \pm 3,525$ | $0.9 \%$ | 58 |
| 2012 | 158 | 5,309 | 838,822 | 24,637 | 813 | $116^{\text {c }}$ | $25,566 \pm 6,020$ | $3.0 \%$ | 162 |
| 2013 | 135 | 5,749 | 777,735 | 11,837 | $1,508^{\mathrm{c}}$ | - | - | - | - |
| 2014 | 198 | 5,831 | $1,154,538$ | $22,504^{\mathrm{c}}$ | - | - | - | - | - |
| Avg $^{\mathrm{b}}$ | 169 | 5,769 | 969,130 | 13,646 | 1,653 | 90 | 15,380 | $2.3 \%$ | 129 |

a Data provided by Hillman et al. 2015
${ }^{\text {b }}$ 10-year average of complete brood estimates, BY2003-2012
${ }^{\text {c }}$ Pooled estimate



Figure 14. Relationships between estimated egg deposition and total emigrants produced, egg-to-emigrant survival, and emigrants per redd for Nason Creek summer Steelhead, BY 2003 to 2012. *2012 brood denoted by red border.

### 3.4.4 Coho Yearlings (BY2013)

Limited abundance of BY2013 coho yearlings did not provide any opportunities to perform any efficiency trials in the spring of 2015. In lieu of a species-specific model, a pooled estimate using releases of marked hatchery-origin coho smolts was applied to wild coho smolts. In the
spring of 2015, we estimated that 91 ( $\pm 711 ; 95 \%$ CI) emigrated past the trap (Table 11). This gave us a total BY2013 emigrant estimate of 161 ( $\pm 714$; 95\% CI).

Table 10. Estimated egg-to-emigrant survival and smolts-per-redd production for Nason Creek coho salmon.

| Brood Year | No. of Redds | Fecundity | Est. Egg Deposition | No. of Emigrants |  |  | Egg-toEmigrant | Emigrants per Redd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Age-0 ${ }^{\text {a }}$ | Age-1 | $\begin{gathered} \text { Total } \pm 95 \% \\ \text { CI } \\ \hline \end{gathered}$ |  |  |
| 2003 | 6 | 2,458 | 14,748 | DNOT | 394 | - | - | - |
| 2004 | 35 | 3,084 | 107,940 | 204 | 56 | $260 \pm 155$ | 0.2\% | 7 |
| 2005 | 41 | 2,866 | 117,506 | 27 | 910 | $937 \pm 347$ | 0.8\% | 23 |
| 2006 | 4 | 3,126 | 12,504 | 7 | 0 | $7 \pm 10$ | 0.1\% | 2 |
| 2007 | 10 | 2,406 | 24,060 | 14 | 136 | $150 \pm 104$ | 0.6\% | 15 |
| 2008 | 3 | 3,275 | 9,825 | 50 | 0 | $50 \pm 57$ | 0.5\% | 17 |
| 2009 | 14 | 2,691 | 37,674 | 471 | 237 | $708 \pm 478$ | 1.9\% | 51 |
| 2010 | 8 | 3,411 | 27,288 | 27 | 437 | $464 \pm 231$ | 1.7\% | 58 |
| 2011 | 89 | 3,114 | 277,146 | 1,018 | 1,387 | 2,405 $\pm 612$ | 0.9\% | 27 |
| 2012 | 21 | 2,752 | 57,792 | 46 | 434 | $480 \pm 237$ | 0.8\% | 23 |
| 2013 | 0 | 2,973 | 0 | 70 | 91 | $161 \pm 714$ | NA | NA |
| 2014 | 16 | 2,992 | 47,872 | 84 | - | - | - | - |
| Avg. ${ }^{\text {b }}$ | 23 | 2,970 | 67,174 | 193 | 369 | 562 | 0.8\% | 25 |

[^17]


Figure 15. Relationships between estimated egg deposition and total emigrants produced, egg-to-emigrant survival, and emigrants per redd for Nason Creek naturally-produced coho, BY 2004 to 2012.

### 3.4.5 Coho Subyearlings (BY2014)

A total of only three coho subyearling parr did not allow us to make any attempts to build a species/age specific a regression model at the new trap location. The subyearling spring chinook flow-efficiency regression model was used to estimate subyearling coho parr emigrants. We estimated that $84( \pm 70 ; 95 \% \mathrm{CI})$ emigrated past the trap in the fall of 2015 (Table 11).

### 3.5 PIT Tagging

During the 2015 trapping season, we PIT tagged 361 wild spring Chinook, 383 steelhead, and 2 naturally produced coho (Table 12). All tagging files were submitted to the PTAGIS database. One shed PIT tag (implanted in steelhead parr) was recovered in holding boxes where fish had been held for 24-72 hours after tagging.

Table 11. Number of PIT tagged coho, Chinook, and steelhead with shed rates at the Nason Creek rotary trap in 2015.

| Species/Stage | Year-to- <br> date Catch | Year-to- <br> date PIT <br> Tagged | No. of Shed <br> Tags | Percent <br> Shed Tags |
| :--- | :---: | :---: | :---: | :---: |
| Chinook Yearling Smolt | 152 | 142 | 0 | $0.00 \%$ |
| Chinook Subyearling Parr (Mar 1 to June 30) | 111 | 28 | 0 | $0.00 \%$ |
| Chinook Subyearling Parr (July 1 to Nov 30) | 201 | 191 | 0 | $0.00 \%$ |
| Steelhead Parr | 388 | 371 | 1 | $0.27 \%$ |
| Steelhead Smolt | 12 | 12 | 0 | $0.00 \%$ |


| Coho Yearling Smolt | 2 | 2 | 0 | $0.00 \%$ |
| :--- | :--- | :--- | :---: | :---: |
| Coho Subyearling Parr | 5 | 0 | - | - |

* Counts do not include fish with $\mathrm{FL}<50 \mathrm{~mm}$ (fry).

During remote tagging efforts in the fall of 2014, 1,893 spring Chinook were PIT tagged by YNFRM and WDFW personnel. Of the total tagged, $78 \%$ were held overnight to determine tag retention. Shed rate for this tagging effort was $0.07 \%$.

### 3.6 Incidental Species

Along with wild spring Chinook, wild steelhead/rainbow trout, and naturally produced coho, other resident fish species captured at the Nason Creek rotary trap and included in Table 13 are: bull trout Salvelinus confluentus, cutthroat trout Oncorhynchus clarki, fathead minnow Pimephales promelas, longnose dace Rhinichthys cataractae, northern pikeminnow Ptychocheilus oregonensis, redside shiner Richardsonius balteatus, sculpin Cottus sp., sucker Catostomus sp., summer sockeye salmon fry Oncorhynchus nerka, and mountain whitefish Prosopium williamsoni.

Table 12. Summary of length and weight sampling of incidental species captured at the Nason Creek rotary trap in 2015.

| Species | Total Count | Length (mm) |  |  | Weight (g) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | N | SD | Mean | N | SD |
| Bull Trout | 1 | 180 | 1 | - | 50.1 | 1 | - |
| Cutthroat Trout | 1 | 168 | 1 | - | 45.3 | 1 | - |
| Fathead Minnow | 2 | 46 | 2 | 12.0 | 1.1 | 2 | 0.9 |
| Longnose Dace | 117 | 92 | 117 | 24.8 | 11.7 | 116 | 6.6 |
| Northern Pikeminnow | 11 | 142 | 11 | 78.9 | 58.4 | 11 | 78.8 |
| Redside Shiner | 8 | 58 | 8 | 13.8 | 2.8 | 7 | 1.1 |
| Sculpin | 81 | 78 | 81 | 38.7 | 12.3 | 78 | 17.3 |
| Sucker | 39 | 120 | 39 | 91.4 | 20.7 | 34 | 58.5 |
| Summer Sockeye Fry | 2 | 32 | 2 | 8.5 | 0.5 | 1 | - |
| Whitefish Fry | 4 | 40 | 4 | 9.3 | 0.8 | 3 | 0.1 |
| Whitefish | 21 | 97 | 21 | 68.8 | 25.0 | 20 | 65.5 |

### 3.7 ESA Compliance

The Nason Creek smolt trap was operated under consultation with NMFS and USFWS. Total numbers of UCR spring Chinook and UCR summer steelhead that were captured or handled (indirect take) at the trap were less than the maximum permitted ( $20 \%$ ) for each species. Lethal take was well below the allowable level of $2 \%$ for wild summer steelhead, hatchery summer steelhead, and bull trout (Table 14). Final spring Chinook lethal take for 2015 was at the $2 \%$ maximum. Exceedance of this maximum in early March was addressed in a memo sent to NMFS (See Appendix D). Stream temperatures did not exceed $18^{\circ} \mathrm{C}$ at any time in which fish were being handled.

Table 13. Summary of ESA species and coho salmon mortality at the Nason Creek rotary trap.

| Species/Stage/Brood Year | Total Collected | Total Mortality | \% Mortality |
| :--- | :---: | :---: | :---: |
| Spring Chinook Yearling (BY2013) | 152 |  | $3.29 \%$ |
| Spring Chinook Subyearling (BY 2014) | 548 | $9^{*}$ | $1.64 \%$ |
| Total Wild Spring Chinook | $\mathbf{7 0 0}$ | $\mathbf{1 4}$ | $\mathbf{2 . 0 0 \%}$ |
| Total Hatchery Spring Chinook | $\mathbf{7 1 4}$ | $\mathbf{0}$ | $\mathbf{0 . 0 0 \%}$ |
| Steelhead Age-0 (BY2015) | 182 | 1 | $0.55 \%$ |
| Steelhead Age-1 (BY2014) | 233 | 1 | $0.43 \%$ |
| Steelhead Age-2 (BY2013) | 28 | 0 | $0.00 \%$ |
| Steelhead Age-3 (BY2012) | 1 | 0 | $0.00 \%$ |
| Total Wild Summer Steelhead | $\mathbf{4 4 4}$ | $\mathbf{2}$ | $\mathbf{0 . 4 5 \%}$ |
| Total Hatchery Summer Steelhead | $\mathbf{4 4 8}$ | $\mathbf{1}$ | $\mathbf{0 . 2 2 \%}$ |
| Total Bull Trout | $\mathbf{1}$ | $\mathbf{0}$ | $\mathbf{0 . 0 0 \%}$ |
| Coho Yearling (BY2013) | 2 | 0 | $0.00 \%$ |
| Coho Subyearling (BY2014) | 5 | 1 | $\mathbf{2 0 . 0 0 \%}$ |
| Total Naturally-Produced Coho | $\mathbf{7}$ | $\mathbf{1}$ | $\mathbf{1 4 . 2 9 \%}$ |
| Majority occurring during incident detailed in Appendix D |  |  |  |

[^18]
### 4.0 DISCUSSION

Operations in 2015 marked the first full season of continuous trapping at the Bolser site. Preliminary trapping this new site has achieved the goal of minimizing interactions with the public; we have yet to encounter any act of vandalism or tampering with the trap since the move. Aside from the benefit of added safety to the public and captured fish, relocation of the Nason Creek trap was intended to improve the quality of data collected via simplified trapping regime and favorable channel morphology. Initial subyearling Chinook releases in the fall of 2014 suggested that the flow-efficiency relationship was statistically significant at the flows tested ( $r^{2}$ $=0.63, p=0.007$ ). However, in three of the contributing trials, a stoppage or inconsistent operation during the recapture period dictated that they be omitted from any expansions performed (non-continuous operation of the trap in the 3-day recapture period is a violation of our estimation protocol). Although the flow-efficiency regression was ultimately rendered unusable, subyearling Chinook efficiency trials in 2014 were an indication that a consistent flowefficiency relationship is present at the new site.

Attempts to further develop our flow-efficiency models in 2015 were largely prevented by extreme low spring/summer and high fall flow conditions, as well as low fish abundance. Steelhead and Chinook mark-group releases were generally small ( $n \leq 26$ ), providing little chance for recaptures given potentially low trap efficiency. A single large release of 138 subyearling spring Chinook on November 3 failed to produce any recaptures, initially suggesting a trap efficiency of less than $1.0 \%$. Later examination of daily subyearling spring Chinook catch showed that the release was performed concurrently with a significant drop in abundance, from 89 to 6 fish captured. The release also coincided with a rapidly decreasing hydrograph following a significant peak in discharge. The precipitous drop in catch may have resulted in a lack of active migration, with the spring Chinook subyearlings becoming less prone to downstream displacement as flows subsided. The suspected non-migratory behavior of spring Chinook subyearlings in Nason Creek during that period likely contributed to a lack of recaptures despite the large mark-group size. However, given that the trial occurred during the recognized subyearling spring Chinook migratory period and lacked any violations of release or trapping protocols, it was deemed valid.

With viable regression models unavailable for all species/stages, pooled estimates were predominantly used. These estimates were used as a means to produce some form of emigrant estimate, albeit with a higher degree of bias. All pooled estimates reported are considered provisional, and will be recalculated as viable flow-efficiency regressions are developed.

## Spring Chinook

Nason Creek spring Chinook egg-to-emigrant survival rates are generally lower than those of the Chiwawa River and White River populations (Figure 16). However, the 2013 Nason Creek spring Chinook brood deviated from this trend markedly, with an survival rate exceeding those of the other two tributaries. Whereas the Chiwawa River and White River populations saw egg-to-emigrant survival rates typical of their corresponding estimated egg depositions in 2013, Nason Creek produced an outlier value (Figures 13 \& 17). The total BY2013 spring Chinook estimate (excluding the non-trapping period) of 50,703 ( $\pm 38,852 ; 95 \% \mathrm{CI}$ ) emigrants greatly exceeded the corresponding 11-year average ( $n=23,211$ ).


Figure 16. Comparison of wild spring Chinook abundance estimates (BY2007-2013) made at the White River, Nason Creek, and Chiwawa River smolt traps. *Non-trapping estimates not included.


Figure 17. Comparison of egg-to-emigrant survival (BY 2007-2013) and egg deposition for Nason Creek, Chiwawa River, and White River spring Chinook. *Non-trapping estimates not included.

Though possible that the Nason Creek population alone saw above-average survival, it is likely that some degree of overestimation by our modeled and pooled estimates occurred. Composed
primarily of smaller ( $n \leq 96$ ) trials, the weighted (mark-group size) model was heavily influenced by the aforementiomned large ( $n=138$ ) release in 2015 that did not produce any recaptures. Because the unsuccessful trial was performed at the high end of the discharge range tested, it decreased the slope of the regression, and therefore the trap efficiencies used to expand catch at elevated flows. Additional trials at higher flows will mitigate the effect of this subyearling release outlier and likely produce a lower emigrant estimate when recalculated. Overestimation of the yearling pooled estimate was also likely influenced by a lack of consistent releases throughout the migratory period. We expect that eventual recalculation of BY2013 yearlings will also contribute to a lowering of the overall emigrant estimate.

The non-trapping period estimate of $6,822( \pm 9,035 ; 95 \% \mathrm{CI})$ BY2013 migrants suggests that movement out of the system was present in the winter, but at a much lower rate in comparison to the fall. Winter emigration for the 2013 spring Chinook brood accounted for $11.9 \%$ of the total estimate, whereas fall subyearling migrants made up a total of $76.0 \%$. Yearling spring emigrants composed a slightly larger proportion than non-trapping period, with $12.1 \%$ of the total run. Upon eventual recalculation of the BY2013 trapping estimates, proportion of non-trapping period to total run will likely increase as the smolt trap-derived estimates decrease. Although detections during the winter confirm movement, they are too few and infrequent to determine fine-scale temporal trends in emigration and/or relation to environmental conditions.

## Summer Steelhead

The pooled estimate used to expand 2015 steelhead migrants was based on 13 mark-groups; a total of 116 fish released, and 1 recapture. Consequently, the model tended to overestimate emigrant abundance as an efficiency of $0.86 \%$ was used to expand all daily catch. With no prior mark-group releases at this location, we are unsure if the low efficiency observed is accurate, or the product of the abnormally low water-year and its potential effects on steelhead migratory behavior. Comparisons of yearling Chinook and hatchery coho efficiencies at the new trap site to those of the old show they are comparatively lower, but not to the degree seen in 2015 summer steelhead migrants.

The total estimate of $25,566( \pm 6,020 ; 95 \%$ CI) BY2012 steelhead exceeded the 10-year mean of 15,380 emigrants, and was the second highest estimate in the past 10 broods. Although the model used to expand age- 3 fish was admittedly skewed toward overestimation, their contribution to the overall estimate was small ( $n=116$ ), and therefore did not impact it greatly. Both models used to calculate the bulk of the estimate (age-1 and age- 2 ) were satistically robust ( $\alpha \leq 0.05$ ); the product of trapping at the former site. The above-average emigrant survival and emigrants per redd of the 2012 brood despite relatively low egg deposition is characteristic of Nason Creek. In previous years, the highest rates of survival have corresponded to the lowest levels of spawner success, suggesting denstiy-dependence.

The migratory timing of summer steelhead captured in 2015 was typical of what we have previously seen in Nason Creek. Of the steelhead caught in the spring migratory period, $81.5 \%$ were were age-1, with age-2 (5.4\%) and age-3 (0.4\%) classes constituting a small portion of the total. The majority of the summer/fall non-migratory period was not trapped as a consequence of low flows. This period is normally dominated by young-of-the-year fry and parr.

## Coho

A poor return of adult coho in 2013 required exhaustive measures to collect program broodstock, including increased retention at Tumwater Dam (Kamphaus et al. 2016). As a result, a limited number of adult coho $(n=32)$ were allowed to pass into the upper-basin. Spawner escapement into Nason Creek was estimated at zero fish, with no redds documented during surveys in the fall of 2013. We attribute the capture of natural-origin coho to surveyor error, which may have lead to one or more redds to go unseen.

The BY2013 naturally-produced coho estimate of 161 ( $\pm 714 ; 95 \%$ CI) was likely overestimated to some degree by the under-developed models used for expansion. Despite the likely overestimation, the BY2013 estimate was less than the 10-year mean emigrant abundace ( $n=$ 562), and the third lowest estimate thus far at Nason Creek. We assume that the comparatively low estimate is a reflection of the poor spawner escapement of 2013. Recalculation of BY2013 emigrants will likely produce and even lower emigrant abudance.

## 2016 Trap Operations at Nason Creek

Pooled estimates have been used here, and in previous reports as an alternative when regression analysis is not feasible. However, this has proven problematic as each method requires a different efficiency-testing strategy. While flow-efficiency modeling can be built by gauging efficiency at specific flows over multiple years, a pooled estimate is based on regular releases over discrete strata. Pooled estimates based on few, unevenly-spaced releases will utimately be skewed toward the efficiencies of the discrete periods tested, not the entire migratory period. Recognizing the necessity to produce viable models depite potentially low emigrant abundaces in 2016, we have revised our system of efficiency trials to accommodate both pooled, and regression models. Along with the accustomed targeting of specific flows, regular releases at even intervals will occur throughout the year. Regardless of mark group size, or flows tested, migratory juveniles will be transported every three to four days upstream to be released. In doing so, we will ensure that estimates made with either methodology are as sound as possible.

Additionally, we will verify that the location of our upstream release point upholds smolt trapping assumption 3: that marked fish are randomly dispersed in the population prior to recapture. Currently, marked fish are released evenly on both sides of the creek to eliminate the potential bias of a single release point on one bank. In 2016, pre-release scans of both right, and left-bank release-groups will test if recapture probability differs depending on the side of the channel. In the event that recapture rates are markedly different between the two sites, we will pursue a different release point.

### 5.0 LITERATURE CITED

CBFWA (Columbia Basin Fish and Wildlife Authority). 1999. PIT tag marking procedures manual, version 2.0. Columbia Basin Fish and Wildlife Authority, Portland OR.

Everhart, W.H. and W.D. Youngs. 1981. Principles of Fishery Science, second edition. Comstock Publishing Associates, a division of Cornell University Press, Ithica and London.

Hillman, T.W. 2004. Monitoring strategy for the Upper Columbia Basin: Draft report February 1, 2004. Prepared for Upper Columbia Regional Technical Team, Wenatchee, Washington.

Hillman, T.W., P. Graf, B. Ishida, M. Johnson, C. Kamphaus, M. Miller, C. Moran, A. Murdoch, T. Pearsons, M. Tonseth, and C. Willard. 2015. Monitoring and Evaluation of the Chelan and Grant County PUD's Hatchery Programs: 2014 Annual Report. Prepared for The Habitat Conservation Plan Hatchery Committee and the Priest Rapids Coordinating Committee Hatchery Sub Committee. Wenatchee and Ephrata, WA.

Kamphaus, C.M., R. Alford, T. Jeffris, B. Ishida, and K. Mott. 2016. Mid-Columbia Coho Reintroduction Feasibility Study: 2013 Annual Report. Prepared for Bonneville Power Adimistration, Portland, Oregon, Public Utility District No. 1 of Chelan County, Wenatchee, Washington, and Public Utility District No. 2 of Grant County, Ephrata, Washington.

Murdoch, A. T. Miller, B. L. Truscott, C. Snow, C. Frady, K. Ryding, J. Arteburn and D. Hathaway. 2012. Upper Columbia Spring Chinook Salmon and Steelhead Juvenile and Adult Abundance, Productivity and Spatial Structure Monitoring. BPA Project 2010-034-00.

Murdoch, A., and K. Petersen. 2000. Freshwater Production and Emigration of Juvenile Spring Chinook from the Chiwawa River in 2000. Washington State Department of Fish and Wildlife

PTAGIS (Columbia Basin PIT Tag Information System). 2015. Interrogation Site Metadata: http://www.ptagis.org/sites/interrogation-site-metadata?IntSiteCode=NAL

Seber, G.A.F. 1982. The Estimation of Animal Abundance and Related Parameters, $2^{\text {nd }}$ edition. Edward Arnold: London

Tussing, S.P. 2008. A Field Manual of Scientific Protocols for Downstream Migrant Trapping within the Upper Columbia Monitoring Strategy: 2008 Working Version 1.0. Prepared for Bonneville Power Administration's Integrated Status and Effectiveness Monitoring Program.

UCRTT (Upper Columbia Regional Technical Team). 2001. A Strategy to Protect and Restore Salmonid Habitat in the Upper Columbia Region, a Discussion Draft Report. Upper Columbia Salmon Recovery Board.

USFS (United States Forest Service). 1996. Nason Creek Stream Survey Report.

WDOE (Washington State Department of Ecology). 2015. River and Stream Flow Monitoring: https://fortress.wa.gov/ecy/wrx/wrx/flows/station.asp?sta=45J070

YNFRM (Yakama Nation Fisheries Resource Management). 2010. Mid-Columbia Coho Restoration Master Plan. Prepared for: Northwest Power and Conservation Council, Portland OR.

APPENDIX A. Daily Stream Discharge and Stream Temperature

|  | Stream |  | 2/10/2015 | 804 | 3.7 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Date | Discharge | Temperature | 2/11/2015 | 756 | 3.5 |
|  | (CFS) | $\left({ }^{\circ} \mathrm{C}\right)$ | 2/12/2015 | 675 | 3.8 |
| 1/1/2015 |  | 0.0 | 2/13/2015 | 674 | 3.9 |
| 1/2/2015 |  | 0.0 | 2/13/2015 | 674 | 3.9 |
| 1/3/2015 |  | 0.0 | 2/14/2015 | 677 | 3.9 |
| 1/4/2015 |  | 0.0 | 2/15/2015 | 653 | 3.1 |
| 1/5/2015 |  | 0.2 | 2/16/2015 | 587 | 2.6 |
| 1/6/2015 | 1110 | 1.8 | 2/17/2015 | 536 | 2.6 |
| 1/7/2015 | 723 | 2.2 | 2/18/2015 | 492 | 2.6 |
| 1/8/2015 | 607 | 1.9 | 2/19/2015 | 463 | 3.6 |
| 1/9/2015 | 534 | 2.4 | 2/20/2015 | 447 | 3.9 |
| 1/10/2015 | 485 | 2.4 | 2/21/2015 | 422 | 3.3 |
| 1/11/2015 | 444 | 2.6 | 2/22/2015 | 387 | 2.6 |
| 1/12/2015 | 402 | 2.6 | 2/23/2015 | 357 | 1.9 |
| 1/13/2015 | 368 | 2.3 | 2/24/2015 | 341 | 2.5 |
| 1/14/2015 | 343 | 2.1 | 2/25/2015 | 323 | 3.4 |
| 1/15/2015 | 319 | 1.7 | 2/26/2015 | 312 | 4.2 |
| 1/16/2015 | 311 | 1.3 | 2/27/2015 | 317 | 4.0 |
| 1/17/2015 | 296 | 1.1 | 2/28/2015 | 295 | 3.3 |
| 1/18/2015 | 338 | 0.5 | 3/1/2015 | 276 | 2.7 |
| 1/19/2015 | 375 | 2.0 | 3/2/2015 | 264 | 3.2 |
| 1/20/2015 | 318 | 1.6 | 3/3/2015 | 247 | 2.4 |
| 1/21/2015 | 285 | 0.8 | 3/4/2015 | 238 | 2.2 |
| 1/22/2015 | 272 | 1.7 | 3/5/2015 | 232 | 2.8 |
| 1/23/2015 | 286 | 2.5 | 3/6/2015 | 225 | 3.9 |
| 1/24/2015 | 691 | 2.5 | 3/7/2015 | 224 | 4.3 |
| 1/25/2015 | 781 | 2.7 | 3/8/2015 | 226 | 4.3 |
| 1/26/2015 | 673 | 2.5 | 3/9/2015 | 227 | 4.5 |
| 1/27/2015 | 632 | 2.8 | 3/10/2015 | 231 | 4.4 |
| 1/28/2015 | 613 | 2.9 | 3/11/2015 | 237 | 5.2 |
| 1/29/2015 | 556 | 2.4 | 3/12/2015 | 285 | 5.8 |
| 1/30/2015 | 503 | 2.1 | 3/13/2015 | 303 | 4.9 |
| 1/31/2015 | 463 | 2.2 | 3/14/2015 | 526 | 5.3 |
| 2/1/2015 | 433 | 2.3 | 3/15/2015 | 733 | 3.9 |
| 2/2/2015 | 417 | 2.2 | 3/16/2015 | 624 | 4.0 |
| 2/3/2015 | 438 | 2.8 | 3/17/2015 | 517 | 4.2 |
| 2/4/2015 | 392 | 2.8 | 3/18/2015 | 457 | 4.9 |
| 2/5/2015 | 404 | 2.4 | 3/19/2015 | 422 | 4.8 |
| 2/6/2015 | 701 | 2.8 | 3/20/2015 | 402 | 5.3 |
| 2/7/2015 | 832 | 3.1 | 3/21/2015 | 434 | 5.5 |
| 2/8/2015 | 929 | 3.2 | 3/22/2015 | 426 | 4.2 |
| 2/9/2015 | 829 | 3.6 | 3/23/2015 | 389 | 4.5 |


| 3/24/2015 | 366 | 5.1 | 5/8/2015 | 297 | 8.9 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 3/25/2015 | 368 | 4.7 | 5/9/2015 | 307 | 9.2 |
| 3/26/2015 | 506 | 5.7 | 5/10/2015 | 334 | 9.2 |
| 3/27/2015 | 488 | 5.8 | 5/11/2015 | 371 | 10.0 |
| 3/28/2015 | 632 | 5.9 | 5/12/2015 | 408 | 7.9 |
| 3/29/2015 | 575 | 5.6 | 5/13/2015 | 416 | 7.5 |
| 3/30/2015 | 537 | 6.1 | 5/14/2015 | 418 | 8.0 |
| 3/31/2015 | 550 | 5.8 | 5/15/2015 | 379 | 9.1 |
| 4/1/2015 | 486 | 4.8 | 5/16/2015 | 374 | 9.9 |
| 4/2/2015 | 435 | 4.7 | 5/17/2015 | 373 | 8.4 |
| 4/3/2015 | 401 | 4.4 | 5/18/2015 | 392 | 10.0 |
| 4/4/2015 | 372 | 4.7 | 5/19/2015 | 421 | 10.4 |
| 4/5/2015 | 347 | 4.2 | 5/20/2015 | 437 | 10.8 |
| 4/6/2015 | 325 | 4.1 | 5/21/2015 | 435 | 10.8 |
| 4/7/2015 | 308 | 4.4 | 5/22/2015 | 421 | 10.4 |
| 4/8/2015 | 291 | 5.3 | 5/23/2015 | 416 | 11.4 |
| 4/9/2015 | 281 | 5.7 | 5/24/2015 | 409 | 11.5 |
| 4/10/2015 | 271 | 5.7 | 5/25/2015 | 378 | 10.9 |
| 4/11/2015 | 282 | 5.7 | 5/26/2015 | 337 | 10.3 |
| 4/12/2015 | 277 | 4.3 | 5/27/2015 | 310 | 11.5 |
| 4/13/2015 | 263 | 4.8 | 5/28/2015 | 315 | 12.1 |
| 4/14/2015 | 256 | 5.5 | 5/29/2015 | 330 | 11.9 |
| 4/15/2015 | 239 | 5.3 | 5/30/2015 | 365 | 12.7 |
| 4/16/2015 | 235 | 6.2 | 5/31/2015 | 310 | 12.2 |
| 4/17/2015 | 251 | 7.3 | 6/1/2015 | 272 | 11.9 |
| 4/18/2015 | 272 | 7.8 | 6/2/2015 | 257 | 11.2 |
| 4/19/2015 | 282 | 8.0 | 6/3/2015 | 236 | 11.8 |
| 4/20/2015 | 311 | 8.3 | 6/4/2015 | 218 | 12.6 |
| 4/21/2015 | 359 | 8.2 | 6/5/2015 | 205 | 13.8 |
| 4/22/2015 | 386 | 7.2 | 6/6/2015 | 200 | 15.0 |
| 4/23/2015 | 337 | 6.0 | 6/7/2015 | 198 | 15.9 |
| 4/24/2015 | 320 | 5.5 | 6/8/2015 | 192 | 16.5 |
| 4/25/2015 | 295 | 5.7 | 6/9/2015 | 182 | 16.3 |
| 4/26/2015 | 274 | 5.9 | 6/10/2015 | 168 | 16.1 |
| 4/27/2015 | 269 | 7.9 | 6/11/2015 | 154 | 15.6 |
| 4/28/2015 | 305 | 8.7 | 6/12/2015 | 145 | 14.6 |
| 4/29/2015 | 335 | 8.1 | 6/13/2015 | 134 | 13.8 |
| 4/30/2015 | 317 | 7.7 | 6/14/2015 | 124 | 14.4 |
| 5/1/2015 | 316 | 8.6 | 6/15/2015 | 116 | 14.9 |
| 5/2/2015 | 338 | 8.5 | 6/16/2015 | 109 | 16.0 |
| 5/3/2015 | 329 | 8.2 | 6/17/2015 | 104 | 16.6 |
| 5/4/2015 | 340 | 8.4 | 6/18/2015 | 100 | 16.0 |
| 5/5/2015 | 370 | 7.9 | 6/19/2015 | 97.2 | 15.4 |
| 5/6/2015 | 330 | 6.6 | 6/20/2015 | 95.1 | 15.2 |
| 5/7/2015 | 299 | 7.8 | 6/21/2015 | 90.3 | 15.2 |


| 6/22/2015 | 85.9 | 15.6 | 8/6/2015 | 33.7 | 18.2 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 6/23/2015 | 82.1 | 16.4 | 8/7/2015 | 34.1 | 18.2 |
| 6/24/2015 | 79.6 | 16.7 | 8/8/2015 | 33.1 | 18.8 |
| 6/25/2015 | 76.9 | 17.2 | 8/9/2015 | 32.3 | 18.8 |
| 6/26/2015 | 74 | 19.2 | 8/10/2015 | 32.2 | 19.9 |
| 6/27/2015 | 72.1 | 20.0 | 8/11/2015 | 31.8 | 18.7 |
| 6/28/2015 | 70.2 | 20.6 | 8/12/2015 | 31.7 | 19.1 |
| 6/29/2015 | 71.8 | 21.1 | 8/13/2015 | 30.4 | 20.4 |
| 6/30/2015 | 70.5 | 21.2 | 8/14/2015 | 30.3 | 19.8 |
| 7/1/2015 | 66.2 | 21.1 | 8/15/2015 | 31.9 | 17.6 |
| 7/2/2015 | 63.8 | 21.2 | 8/16/2015 | 33.4 | 16.9 |
| 7/3/2015 | 61.1 | 21.3 | 8/17/2015 | 32.2 | 17.7 |
| 7/4/2015 | 58.8 | 21.3 | 8/18/2015 | 31.2 | 18.2 |
| 7/5/2015 | 56.8 | 20.9 | 8/19/2015 | 30 | 18.9 |
| 7/6/2015 | 55.2 | 20.6 | 8/20/2015 | 28.9 | 19.2 |
| 7/7/2015 | 53.5 | 20.3 | 8/21/2015 | 28.5 | 18.0 |
| 7/8/2015 | 52.5 | 20.8 | 8/22/2015 | 28.7 | 16.4 |
| 7/9/2015 | 50.9 | 21.3 | 8/23/2015 | 28.6 | 16.0 |
| 7/10/2015 | 49.7 | 20.7 | 8/24/2015 | 28 | 16.8 |
| 7/11/2015 | 49.5 | 18.8 | 8/25/2015 | 27.5 | 17.0 |
| 7/12/2015 | 50.2 | 17.8 | 8/26/2015 | 27.5 | 17.1 |
| 7/13/2015 | 48.9 | 18.4 | 8/27/2015 | 27 | 17.8 |
| 7/14/2015 | 47.8 | 18.8 | 8/28/2015 | 27.1 | 17.9 |
| 7/15/2015 | 46.5 | 18.7 | 8/29/2015 | 29 | 16.5 |
| 7/16/2015 | 45.3 | 18.4 | 8/30/2015 | 37.1 | 15.5 |
| 7/17/2015 | 44.8 | 18.5 | 8/31/2015 | 49.4 | 14.1 |
| 7/18/2015 | 43.9 | 19.5 | 9/1/2015 | 43.9 | 14.2 |
| 7/19/2015 | 42.7 | 20.9 | 9/2/2015 | 47.7 | 14.5 |
| 7/20/2015 | 41.1 | 21.3 | 9/3/2015 | 48.1 | 13.3 |
| 7/21/2015 | 40.1 | 19.7 | 9/4/2015 | 42.2 | 12.8 |
| 7/22/2015 | 39.7 | 18.4 | 9/5/2015 | 37.1 | 12.9 |
| 7/23/2015 | 39.6 | 18.3 | 9/6/2015 | 38.6 | 12.7 |
| 7/24/2015 | 39.3 | 18.1 | 9/7/2015 | 48 | 13.4 |
| 7/25/2015 | 40 | 17.3 | 9/8/2015 | 40.4 | 13.9 |
| 7/26/2015 | 42.7 | 17.2 | 9/9/2015 | 37.2 | 14.6 |
| 7/27/2015 | 41.5 | 17.0 | 9/10/2015 | 34.7 | 15.3 |
| 7/28/2015 | 40.2 | 17.7 | 9/11/2015 | 33 | 15.5 |
| 7/29/2015 | 38.8 | 18.9 | 9/12/2015 | 32 | 15.9 |
| 7/30/2015 | 37.2 | 19.5 | 9/13/2015 | 30.6 | 16.3 |
| 7/31/2015 | 35.9 | 19.8 | 9/14/2015 | 30.1 | 13.8 |
| 8/1/2015 | 34.7 | 20.0 | 9/15/2015 | 30.5 | 12.1 |
| 8/2/2015 | 33.9 | 20.0 | 9/16/2015 | 30.8 | 11.9 |
| 8/3/2015 | 33.1 | 19.3 | 9/17/2015 | 31.4 | 11.8 |
| 8/4/2015 | 33.7 | 18.8 | 9/18/2015 | 34.3 | 12.7 |
| 8/5/2015 | 33.2 | 18.5 | 9/19/2015 | 34.2 | 12.8 |


| 9/20/2015 | 33.4 | 13.7 | 11/4/2015 | 333 | 4.6 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 9/21/2015 | 38.1 | 14.1 | 11/5/2015 | 280 | 5.4 |
| 9/22/2015 | 38 | 12.2 | 11/6/2015 | 249 | 5.3 |
| 9/23/2015 | 33.8 | 11.6 | 11/7/2015 | 228 | 6.0 |
| 9/24/2015 | 32.5 | 12.5 | 11/8/2015 | 263 | 6.1 |
| 9/25/2015 | 32 | 13.0 | 11/9/2015 | 245 | 5.3 |
| 9/26/2015 | 32 | 12.7 | 11/10/2015 |  |  |
| 9/27/2015 | 31.7 | 10.6 | 11/11/2015 |  |  |
| 9/28/2015 | 31.3 | 10.0 | 11/12/2015 |  |  |
| 9/29/2015 | 30.8 | 9.9 | 11/13/2015 | 1450 | 4.2 |
| 9/30/2015 | 30.5 | 10.3 | 11/14/2015 | 2250 | 5.3 |
| 10/1/2015 | 30.1 | 10.9 | 11/15/2015 | 1220 | 5.0 |
| 10/2/2015 | 29.7 | 11.5 | 11/16/2015 |  |  |
| 10/3/2015 | 29.5 | 12.1 | 11/17/2015 |  | 2.6 |
| 10/4/2015 | 30 | 11.3 | 11/18/2015 |  | 3.5 |
| 10/5/2015 | 30.1 | 10.4 | 11/19/2015 | 1410 | 3.9 |
| 10/6/2015 | 29.9 |  | 11/20/2015 | 938 | 2.9 |
| 10/7/2015 | 31.9 | 11.1 | 11/21/2015 | 728 | 2.1 |
| 10/8/2015 | 46.9 | 11.6 | 11/22/2015 | 607 | 2.0 |
| 10/9/2015 | 42.8 | 12.2 | 11/23/2015 | 520 | 2.1 |
| 10/10/2015 | 42.2 | 12.3 | 11/24/2015 | 457 | 2.8 |
| 10/11/2015 | 111 | 11.0 | 11/25/2015 | 391 | 2.1 |
| 10/12/2015 | 78.3 | 9.9 | 11/26/2015 | 343 | 0.8 |
| 10/13/2015 | 56.2 | 11.4 | 11/27/2015 | 313 | 0.4 |
| 10/14/2015 | 53.4 | 9.6 | 11/28/2015 | 288 | 0.1 |
| 10/15/2015 | 47.9 | 8.8 | 11/29/2015 | 273 | 0.0 |
| 10/16/2015 | 44.8 | 8.5 | 11/30/2015 | 251 | 0.2 |
| 10/17/2015 | 43 | 9.1 | 12/1/2015 | 234 | 0.4 |
| 10/18/2015 | 43.8 | 10.6 | 12/2/2015 | 226 | 0.8 |
| 10/19/2015 | 46.8 | 10.9 | 12/3/2015 | 222 | 0.7 |
| 10/20/2015 | 46 | 10.4 | 12/4/2015 | 210 | 2.0 |
| 10/21/2015 | 45.5 | 9.3 | 12/5/2015 | 203 | 1.7 |
| 10/22/2015 | 43 | 8.9 | 12/6/2015 | 198 | 1.5 |
| 10/23/2015 | 41.7 | 7.8 | 12/7/2015 |  |  |
| 10/24/2015 | 40.7 | 7.0 | 12/8/2015 | 848 | 1.5 |
| 10/25/2015 | 40.9 | 7.4 | 12/9/2015 | 2730 | 1.6 |
| 10/26/2015 | 42.8 | 8.7 | 12/10/2015 | 1370 | 2.3 |
| 10/27/2015 | 45.7 | 8.2 | 12/11/2015 | 915 | 2.9 |
| 10/28/2015 |  |  | 12/12/2015 |  |  |
| 10/29/2015 | 54.9 | 8.5 | 12/13/2015 |  |  |
| 10/30/2015 | 338 | 8.3 | 12/14/2015 | 551 | 2.7 |
| 10/31/2015 | 1800 | 7.6 | 12/15/2015 | 486 | 2.5 |
| 11/1/2015 | 1430 | 6.8 | 12/16/2015 | 444 | 2.5 |
| 11/2/2015 | 745 | 6.2 | 12/17/2015 | 409 | 1.0 |
| 11/3/2015 | 460 | 5.6 | 12/18/2015 | 387 | 0.7 |


| $12 / 19 / 2015$ | 357 | 1.3 |
| :--- | :--- | :--- |
| $12 / 20 / 2015$ | 332 | 1.5 |
| $12 / 21 / 2015$ | 318 | 0.8 |
| $12 / 22 / 2015$ | 298 | 1.2 |
| $12 / 23 / 2015$ | 285 | 1.1 |
| $12 / 24 / 2015$ | 269 | 1.0 |
| $12 / 25 / 2015$ | 248 | 1.3 |
| $12 / 26 / 2015$ | 232 | 0.7 |
| $12 / 27 / 2015$ | 225 | 0.3 |
| $12 / 28 / 2015$ | 217 | 0.7 |
| $12 / 29 / 2015$ | 207 | 1.2 |
| $12 / 30 / 2015$ | 197 | 0.8 |
| $12 / 31 / 2015$ | 184 | 0.1 |

## APPENDIX B. Daily Trap Operation

| Date | Trap Status | Comments | $\begin{aligned} & 4 / 10 / 2015 \\ & 4 / 11 / 2015 \end{aligned}$ | Op. Op. |
| :---: | :---: | :---: | :---: | :---: |
| 3/1/2015 | Op. |  | 4/12/2015 | Op. |
| 3/2/2015 | Op. |  | 4/13/2015 | Op. |
| 3/3/2015 | Op. |  | 4/14/2015 | Op. |
| 3/4/2015 | Op. |  | 4/15/2015 | Op. |
| 3/5/2015 | Op. |  | 4/16/2015 | Op. |
| 3/6/2015 | Op. |  | 4/17/2015 | Op. |
| 3/7/2015 | Op. |  | 4/18/2015 | Op. |
| 3/8/2015 | Op. |  | 4/19/2015 | Op. |
| 3/9/2015 | Op. |  | 4/20/2015 | Op. |
| 3/10/2015 | Op. |  | 4/21/2015 | Op. |
| 3/11/2015 | Op. |  | 4/22/2015 | Op. |
| 3/12/2015 | Op. |  | 4/23/2015 | Op. |
| 3/13/2015 | Op. |  | 4/24/2015 | Op. |
| 3/14/2015 | No Op. | Stopped - debris | 4/25/2015 | Op. |
| 3/15/2015 | No Op. | Stopped - debris | 4/26/2015 | Op. |
| 3/16/2015 | Op. |  | 4/27/2015 | Op. |
| 3/17/2015 | Op. |  | 4/28/2015 | Op. |
| 3/18/2015 | Op. |  | 4/29/2015 | Op. |
| 3/19/2015 | Op. |  | 4/30/2015 | Op. |
| 3/20/2015 | Op. |  | 5/1/2015 | Op. |
| 3/21/2015 | Op. |  | 5/2/2015 | Op. |
| 3/22/2015 | Op. |  | 5/3/2015 | Op. |
| 3/23/2015 | Op. |  | 5/4/2015 | Op. |
| 3/24/2015 | Op. |  | 5/5/2015 | Op. |
| 3/25/2015 | Op. |  | 5/6/2015 | Op. |
| 3/26/2015 | Op. |  | 5/7/2015 | Op. |
| 3/27/2015 | Op. |  | 5/8/2015 | Op. |
| 3/28/2015 | Op. |  | 5/9/2015 | Op. |
| 3/29/2015 | Op. |  | 5/10/2015 | Op. |
| 3/30/2015 | Op. |  | 5/11/2015 | Op. |
| 3/31/2015 | Op. |  | 5/12/2015 | Op. |
| 4/1/2015 | Op. |  | 5/13/2015 | Op. |
| 4/2/2015 | Op. |  | 5/14/2015 | Op. |
| 4/3/2015 | Op. |  | 5/15/2015 | Op. |
| 4/4/2015 | Op. |  | 5/16/2015 | Op. |
| 4/5/2015 | Op. |  | 5/17/2015 | Op. |
| 4/6/2015 | Op. |  | 5/18/2015 | Op. |
| 4/7/2015 | Op. |  | 5/19/2015 | Op. |
| 4/8/2015 | Op. |  | 5/20/2015 | Op. |
| 4/9/2015 | Op. |  |  |  |


| 5/21/2015 | Op. |  |
| :---: | :---: | :---: |
| 5/22/2015 | Op. |  |
| 5/23/2015 | Op. |  |
| 5/24/2015 | Op. |  |
| 5/25/2015 | Op. |  |
| 5/26/2015 | Op. |  |
| 5/27/2015 | Op. |  |
| 5/28/2015 | Op. |  |
| 5/29/2015 | Op. |  |
| 5/30/2015 | Op. |  |
| 5/31/2015 | Op. |  |
| 6/1/2015 | Op. |  |
| 6/2/2015 | Op. |  |
| 6/3/2015 | Op. |  |
| 6/4/2015 | Op. |  |
| 6/5/2015 | Op. |  |
| 6/6/2015 | Op. |  |
| 6/7/2015 | Op. |  |
| 6/8/2015 | Op. |  |
| 6/9/2015 | Op. |  |
| 6/10/2015 | Op. |  |
| 6/11/2015 | Op. |  |
| 6/12/2015 | Op. |  |
| 6/13/2015 | Op. |  |
| 6/14/2015 | Op. |  |
| 6/15/2015 | Op. |  |
| 6/16/2015 | Op. |  |
| 6/17/2015 | No Op. | Stopped - debris |
| 6/18/2015 | Op. |  |
| 6/19/2015 | Op. |  |
| 6/20/2015 | Op. |  |
| 6/21/2015 | Op. |  |
| 6/22/2015 | Op. |  |
| 6/23/2015 | Op. |  |
| 6/24/2015 | Op. |  |
| 6/25/2015 | Op. |  |
| 6/26/2015 | Op. |  |
| 6/27/2015 | Op. |  |
| 6/28/2015 | Op. |  |
| 6/29/2015 | Op. |  |
| 6/30/2015 | Op. |  |
| 7/1/2015 | Op. |  |
| 7/2/2015 | No Op. | Stopped - bed contact |
| 7/3/2015 | Op. |  |


| 7/4/2015 | Op. |  |
| :---: | :---: | :---: |
| 7/5/2015 | Op. |  |
| 7/6/2015 | Op. |  |
| 7/7/2015 | Op. |  |
| 7/8/2015 | No Op. | Stopped - bed contact |
| 7/9/2015 | No Op. | Stopped - bed contact |
| 7/10/2015 | Op. |  |
| 7/11/2015 | Op. |  |
| 7/12/2015 | Op. |  |
| 7/13/2015 | Op. |  |
| 7/14/2015 | No Op. | Stopped - bed contact |
| 7/15/2015 | No Op. | Stopped - bed contact |
| 7/16/2015 | No Op. | Stopped - low flow |
| 7/17/2015 | No Op. | Stopped - low flow |
| 7/18/2015 | No Op. | Pulled - low water |
| 7/19/2015 | No Op. | Pulled - low water |
| 7/20/2015 | No Op. | Pulled - low water |
| 7/21/2015 | No Op. | Pulled - low water |
| 7/22/2015 | No Op. | Pulled - low water |
| 7/23/2015 | No Op. | Pulled - low water |
| 7/24/2015 | No Op. | Pulled - low water |
| 7/25/2015 | No Op. | Pulled - low water |
| 7/26/2015 | No Op. | Pulled - low water |
| 7/27/2015 | No Op. | Pulled - low water |
| 7/28/2015 | No Op. | Pulled - low water |
| 7/29/2015 | No Op. | Pulled - low water |
| 7/30/2015 | No Op. | Pulled - low water |
| 7/31/2015 | No Op. | Pulled - low water |
| 8/1/2015 | No Op. | Pulled - low water |
| 8/2/2015 | No Op. | Pulled - low water |
| 8/3/2015 | No Op. | Pulled - low water |
| 8/4/2015 | No Op. | Pulled - low water |
| 8/5/2015 | No Op. | Pulled - low water |
| 8/6/2015 | No Op. | Pulled - low water |
| 8/7/2015 | No Op. | Pulled - low water |
| 8/8/2015 | No Op. | Pulled - low water |
| 8/9/2015 | No Op. | Pulled - low water |
| 8/10/2015 | No Op. | Pulled - low water |
| 8/11/2015 | No Op. | Pulled - low water |
| 8/12/2015 | No Op. | Pulled - low water |
| 8/13/2015 | No Op. | Pulled - low water |
| 8/14/2015 | No Op. | Pulled-low water |


| 8/15/2015 | No Op. | Pulled - low water | 9/29/2015 | No Op. | Pulled - low water |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8/16/2015 | No Op. | Pulled - low water | 9/30/2015 | No Op. | Pulled - low water |
| 8/17/2015 | No Op. | Pulled - low water | 10/1/2015 | No Op. | Pulled - low water |
| 8/18/2015 | No Op. | Pulled - low water | 10/2/2015 | No Op. | Pulled - low water |
| 8/19/2015 | No Op. | Pulled - low water | 10/3/2015 | No Op. | Pulled - low water |
| 8/20/2015 | No Op. | Pulled - low water | 10/4/2015 | No Op. | Pulled - low water |
| 8/21/2015 | No Op. | Pulled - low water | 10/5/2015 | No Op. | Pulled - low water |
| 8/22/2015 | No Op. | Pulled - low water | 10/6/2015 | No Op. | Pulled - low water |
| 8/23/2015 | No Op. | Pulled - low water | 10/7/2015 | No Op. | Pulled - low water |
| 8/24/2015 | No Op. | Pulled - low water | 10/8/2015 | No Op. | Pulled - low water |
| 8/25/2015 | No Op. | Pulled - low water | 10/9/2015 | No Op. | Pulled - low water |
| 8/26/2015 | No Op. | Pulled - low water | 10/10/2015 | No Op. | Pulled - low water |
| 8/27/2015 | No Op. | Pulled-low water | 10/11/2015 | No Op. | Pulled - low water |
| 8/28/2015 | No Op. | Pulled - low water | 10/12/2015 | No Op. | Stopped - low flow |
| 8/29/2015 | No Op. | Pulled - low water | 10/13/2015 | No Op. | Pulled - low water |
| 8/30/2015 | No Op. | Pulled - low water | 10/14/2015 | No Op. | Pulled - low water |
| 8/31/2015 | No Op. | Pulled - low water | 10/15/2015 | No Op. | Pulled - low water |
| 9/1/2015 | No Op. | Pulled-low water | 10/16/2015 | No Op. | Pulled - low water |
| 9/2/2015 | No Op. | Pulled - low water | 10/17/2015 | No Op. | Pulled - low water |
| 9/3/2015 | No Op. | Stopped - low flow | 10/18/2015 | No Op. | Pulled - low water |
| 9/4/2015 | No Op. | Pulled - low water | 10/19/2015 | No Op. | Pulled - low water |
| 9/5/2015 | No Op. | Pulled-low water | 10/20/2015 | No Op. | Pulled - low water |
| 9/6/2015 | No Op. | Pulled - low water | 10/21/2015 | Op. |  |
| 9/7/2015 | No Op. | Pulled - low water | 10/22/2015 | Op. |  |
| 9/8/2015 | No Op. | Pulled-low water | 10/23/2015 | Op. |  |
| 9/9/2015 | No Op. | Pulled - low water | 10/24/2015 | No Op. | Stopped - low flow |
| 9/10/2015 | No Op. | Pulled - low water | 10/25/2015 | No Op. | Pulled - low water |
| 9/11/2015 | No Op. | Pulled - low water | 10/26/2015 | No Op. | Pulled - low water |
| 9/12/2015 | No Op. | Pulled-low water | 10/27/2015 | No Op. | Pulled - low water |
| 9/13/2015 | No Op. | Pulled - low water | 10/28/2015 | No Op. | Pulled - low water |
| 9/14/2015 | No Op. | Pulled - low water | 10/29/2015 | No Op. | Pulled - low water |
| 9/15/2015 | No Op. | Pulled-low water | 10/30/2015 | No Op. | Stopped - low flow |
| 9/16/2015 | No Op. | Pulled-low water | 10/31/2015 | No Op. | Pulled - low water |
| 9/17/2015 | No Op. | Pulled-low water | 11/1/2015 | No Op. | Pulled - low water |
| 9/18/2015 | No Op. | Pulled - low water | 11/2/2015 | Op. |  |
| 9/19/2015 | No Op. | Pulled - low water | 11/3/2015 | Op. |  |
| 9/20/2015 | No Op. | Pulled - low water | 11/4/2015 | Op. |  |
| 9/21/2015 | No Op. | Pulled-low water | 11/5/2015 | Op. |  |
| 9/22/2015 | No Op. | Pulled-low water | 11/6/2015 | Op. |  |
| 9/23/2015 | No Op. | Pulled - low water | 11/7/2015 | Op. |  |
| 9/24/2015 | No Op. | Pulled-low water | 11/8/2015 | Op. |  |
| 9/25/2015 | No Op. | Pulled-low water | 11/9/2015 | Op. |  |
| 9/26/2015 | No Op. | Pulled - low water | 11/10/2015 | Op. |  |
| 9/27/2015 | No Op. | Pulled - low water | 11/11/2015 | Op. |  |
| 9/28/2015 | No Op. | Pulled - low water | 11/12/2015 | Op. |  |


| $11 / 13 / 2015$ | No Op. | Pulled - high water |
| :---: | :---: | :---: |
| $11 / 14 / 2015$ | No Op. | Pulled - high water |
| $11 / 15 / 2015$ | No Op. | Pulled - high water |
| $11 / 16 / 2015$ | Op. |  |
| $11 / 17 / 2015$ | Op. |  |
| $11 / 18 / 2015$ | No Op. | Pulled - high water |
| $11 / 19 / 2015$ | No Op. | Pulled - high water |
| $11 / 20 / 2015$ | Op. |  |
| $11 / 21 / 2015$ | Op. |  |
| $11 / 22 / 2015$ | Op. |  |
| $11 / 23 / 2015$ | Op. |  |
| $11 / 24 / 2015$ | Op. |  |
| $11 / 25 / 2015$ | Op. |  |
| $11 / 26 / 2015$ | Op. |  |
| $11 / 27 / 2015$ | Op. |  |
| $11 / 28 / 2015$ | No Op. | Stopped - ice |
| $11 / 29 / 2015$ | No Op. | Stopped - ice |
| $11 / 30 / 2015$ | No Op. | Stopped - ice |

## APPENDIX C. Regression Models

Model: Chinook Yearlings (Spring '06-'14) Back Position, $\left(r^{2}=0.15 ; p=0.03\right)$

|  |  |  | Date | Trap <br> Position | Mark | Recap | Trap <br> Efficiency <br> $(\mathrm{R}+1) / \mathrm{M}$ | ASIN <br> Transform |
| :--- | :--- | :---: | :--- | :---: | :---: | :---: | :---: | :---: |
| Origin/Species/Stage | Age |  |  |  |  |  |  |  |
| Wild Chinook Smolt | $1+$ | $3 / 31 / 2007$ | Back | 40 | 2 | 0.08 | 0.28 | 869 |
| Wild Chinook Smolt | $1+$ | $4 / 6 / 2006$ | Back | 42 | 9 | 0.24 | 0.51 | 264 |
| Wild Chinook Smolt | $1+$ | $4 / 14 / 2010$ | Back | 42 | 4 | 0.12 | 0.35 | 173 |
| Wild Chinook Smolt | $1+$ | $3 / 31 / 2012$ | Back | 43 | 5 | 0.14 | 0.38 | 250 |
| Wild Chinook Smolt | $1+$ | $4 / 3 / 2007$ | Back | 46 | 1 | 0.04 | 0.21 | 656 |
| Wild Chinook Smolt | $1+$ | $4 / 19 / 2012$ | Back | 48 | 7 | 0.17 | 0.42 | 434 |
| Wild Chinook Smolt | $1+$ | $4 / 10 / 2007$ | Back | 53 | 4 | 0.09 | 0.31 | 966 |
| Wild Chinook Smolt | $1+$ | $4 / 21 / 2009$ | Back | 53 | 0 | 0.02 | 0.14 | 732 |
| Wild Chinook Smolt | $1+$ | $4 / 13 / 2012$ | Back | 53 | 4 | 0.09 | 0.31 | 358 |
| Wild Chinook Smolt | $1+$ | $4 / 16 / 2012$ | Back | 53 | 7 | 0.15 | 0.40 | 443 |
| Wild Chinook Smolt | $1+$ | $4 / 24 / 2008$ | Back | 57 | 8 | 0.158 | 0.409 | 210 |
| Wild Chinook Smolt | $1+$ | $4 / 23 / 2012$ | Back | 58 | 1 | 0.034 | 0.187 | 1380 |
| Wild Chinook Smolt | $1+$ | $4 / 24 / 2006$ | Back | 59 | 3 | 0.068 | 0.263 | 368 |
| Wild Chinook Smolt | $1+$ | $3 / 23 / 2007$ | Back | 59 | 7 | 0.136 | 0.377 | 876 |
| Wild Chinook Smolt | $1+$ | $3 / 17 / 2007$ | Back | 64 | 7 | 0.125 | 0.361 | 936 |
| Wild Chinook Smolt | $1+$ | $4 / 18 / 2010$ | Back | 67 | 2 | 0.045 | 0.213 | 330 |
| Wild Chinook Smolt | $1+$ | $4 / 17 / 2008$ | Back | 72 | 13 | 0.194 | 0.457 | 274 |
| Wild Chinook Smolt | $1+$ | $4 / 3 / 2006$ | Back | 81 | 10 | 0.136 | 0.377 | 188 |
| Wild Chinook Smolt | $1+$ | $3 / 20 / 2007$ | Back | 91 | 13 | 0.154 | 0.403 | 1230 |
| Wild Chinook Smolt | $1+$ | $5 / 1 / 2008$ | Back | 102 | 16 | 0.167 | 0.421 | 315 |
| Wild Chinook Smolt | $1+$ | $4 / 28 / 2008$ | Back | 127 | 19 | 0.157 | 0.408 | 271 |
| Wild Chinook Smolt | $1+$ | $4 / 14 / 2008$ | Back | 195 | 40 | 0.21 | 0.476 | 327 |
| Wild Chinook Smolt | $1+$ | $3 / 9 / 2014$ | Back | 65 | 4 | 0.077 | 0.281 | 958 |
| Wild Chinook Smolt | $1+$ | $3 / 13 / 2014$ | Back | 67 | 9 | 0.149 | 0.397 | 566 |

Model: Chinook Subyearling (Fall '06-'13) Back Position, $\left(r^{2}=0.55 ; p=0.001\right)$

| Origin/Species/Stage | Age | Date | Trap <br> Position | Mark | Recap | Trap <br> Efficiency <br> $(\mathrm{R}+1) / \mathrm{M}$ | ASIN <br> Transform | Discharge |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Parr | 0 | $10 / 26 / 2006$ | Back | 183 | 50 | 0.28 | 0.56 | 51 |
| Wild Chinook Parr | 0 | $10 / 30 / 2006$ | Back | 168 | 52 | 0.32 | 0.60 | 63 |
| Wild Chinook Parr | 0 | $11 / 1 / 2010$ | Back | 254 | 42 | 0.17 | 0.42 | 198 |
| Wild Chinook Parr | 0 | $11 / 4 / 2010$ | Back | 287 | 49 | 0.17 | 0.43 | 215 |
| Wild Chinook Parr | 0 | $11 / 7 / 2010$ | Back | 168 | 32 | 0.20 | 0.46 | 241 |
| Wild Chinook Parr | 0 | $11 / 13 / 2010$ | Back | 185 | 35 | 0.19 | 0.46 | 131 |

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| Wild Chinook Parr | 0 | $11 / 3 / 2012$ | Back | 201 | 25 | 0.13 | 0.37 | 402 |
| :--- | :--- | :---: | :--- | :--- | :--- | :--- | :--- | :--- |
| Wild Chinook Parr | 0 | $11 / 7 / 2012$ | Back | 233 | 27 | 0.12 | 0.35 | 394 |
| Wild Chinook Parr | 0 | $11 / 11 / 2012$ | Back | 328 | 87 | 0.27 | 0.54 | 217 |
| Wild Chinook Parr | 0 | $11 / 15 / 2012$ | Back | 195 | 34 | 0.18 | 0.44 | 213 |
| Wild Chinook Parr | 0 | $9 / 30 / 2013$ | Back | 171 | 12 | 0.08 | 0.28 | 542 |
| Wild Chinook Parr | 0 | $10 / 2 / 2013$ | Back | 213 | 43 | 0.21 | 0.47 | 328 |
| Wild Chinook Parr | 0 | $10 / 3 / 2013$ | Back | 181 | 41 | 0.23 | 0.50 | 296 |
| Wild Chinook Parr | 0 | $10 / 7 / 2013$ | Back | 242 | 31 | 0.13 | 0.37 | 233 |
| Wild Chinook Parr | 0 | $10 / 9 / 2013$ | Back | 203 | 40 | 0.20 | 0.47 | 303 |
| Wild Chinook Parr | 0 | $11 / 27 / 2013$ | Back | 241 | 55 | 0.23 | 0.50 | 182 |

Model: Chinook Subyearling (Fall '06-'13) Forward Position, $\left(r^{2}=0.16 ; p=0.02\right)$

|  |  |  |  | Trap |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Origin/Species/Stage | Age | Date | Mark | Recap | Trap <br> Efficiency <br> $($ R+1 $/ \mathrm{M}$ | ASIN <br> Transform | Discharge |  |
| Wild Chinook Parr | 0 | $7 / 13 / 2006$ | Back | 52 | 8 | 0.17 | 0.43 | 171 |
| Wild Chinook Parr | 0 | $7 / 17 / 2006$ | Back | 138 | 15 | 0.12 | 0.35 | 129 |
| Wild Chinook Parr | 0 | $7 / 20 / 2006$ | Back | 74 | 5 | 0.08 | 0.29 | 113 |
| Wild Chinook Parr | 0 | $7 / 28 / 2006$ | Back | 54 | 5 | 0.11 | 0.34 | 91 |
| Wild Chinook Parr | 0 | $7 / 31 / 2006$ | Back | 99 | 7 | 0.08 | 0.29 | 79 |
| Wild Chinook Parr | 0 | $9 / 18 / 2006$ | Back | 55 | 10 | 0.20 | 0.46 | 46 |
| Wild Chinook Parr | 0 | $7 / 31 / 2008$ | Back | 60 | 15 | 0.27 | 0.54 | 121 |
| Wild Chinook Parr | 0 | $8 / 12 / 2008$ | Back | 103 | 2 | 0.03 | 0.17 | 85.6 |
| Wild Chinook Parr | 0 | $8 / 22 / 2008$ | Back | 75 | 11 | 0.16 | 0.41 | 97 |
| Wild Chinook Parr | 0 | $8 / 28 / 2008$ | Back | 72 | 7 | 0.11 | 0.34 | 81.9 |
| Wild Chinook Parr | 0 | $10 / 9 / 2008$ | Back | 110 | 22 | 0.21 | 0.48 | 63.5 |
| Wild Chinook Parr | 0 | $10 / 27 / 2008$ | Back | 51 | 12 | 0.26 | 0.53 | 56.1 |
| Wild Chinook Parr | 0 | $10 / 30 / 2008$ | Back | 84 | 15 | 0.19 | 0.45 | 53 |
| Wild Chinook Parr | 0 | $11 / 6 / 2008$ | Back | 78 | 8 | 0.12 | 0.35 | 77.7 |
| Wild Chinook Parr | 0 | $11 / 10 / 2008$ | Back | 88 | 0 | 0.01 | 0.11 | 309 |
| Wild Chinook Parr | 0 | $7 / 14 / 2009$ | Back | 86 | 2 | 0.04 | 0.19 | 193 |
| Wild Chinook Parr | 0 | $7 / 15 / 2009$ | Back | 105 | 4 | 0.05 | 0.22 | 179 |
| Wild Chinook Parr | 0 | $7 / 17 / 2009$ | Back | 122 | 8 | 0.07 | 0.28 | 157 |
| Wild Chinook Parr | 0 | $7 / 20 / 2009$ | Back | 89 | 2 | 0.03 | 0.19 | 135 |
| Wild Chinook Parr | 0 | $8 / 17 / 2009$ | Back | 73 | 1 | 0.03 | 0.17 | 58 |
| Wild Chinook Parr | 0 | $9 / 10 / 2009$ | Back | 56 | 7 | 0.14 | 0.39 | 60 |
| Wild Chinook Parr | 0 | $8 / 8 / 2010$ | Back | 58 | 1 | 0.03 | 0.19 | 85 |
| Wild Chinook Parr | 0 | $8 / 11 / 2010$ | Back | 114 | 8 | 0.08 | 0.29 | 77 |
| Wild Chinook Parr | 0 | $9 / 11 / 2010$ | Back | 68 | 9 | 0.15 | 0.39 | 75 |
| Wild Chinook Parr | 0 | $10 / 12 / 2010$ | Back | 216 | 42 | 0.20 | 0.46 | 126 |
| Wild Chinook Parr | 0 | $10 / 15 / 2010$ | Back | 192 | 37 | 0.20 | 0.46 | 95 |
| Wild Chinook Parr | 0 | $10 / 18 / 2010$ | Back | 193 | 36 | 0.19 | 0.45 | 81 |
| Wild Chinook Parr | 0 | $10 / 22 / 2010$ | Back | 92 | 18 | 0.21 | 0.47 | 69 |
| Wild Chinook Parr | 0 | $10 / 25 / 2010$ | Back | 60 | 7 | 0.13 | 0.37 | 78 |
| Wild Chinook Parr | 0 | $10 / 29 / 2010$ | Back | 127 | 0 | 0.01 | 0.09 | 95.1 |
| Wild Chinook Parr | 0 | $8 / 19 / 2011$ | Back | 106 | 5 | 0.06 | 0.24 | 123 |
|  |  |  |  |  |  |  |  |  |

Model: Chinook Subyearling (Fall '14-'15) Bolser Site ( $r^{2}=0.36 ; p=0.09$ )

| Origin/Species/Stage | Age | Date | Trap <br> Position | Mark | Recap | Trap <br> Efficiency <br> $(\mathrm{R}+1) / \mathrm{M}$ | ASIN <br> Transform | Discharge |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Parr | 0 | $7 / 14 / 2014$ | Back | 89 | 7 | 0.09 | 0.30 | 171 |
| Wild Chinook Parr | 0 | $7 / 21 / 2014$ | Back | 74 | 4 | 0.07 | 0.26 | 129 |
| Wild Chinook Parr | 0 | $7 / 27 / 2014$ | Back | 72 | 4 | 0.07 | 0.27 | 113 |
| Wild Chinook Parr | 0 | $10 / 27 / 2014$ | Back | 71 | 3 | 0.06 | 0.24 | 91 |
| Wild Chinook Parr | 0 | $10 / 30 / 2014$ | Back | 70 | 5 | 0.09 | 0.30 | 79 |
| Wild Chinook Parr | 0 | $11 / 1 / 2014$ | Back | 96 | 6 | 0.07 | 0.27 | 46 |
| Wild Chinook Parr | 0 | $11 / 3 / 2015$ | Back | 138 | 0 | 0.01 | 0.09 | 121 |

Model: Summer Steelhead Back Position ('07-'14), $\left(r^{2}=0.35 ; p=2.90 \mathrm{E}-05\right)$

| Origin/Species/Stage | Age | Date | Trap <br> Position | Mark | Recap | Trap <br> Efficiency <br> $($ R +1$) /$ M | ASIN <br> Transform | Discharge |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Steelhead Parr/Smolt | $1+$ | $3 / 20 / 2007$ | Back | 55 | 1 | 0.04 | 0.19 | 1230 |
| Wild Steelhead Parr/Smolt | $1+$ | $3 / 31 / 2007$ | Back | 56 | 4 | 0.09 | 0.30 | 869 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 10 / 2007$ | Back | 60 | 8 | 0.15 | 0.40 | 966 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 1 / 2007$ | Back | 52 | 2 | 0.06 | 0.24 | 783 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 9 / 2007$ | Back | 71 | 9 | 0.14 | 0.38 | 842 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 12 / 2007$ | Back | 65 | 8 | 0.14 | 0.38 | 704 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 14 / 2007$ | Back | 61 | 5 | 0.10 | 0.32 | 687 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 21 / 2007$ | Back | 67 | 4 | 0.07 | 0.28 | 751 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 14 / 2008$ | Back | 149 | 46 | 0.32 | 0.60 | 327 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 17 / 2008$ | Back | 75 | 3 | 0.05 | 0.23 | 274 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 28 / 2008$ | Back | 74 | 11 | 0.16 | 0.41 | 271 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 1 / 2008$ | Back | 176 | 29 | 0.17 | 0.43 | 315 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 12 / 2008$ | Back | 55 | 8 | 0.16 | 0.42 | 663 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 15 / 2008$ | Back | 57 | 1 | 0.04 | 0.19 | 1390 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 9 / 2008$ | Back | 142 | 20 | 0.15 | 0.39 | 938 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 12 / 2008$ | Back | 83 | 10 | 0.13 | 0.37 | 823 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 16 / 2008$ | Back | 81 | 8 | 0.11 | 0.34 | 1140 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 20 / 2010$ | Back | 121 | 11 | 0.10 | 0.32 | 675 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 22 / 2010$ | Back | 121 | 10 | 0.09 | 0.31 | 726 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 20 / 2010$ | Back | 128 | 11 | 0.09 | 0.31 | 926 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 5 / 2011$ | Back | 52 | 1 | 0.04 | 0.20 | 761 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 22 / 2011$ | Back | 84 | 3 | 0.05 | 0.22 | 1540 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 12 / 2012$ | Back | 69 | 5 | 0.09 | 0.30 | 1170 |
| Wild Steelhead Parr/Smolt | $1+$ | $7 / 26 / 2012$ | Back | 63 | 4 | 0.08 | 0.29 | 278 |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 22 / 2013$ | Back | 66 | 6 | 0.11 | 0.33 | 520 |
|  |  |  | 59 |  |  |  |  |  |

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| Wild Steelhead Parr/Smolt | $1+$ | $4 / 26 / 2013$ | Back | 50 | 2 | 0.06 | 0.25 | 642 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Steelhead Parr/Smolt | $1+$ | $4 / 30 / 2013$ | Back | 54 | 2 | 0.06 | 0.24 | 778 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 8 / 2013$ | Back | 62 | 0 | 0.02 | 0.13 | 2170 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 19 / 2013$ | Back | 122 | 15 | 0.13 | 0.37 | 1130 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 22 / 2013$ | Back | 58 | 4 | 0.09 | 0.30 | 1080 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 26 / 2013$ | Back | 79 | 3 | 0.05 | 0.23 | 724 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 30 / 2013$ | Back | 92 | 7 | 0.09 | 0.30 | 849 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 3 / 2013$ | Back | 71 | 6 | 0.10 | 0.32 | 962 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 7 / 2013$ | Back | 94 | 4 | 0.05 | 0.23 | 1420 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 13 / 2013$ | Back | 64 | 2 | 0.05 | 0.22 | 745 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 17 / 2013$ | Back | 115 | 5 | 0.05 | 0.23 | 883 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 29 / 2013$ | Back | 60 | 12 | 0.22 | 0.48 | 730 |
| Wild Steelhead Parr/Smolt | $1+$ | $7 / 7 / 2013$ | Back | 75 | 9 | 0.13 | 0.37 | 325 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 5 / 2014$ | Back | 55 | 3 | 0.07 | 0.27 | 1260 |
| Wild Steelhead Parr/Smolt | $1+$ | $5 / 20 / 2014$ | Back | 57 | 0 | 0.02 | 0.13 | 1490 |
| Wild Steelhead Parr/Smolt | $1+$ | $6 / 3 / 2014$ | Back | 75 | 1 | 0.03 | 0.16 | 1610 |

Model: 2013 Summer Steelhead Back Position (In-yr.), ( $r^{2}=0.15 ; p=0.05$ )

|  |  | Date | Trap <br> Position | Mark | Recap | Trap <br> Efficiency <br> $($ R+1 $/ \mathrm{M}$ | ASIN <br> Transform | Discharge |
| :--- | :--- | :---: | :--- | :---: | :---: | :---: | :---: | :---: |
| Origin/Species/Stage | Age |  |  |  |  |  |  |  |
| Wild Chinook Smolt | $1+$ | $3 / 31 / 2007$ | Back | 40 | 2 | 0.08 | 0.28 | 869 |
| Wild Chinook Smolt | $1+$ | $4 / 6 / 2006$ | Back | 42 | 9 | 0.24 | 0.51 | 264 |
| Wild Chinook Smolt | $1+$ | $4 / 14 / 2010$ | Back | 42 | 4 | 0.12 | 0.35 | 173 |
| Wild Chinook Smolt | $1+$ | $3 / 31 / 2012$ | Back | 43 | 5 | 0.14 | 0.38 | 250 |
| Wild Chinook Smolt | $1+$ | $4 / 3 / 2007$ | Back | 46 | 1 | 0.04 | 0.21 | 656 |
| Wild Chinook Smolt | $1+$ | $4 / 19 / 2012$ | Back | 48 | 7 | 0.17 | 0.42 | 434 |
| Wild Chinook Smolt | $1+$ | $4 / 10 / 2007$ | Back | 53 | 4 | 0.09 | 0.31 | 966 |
| Wild Chinook Smolt | $1+$ | $4 / 21 / 2009$ | Back | 53 | 0 | 0.02 | 0.14 | 732 |
| Wild Chinook Smolt | $1+$ | $4 / 13 / 2012$ | Back | 53 | 4 | 0.09 | 0.31 | 358 |
| Wild Chinook Smolt | $1+$ | $4 / 16 / 2012$ | Back | 53 | 7 | 0.15 | 0.40 | 443 |
| Wild Chinook Smolt | $1+$ | $4 / 24 / 2008$ | Back | 57 | 8 | 0.158 | 0.409 | 210 |
| Wild Chinook Smolt | $1+$ | $4 / 23 / 2012$ | Back | 58 | 1 | 0.034 | 0.187 | 1380 |
| Wild Chinook Smolt | $1+$ | $4 / 24 / 2006$ | Back | 59 | 3 | 0.068 | 0.263 | 368 |
| Wild Chinook Smolt | $1+$ | $3 / 23 / 2007$ | Back | 59 | 7 | 0.136 | 0.377 | 876 |
| Wild Chinook Smolt | $1+$ | $3 / 17 / 2007$ | Back | 64 | 7 | 0.125 | 0.361 | 936 |
| Wild Chinook Smolt | $1+$ | $4 / 18 / 2010$ | Back | 67 | 2 | 0.045 | 0.213 | 330 |
| Wild Chinook Smolt | $1+$ | $4 / 17 / 2008$ | Back | 72 | 13 | 0.194 | 0.457 | 274 |
| Wild Chinook Smolt | $1+$ | $4 / 3 / 2006$ | Back | 81 | 10 | 0.136 | 0.377 | 188 |
| Wild Chinook Smolt | $1+$ | $3 / 20 / 2007$ | Back | 91 | 13 | 0.154 | 0.403 | 1230 |
| Wild Chinook Smolt | $1+$ | $5 / 1 / 2008$ | Back | 102 | 16 | 0.167 | 0.421 | 315 |
| Wild Chinook Smolt | $1+$ | $4 / 28 / 2008$ | Back | 127 | 19 | 0.157 | 0.408 | 271 |


| Wild Chinook Smolt | $1+$ | $4 / 14 / 2008$ | Back | 195 | 40 | 0.21 | 0.476 | 327 |
| :--- | :--- | :---: | :--- | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Smolt | $1+$ | $3 / 9 / 2014$ | Back | 65 | 4 | 0.077 | 0.281 | 958 |
| Wild Chinook Smolt | $1+$ | $3 / 13 / 2014$ | Back | 67 | 9 | 0.149 | 0.397 | 566 |

Model: Spring Chinook 2010-2014 Non-Trapping Period Array (NAL) Efficiency, ( $r^{2}=0.61 ; p$ $=0.0002$ )

| Origin/Species/Stage | Age | Date | Mark | Detections | Trap <br> Efficiency <br> $(\mathrm{R}+1) / \mathrm{M}$ | ASIN <br> Transform | Discharge |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Parr | 0 | $11 / 4 / 2010$ | 254 | 95 | 0.38 | 0.66 | 224 |
| Wild Chinook Parr | 0 | $11 / 7 / 2010$ | 287 | 70 | 0.25 | 0.52 | 248 |
| Wild Chinook Parr | 0 | $11 / 10 / 2010$ | 168 | 74 | 0.45 | 0.73 | 169 |
| Wild Chinook Parr | 0 | $11 / 13 / 2010$ | 74 | 41 | 0.57 | 0.85 | 140 |
| Wild Chinook Parr | 0 | $11 / 18 / 2010$ | 185 | 22 | 0.12 | 0.36 | 278 |
| Wild Chinook Parr | 0 | $11 / 3 / 2012$ | 201 | 21 | 0.11 | 0.34 | 384 |
| Wild Chinook Parr | 0 | $11 / 7 / 2012$ | 233 | 31 | 0.14 | 0.38 | 378 |
| Wild Chinook Parr | 0 | $11 / 11 / 2012$ | 328 | 66 | 0.20 | 0.47 | 223 |
| Wild Chinook Parr | 0 | $11 / 15 / 2012$ | 195 | 68 | 0.35 | 0.64 | 219 |
| Wild Chinook Parr | 0 | $11 / 4 / 2013$ | 130 | 51 | 0.40 | 0.68 | 130 |
| Wild Chinook Parr | 0 | $11 / 8 / 2013$ | 106 | 39 | 0.38 | 0.66 | 148 |
| Wild Chinook Parr | 0 | $3 / 9 / 2014$ | 65 | 4 | 0.08 | 0.28 | 880 |
| Wild Chinook Parr | 0 | $3 / 13 / 2014$ | 67 | 5 | 0.09 | 0.30 | 541 |
| Wild Chinook Parr | 0 | $11 / 4 / 2014$ | 114 | 5 | 0.05 | 0.23 | 370 |
| Wild Chinook Parr | 0 | $11 / 1 / 2014$ | 96 | 5 | 0.06 | 0.25 | 583 |
| Wild Chinook Parr | 0 | $11 / 10 / 2014$ | 78 | 8 | 0.12 | 0.35 | 398 |

APPENDIX D. Historical Morphometric Data

Spring Chinook (2004-2015)


2012 Nason Creek Rotary Trap Report

| 2015 | 2013 | Hatchery Chinook Yearling Smolt | 136 | 284 | 12.3 | 29.5 | 284 | 8.8 | 1.1 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

Summer Steelhead (2004-2015)

| Trap Year | Brood <br> Year | Age | Origin/Species | Fork Length (mm) |  |  | Weight (g) |  |  | K- <br> factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Mean | n | SD | Mean | n | SD |  |
| 2004 | 2004 | 0 | Wild Summer Steelhead | 67 | 358 | 10 | 3.5 | 279 | 1.5 | 1.2 |
| 2004 | 2003 | 1 | Wild Summer Steelhead | 101.7 | 394 | 23.2 | 13.2 | 366 | 27.3 | 1.3 |
| 2004 | 2002 | 2 | Wild Summer Steelhead | 161.6 | 146 | 19.8 | 43.4 | 141 | 15.5 | 1 |
| 2004 | 2001 | 3 | Wild Summer Steelhead | 201.6 | 43 | 11.2 | 76 | 43 | 21.2 | 0.9 |
| 2004 | 2003 | 1 | Hat. Summer Steelhead | 182.8 | 523 | 22.4 | 62.1 | 497 | 21.2 | 1 |
| 2005 | 2005 | 0 | Wild Summer Steelhead | 54.1 | 649 | 15.7 | 2.2 | 616 | 3.2 | 1.4 |
| 2005 | 2004 | 1 | Wild Summer Steelhead | 93.6 | 585 | 25.6 | 10.8 | 575 | 10.1 | 1.3 |
| 2005 | 2003 | 2 | Wild Summer Steelhead | 153.5 | 103 | 21.2 | 38.1 | 102 | 16.4 | 1.1 |
| 2005 | 2002 | 3 | Wild Summer Steelhead | 144 | 1 | - | 43.2 | 1 | - | 1.4 |
| 2005 | 2004 | 1 | Hat. Summer Steelhead | 188.2 | 343 | 21.2 | 66 | 343 | 24 | 1 |
| 2006 | 2006 | 0 | Wild Summer Steelhead | 66.3 | 180 | 5.8 | 2.5 | 180 | 1 | 0.9 |
| 2006 | 2005 | 1 | Wild Summer Steelhead | 85.2 | 877 | 18.7 | 6.7 | 877 | 6.6 | 1.1 |
| 2006 | 2004 | 2 | Wild Summer Steelhead | 155.9 | 106 | 26.8 | 36.1 | 105 | 13.5 | 1 |
| 2006 | 2003 | 3 | Wild Summer Steelhead | 197 | 2 | - | 73.5 | 2 | - | 1 |
| 2006 | 2005 | 1 | Hat. Summer Steelhead | - | - | - | - | - | - |  |
| 2007 | 2007 | 0 | Wild Summer Steelhead | 54.2 | 329 | 11.7 | 2 | 328 | 1.4 | 1.3 |
| 2007 | 2006 | 1 | Wild Summer Steelhead | 82.7 | 1,330 | 16.8 | 7.2 | 1,329 | 6.3 | 1.3 |
| 2007 | 2005 | 2 | Wild Summer Steelhead | 143.8 | 102 | 20.6 | 31.4 | 102 | 11.9 | 1.1 |
| 2007 | 2004 | 3 | Wild Summer Steelhead | 143 | 1 | - | 26.8 | 1 | - | 0.9 |
| 2007 | 2006 | 1 | Hat. Summer Steelhead | 149.3 | 3 | 47 | 33.1 | 3 | 29.1 | 1 |
| 2008 | 2008 | 0 | Wild Summer Steelhead | 52.9 | 930 | 11.1 | 1.7 | 930 | 1.2 | 1.1 |
| 2008 | 2007 | 1 | Wild Summer Steelhead | 84.5 | 1,876 | 17.1 | 7.4 | 1,874 | 6.6 | 1.2 |
| 2008 | 2006 | 2 | Wild Summer Steelhead | 149.9 | 122 | 22.9 | 36 | 122 | 15.5 | 1.1 |
| 2008 | 2005 | 3 | Wild Summer Steelhead | 180.3 | 13 | 18.9 | 57.4 | 13 | 16.4 | 1 |
| 2008 | 2007 | 1 | Hat. Summer Steelhead | 179.4 | 389 | 16.5 | 55.9 | 388 | 14.8 | 1 |
| 2009 | 2009 | 0 | Wild Summer Steelhead | 55.6 | 843 | 10.5 | 2.2 | 688 | 1.1 | 1.3 |
| 2009 | 2008 | 1 | Wild Summer Steelhead | 82.6 | 452 | 18.6 | 7.1 | 447 | 5.5 | 1.3 |
| 2009 | 2007 | 2 | Wild Summer Steelhead | 156.9 | 72 | 22 | 40.9 | 72 | 15.5 | 1.1 |
| 2009 | 2006 | 3 | Wild Summer Steelhead | 195 | 3 | 5 | 73 | 3 | 6.7 | 1 |
| 2009 | 2008 | 1 | Hat. Summer Steelhead | 183.1 | 280 | 16.7 | 60.8 | 280 | 18.2 | 1 |
| 2010 | 2010 | 0 | Wild Summer Steelhead | 55 | 1,287 | 11.1 | 2.5 | 917 | 1.3 | 1.5 |
| 2010 | 2009 | 1 | Wild Summer Steelhead | 89.8 | 1,079 | 19.1 | 9 | 1,072 | 7.1 | 1.2 |
| 2010 | 2008 | 2 | Wild Summer Steelhead | 144.9 | 87 | 25.1 | 35 | 87 | 17.4 | 1.2 |
| 2010 | 2007 | 3 | Wild Summer Steelhead | 184 | 8 | 12.2 | 61.9 | 8 | 10.2 | 1 |

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2012 Nason Creek Rotary Trap Report

| 2010 | 2009 | 1 | Hat. Summer Steelhead | 183.5 | 531 | 19.5 | 61.3 | 526 | 19.6 | 1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2011 | 2011 | 0 | Wild Summer Steelhead | 43.5 | 1,093 | 10.1 | 1.1 | 783 | 0.9 | 1.3 |
| 2011 | 2010 | 1 | Wild Summer Steelhead | 75.7 | 818 | 18.5 | 5.5 | 811 | 5.7 | 1.3 |
| 2011 | 2009 | 2 | Wild Summer Steelhead | 144.8 | 27 | 41.3 | 42.1 | 27 | 62.1 | 1.4 |
| 2011 | 2008 | 3 | Wild Summer Steelhead | - | - | - | - | - | - | - |
| 2011 | 2010 | 1 | Hat. Summer Steelhead | 180.7 | 464 | 17 | 59.1 | 464 | 17.6 | 1 |
| 2012 | 2012 | 0 | Wild Summer Steelhead | 55.1 | 589 | 14.2 | 2.6 | 402 | 1.2 | 1.6 |
| 2012 | 2011 | 1 | Wild Summer Steelhead | 84.7 | 747 | 17.4 | 7.6 | 741 | 5.7 | 1.3 |
| 2012 | 2010 | 2 | Wild Summer Steelhead | 127.1 | 132 | 27 | 23.7 | 132 | 14.5 | 1.2 |
| 2012 | 2009 | 3 | Wild Summer Steelhead | 161 | 4 | 32 | 40.5 | 4 | 15.6 | 1 |
| 2012 | 2011 | 1 | Hat. Summer Steelhead | 154.8 | 318 | 20.9 | 37.7 | 318 | 14 | 1 |
| 2013 | 2013 | 0 | Wild Summer Steelhead | 56.1 | 878 | 11.3 | 2.1 | 777 | 1.1 | 1.2 |
| 2013 | 2012 | 1 | Wild Summer Steelhead | 44.5 | 1,777 | 14.7 | 5.4 | 1,772 | 4.2 | 1.2 |
| 2013 | 2011 | 2 | Wild Summer Steelhead | 144.7 | 21 | 15.7 | 36.1 | 21 | 10.2 | 1 |
| 2013 | 2010 | 3 | Wild Summer Steelhead | - | - | - | - | - | - | - |
| 2013 | 2012 | 1 | Hat. Summer Steelhead | 166.2 | 365 | 21.4 | 49.2 | 363 | 18.2 | 1.1 |
| 2014 | 2014 | 0 | Wild Summer Steelhead | 49.6 | 490 | 12.8 | 1.7 | 389 | 1.1 | 1.4 |
| 2014 | 2013 | 1 | Wild Summer Steelhead | 82.2 | 745 | 13.6 | 6.3 | 745 | 3.5 | 1.1 |
| 2014 | 2012 | 2 | Wild Summer Steelhead | 145.1 | 30 | 16.5 | 33 | 30 | 13.4 | 1.1 |
| 2014 | 2011 | 3 | Wild Summer Steelhead | - | - | - | - | - | - | - |
| 2014 | 2013 | 1 | Hat. Summer Steelhead | 173.4 | 632 | 18.7 | 52.6 | 633 | 15.9 | 1.0 |
| 2015 | 2015 | 0 | Wild Summer Steelhead | 70 | 182 | 15.5 | 4.3 | 176 | 2.0 | 1.1 |
| 2015 | 2014 | 1 | Wild Summer Steelhead | 88 | 233 | 20.2 | 8.3 | 233 | 6.7 | 1.0 |
| 2015 | 2013 | 2 | Wild Summer Steelhead | 149 | 14 | 13.5 | 33.7 | 14 | 8.2 | 1.0 |
| 2015 | 2012 | 3 | Wild Summer Steelhead | 191 | 1 | - | 73.8 | 1 | - | 1.1 |
| 2015 | 2014 | 1 | Hat. Summer Steelhead | 175 | 273 | 15.2 | 51.3 | 273 | 12.5 | 0.9 |

Coho (2007-2015)

| Trap Year | Brood Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | Kfactor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Mean | n | SD | Mean | n | SD |  |
| 2004 | 2002 | Nat. Orig. Coho Yearling Smolt | - | - | - | - | - | - | - |
| 2004 | 2003 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - | - | - |
| 2004 | 2003 | Nat. Orig. Coho Subyearling Parr | - | - | - | - | - | - | - |
| 2004 | 2002 | Hatchery Coho Yearling Smolt | 136.6 | 847 | 12.8 | 27.4 | 820 | 7.5 | 1.1 |
| 2005 | 2003 | Nat. Orig. Coho Yearling Smolt | 114.4 | 17 | 8.8 | 16.2 | 17 | 3.6 | 1.1 |
| 2005 | 2004 | Nat. Orig. Coho Subyearling Fry | 49.1 | 9 | 10.4 | 1.3 | 9 | 0.8 | 1.1 |
| 2005 | 2004 | Nat. Orig. Coho Subyearling Parr | 76.7 | 9 | 12.8 | 4.9 | 9 | 2.7 | 1.1 |
| 2005 | 2003 | Hatchery Coho Yearling Smolt | 137.3 | 689 | 11.3 | 28.6 | 690 | 7.2 | 1.1 |
| 2006 | 2004 | Nat. Orig. Coho Yearling Smolt | - | - | - | - | - | - | - |
| 2006 | 2005 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - | - | - |
| 2006 | 2005 | Nat. Orig. Coho Subyearling Parr | 71 | 4 | 13.6 | 3.8 | 4 | 2.9 | 1.1 |
|  |  |  | 64 |  |  |  |  |  |  |

2012 Nason Creek Rotary Trap Report

| 2006 | 2004 | Hatchery Coho Yearling Smolt | - | - | - | - | - | - |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2007 | 2005 | Nat. Orig. Coho Yearling Smolt | 92.9 | 36 | 12.5 | 8.7 | 36 | 4 | 1.1 |
| 2007 | 2006 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - | - |  |
| 2007 | 2006 | Nat. Orig. Coho Subyearling Parr | 83 | 1 | - | 6.2 | 1 | - | 1.1 |
| 2007 | 2005 | Hatchery Coho Yearling Smolt | 116 | 2 | - | 16.8 | 2 | - | 1.1 |
| 2008 | 2006 | Nat. Orig. Coho Yearling Smolt | - | - | - | - | - | - | - |
| 2008 | 2007 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - | - | - |
| 2008 | 2007 | Nat. Orig. Coho Subyearling Parr | 87 | 1 | - | 6.4 | 1 | - | 1 |
| 2008 | 2006 | Hatchery Coho Yearling Smolt | 130.2 | 843 | 10.4 | 23.6 | 843 | 6.2 | 1.1 |
| 2009 | 2007 | Nat. Orig. Coho Yearling Smolt | 103 | 4 | 9.7 | 11.7 | 4 | 3.4 | 1.1 |
| 2009 | 2008 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - |  |  |
| 2009 | 2008 | Nat. Orig. Coho Subyearling Parr | 79.6 | 5 | 20.1 | 6.6 | 5 | 4.8 | 1.3 |
| 2009 | 2007 | Hatchery Coho Yearling Smolt | 135.3 | 625 | 8.9 | 26.2 | 579 | 5.2 | 1.1 |
| 2010 | 2008 | Nat. Orig. Coho Yearling Smolt | - | - | - | - | - | - | - |
| 2010 | 2009 | Nat. Orig. Coho Subyearling Fry | 48 | 2 | - | 1.3 | 2 | - | 1.2 |
| 2010 | 2009 | Nat. Orig. Coho Subyearling Parr | 83.6 | 27 | 8.6 | 6.7 | 27 | 2.4 | 1.1 |
| 2010 | 2008 | Hatchery Coho Yearling Smolt | 130 | 1,051 | 10.1 | 23.8 | 1,049 | 5.3 | 1.1 |
| 2011 | 2009 | Nat. Orig. Coho Yearling Smolt | 100.2 | 14 | 12.7 | 11.3 | 14 | 3.9 | 1.1 |
| 2011 | 2010 | Nat. Orig. Coho Subyearling Fry | - | - | - | - | - | - |  |
| 2011 | 2010 | Nat. Orig. Coho Subyearling Parr | 64.7 | 3 | 10.8 | 3 | 3 | 1.5 | 1.1 |
| 2011 | 2009 | Hatchery Coho Yearling Smolt | 124.6 | 969 | 8.6 | 21 | 969 | 4.8 | 1.1 |
| 2012 | 2010 | Nat. Orig. Coho Yearling Smolt | 102.1 | 17 | 9.1 | 11.9 | 17 | 3 | 1.1 |
| 2012 | 2011 | Nat. Orig. Coho Subyearling Fry | 36 | 1 | - | - | - | - |  |
| 2012 | 2011 | Nat. Orig. Coho Subyearling Parr | 78.4 | 84 | 9.3 | 5 | 84 | 2.1 | 1 |
| 2012 | 2010 | Hatchery Coho Yearling Smolt | 126.2 | 1,684 | 7.6 | 21.5 | 1,684 | 5.5 | 1.1 |
| 2013 | 2011 | Nat. Orig. Coho Yearling Smolt | 97 | 81 | 10 | 10 | 81 | 3.1 | 1.1 |
| 2013 | 2012 | Nat. Orig. Coho Subyearling Fry | 47.3 | 3 | 1 | 1 | 3 | 1 | 0.9 |
| 2013 | 2012 | Nat. Orig. Coho Subyearling Parr | 87.8 | 4 | 3.8 | 6.6 | 4 | 1 | 1 |
| 2013 | 2011 | Hatchery Coho Yearling Smolt | 130.1 | 982 | 8.5 | 23.3 | 977 | 4.9 | 1.1 |
| 2014 | 2012 | Nat. Orig. Coho Yearling Smolt | 96.3 | 20 | 9.8 | 9.9 | 20 | 3 | 1.1 |
| 2014 | 2013 | Nat. Orig. Coho Subyearling Fry | 36 | 1 | - | - | - | - |  |
| 2014 | 2013 | Nat. Orig. Coho Subyearling Parr | 73 | 3 | 22.5 | 5.9 | 3 | 4.7 | 1.5 |
| 2014 | 2012 | Hatchery Coho Yearling Smolt | 127 | 1,203 | 9.7 | 21.7 | 1,207 | 5 | 1.1 |
| 2015 | 2013 | Nat. Orig. Coho Yearling Smolt | 109 | 2 | 4.9 | 12.0 | 2 | 0.1402 | 0.9 |
| 2015 | 2014 | Nat. Orig. Coho Subyearling Fry | 47 | 7 | 13.7 | 1.4 | 7 | 1.4511 | 0.9 |
| 2015 | 2014 | Nat. Orig. Coho Subyearling Parr | 69 | 3 | 7.0 | 4.0 | 3 | 1.2583 | 1.2 |
| 2015 | 2013 | Hatchery Coho Yearling Smolt | 131 | 952 | 9.9 | 23.3 | 952 | 4.7946 | 1.0 |



Columbia River
Honor. Protect Restore.

## office

post us hey 97
Prinutin wa cosen
PHONE
(500) Seas-9619 Ex 100
fax
Emall
nebermbarmity-ran gov
WEB

To: Craig Busack
CC: Michelle Guay. Tom Scribner, Keely Murdoch, Bryan Ishida
From: Bryan Ishida
Date: March 15. 2015
RE: Nason Creek Smolt Trap Mortalities - 3/15/15

Dear Mr. Busack,
On March 15, 2015 YN FRM personnel arrived at the Nason rotary smolt trap at 9:00am to find it stopped by a $5^{\prime \prime} \times 6^{\prime \prime} \times 6^{\prime \prime}$ pressure-treated beam that had become wedged between the cone and the starboard pontoon. The halted rotation subsequently caused a small-diameter debris blockage at the rear of the cone preventing movement of fish and additional debris into the livebox. As a result, six wild spring Chinook subyearling fry and four wild spring Chinook yearling smolt mortalities were incurred. With a total of only 37 wild spring Chinook captured since trapping began on March I, our mortality rate for the species is currently at $27 \%$. The increase in debris load is attributed to a rapid spike in discharge level brought on by heavy rains. At the time of the stoppage, spring night operations (personnel on-site during hours of operation) had not yet commenced.

This event occurred during initial spring operations at the new Nason Creek smolt trap site (rkm 0.3). Due to its location on the outside of a channel bend, this new location appears to be more susceptible to debris stoppages than the previously-used site (rkm 0.9). In order to prevent further such instances, we will increase the duration of our night operations schedule to include highwater events prior to the scheduled May Istart as needed. Upon initial onset of elevated spring flows, we will begin night operations and continue until discharge levels have subsided. Discharge data from the upstream Department of Ecology gauge and snowpack data from nearby snow telemetry (SNOTEL) sites will be used to will be used to predict trends in flow and guide trap operations. Additionally, the Nason Creek smolt trap will also be manned during fall freshets to mitigate the increased stoppage potential at the new site. We will increase our vigilance in the monitoring of high-water events and take the necessary precautions to prevent any further loss of ESA-listed species. Please feel free to contact me with any questions regarding this event.

Sincerely.
Bryan Ishida

## Appendix L

Fish Trapping at the White River Smolt Trap during 2015

# Population Estimates for Juvenile Spring Chinook Salmon in White River, WA 

## 2015 Annual Final Report

Prepared by:<br>Bryan Ishida<br>Cory Kamphaus<br>Keely Murdoch<br>YAKAMA NATION<br>FISHERIES RESOURCE MANAGEMENT<br>Toppenish, WA 98948<br><br>Prepared for:<br>Public Utility District No. 2 of Grant County<br>Ephrata, Washington 98823


#### Abstract

In 2007, Yakama Nation Fisheries Resource Management began monitoring emigration of Endangered Species Act (ESA) listed Upper Columbia River (UCR) spring Chinook salmon in the White River to provide abundance and freshwater survival estimates. This report summarizes data collected between March 1 and November 30, 2015. We used a 1.5 m rotary screw trap to collect 196 juvenile spring Chinook; 2 precocial parr, 11 fry, 151 subyearling parr, and 32 yearling smolts. Daily counts at the trap were expanded via regression analysis derived from mark and recapture trials. We estimated that 3,023 ( $\pm$ 2,728; 95\% CI) BY2013 wild spring Chinook smolts and 1,449 ( $\pm 421$; $95 \%$ CI) BY2014 wild spring Chinook parr emigrated past the White River trap. Combined with data collected in 2014, this gives us a total estimate of $5,484( \pm 2,836$; $95 \%$ CI) BY2012 emigrants. Using spring Chinook spawning ground data collected by Washington Department of Fish and Wildlife (WDFW) in 2013, we estimated egg-toemigrant survival of BY2013 spring Chinook to be $2.2 \%$ ( 102 smolts-per-redd).


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Figure 3. Mean daily water temperatures at the White River DOE stream monitoring station at Sears Creek Bridge in 2015.

Figure 4. Daily catch of yearling spring Chinook smolt with mean daily stream discharge at the White River rotary trap, March 1 to June 30, 2015......................................... 13

Figure 5. Daily catch of wild subyearling spring Chinook with mean daily stream discharge at the White River rotary trap, July 1 to November 30, 201513

Figure 6. Relationships between estimated egg deposition and total emigrants produced, egg-to-emigrant survival, and emigrants per redd for White River spring Chinook, BY 2007 to 2013. *BY2013 values denoted by red border.

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### 1.0 INTRODUCTION

White River spring Chinook salmon (tkwínat) Oncorhynchus tshawytscha are part of the Upper Columbia River (UCR) spring Chinook salmon Evolutionarily Significant Unit (ESU) which was listed as endangered under the Endangered Species Act (ESA) in 1999. Due to critically low abundance, a captive broodstock program was operated in the White River between 1997 and 2015 as a risk aversion measure. Determining freshwater productivity of spring Chinook salmon in the White River is an essential component to overall population monitoring and will help contribute to the body of knowledge needed to evaluate if further supplementation in the White River is warranted. In 2007, Public Utility District No. 2 of Grant County (GCPUD) contracted the Yakama Nation (YN) to operate a rotary trap in the White River. Fish trap operations were conducted in compliance with ESA consultation specifically to address abundance and productivity of spring Chinook salmon in the White River.

Within this document, we will report:

1) Juvenile abundance and productivity of spring Chinook salmon in the White River.
2) Emigration timing of spring Chinook salmon emigrating from the White River.

Data presented will be directly used to address Objective 2 in the Monitoring and Evaluation Plan for PUD Hatchery Programs (Hillman et al. 2013) on a 5-year analytic cycle:

## Objective 2: Determine if the proportion of hatchery fish on the spawning grounds affects the freshwater productivity of supplemented stocks (Hillman et al. 2013).

In the fall of 2005, Washington State Department of Fish and Wildlife (WDFW) began smolt trapping in the lower reach of the White River in order to provide an estimate of juvenile spring Chinook salmon production. No trapping was conducted in 2006 as there was a transition between trap operators. In 2007, YN resumed trap operations on the White River for nine months of the year. This document reports data collected between March 1 and November 30, 2015 and provides emigration estimates for spring Chinook salmon yearlings (BY2013) and subyearlings (BY2014) during that time period. Data generated from this project was used to calculate annual egg-to-emigrant survival.

### 1.1 Watershed Description

The White River drainage encompasses 99,956 acres and originates in alpine glaciers and perennial snow fields (Figure 1; USFS 2004). Elevations in the drainage vary from 1,868 ft. at the Lake Wenatchee surface to $8,575 \mathrm{ft}$. at Clark Mountain (Andonaegui 2001). As one of two primary tributaries to Lake Wenatchee, the White River flows in a south-easterly direction for 42.9 river kilometers (RK) before emptying into the lake. Precipitation ranges from 79 cm at the
mouth to more than 356 cm in the head waters (Andonaegui 2001). Due to its glacial origins, peak runoff for the White River typically occurs between April and July with occasional high


Figure 1. Map of the Wenatchee River subbasin with White River rotary trap location.
flows caused by rain-on-snow events in the fall and winter months. Water temperatures in this watershed tend to be cooler than other tributaries to the upper Wenatchee River subbasin. As of September 2002, Washington State Department of Ecology (WDOE) began operating a stream monitoring station at RK 9.9 of White River. Operation of this station by WDOE is currently maintained with funding provided by GCPUD. In 2015, daily mean stream discharge ranged from 87 cfs to 4,280 cfs (Figure 2) while mean daily stream temperatures ranged from $0.2^{\circ} \mathrm{C}$ to
$15.9^{\circ} \mathrm{C}$ (Figure 3). Discharge and temperature data provided by WDOE should be considered provisional and are presented in Appendix A.

The White River drainage has had minimal riparian harvest from the 1950's to the present on federally owned land. Turn of the century settlement and land clearing have impacted the


Figure 2. Mean daily stream discharge at the White River DOE stream monitoring station at Sears Creek Bridge in 2015.


Figure 3. Mean daily water temperatures at the White River DOE stream monitoring station at Sears Creek Bridge in 2015.
riparian reserve network up to the Napeequa confluence, yet, riparian areas in the mainstem below Panther Creek remain in fair condition (USFS 2004). In the remainder of the watershed woody debris recruitment, shade, aquatic habitat connectivity, and riparian vegetation appear to be in good condition. Current habitat concerns pertaining to the development of homes and vacation retreats on private lands do exist. Rip-rapping, channel constriction, and stream degradation are considered minor in the watershed. Public ownership comprises $78 \%$ of the drainage area; more than half of public land is located within the Glacier Peak Wilderness. The remaining $22 \%$ of the drainage is in private ownership (USFS 2004).

Downstream of White River Falls are key spawning grounds for spring Chinook salmon (tkwínat) Oncorhynchus tshawytscha, sockeye salmon (kálux) O. nerka, and bull trout Salvilinus confluentus. Two large tributaries to the White River, Napeequa River and Panther Creek, are also known to support populations of anadromous salmonids (Mullen et al. 1992). For a complete list of known fish species encountered in the White River see (3.4 Incidental Species).

### 2.0 METHODS

### 2.1 Trapping Equipment and Operation

In 2015, a 1.5 m diameter cone rotary trap was operated in a single position at all discharge levels. This revised trapping regime was implemented in 2013 to simplify data analysis by eliminating obsolete trap positions that generated very little data. Past attempts at developing a high flow position generated very few efficiency trials resulting in limited trap efficiency data. Operating season-long at a single position, the trap was suspended from a river-spanning cable from which its position could be adjusted perpendicular to stream flow by hand powered winches anchored on a tree on the river-right bank.

The trap was operated 24 hours per day, seven days per week for the majority of the season. During spring snowmelt, operations only occurred during hours of darkness to minimize trap damage and subsequent capture mortality; still enabling sampling during the hours of peak fish movement. When trap operations were suspended, the cone was raised to avoid damage by debris.

During all ranges of river discharge, fish were removed daily. Additional trap checks were necessary during periods of high discharge in the spring and in the autumn due to increased leaf litter. Debris in the live-box was removed continually by a rotating drum screen, located at the rear of the holding box and hydraulically powered by the cone. A record of daily trap operations is provided in Appendix B.

### 2.2 Biological Sampling

Trap operating procedures and techniques followed a standardized, basin-wide monitoring plan developed by the Upper Columbia Regional Technical Team (UCRTT) for the Upper Columbia Salmon Recovery Board (UCSRB; Hillman 2004), which was adapted from Murdoch \& Petersen (2000).

Captured fish were transferred from the rotary trap's live box using five-gallon plastic buckets with lids to a stream-side, portable sampling station. Fish were anesthetized in a solution of MS222 to facilitate sampling and reduce handling stress. Fork length (FL) and weight were recorded for all fish, except large numbers of sockeye (Oncorhynchus nerka) fry. For these fish, a representative sample of 25 individuals was measured and weighed while the remaining fish were enumerated and released. Weight was measured to the nearest 0.1 g with a portable digital scale while FL was recorded to the nearest 1.0 mm using a trough-type measuring board. These data were used to calculate a Fulton-type condition factor (K-factor) for each target species using the formula:

$$
\mathrm{K}=\left(\mathrm{W} / \mathrm{L}^{3}\right) \times 100,000
$$

Where $\mathrm{K}=$ Fulton-type condition metric, $\mathrm{W}=$ weight in grams, $\mathrm{L}=$ fork length in millimeters and 100,000 is a scaling constant.

Portable aerators were used to oxygenate holding water during sampling. All fish were allowed to fully recover from anesthesia before being released. Spring Chinook salmon were classified as either natural or hatchery origin by the presence/absence of coded wire tags (CWT's). Developmental stages (fry, parr, transitional or smolt) were visually identified and assigned to each individual sampled. Transitional juveniles were identified as having both parr and smolt characteristics; visible parr marks, semi-transparent fin coloration along with silvery coloration throughout body. Smolts were identified by a strong silvery coloration over entire body and faint or absent parr marks. Fry were defined as newly emerged fish with or without a visible yolk sac and a FL measuring < 50 mm . Age- 0 spring Chinook salmon captured before July 1 were considered 'fry' and excluded from population estimates due to the inconclusive nature of their movement (i.e. active emigration or local distribution in-stream). Age-0 spring Chinook salmon captured after 1 July were considered subyearling emigrants and included in the population estimate (UCRTT, 2001).

Tissue samples were taken from spring Chinook salmon and steelhead (small, upper lobe caudal fin clip) and applied to blotter sheets. Samples from both species were provided to WDFW for reproductive success analysis. Scale samples were also collected from all steelhead captured. Scale samples were submitted to WDFW for age analysis. Bull trout tissue samples were not collected in 2015.

During periods when the trap operations were suspended (e.g. - high discharge, high debris and/or mechanical problems), passage estimates were generated to account for emigrants during these time periods. This estimate was calculated using the average number of fish captured three days prior and three days after the break in operation (Hillman et al., 2013; Snow et al., 2013).

### 2.3 Mark-Recapture Trials

Groups of marked spring Chinook salmon were used for trap efficiency trials. Fish were marked by insertion of a passive integrated transponder (PIT) tag into the abdominal cavity. Ideally, marked groups of fish would be released over a broad range of stream discharges in order to determine a trap efficiency-discharge relationship. (See 2.4 Data Analysis). However, due to low abundance and limited holding time of ESA listed species (reducing the ability to meet trials size requirements on a more consistent basis), marked groups were released whenever the minimum sample size ( $\geq 20$ ) was obtained. Mark-recapture (M-R) trials followed the protocol described in Hillman (2004). Although the protocol suggests a minimum sample size of 100 fish for each mark-group, the limited abundance of juvenile emigrants from the White River required that efficiency trials be completed with much smaller sample sizes. YN's continued goal is to increase individual mark-group sizes, when possible, to meet the standard described above.

Number of wild fish included in a marked group was maximized by combining catches from three days of trapping. Fish were held up to 72 hours prior to release in holding boxes located on the river-left bank. Fish to be used in efficiency trials were then transported in five gallon buckets $\sim 1.0$ RK upstream to the release location at Sears Creek Bridge (RK 10.3). All mark groups are released by hand at nautical twilight.

Each M-R trial was conducted over a three-day ( 72 hour) period to allow time for passage or capture. Completed trials were only considered invalid if an interruption to trapping occurred or proper pre-release procedures were not followed. Trials resulting in zero recaptures were included in the efficiency regression as allowed by the new method of observed trap efficiency calculation (See equation 3 in 2.5.1 Estimate of Abundance).

### 2.3.1 Marking and PIT tagging

All spring Chinook and summer steelhead juveniles with FL of $\geq 60 \mathrm{~mm}$ were PIT tagged unless the health of an individual was in question (e.g.- fungus). Once anesthetized, each fish was examined for external wounds or descaling and scanned for the presence of a previously implanted PIT tag. If a tag was not detected, a pre-loaded 12mm Digital Angel 134.2 kHz type TX 1411ST PIT tag was inserted into the body cavity using a Biomark MK-25 Rapid Implant Gun. Each unique tag code was electronically recorded with an appropriate tagging date, release date, tagging personnel and biological data. These data were entered into $\mathrm{P}_{3}$ and submitted to the PIT Tag Information System (PTAGIS) at the end of each month. Tagging methods were consistent with methodology described in the PIT Tag Marking Procedures Manual (CBFWA 1999) as well as with 2008 ISEMP protocols (Tussing 2008).

After marking and/or PIT tagging, fish were held for a minimum of 24-hours to a) ensure complete recovery, b) assess tagging mortality and c) determine tag-shed rate. Fish that were not to be used in an efficiency trial were released downstream of the smolt trap.

### 2.4 Data Analysis

### 2.4.1 Estimate of Abundance

Seasonal juvenile migration, N , was estimated as the sum of daily migrations, $N_{i}$, i.e., $N=\sum_{i} N_{i}$, and daily migration was calculated from catch and efficiency:

$$
\begin{equation*}
\hat{N}_{i}=\frac{C_{i}}{\hat{e}_{i}} \tag{1}
\end{equation*}
$$

where $C_{i}=$ number of fish caught in period $I$;

$$
\hat{e}_{i}=\text { trap efficiency estimated from the flow-efficiency relationship, } \sin ^{2}\left(b_{0}+b_{1} \text { flow } w_{i}\right)
$$

where $b_{0}$ is estimated intercept and $b_{1}$ is the estimated slope of the regression.

The regression parameters $b_{0}$ and $b_{l}$ are estimated using linear regression for the model:

$$
\begin{equation*}
\arcsin \left(\sqrt{e_{k}^{\text {obs }}}\right)=\beta_{0}+\beta_{1} \text { flow }_{k}+\varepsilon \tag{2}
\end{equation*}
$$

where $e_{k}^{\text {obs }}=$ observed trap efficiency of Eq. 2 for trapping period $k$;

$$
\begin{aligned}
& \beta_{0}=\text { intercept of the regression model; } \\
& \beta_{1}=\text { slope parameter; } \\
& \varepsilon=\text { error with mean } 0 \text { and variance } \sigma^{2} .
\end{aligned}
$$

In Equation 2, the observed trap efficiency, $e_{k}^{o b s}$, is calculated as follows,

$$
\begin{equation*}
e_{k}^{o b s}=\frac{r_{k}+1}{m} \tag{3}
\end{equation*}
$$

The estimated variance of seasonal migration is calculated from daily estimates as:

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)=\underbrace{\sum_{i} \operatorname{Var}\left(N_{i}\right)}_{\text {Part } \mathrm{A}}+\underbrace{\sum_{i} \sum_{j} \operatorname{Cov}\left(N_{i}, N_{j}\right)}_{\text {Part B }},
$$

or,

$$
\begin{equation*}
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)=\underbrace{\sum_{i} \operatorname{Var}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}\right)}_{\text {Part } A}+\underbrace{\sum_{i} \sum_{j} \operatorname{Cov}\left(\frac{\left(C_{i}+1\right)}{\hat{e}_{i}}, \frac{\left(C_{j}+1\right)}{\hat{e}_{j}}\right)}_{\text {Part B }} . \tag{4}
\end{equation*}
$$

Part A of equation 4 is the variance of daily estimates. Part B is the between-day covariance. Note that the between-day covariance exists only for days that use the same trap efficiency model. If, for example, day 1 is estimated with one trap efficiency model, and day 2 estimated from a different model, then there is no covariance between day 1 and day 2 . The full expression for the estimated variance:

$$
\begin{aligned}
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}\right)= & \underbrace{\sum_{i} \hat{N}_{i}^{2}\left(\frac{N_{i} \hat{e}_{i}\left(1-\hat{e}_{i}\right)}{\left(C_{i}+1\right)^{2}}+\frac{4\left(1-\hat{e}_{i}\right)}{\hat{e}_{i}} \operatorname{Var}\left(b_{0}+b_{1} \text { flow }_{i}\right)\right)}_{\text {PartA }}+ \\
& \underbrace{\sum_{i} \sum_{j} 4\left(\hat{N}_{i}\left(1-\hat{e}_{i}\right)\right)\left(\hat{N}_{j}\left(1-\hat{e}_{j}\right)\right) \cdot\left[\hat{\operatorname{Var}}\left(b_{0}\right)+\text { flow }_{i} \text { flow }_{j} \hat{\operatorname{Var}}\left(b_{1}\right)\right]}_{\text {PartB }}
\end{aligned}
$$

where $\operatorname{Vâ} r\left(b_{0}+b_{1} f l o w_{i}\right)=\operatorname{MS} E\left(1+\frac{1}{n}+\frac{\left(\text { flow }_{i}-\overline{\text { flow }}\right)^{2}}{(n-1) s_{\text {flow }}^{2}}\right)$, and $\hat{\operatorname{Var}}\left(b_{0}\right)$ and $\hat{\operatorname{Var}}\left(b_{1}\right)$ are
obtained from regression results. In Excel, the standard error (SE) of the coefficients is provided. The variance is calculated as the square of the standard error, $S E^{2}$.

In cases when there was no significant flow-efficiency relationship (i.e., low correlation), then a pooled, or average trap efficiency will suffice for the stratum. The estimator is calculated as follows:

$$
\hat{\bar{e}}=\frac{\sum_{j=1}^{k} r_{j}}{\sum_{j=1}^{k} m_{j}}
$$

where $\hat{\bar{e}}=$ the average or pooled trap efficiency for the stratum;
$m_{j}=$ the number of smolts marked and released in efficiency trial $j$ for the stratum;
$r_{j}=$ the number of smolts recaptured out of $m_{j}$ marked fish in efficiency trial $j$.

Abundance for a trapping period is estimated as:

$$
\hat{N}_{i}^{\text {pooled }}=\frac{C_{i}}{\hat{\bar{e}}},
$$

,and total stratum abundance is:

$$
N^{\text {pooled }}=\sum_{i} \hat{N}_{i}^{\text {pooled }}
$$

The variance of seasonal abundance takes into account the variability in catch numbers that are a result of binomial sampling (Part A), the pooled variance of trap efficiency, $\hat{\bar{e}}$ (Part B), and the
covariance in daily estimates that arises from using a common estimate of efficiency across all trapping days (Part C):

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}^{\text {pooled }}\right)=\underbrace{\left(\sum_{i} \frac{\hat{N}_{i}(1-\hat{\bar{e}})}{\hat{\bar{e}}}\right)}_{\text {PartA }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}} \sum_{i} \hat{N}_{i}^{2}}_{\text {PartB }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}} \sum_{i} \sum_{j} \hat{N}_{i} \hat{N}_{j}}_{\text {Part }} .
$$

The Part B and Part C terms are combined in the calculation as a new Part B:

$$
\operatorname{Var}\left(\sum_{i=1}^{n} \hat{N}_{i}^{\text {pooled }}\right)=\underbrace{\left(\sum_{i} \frac{\hat{N}_{i}(1-\hat{\bar{e}})}{\hat{\bar{e}}}\right)}_{\text {Part }}+\underbrace{\frac{\operatorname{Var}(\hat{\bar{e}})}{\hat{\bar{e}}^{2}}\left[\sum_{i} \hat{N}_{i}^{2}+\sum_{i} \sum_{j} \hat{N}_{i} \hat{N}_{j}\right]}_{\text {PartB }} .
$$

The variance of $\hat{\bar{e}}$ is calculated as:

$$
\operatorname{Vâr}(\hat{\bar{e}})=\operatorname{Var}\left(\frac{\sum_{k=1}^{n} r_{k}}{\sum_{k=1}^{n} m_{k}}\right)=\frac{\sum_{k=1}^{n}\left(r_{k}-\hat{\bar{e}}_{k} m_{k}\right)^{2}}{\bar{m}^{2} n(n-1)}
$$

where $\bar{m}$ is the average release size across all efficiency trial, $\frac{\sum_{k=1}^{n} m_{k}}{n}$.
Confidence intervals were calculated using the following formulas:

$$
95 \% \text { confidence interval }=1.96 \times \sqrt{\sum \operatorname{var}}\left[\hat{N}_{i}\right]
$$

The single M-R estimator of abundance carries a set of well documented assumptions (Everhart and Youngs 1981; Seber 1982),

1. The population is closed to mortality.
2. The probability of capturing a marked or unmarked fish is equal.
3. Marked fish were randomly dispersed in the population prior to recapture.
4. Marking does not affect probabilities of capture.
5. Marks were not lost between the time of release and recapture.
6. All marks are reported upon recapture.
7. The number of fish in the trap, C , is fully enumerated and known without error.

### 3.0 RESULTS

### 3.1 Dates of Operation

In 2015, we operated a 1.5 m rotary trap between March 1 and November 30. During this period, the trap operated 24 hours per day, 7 days per week barring inoperable environmental conditions (i.e. heavy debris loads or high discharge), mechanical malfunctions, or periods of suspended trapping due to issues relating to lapsed liability insurance. Trapping was interrupted or intentionally suspended for a total of 49 days (Table 1).

Table 1. Summary of White River smolt trap operation, 2015.

| Trap Status | Description | Days |
| :--- | :--- | :---: |
| Operating | Continuous data collection | 226 |
| Interrupted | Interrupted by debris, ice, tampering, or improper positioning | 7 |
| Pulled | Intentionally pulled due to flooding risk or administrative reasons | 42 |

### 3.2 Daily Captures and Biological Sampling

### 3.2.1 Wild Spring Chinook Yearlings (BY2013)

A total of 32 wild yearling Chinook smolts were collected between March 1 and June 30, with peak catch occurring on April 9 ( $n=4$; Figure 4). Mean fork-length (FL) was 103 mm ( $n=32$; $S D=6.9)$ and mean weight was $13.0 \mathrm{~g}(n=31 ; S D=2.8)$; see Table 2. PIT tags were implanted into 32 smolts. Genetic samples were also taken from the same 32 fish. An additional two suspected BY2013 Chinook were captured after July 1. Mean FL for these fish was 145 mm ( $n=$ $2 ; S D=13.4)$ and mean weight was $35.15 \mathrm{~g}(n=2 ; S D=11.4)$; see Table 2 . These fish were identified as precocial parr by their large size, timing of capture, and release of milt during handling. All precocial parr were excluded from emigration estimates. There were no BY2013 spring Chinook mortalities incurred (See 3.4 ESA Compliance).

### 3.2.2 Wild Spring Chinook Subyearlings (BY2014)

Spring Chinook fry were captured at the trap between March 15 and June 13 ( $n=11$ ). During this period there were no fry trapping mortalities incurred. A total of 151 wild subyearling Chinook parr were collected between July 13 and November 30, with peak catch occurring on September 5 ( $n=15$; Figure 5). The mean FL for subyearling parr was $96 \mathrm{~mm}(n=151 ; S D=$ $7.4)$ and the mean weight was $9.9 \mathrm{~g}(n=148 ; S D=2.3)$; see Table 2. PIT tags were implanted into a total of 149 subyearling Chinook parr. Genetic samples were taken from 150 parr. One Chinook parr was not tagged due to a visible external injury. Additionally, one tag was shed during the 24 hr holding period (Table 4). There were no BY2014 spring Chinook mortalities during the 2015 trapping season (See 3.4 ESA Compliance).


Figure 4. Daily catch of yearling spring Chinook smolt with mean daily stream discharge at the White River rotary trap, March 1 to June 30, 2015.


Figure 5. Daily catch of wild subyearling spring Chinook with mean daily stream discharge at the White River rotary trap, July 1 to November 30, 2015.

### 3.2.3 Hatchery Spring Chinook Yearlings (BY2013)

Hatchery-origin yearling Chinook released downstream of the smolt trap are sometimes caught in the summer months as precocial parr. Direct releases of BY2013 spring Chinook were not performed in the White River or in close proximity to is confluence with Lake Wenatchee (netpen-rearing). There were no hatchery-origin spring Chinook captured at the smolt trap in 2015. Hatchery fish captured at the trap are identified by the presence of CWT tags.

Table 2. Summary of length and weight sampling of juvenile spring Chinook captured at the White River rotary trap in 2015.

| Brood Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | $\begin{aligned} & \text { K- } \\ & \text { factor } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | n | SD | Mean | n | SD |  |
| 2013 | Wild Yearling Smolt | 103 | 32 | 6.9 | 13.0 | 31 | 2.8 | 1.14 |
| 2013 | Wild Yearling Precocial Parr | 145 | 2 | 13.4 | 35.2 | 2 | 11.4 | 1.14 |
| 2014 | Wild Subyearling Fry | 38 | 11 | 3.3 | 0.5 | 10 | 0.2 | 0.86 |
| 2014 | Wild Subyearling Parr | 96 | 151 | 7.4 | 9.9 | 148 | 2.3 | 1.11 |

### 3.3 Trap Efficiency Calibration and Population Estimates

### 3.3.1 Wild Spring Chinook Yearlings (BY 2013)

Due to low abundance, no BY2013 natural yearling Chinook efficiency trials were performed in 2015. A composite regression model using previous year's (2008-2012) efficiency trials showed statistical significance ( $r^{2}=0.57 ; p=0.001$ ) for a flow-efficiency relationship and was used to calculate yearling abundance. Use of a single spring trapping position allowed this regression to be applied to all yearling Chinook captured in 2015. Weighting of this regression via an R script (provided by WDFW) did not affect calculation parameters greatly and yielded the same r-square and $p$-values. In the fall of 2014, we estimated that 2,461 ( $\pm 779 ; 95 \% \mathrm{CI})$ BY2013 subyearlings emigrated past the trap. In the spring of 2015, we estimated that $3,023( \pm 2,728 ; 95 \% \mathrm{CI})$ emigrated past the trap. Combining the two estimates, total BY2013 wild spring Chinook emigrants was 5,484 ( $\pm 2,836$; $95 \%$ CI; Table 3 ).

### 3.3.2 Wild Spring Chinook Subyearling (BY 2014)

Low parr abundance presented only one opportunity to perform a mark-group release in 2015. Despite being smaller $(n=39)$ than the previously-set minimum mark group size of 50 parr, the efficiency trial was performed due to the low cfs being tested ( 89.5 cfs ). Our current strategy to improve the flow-efficiency model includes targeting mark-group releases at discharge levels where data is currently lacking. The updated multi-year composite regression was applied to BY2014 subyearling emigrants. The regression was comprised of all trails conducted fulfilling the minimum number marked ( $n \geq 20$ ) including efforts in which zero recaptured were made (Appendix C). Mark-groups in which validity of the trial could be called into question (suspected trap stoppage or improper pre-release handling of the mark group) were removed. The weighted regression was not significant $\left(r^{2}=0.12 ; p=0.086\right)$ at our accepted limit $(\alpha=$ $0.05)$. However, after comparison with a pooled method and considerations of the pooled
estimate limitations, we decided to use the regression model despite its slightly higher $p$-value. This single regression was the only model required to estimate total subyearling migration due to the fact only one fall trapping position was used in 2015. We estimated that in 2015, 1,449 $\pm$ $421 ; 95 \%$ CI) spring Chinook subyearling parr moved past the trap (Table 3).

Table 3. Estimated egg-to-emigrant survival and emigrants per redd for White River spring Chinook

| Brood <br> Year | No. of Redds ${ }^{\text {a }}$ | Fecundity ${ }^{\text {b }}$ | No. of Eggs | No. of Emigrants |  |  | Egg-to <br> Emigrant | Emigrants per Redd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Age-0 ${ }^{\text {c }}$ | Age-1 | $\begin{gathered} \text { Total } \pm 95 \% \\ \text { CI } \end{gathered}$ |  |  |
| 2005 | 86 | 4,327 | 372,122 | DNOT ${ }^{\text {d }}$ | 4,856 | - |  | - |
| 2006 | 31 | 4,324 | 134,044 | 642 | 2,004 | $2,646 \pm 1,597$ | 2.0\% | 85 |
| 2007 | 20 | 4,441 | 88,820 | 2,293 | 3,399 | $\begin{gathered} 5,692 \pm \\ 2,214 \end{gathered}$ | 6.4\% | 285 |
| 2008 | 31 | 4,592 | 142,352 | 5,552 | 5,193 | $\begin{gathered} 10,745 \pm \\ 3,837 \end{gathered}$ | 7.5\% | 347 |
| 2009 | 54 | 4,573 | 246,942 | 2,485 | 2,939 | 5,424 $\pm 2,522$ | 2.2\% | 100 |
| 2010 | 33 | 4,314 | 142,362 | 1,859 | 4,121 | $\begin{gathered} 5,980 \pm \\ 3,455 \end{gathered}$ | 4.2\% | 181 |
| 2011 | 20 | 4,385 | 87,700 | 3,128 | 1,659 | $4,787 \pm 2,022$ | 5.5\% | 239 |
| 2012 | 86 | 4,223 | 363,178 | 3,905 | 3,995 | $7,900 \pm 3,898$ | 2.2\% | 92 |
| 2013 | 54 | 4,716 | 254,664 | 2,461 | 3,023 | $5,484 \pm 2,836$ | 2.2\% | 102 |
| 2014 | 26 | 4,045 | 105,170 | 1,449 | - | - | - | - |
| Avg | 41 | 4,446 | 182,508 | 2,791 | 3,292 | 6,082 | 4.0\% | 179 |

[^19]


Figure 6. Relationships between estimated egg deposition and total emigrants produced, egg-to-emigrant survival, and emigrants per redd for White River spring Chinook, BY 2007 to 2013. *BY2013 values denoted by red border.

### 3.4 PIT Tagging

In 2015, a total of 185 spring Chinook and 6 steelhead were PIT tagged at the trap. PIT tag retention after 24 hours of observation was $100 \%$ for all species/stages, with the exception of wild spring Chinook parr (Table 4). There no tagging mortalities (Table 6).

Table 4. Number of PIT tagged spring Chinook and steelhead with shed rates at the White River rotary trap in 2015.

| Brood <br> Year | Species/Stage | Total <br> Catch | Total PIT <br> Tagged | Percent <br> Tagged | Percent <br> Tags Shed |
| :---: | :--- | :---: | :---: | :---: | :---: |
| 2013 | Yearling Chinook Smolt | 32 | 32 | $100.00 \%$ | $0.0 \%$ |
| 2013 | Yearling Chinook Precocial Parr | 2 | 2 | $100.00 \%$ | $0.0 \%$ |
| 2014 | Subyearling Chinook Parr | 151 | 149 | $98.68 \%$ | $0.7 \%$ |
| $*$ | Steelhead Parr | 6 | 6 | $100.00 \%$ | $0.0 \%$ |

* Brood year unknown


### 3.5 Incidental Species

Incidental species were enumerated and sampled for length and weight (Table 5). Incidental species included: bull trout Salvelinus confluentus, eastern brook trout Salvelinus fontilalis, longnose dace Rhinichthys cataractae, mountain whitefish Prosopium williamsoni, northern pikeminnow Ptychocheilus oregonensis, steelhead/rainbow trout (shúshaynsh) Oncorhynchus
mykiss, redside shiner Richardsonius balteatus, sculpin Cottus sp., sockeye salmon Oncorhynchus nerka, sucker Catostomus sp., and westslope cutthroat Oncorhynchus clarkii lewisi.

Table 5. Summary of length and weight sampling of incidental species captured at the White River rotary trap in 2015.

| Species | Total Count | Fork Length (mm) |  |  | Weight (g) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | $n$ | SD | Mean | $n$ | SD |
| Bull Trout Fry | 1 | 28 | 1 | - | - | - | - |
| Bull Trout Parr | 8 | 147 | 8 | 56.3 | 43.0 | 8 | 34.7 |
| Eastern Brook Trout | 1 | 245 | 1 | - | 145 | 1 | - |
| Longnose Dace | 12 | 59 | 12 | 22.8 | 4.0 | 9 | 1.9 |
| Mountain Whitefish | 93 | 87 | 93 | 36.8 | 9.5 | 88 | 20.4 |
| Northern Pikeminnow | 37 | 128 | 37 | 47.7 | 35.8 | 35 | 53.0 |
| Rainbow Trout/Steelhead Parr | 6 | 158 | 5 | 54.5 | 52.3 | 5 | 38.6 |
| Redside Shiner | 147 | 73 | 147 | 16.2 | 5.6 | 142 | 3.0 |
| Sculpin | 172 | 45 | 170 | 22.3 | 3.1 | 97 | 4.4 |
| Sockeye - Kokanee | 5 | 203 | 5 | 9.1 | 90.1 | 5 | 10.5 |
| Sockeye Fry | 7,212 | 28 | 1,200 | 1.2 | - | - | - |
| Sockeye Parr | 5 | 73 | 5 | 10.5 | 3.7 | 5 | 1.5 |
| Sucker | 37 | 140 | 37 | 104.6 | 90.0 | 28 | 107.5 |
| Westslope Cutthroat | 30 | 221 | 30 | 34.9 | 103.3 | 28 | 50.7 |

### 3.6 ESA Compliance

There were no ESA species mortalities incurred in 2015 (Table 6). All fish handled were inspected prior to tagging or further sampling, with only one wild spring Chinook parr warranting immediate release (injury).

Table 6. Summary of White River ESA listed species catch and mortality in 2015.

| Species/Stage | Total Catch | Total Mortality | Total \% <br> Mortality |
| :--- | :---: | :---: | :---: |
| Yearling Chinook Smolt | 32 | 0 | $0.00 \%$ |
| Yearling Chinook Precocial Parr | 2 | 0 | $0.00 \%$ |
| Subyearling Chinook Parr | 151 | 0 | $0.00 \%$ |
| Subyearling Chinook Fry | 11 | 0 | $0.00 \%$ |
|  | $\mathbf{1 9 6}$ | $\mathbf{0}$ | $\mathbf{0 . 0 0 \%}$ |
| Bull Trout | 9 | 0 | $0.00 \%$ |
| Steelhead/Rainbow Trout Wild Spring Chinook | 6 | 0 | $0.00 \%$ |

### 4.0 DISCUSSION

Estimations of White River yearling (BY2013) and subyearling (BY2014) wild spring Chinook emigrants in 2015 were calculated using multi-year compounded regressions. Given the overall low abundance of White River spring Chinook, the ability to use mark-group releases over multiple years provides the most effective means of expansion, when pooled and year-specific regression models are impracticable.

Using the multi-year yearling Chinook model, we estimated that 3,023 BY2013 spring Chinook emigrated past the trap in the spring of 2015. This estimation of smolt migrants falls below the 8 -year yearling average ( $n=3,292$ ), despite above average redd counts. Combined with the previous estimate of 2,482 subyearling emigrants, the total emigrant expansion of 5,505 BY2013 was also below the 8 -year average ( $n=6,085$ ). Although above-average egg deposition produced a below average estimate of emigrant abundance, the observed BY2013 egg-toemigrant survival rate ( $2.2 \%$ ) was consistent with the inverse relationship between total egg deposition and egg-to-emigrant survival previously observed in the White River. This suggest that the total estimated emigrants for BY2013 although potentially low, is not necessarily atypical of the system.

Base flows extending into mid-September and a brief increase in subyearling catch provided an important opportunity to expand the breadth of our subyearling regression. Efficiency trials at the high and low ends of the hydrograph are generally unfeasible due to inadequate mark-group sizes; active emigration is low at base flows and trap efficiency is extremely low at very high flows. A single, yet significant mark-group release effectively set the lower bound of the subyearling regression at 89.5 cfs , a flow representative of near-base discharge on the White River. Using the improved regression model, we estimated that 1,449 BY2014 parr emigrated past the trap in 2015.

Compared to other upper-Wenatchee River tributaries (Chiwawa River and Nason Creek), the White River was the only tributary that did not have an increasing trend in egg-to-emigrant survival for the 2013 brood (Figure 7). Egg-to-emigrant survival in Nason Creek is generally lower than that of the White River, yet the inverse is seen in BY2013. This deviation from previous trends is attributed in part to a suspected overestimate of Nason Creek emigrant abundance; the product of new regression models skewed by limited trials (Ishida et al. 2016). We speculate that the eventual recalculation of the estimate may decrease Nason Creek's BY2013 estimated survival markedly.


Figure 7. Comparison of wild spring Chinook abundance estimates (BY2007-2013) made at the White R., Nason Cr., and Chiwawa R. smolt traps. Chiwawa R. data provided by Hillman et al. (2015).

In 2016 we will continue to use the methodologies described in this report. Our priority will again be the strengthening of both our subyearling and yearling models through efficiency trials first vetted for adequate size and potential redundancies prior to release. While limited by the low abundance of White River spring Chinook, we remain confident that improvement to our models, albeit potentially slow, will persist through our refined methodologies.

### 5.0 LITERATURE CITED

Andonaegui, C. 2001. Salmon, Steelhead, and Bull Trout Habitat Limiting Factors for the Wenatchee Subbasin (Water Resource Inventory Area 45) and Portions of WRIA 40 within Chelan County (Squilchuck, Stemilt and Colockum drainages). Final draft report. WSCC.

CBFWA (Columbia Basin Fish and Wildlife Authority). 1999. PIT Tag Marking Procedures Manual, Version 2.0. Columbia Basin Fish and Wildlife Authority, Portland OR.

Everhart, W.H. and W.D. Youngs. 1953. Principles of Fishery Science, second edition. Comstock Publishing Associates, a division of Cornell University Press, Ithica and London.

Hillman, T.W. 2004. Monitoring strategy for the Upper Columbia Basin: Draft report February 1, 2004. Prepared for Upper Columbia Regional Technical Team, Wenatchee, Washington.

Hillman, T.W., P. Graf, B. Ishida, M. Johnson, C. Kamphaus, M. Miller, C. Moran, A. Murdoch, T. Pearsons, M. Tonseth, and C. Willard. 2015. Monitoring and Evaluation of the Chelan and Grant County PUD's Hatchery Programs: 2014 Annual Report. Prepared for The Habitat Conservation Plan Hatchery Committee and the Priest Rapids Coordinating Committee Hatchery Sub Committee. Wenatchee and Ephrata, WA.

Ishida, B., C. Kamphaus, and K. Murdoch. 2016. Population Estimates for Juvenile Salmonids in Nason Creek, WA: 2015 Annual Report. Prepared for Public Utility District No. 2 of Grant County and U.S. Department of Energy Bonneville Power Administration. Ephrata, WA and Portland, OR.

Marshall, A. R., and S. Young. 1994. Genetic Analysis of Upper Columbia Spring and Summer Chinook Salmon for the Rock Island Hatchery Evaluation Program. Final report, Washington Department of Fisheries, Olympia.

Mullan, J.W., K.R. Williams, G. Rhodus, T.W. Hillman and J.D. McIntyre. 1992. Production and Habitat of Salmonids in Mid Columbia River Tributaries. Monograph 1, USFWS, Leavenworth, Washington.

Murdoch, A., and K. Petersen. 2000. Freshwater Production and Emigration of Juvenile Spring Chinook from the Chiwawa River in 2000. Washington State Department of Fish and Wildlife.

Pearsons, T.N., and R.B. Langshaw. 2009. Monitoring and Evaluation Plan for Grant County PUD's Salmon and Steelhead Supplementation Programs. Grant County Public Utility District, Ephrata, WA.

Snow, C., C. Frady, and A. Repp. 2013. Monitoring and Evaluation of Wells and Methow Hatchery Programs: 2012 Annual Report. Prepared for Douglas County Public Utility District and Wells Habitat Conservation Plan Hatchery Committee.

Tussing, S.P. 2008. A Field Manual of Scientific Protocols for Downstream Migrant Trapping within the Upper Columbia Monitoring Strategy: 2008 Working Version 1.0. Prepared for Bonneville Power Administration's Integrated Status and Effectiveness Monitoring Program.

Upper Columbia RTT. 2001. A Strategy to Protect and Restore Salmonid Habitat in the Upper Columbia Region, a Discussion Draft Report. Upper Columbia Salmon Recovery Board.

United States Forest Service. 2004. White River Survey Report.
Washington Department of Ecology (WDOE). 2012. River and Stream Flow Monitoring. https://fortress.wa.gov/ecy/wrx/wrx/flows/station.asp?sta=45J0

APPENDIX A: White River Temperature and Discharge Data

| Date | Stream Discharge (CFS) | Water Temperature ( ${ }^{\circ} \mathrm{C}$ ) | 4/7/2015 | 594 | 4.2 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 4/8/2015 | 560 | 4.6 |
|  |  |  | 4/9/2015 | 536 | 5.4 |
| 3/1/2015 | 488 | 2.8 | 4/10/2015 | 516 | 5.6 |
| 3/2/2015 | 464 | 3.3 | 4/11/2015 | 536 | 5.0 |
| 3/3/2015 | 435 | 2.6 | 4/12/2015 | 503 | 3.8 |
| 3/4/2015 | 418 | 2.3 | 4/13/2015 | 475 | 4.7 |
| 3/5/2015 | 405 | 2.9 | 4/14/2015 | 453 | 5.4 |
| 3/6/2015 | 395 | 3.7 | 4/15/2015 | 428 | 5.2 |
| 3/7/2015 | 395 | 4.0 | 4/16/2015 | 426 | 5.9 |
| 3/8/2015 | 406 | 4.0 | 4/17/2015 | 488 | 6.7 |
| 3/9/2015 | 423 | 4.1 | 4/18/2015 | 561 | 6.9 |
| 3/10/2015 | 443 | 4.1 | 4/19/2015 | 600 | 7.1 |
| 3/11/2015 | 468 | 4.7 | 4/20/2015 | 714 | 7.4 |
| 3/12/2015 | 561 | 5.1 | 4/21/2015 | 881 | 7.2 |
| 3/13/2015 | 586 | 4.7 | 4/22/2015 | 952 | 6.3 |
| 3/14/2015 | 1110 | 4.6 | 4/23/2015 | 804 | 5.2 |
| 3/15/2015 | 1390 | 3.4 | 4/24/2015 | 721 | 4.9 |
| 3/16/2015 | 1210 | 3.8 | 4/25/2015 | 642 | 5.2 |
| 3/17/2015 | 1040 | 4.2 | 4/26/2015 | 586 | 5.5 |
| 3/18/2015 | 945 | 4.7 | 4/27/2015 | 576 | 7.2 |
| 3/19/2015 | 883 | 4.7 | 4/28/2015 | 689 | 7.4 |
| 3/20/2015 | 852 | 5.0 | 4/29/2015 | 772 | 6.9 |
| 3/21/2015 | 964 | 4.9 | 4/30/2015 | 709 | 7.0 |
| 3/22/2015 | 896 | 3.9 | 5/1/2015 | 732 | 7.4 |
| 3/23/2015 | 825 | 4.3 | 5/2/2015 | 823 | 7.0 |
| 3/24/2015 | 762 | 4.7 | 5/3/2015 | 828 | 6.9 |
| 3/25/2015 | 742 | 4.4 | 5/4/2015 | 904 | 7.4 |
| 3/26/2015 | 845 | 5.5 | 5/5/2015 | 977 | 6.2 |
| 3/27/2015 | 918 | 5.4 | 5/6/2015 | 794 | 5.3 |
| 3/28/2015 | 1200 | 5.7 | 5/7/2015 | 706 | 6.6 |
| 3/29/2015 | 1060 | 5.2 | 5/8/2015 | 726 | 7.6 |
| 3/30/2015 | 1020 | 5.7 | 5/9/2015 | 810 | 7.7 |
| 3/31/2015 | 1090 | 5.3 | 5/10/2015 | 962 | 7.8 |
| 4/1/2015 | 970 | 4.4 | 5/11/2015 | 1190 | 8.0 |
| 4/2/2015 | 870 | 4.3 | 5/12/2015 | 1370 | 6.4 |
| 4/3/2015 | 796 | 4.3 | 5/13/2015 | 1160 | 6.5 |
| 4/4/2015 | 732 | 4.8 | 5/14/2015 | 1160 | 6.8 |
| 4/5/2015 | 681 | 4.0 | 5/15/2015 | 1150 | 7.5 |
| 4/6/2015 | 633 | 4.0 | 5/16/2015 | 1350 | 8.1 |

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| 5/17/2015 | 1440 | 6.8 | 7/1/2015 | 699 | 14.4 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 5/18/2015 | 1520 | 8.1 | 7/2/2015 | 612 | 14.7 |
| 5/19/2015 | 1730 | 7.9 | 7/3/2015 | 587 | 15.2 |
| 5/20/2015 | 1920 | 8.2 | 7/4/2015 | 559 | 15.2 |
| 5/21/2015 | 1940 | 7.9 | 7/5/2015 | 525 | 15.0 |
| 5/22/2015 | 1920 | 8.0 | 7/6/2015 | 473 | 14.8 |
| 5/23/2015 | 2070 | 8.5 | 7/7/2015 | 449 | 14.7 |
| 5/24/2015 | 2050 | 8.6 | 7/8/2015 | 444 | 15.0 |
| 5/25/2015 | 1750 | 7.6 | 7/9/2015 | 436 | 15.2 |
| 5/26/2015 | 1430 | 7.6 | 7/10/2015 | 430 | 14.9 |
| 5/27/2015 | 1390 | 8.6 | 7/11/2015 | 418 | 14.2 |
| 5/28/2015 | 1650 | 9.2 | 7/12/2015 | 346 | 13.6 |
| 5/29/2015 | 1970 | 9.4 | 7/13/2015 | 310 | 13.6 |
| 5/30/2015 | 1950 | 9.6 | 7/14/2015 | 283 | 13.7 |
| 5/31/2015 | 1620 | 8.8 | 7/15/2015 | 271 | 13.8 |
| 6/1/2015 | 1450 | 9.0 | 7/16/2015 | 273 | 14.0 |
| 6/2/2015 | 1460 | 8.6 | 7/17/2015 | 259 | 14.0 |
| 6/3/2015 | 1220 | 8.4 | 7/18/2015 | 257 | 14.7 |
| 6/4/2015 | 1150 | 9.1 | 7/19/2015 | 288 | 15.5 |
| 6/5/2015 | 1220 | 10.1 | 7/20/2015 | 314 | 15.9 |
| 6/6/2015 | 1450 | 10.8 | 7/21/2015 | 293 | 15.2 |
| 6/7/2015 | 1650 | 11.1 | 7/22/2015 | 250 | 14.3 |
| 6/8/2015 | 1710 | 11.1 | 7/23/2015 | 226 | 14.4 |
| 6/9/2015 | 1520 | 10.9 | 7/24/2015 | 222 | 14.0 |
| 6/10/2015 | 1300 | 10.8 | 7/25/2015 | 254 | 13.7 |
| 6/11/2015 | 1190 | 11.1 | 7/26/2015 | 221 | 13.1 |
| 6/12/2015 | 1040 | 10.0 | 7/27/2015 | 194 | 12.9 |
| 6/13/2015 | 762 | 9.1 | 7/28/2015 | 180 | 13.7 |
| 6/14/2015 | 678 | 10.3 | 7/29/2015 | 190 | 14.5 |
| 6/15/2015 | 664 | 10.8 | 7/30/2015 | 202 | 15.0 |
| 6/16/2015 | 715 | 11.8 | 7/31/2015 | 215 | 15.1 |
| 6/17/2015 | 746 | 12.1 | 8/1/2015 | 220 | 15.2 |
| 6/18/2015 | 701 | 11.3 | 8/2/2015 | 218 | 14.9 |
| 6/19/2015 | 657 | 10.9 | 8/3/2015 | 220 | 14.2 |
| 6/20/2015 | 582 | 10.6 | 8/4/2015 | 211 | 14.1 |
| 6/21/2015 | 536 | 10.9 | 8/5/2015 | 198 | 14.3 |
| 6/22/2015 | 501 | 11.3 | 8/6/2015 | 186 | 13.8 |
| 6/23/2015 | 527 | 12.1 | 8/7/2015 | 180 | 14.0 |
| 6/24/2015 | 537 | 12.2 | 8/8/2015 | 189 | 14.4 |
| 6/25/2015 | 586 | 12.8 | 8/9/2015 | 191 | 14.6 |
| 6/26/2015 | 680 | 13.9 | 8/10/2015 | 197 | 14.7 |
| 6/27/2015 | 732 | 14.2 | 8/11/2015 | 208 | 14.2 |
| 6/28/2015 | 803 | 14.4 | 8/12/2015 | 212 | 14.3 |
| 6/29/2015 | 930 | 14.7 | 8/13/2015 | 230 | 15.0 |
| 6/30/2015 | 815 | 14.5 | 8/14/2015 | 228 | 14.2 |

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| 8/15/2015 | 206 | 13.4 | 9/29/2015 | 94.7 | 8.2 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8/16/2015 | 173 | 13.1 | 9/30/2015 | 92.3 | 8.6 |
| 8/17/2015 | 167 | 13.5 | 10/1/2015 | 95.7 | 9.1 |
| 8/18/2015 | 173 | 13.9 | 10/2/2015 | 98.4 | 9.4 |
| 8/19/2015 | 195 | 14.5 | 10/3/2015 | 97.4 | 9.9 |
| 8/20/2015 | 204 | 14.8 | 10/4/2015 | 91.5 | 9.7 |
| 8/21/2015 | 190 | 14.3 | 10/5/2015 | 89 | 9.1 |
| 8/22/2015 | 149 | 12.6 | 10/6/2015 | 86.6 | 8.6 |
| 8/23/2015 | 138 | 12.5 | 10/7/2015 | 116 | 9.1 |
| 8/24/2015 | 136 | 13.1 | 10/8/2015 | 193 | 9.7 |
| 8/25/2015 | 133 | 13.1 | 10/9/2015 | 142 | 10.1 |
| 8/26/2015 | 134 | 13.3 | 10/10/2015 | 429 | 10.1 |
| 8/27/2015 | 149 | 14.0 | 10/11/2015 | 650 | 8.7 |
| 8/28/2015 | 157 | 13.6 | 10/12/2015 | 268 | 8.5 |
| 8/29/2015 | 196 | 12.4 | 10/13/2015 | 273 | 9.3 |
| 8/30/2015 | 416 | 11.1 | 10/14/2015 | 197 | 8.2 |
| 8/31/2015 | 328 | 10.6 | 10/15/2015 | 161 | 7.3 |
| 9/1/2015 | 359 | 10.6 | 10/16/2015 | 142 | 7.5 |
| 9/2/2015 | 376 | 10.8 | 10/17/2015 | 144 | 8.0 |
| 9/3/2015 | 249 | 10.1 | 10/18/2015 | 153 | 9.0 |
| 9/4/2015 | 173 | 9.6 | 10/19/2015 | 172 | 9.3 |
| 9/5/2015 | 136 | 9.8 | 10/20/2015 | 140 | 9.0 |
| 9/6/2015 | 126 | 9.8 | 10/21/2015 | 123 | 8.0 |
| 9/7/2015 | 126 | 10.8 | 10/22/2015 | 111 | 7.7 |
| 9/8/2015 | 115 | 11.4 | 10/23/2015 | 101 | 6.8 |
| 9/9/2015 | 129 | 11.7 | 10/24/2015 | 95.4 | 6.0 |
| 9/10/2015 | 126 | 12.2 | 10/25/2015 | 92.5 | 6.2 |
| 9/11/2015 | 147 | 12.4 | 10/26/2015 | 107 | 7.2 |
| 9/12/2015 | 160 | 12.8 | 10/27/2015 | 111 | 7.5 |
| 9/13/2015 | 162 | 13.1 | 10/28/2015 | 102 | 7.0 |
| 9/14/2015 | 134 | 11.5 | 10/29/2015 | 174 | 7.3 |
| 9/15/2015 | 108 | 10.4 | 10/30/2015 | 675 | 7.4 |
| 9/16/2015 | 92.8 | 10.2 | 10/31/2015 | 3580 | 6.7 |
| 9/17/2015 | 89.5 | 10.0 | 11/1/2015 | 2110 | 5.4 |
| 9/18/2015 | 110 | 10.5 | 11/2/2015 | 1160 | 5.2 |
| 9/19/2015 | 103 | 10.8 | 11/3/2015 | 809 | 5.0 |
| 9/20/2015 | 181 | 11.2 | 11/4/2015 | 633 | 4.2 |
| 9/21/2015 | 278 | 11.4 | 11/5/2015 | 551 | 4.8 |
| 9/22/2015 | 154 | 10.0 | 11/6/2015 | 480 | 4.6 |
| 9/23/2015 | 123 | 9.2 | 11/7/2015 | 483 | 5.5 |
| 9/24/2015 | 115 | 9.9 | 11/8/2015 | 589 | 5.6 |
| 9/25/2015 | 147 | 10.6 | 11/9/2015 | 525 | 4.6 |
| 9/26/2015 | 167 | 10.3 | 11/10/2015 | 462 | 3.9 |
| 9/27/2015 | 114 | 9.0 | 11/11/2015 | 459 | 4.4 |
| 9/28/2015 | 101 | 8.2 | 11/12/2015 | 426 | 3.8 |

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| $11 / 13 / 2015$ | 3220 | 3.7 |
| :---: | :---: | :---: |
| $11 / 14 / 2015$ | 3970 | 4.6 |
| $11 / 15 / 2015$ | 2290 | 4.8 |
| $11 / 16 / 2015$ | 1560 | 3.5 |
| $11 / 17 / 2015$ | 3740 | 2.5 |
| $11 / 18 / 2015$ | 4280 | 3.4 |
| $11 / 19 / 2015$ | 2310 | 3.8 |
| $11 / 20 / 2015$ | 1670 | 2.9 |
| $11 / 21 / 2015$ | 1320 | 2.5 |
| $11 / 22 / 2015$ | 1100 | 2.5 |
| $11 / 23 / 2015$ | 958 | 2.7 |
| $11 / 24 / 2015$ | 850 | 3.2 |
| $11 / 25 / 2015$ | 740 | 2.8 |
| $11 / 26 / 2015$ | 662 | 1.5 |
| $11 / 27 / 2015$ | 606 | 1.2 |
| $11 / 28 / 2015$ | 560 | 0.9 |
| $11 / 29 / 2015$ | 525 | 1.0 |
| $11 / 30 / 2015$ | 486 | 1.1 |

## APPENDIX B: Daily Trap Operation Status

| Date | Trap Status | Comments | $\begin{gathered} 4 / 9 / 2015 \\ 4 / 10 / 2015 \end{gathered}$ | Op. <br> Op. |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 3/1/2015 | No Op. | Pulled-administrative | 4/11/2015 | Op. |  |
| 3/2/2015 | No Op. | Pulled-administrative | 4/12/2015 | Op. |  |
| 3/3/2015 | No Op. | Pulled-administrative | 4/13/2015 | Op. |  |
| 3/4/2015 | No Op. | Pulled-administrative | 4/14/2015 | Op. |  |
| 3/5/2015 | No Op. | Pulled-administrative | 4/15/2015 | Op. |  |
| 3/6/2015 | No Op. | Pulled-administrative | 4/16/2015 | Op. |  |
| 3/7/2015 | No Op. | Pulled-administrative | 4/17/2015 | Op. |  |
| 3/8/2015 | No Op. | Pulled-administrative | 4/18/2015 | Op. |  |
| 3/9/2015 | No Op. | Pulled-administrative | 4/19/2015 | Op. |  |
| 3/10/2015 | Op. |  | 4/20/2015 | Op. |  |
| 3/11/2015 | Op. |  | 4/21/2015 | Op. |  |
| 3/12/2015 | Op. |  | 4/22/2015 | Op. |  |
| 3/13/2015 | Op. |  | 4/23/2015 | Op. |  |
| 3/14/2015 | Op. |  | 4/24/2015 | Op. |  |
| 3/15/2015 | No Op. | Stopped-debris | 4/25/2015 | Op. |  |
| 3/16/2015 | Op. |  | 4/26/2015 | Op. |  |
| 3/17/2015 | Op. |  | 4/27/2015 | Op. |  |
| 3/18/2015 | Op. |  | 4/28/2015 | Op. |  |
| 3/19/2015 | Op. |  | 4/29/2015 | Op. |  |
| 3/20/2015 | Op. |  | 4/30/2015 | Op. |  |
| 3/21/2015 | Op. |  | 5/1/2015 | Op. |  |
| 3/22/2015 | Op. |  | 5/2/2015 | Op. |  |
| 3/23/2015 | Op. |  | 5/3/2015 | Op. |  |
| 3/24/2015 | Op. |  | 5/4/2015 | Op. |  |
| 3/25/2015 | Op. |  | 5/5/2015 | Op. |  |
| 3/26/2015 | Op. |  | 5/6/2015 | Op. |  |
| 3/27/2015 | Op. |  | 5/7/2015 | Op. |  |
| 3/28/2015 | Op. |  | 5/8/2015 | Op. |  |
| 3/29/2015 | Op. |  | 5/9/2015 | Op. |  |
| 3/30/2015 | Op. |  | 5/10/2015 | Op. |  |
| 3/31/2015 | Op. |  | 5/11/2015 | Op. |  |
| 4/1/2015 | Op. |  | 5/12/2015 | No Op. | Stopped-debris |
| 4/2/2015 | Op. |  | 5/13/2015 | Op. |  |
| 4/3/2015 | Op. |  | 5/14/2015 | Op. |  |
| 4/4/2015 | Op. |  | 5/15/2015 | Op. |  |
| 4/5/2015 | Op. |  | 5/16/2015 | Op. |  |
| 4/6/2015 | Op. |  | 5/17/2015 | Op. |  |
| 4/7/2015 | Op. |  | 5/18/2015 | Op. |  |
| 4/8/2015 | Op. |  | 5/19/2015 | Op. |  |
|  |  |  | 5/20/2015 | Op. |  |


| 5/21/2015 | Op. |  | 7/5/2015 | No Op. | Pulled - administrative |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 5/22/2015 | Op. |  | 7/6/2015 | No Op. | Pulled - administrative |
| 5/23/2015 | Op. |  | 7/7/2015 | No Op. | Pulled - administrative |
| 5/24/2015 | Op. |  | 7/8/2015 | No Op. | Pulled-administrative |
| 5/25/2015 | Op. |  | 7/9/2015 | No Op. | Pulled-administrative |
| 5/26/2015 | No Op. | Stopped-tampering | 7/10/2015 | No Op. | Pulled-administrative |
| 5/27/2015 | Op. |  | 7/11/2015 | No Op. | Pulled - administrative |
| 5/28/2015 | Op. |  | 7/12/2015 | No Op. | Pulled - administrative |
| 5/29/2015 | Op. |  | 7/13/2015 | No Op. | Pulled-administrative |
| 5/30/2015 | No Op. | Stopped-debris | 7/14/2015 | No Op. | Pulled-administrative |
| 5/31/2015 | Op. |  | 7/15/2015 | No Op. | Pulled-administrative |
| 6/1/2015 | Op. |  | 7/16/2015 | No Op. | Pulled-administrative |
| 6/2/2015 | Op. |  | 7/17/2015 | No Op. | Pulled-administrative |
| 6/3/2015 | Op. |  | 7/18/2015 | No Op. | Pulled-administrative |
| 6/4/2015 | Op. |  | 7/19/2015 | No Op. | Pulled-administrative |
| 6/5/2015 | Op. |  | 7/20/2015 | No Op. | Pulled-administrative |
| 6/6/2015 | Op. |  | 7/21/2015 | No Op. | Pulled-administrative |
| 6/7/2015 | Op. |  | 7/22/2015 | No Op. | Pulled-administrative |
| 6/8/2015 | Op. |  | 7/23/2015 | No Op. | Pulled-administrative |
| 6/9/2015 | Op. |  | 7/24/2015 | Op. |  |
| 6/10/2015 | Op. |  | 7/25/2015 | Op. |  |
| 6/11/2015 | Op. |  | 7/26/2015 | Op. |  |
| 6/12/2015 | Op. |  | 7/27/2015 | Op. |  |
| 6/13/2015 | Op. |  | 7/28/2015 | Op. |  |
| 6/14/2015 | Op. |  | 7/29/2015 | Op. |  |
| 6/15/2015 | Op. |  | 7/30/2015 | Op. |  |
| 6/16/2015 | Op. |  | 7/31/2015 | Op. |  |
| 6/17/2015 | Op. |  | 8/1/2015 | Op. |  |
| 6/18/2015 | Op. |  | 8/2/2015 | Op. |  |
| 6/19/2015 | Op. |  | 8/3/2015 | Op. |  |
| 6/20/2015 | Op. |  | 8/4/2015 | Op. |  |
| 6/21/2015 | Op. |  | 8/5/2015 | Op. |  |
| 6/22/2015 | Op. |  | 8/6/2015 | Op. |  |
| 6/23/2015 | Op. |  | 8/7/2015 | Op. |  |
| 6/24/2015 | Op. |  | 8/8/2015 | Op. |  |
| 6/25/2015 | Op. |  | 8/9/2015 | Op. |  |
| 6/26/2015 | Op. |  | 8/10/2015 | Op. |  |
| 6/27/2015 | Op. |  | 8/11/2015 | Op. |  |
| 6/28/2015 | Op. |  | 8/12/2015 | Op. |  |
| 6/29/2015 | Op. |  | 8/13/2015 | Op. |  |
| 6/30/2015 | Op. |  | 8/14/2015 | Op. |  |
| 7/1/2015 | Op. |  | 8/15/2015 | Op. |  |
| 7/2/2015 | No Op. | Pulled-administrative | 8/16/2015 | Op. |  |
| 7/3/2015 | No Op. | Pulled-administrative | 8/17/2015 | Op. |  |
| 7/4/2015 | No Op. | Pulled-administrative | 8/18/2015 | Op. |  |


| 8/19/2015 | Op. |  | 10/3/2015 | No Op. | Pulled-high flows |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8/20/2015 | Op. |  | 10/4/2015 | No Op. | Pulled-high flows |
| 8/21/2015 | Op. |  | 10/5/2015 | No Op. | Pulled-high flows |
| 8/22/2015 | Op. |  | 10/6/2015 | No Op. | Pulled-high flows |
| 8/23/2015 | Op. |  | 10/7/2015 | Op. |  |
| 8/24/2015 | Op. |  | 10/8/2015 | Op. |  |
| 8/25/2015 | Op. |  | 10/9/2015 | Op. |  |
| 8/26/2015 | Op. |  | 10/10/2015 | Op. |  |
| 8/27/2015 | Op. |  | 10/11/2015 | No Op. | Stopped-debris |
| 8/28/2015 | Op. |  | 10/12/2015 | Op. |  |
| 8/29/2015 | Op. |  | 10/13/2015 | Op. |  |
| 8/30/2015 | Op. |  | 10/14/2015 | Op. |  |
| 8/31/2015 | Op. |  | 10/15/2015 | Op. |  |
| 9/1/2015 | Op. |  | 10/16/2015 | Op. |  |
| 9/2/2015 | Op. |  | 10/17/2015 | Op. |  |
| 9/3/2015 | Op. |  | 10/18/2015 | Op. |  |
| 9/4/2015 | Op. |  | 10/19/2015 | Op. |  |
| 9/5/2015 | Op. |  | 10/20/2015 | Op. |  |
| 9/6/2015 | Op. |  | 10/21/2015 | Op. |  |
| 9/7/2015 | Op. |  | 10/22/2015 | Op. |  |
| 9/8/2015 | Op. |  | 10/23/2015 | Op. |  |
| 9/9/2015 | Op. |  | 10/24/2015 | Op. |  |
| 9/10/2015 | Op. |  | 10/25/2015 | Op. |  |
| 9/11/2015 | Op. |  | 10/26/2015 | Op. |  |
| 9/12/2015 | No Op. | Stopped-out of position | 10/27/2015 | Op. |  |
| 9/13/2015 | Op. |  | 10/28/2015 | Op. |  |
| 9/14/2015 | Op. |  | 10/29/2015 | Op. |  |
| 9/15/2015 | Op. |  | 10/30/2015 | Op. |  |
| 9/16/2015 | Op. |  | 10/31/2015 | No Op. | Pulled-high flows |
| 9/17/2015 | Op. |  | 11/1/2015 | No Op. | Pulled-high flows |
| 9/18/2015 | Op. |  | 11/2/2015 | Op. |  |
| 9/19/2015 | Op. |  | 11/3/2015 | Op. |  |
| 9/20/2015 | Op. |  | 11/4/2015 | Op. |  |
| 9/21/2015 | Op. |  | 11/5/2015 | Op. |  |
| 9/22/2015 | Op. |  | 11/6/2015 | No Op. | Stopped-debris |
| 9/23/2015 | Op. |  | 11/7/2015 | Op. |  |
| 9/24/2015 | Op. |  | 11/8/2015 | Op. |  |
| 9/25/2015 | Op. |  | 11/9/2015 | Op. |  |
| 9/26/2015 | Op. |  | 11/10/2015 | Op. |  |
| 9/27/2015 | Op. |  | 11/11/2015 | Op. |  |
| 9/28/2015 | Op. |  | 11/12/2015 | Op. |  |
| 9/29/2015 | Op. |  | 11/13/2015 | No Op. | Pulled-high flows |
| 9/30/2015 | Op. |  | 11/14/2015 | No Op. | Pulled-high flows |
| 10/1/2015 | Op. |  | 11/15/2015 | Op. |  |
| 10/2/2015 | No Op. | Pulled-high flows | 11/16/2015 | Op. |  |


| $11 / 17 / 2015$ | Op. |  |
| :---: | :---: | :---: |
|  |  |  |
| $11 / 18 / 2015$ | No Op. | Pulled-high flows |
| $11 / 19 / 2015$ | No Op. | Pulled-high flows |
| $11 / 20 / 2015$ | Op. |  |
| $11 / 21 / 2015$ | Op. |  |
| $11 / 22 / 2015$ | Op. |  |
| $11 / 23 / 2015$ | Op. |  |
| $11 / 24 / 2015$ | Op. |  |
| $11 / 25 / 2015$ | Op. |  |
| $11 / 26 / 2015$ | Op. |  |
| $11 / 27 / 2015$ | Op. |  |
| $11 / 28 / 2015$ | Op. |  |
| $11 / 29 / 2015$ | Op. |  |
| $11 / 30 / 2015$ | Op. |  |

## APPENDIX C: Regression Models

Model: Chinook Yearlings (Spring '08-'15) Back Position, ( $r^{2}=0.569 ; p=0.001$ )

| Origin/Species/Stage | Date | Marked | Recaptured | Trap <br> Efficiency | ASIN <br> Transform | Discharge <br> (cfs) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Yearlings | $4 / 10 / 2008$ | 25 | 2 | 0.120 | 0.354 | 229 |
| Wild Chinook Yearlings | $3 / 26 / 2009$ | 24 | 5 | 0.250 | 0.524 | 191 |
| Wild Chinook Yearlings | $3 / 30 / 2009$ | 34 | 4 | 0.147 | 0.394 | 193 |
| Wild Chinook Yearlings | $4 / 2 / 2009$ | 37 | 10 | 0.297 | 0.577 | 206 |
| Wild Chinook Yearlings | $4 / 5 / 2009$ | 59 | 15 | 0.271 | 0.548 | 205 |
| Wild Chinook Yearlings | $4 / 10 / 2009$ | 36 | 3 | 0.111 | 0.340 | 385 |
| Wild Chinook Yearlings | $3 / 12 / 2010$ | 25 | 1 | 0.080 | 0.287 | 300 |
| Wild Chinook Yearlings | $3 / 16 / 2010$ | 30 | 5 | 0.200 | 0.464 | 278 |
| Wild Chinook Yearlings | $3 / 20 / 2010$ | 21 | 1 | 0.095 | 0.314 | 283 |
| Wild Chinook Yearlings | $4 / 5 / 2010$ | 37 | 1 | 0.054 | 0.235 | 340 |
| Wild Chinook Yearlings | $4 / 9 / 2010$ | 31 | 4 | 0.161 | 0.413 | 310 |
| Wild Chinook Yearlings | $4 / 12 / 2010$ | 58 | 4 | 0.086 | 0.298 | 288 |
| Wild Chinook Yearlings | $4 / 16 / 2010$ | 73 | 2 | 0.041 | 0.204 | 381 |
| Wild Chinook Yearlings | $4 / 14 / 2012$ | 48 | 1 | 0.042 | 0.206 | 527 |

Model: Chinook Subyearlings (Fall '09-'15) Back Position, ( $r^{2}=0.130 ; p=0.086$ )

| Origin/Species/Stage | Date | Marked | Recaptured | Trap <br> Efficiency | ASIN <br> Transform | Discharge <br> $(\mathrm{cfs})$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wild Chinook Subyearlings | $8 / 20 / 2009$ | 20 | 2 | $15.00 \%$ | 0.398 | 311 |
| Wild Chinook Subyearlings | $8 / 29 / 2009$ | 34 | 4 | $14.71 \%$ | 0.394 | 227 |
| Wild Chinook Subyearlings | $10 / 7 / 2009$ | 22 | 2 | $13.64 \%$ | 0.378 | 95 |
| Wild Chinook Subyearlings | $10 / 16 / 2009$ | 34 | 6 | $20.59 \%$ | 0.471 | 134 |
| Wild Chinook Subyearlings | $11 / 17 / 2009$ | 35 | 3 | $11.43 \%$ | 0.345 | 375 |
| Wild Chinook Subyearlings | $11 / 23 / 2009$ | 21 | 0 | $4.76 \%$ | 0.22 | 313 |
| Wild Chinook Subyearlings | $11 / 21 / 2011$ | 39 | 2 | $7.69 \%$ | 0.281 | 172 |
| Wild Chinook Subyearlings | $10 / 4 / 2012$ | 33 | 5 | $18.18 \%$ | 0.441 | 140 |
| Wild Chinook Subyearlings | $10 / 24 / 2012$ | 87 | 6 | $8.05 \%$ | 0.288 | 268 |
| Wild Chinook Subyearlings | $10 / 28 / 2012$ | 36 | 1 | $5.56 \%$ | 0.238 | 711 |
| Wild Chinook Subyearlings | $10 / 31 / 2013$ | 46 | 7 | $17.39 \%$ | 0.43 | 258 |
| Wild Chinook Subyearlings | $11 / 6 / 2013$ | 38 | 9 | $26.32 \%$ | 0.539 | 248 |
| Wild Chinook Subyearlings | $11 / 9 / 2013$ | 40 | 6 | $17.50 \%$ | 0.432 | 251 |
| Wild Chinook Subyearlings | $11 / 13 / 2013$ | 29 | 2 | $10.34 \%$ | 0.327 | 422 |
| Wild Chinook Subyearlings | $11 / 23 / 2013$ | 25 | 3 | $16.00 \%$ | 0.412 | 406 |
| Wild Chinook Subyearlings | $11 / 27 / 2013$ | 24 | 0 | $4.17 \%$ | 0.206 | 335 |
| Wild Chinook Subyearlings | $9 / 17 / 2015$ | 39 | 4 | $12.82 \%$ | 0.366 | 89.5 |

## Appendix D. Historical Morphometric Data

Spring Chinook (2007-2015)

| Trap <br> Year | Brood <br> Year | Origin/Species/Stage | Fork Length (mm) |  |  | Weight (g) |  |  | K- <br> factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Mean | n | SD | Mean | n | SD |  |
| 2007 | 2005 | Wild Yearling Smolt | 93 | 173 | 8.5 | 8.6 | 173 | 2.2 | 1.1 |
| 2007 | 2005 | Wild Yearling Precocial Parr | 123 | 4 | 7.2 | 22.2 | 4 | 5.8 | 1.2 |
| 2007 | 2005 | Hatchery Yearling Smolt* | 76 | 208 | 17.9 | 5.4 | 203 | 4.2 | 1.2 |
| 2007 | 2005 | Hatchery Yearling Precocial Parr | 98 | 20 | 8.7 | 11.1 | 19 | 2.2 | 1.2 |
| 2007 | 2006 | Wild Subyearling Fry | 35 | 7 | 1.6 | - | - | - | - |
| 2007 | 2006 | Wild Subyearling Parr | 95 | 33 | 12.4 | 9.8 | 33 | 4.1 | 1.1 |
| 2008 | 2006 | Wild Yearling Smolt | 100 | 105 | 12.3 | 12.5 | 105 | 13.5 | 1.2 |
| 2008 | 2006 | Wild Yearling Precocial Parr | 126 | 9 | 8.4 | 22.8 | 9 | 4.1 | 1.1 |
| 2008 | 2006 | Hatchery Yearling Smolt | 117 | 229 | 12.7 | 18.7 | 228 | 9.8 | 1.2 |
| 2008 | 2006 | Hatchery Yearling Precocial Parr | 155 | 2 | 15.6 | 47.6 | 2 | 12.6 | 1.3 |
| 2008 | 2007 | Wild Subyearling Fry | 41 | 10 | 4.4 | - | - | - | - |
| 2008 | 2007 | Wild Subyearling Parr | 95 | 202 | 9.1 | 9.4 | 202 | 2.5 | 1.1 |
| 2009 | 2007 | Wild Yearling Smolt | 104 | 275 | 6.4 | 12.5 | 274 | 2.6 | 1.1 |
| 2009 | 2007 | Wild Yearling Precocial Parr | 134 | 5 | 7.0 | 28.5 | 2 | 2.7 | 1.2 |
| 2009 | 2007 | Hatchery Yearling Precocial Parr | 188 | 2 | 17.7 | 81.9 | 2 | 27.1 | 1.2 |
| 2009 | 2008 | Wild Subyearling Fry | 38 | 13 | 2.1 | - | - | - | - |
| 2009 | 2008 | Wild Subyearling Parr | 85 | 507 | 11.8 | 7.2 | 499 | 2.7 | 1.2 |
| 2010 | 2008 | Wild Yearling Smolt | 96 | 345 | 7.1 | 11.2 | 345 | 2.4 | 1.3 |
| 2010 | 2008 | Wild Yearling Precocial Parr | 130 | 15 | 10.3 | 26.4 | 15 | 6.6 | 1.2 |
| 2010 | 2009 | Wild Subyearling Fry | 40 | 31 | 3.6 | - | - | - | - |
| 2010 | 2009 | Wild Subyearling Parr | 87 | 166 | 12.6 | 7.7 | 166 | 3.0 | 1.2 |
| 2011 | 2009 | Wild Yearling Smolt | 99 | 64 | 7.7 | 11.3 | 64 | 2.8 | 1.2 |
| 2011 | 2009 | Wild Yearling Precocial Parr | 137 | 1 | - | 32.3 | 1 | - | 1.3 |
| 2011 | 2009 | Hatchery Yearling Smolt | 127 | 46 | 10.6 | 24.3 | 46 | 6.5 | 1.2 |
| 2011 | 2010 | Wild Subyearling Fry | 37 | 26 | 2.5 | - | - | - | - |
| 2011 | 2010 | Wild Subyearling Parr | 91 | 159 | 13.0 | 9.2 | 159 | 7.1 | 1.2 |
| 2012 | 2010 | Wild Yearling Smolt | 98 | 182 | 7.9 | 10.9 | 179 | 2.8 | 1.2 |
| 2012 | 2010 | Wild Yearling Precocial Parr | 123 | 13 | 12.7 | 22.4 | 13 | 6.5 | 1.2 |
| 2012 | 2011 | Hatchery Subyearling Fry | 84 | 29 | 4.4 | 6.5 | 2 | 2.3 | 1.1 |
| 2012 | 2011 | Hatchery Subyearling Parr | 110 | 25 | 7.4 | 14.6 | 25 | 3.3 | 1.1 |
| 2012 | 2011 | Wild Subyearling Fry | 35 | 18 | 2.7 | - | - | - | - |
| 2012 | 2011 | Wild Subyearling Parr | 91 | 315 | 10.1 | 8.8 | 288 | 2.8 | 1.2 |
| 2013 | 2011 | Wild Yearling Smolt | 103 | 20 | 7.0 | 12.3 | 20 | 3.0 | 1.1 |
| 2013 | 2011 | Wild Yearling Precocial Parr | 111 | 2 | 0.7 | 13.5 | 2 | 3.0 | 1.0 |
| 2013 | 2011 | Hatchery Yearling Precocial Parr | 155 | 4 | 17.4 | 43.4 | 4 | 17.8 | 1.2 |
| 2013 | 2012 | Wild Subyearling Fry | 40 | 77 | 8.1 | - | - | - | - |


| 2013 | 2012 | Wild Subyearling Parr | 84 | 445 | 12.3 | 6.7 | 444 | 4.7 | 1.1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2014 | 2012 | Wild Yearling Smolt | 94 | 43 | 7.0 | 9.4 | 43 | 2.2 | 1.1 |
| 2014 | 2012 | Wild Yearling Precocial Parr | 127 | 7 | 13.0 | 23.2 | 7 | 7.4 | 1.1 |
| 2014 | 2013 | Wild Subyearling Fry | 40 | 22 | 3.8 | - | - | - | - |
| 2014 | 2013 | Wild Subyearling Parr | 86 | 185 | 14.1 | 7.5 | 185 | 3.3 | 1.2 |
| 2015 | 2013 | Wild Yearling Smolt | 103 | 32 | 6.8 | 13.0 | 31 | 2.8 | 1.1 |
| 2015 | 2013 | Wild Yearling Precocial Parr | 145 | 2 | 13.4 | 35.2 | 2 | 11.4 | 1.1 |
| 2015 | 2014 | Wild Subyearling Fry | 38 | 11 | 3.3 | 0.5 | 10 | 0.2 | 0.9 |
| 2015 | 2014 | Wild Subyearling Parr | 96 | 152 | 7.5 | 10.4 | 149 | 6.3 | 1.2 |

${ }^{\text {a }}$ Includes residualized non-precocial smolts caught after June 30
b "Fry" classification based on age despite FL $\geq 50 \mathrm{~mm}$

## Appendix M

Genetic Diversity of Upper Columbia River Summer Chinook Salmon

## Genetic Structure of upper Columbia River Summer Chinook and

 Evaluation of the Effects of Supplementation Programsby

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February 2011


#### Abstract

We investigated genetic relationships among temporally replicated collections of summer Chinook from the Wenatchee River, Methow River, and Okanogan River in the upper Columbia River basin. Samples from the Eastbank Hatchery Wenatchee stock, Eastbank Hatchery - MEOK stock, and Wells Hatchery were also included in the analysis. Samples of natural- and hatchery-origin summer Chinook were analyzed and compared to determine if the supplementation program has had any impacts to the genetic structure of these populations. We also calculated the effective number of breeders for collection locations of natural- and hatchery-origin summer Chinook from 1993 and 2008. In general, population differentiation was not observed among the temporally replicated collection locations. A single collection from the Okanogan River (1993) was the only collection showing statistically significant differences. The effective number of breeders was not statistically different from the early collection in 1993 in comparison to the late collection in 2008. Overall, these analyses revealed a lack of differentiation among the temporal replicates from the same locations and among the collection from different locations, suggesting the populations have been homogenized or that there has been substantial gene flow among populations. Additional comparisons among summer-run and fall-run Chinook populations in the upper Columbia River were conducted to determine if there was any differentiation between Chinook with different run timing. These analyses revealed pairwise $F_{\text {ST }}$ values that were less than 0.01 for the collections of summer Chinook to collections of fall Chinook from Hanford Reach, lower Yakima River, Priest Rapids, and Umatilla. Collections of fall Chinook from Crab Creek, Lyons Ferry Hatchery, Marion Drain, and Snake River had pairwise FsT values that were higher in comparison to the collections of summer Chinook. The consensus clustering analysis did not provide good statistical support to the groupings, but did show relationships among collections based on geographic proximity. Overall the summer and fall run Chinook that have historically been


spawned together were not differentiated while fall Chinook from greater geographic distances were differentiated.

## Introduction

The National Marine Fisheries Service (NMFS) recognizes 15 Evolutionary Significant Units (ESU) for Chinook salmon (Oncorhynchus tshawytscha) (Myers et al. 1998). The summer Chinook from the upper Columbia River are included in the Upper Columbia River Summer- and Fall-Run ESU, which encompasses all late-run (summer and fall), ocean-type Chinook salmon from the mainstem Columbia River and its tributaries (excluding the Snake River) between Chief Joseph and McNary Dams (Waknitz et al. 1995). Waknitz et al. (1995) concluded that due to high total abundance this ESU was not likely to become at risk from extinction. Yet, a majority of natural spawning activity was in the vicinity of Hanford Reach, and it was unclear whether natural production was selfsustaining given the vast summer Chinook artificial propagation efforts (Waknitz et al. 1995). Additionally, the Biological Review Team expressed concern about potential consequences to genetic and life-history traits from an increasing contribution of hatchery fish to total spawning escapement (Waknitz et al. 1995).

Artificial propagation of ocean-type Chinook from the middle/upper Columbia has been continuous since the implementation of the Grand Coulee Fish Maintenance Project (GCFMP) in 1939 (Myers et al. 1998). The US Fish and Wildlife Service established three hatchery programs for summer/fall Chinook during the GCFMP, Leavenworth NFH, Entiat NFH, and Winthrop NFH. The Washington Department of Fisheries (now Washington Department of Fish and Wildlife) followed with hatchery programs at Rocky Reach (1964), Wells Dam (1967), Priest Rapids (1974), and Eastbank (1990) facilities. Currently, only Leavenworth NFH and Winthrop NFH are not producing summer/fall Chinook. Entiat NFH has resumed production of summer/fall Chinook (Wells FH Stock) in 2009 and released their first yearling summer Chinook smolts in 2010. Since

1941, over 200 million ocean-type Chinook salmon have been released into the middle Columbia River Basin (Myers et al. 1998). Initially, the hatchery programs differentiated between early returning fish (i.e., stream-type) and later returning fish (i.e., ocean-type), but no distinction was made regarding the "summer" and "fall" components of the ocean-type stocks (Waknitz et al. 1995). Therefore, all Chinook salmon now migrating above Rock Island Dam descend from not only a mixture between different stocks from the basin, but also a mixture between the endemic summer and fall life histories. While hatchery protocols have been modified of late to maintain discreet summer and fall Chinook hatchery stocks (Utter et al. 1995; see also HGMP), physical evidence and genetic data suggests that summer and fall Chinook may have become homogenized. During the 1970's and 80's, given coded-wire tag recoveries, summer-run Chinook originating from above Rock Island Dam were believed to have spawned extensively with Hanford Reach and Priest Rapids Hatchery fish (Chapman 1994). Stuehrenberg et al. (1995) reported that $10 \%$ of their radio tagged summer Chinook were occupying typical fall-run spawning habitat on the mainstem Columbia river, and $25 \%$ of fall fish released from Priest Rapids were recovered as summers at (or above) Wells Hatchery. Genetic data reported by Marshall et al. (1995) and Waknitz et al. (1995) corroborate these observations, as genetic distances observed between summer and fall Chinook within the Upper Columbia River Summer- and Fall-Run ESU were essentially zero.

In response to the need for evaluation of the supplementation hatchery programs, both a monitoring and evaluation plan (DCPUD 2005; Murdoch and Peven 2005) and the associated analytical framework (Hays et al. 2006) were developed for the Habitat Conservation Plan's Hatchery Committee through the joint effort of the fishery co-managers (CCT, NMFS, USFWS, WDFW, and YN) and Chelan County and Douglas County PUDs. These reports outline 10 objectives to be applied to various species assessing the impacts of hatchery operations mitigating the operation of Wells, Rocky Reach, and Rock Island hydroelectric projects. The present monitoring and evaluation study plan differs
in scope from previous monitoring and evaluation projects proposed by WDFW Molecular Genetics Lab, in that it does not investigate a single watershed, but instead will encompass all summer Chinook stocks from the upper Columbia River including the three supplementation (Wenatchee, Methow, and Okanogan) and the harvest augmentation program (Wells summer Chinook). The objectives of this study were to determine if genetic diversity, population structure, and effective population size have changed in natural spawning populations as a result of the hatchery programs.

## Materials and Methods

## Collections

A total of 2,416 summer Chinook were collected from tributaries in the upper Columbia River basin and were analyzed (Table 1). Two collections of naturalorigin summer Chinook from 1993 (prior to the supplementation program) were taken from the Wenatchee River Basin and were compared to collections of hatchery and natural-origin from 2006 and 2008 that were post-supplementation. Two pre-supplementation collections from the Methow River (1991 and 1993) were compared to post-supplementation collections from 2006 and 2008. Three pre-supplementation collections from the Okanogan River Basin (1991, 1992, and 1993) were compared with post-supplementation collections from 2006 and 2008. A collection of natural-origin summer Chinook from the Chelan River was also analyzed. Additionally, hatchery collections from Eastbank Hatchery (Wenatchee and MEOK stock) and Wells Hatchery were analyzed and compared to the in-river collections. Summer Chinook data (provided by the USFWS) from the Entiat River was also used for comparison. Lastly, data from eight collections of fall Chinook was compared to the collections of summer Chinook.

## Laboratory Analyses

All laboratory analyses were conducted at the WDFW Genetics Laboratory in Olympia, Washington. Genomic DNA was extracted by digesting a small piece of fin tissue using the nucleospin tissue kits obtained from Macherey-Nagel following the recommended conditions in the user manual. Extracted DNA was eluted with a final volume of $100 \mu \mathrm{~L}$.

Genotype information was generated using thirteen microsatellite markers following standard laboratory protocols and analysis methods. Descriptions of the loci assessed in this study and polymerase chain reaction (PCR) conditions are given in Table 2. PCR reactions were run with a thermal profile consisting of: denaturation at $95^{\circ} \mathrm{C}$ for 3 min , denaturation at $95^{\circ} \mathrm{C}$ for 15 sec , anneal for 30 sec at the appropriate temperature for each locus (Table 2), extension at $72^{\circ} \mathrm{C}$ for 1 min, repeat cycle (steps 2-4), final extension at $72^{\circ} \mathrm{C}$ for 30 minutes. PCR products were then processed with an ABI-3730 DNA Analyzer. Genotypes were visualized with a known size standard (GS500LIZ 3730) using GENEMAPPER 3.7 software. Alleles were binned in GENEMAPPER using the standardized allele sizes established for the Chinook GAPS dataset (Seeb et al. 2007).

## Within-collection Statistical Analyses

Allele frequencies were calculated with CONVERT (version 1.3, Glaubitz 2003). Hardy-Weinberg proportions for all loci within each collection were calculated using GENEPOP (version 3.4, Raymond and Rousset 1995). Heterozygosity (observed and expected) was computed for each collection group using GDA (Lewis and Zaykin 2001).

Allelic richness and FIS $_{\text {IS }}$ (Weir and Cockerham 1984) inbreeding coefficient were calculated using FSTAT (version 2.9.3.2, Goudet 2001). Linkage disequilibrium for each pair of loci in each collection was calculated using GENEPOP v 3.4 (10,000 dememorizations, 100 batches, and 5,000 iterations per batch). Pairwise estimates of genetic differentiation between collection groups were
calculated using GENEPOP (version 3.4, Raymond and Rousset 1995).
Statistical significance for the tests of Hardy-Weinberg proportions, linkage disequilibrium, and genotypic differentiation was evaluated using a Bonferroni correction of $p$-values to account for multiple, simultaneous tests (Rice 1989).

## Between-collection Statistical Analyses

Pairwise $\mathrm{F}_{\text {ST }}$ estimates were computed to examine population structure among collections using GENETIX (version 4.03, Belkhir et al. 2001). This estimate uses allelic frequency data and departures from expected heterozygosity to assess differences between pairs of populations.

We used PHYLIP (version 3.5c, Felsenstein 1993) to calculate Cavalli-Sforza and Edwards (1967) pairwise chord distances between collections. Bootstrap calculations were performed using SEQBOOT followed by calculations of genetic distance using GENDIST. The NEIGHBOR-JOINING method of Saitou and Nei (1987) was used to generate the dendrograms and CONSENSE to generate a final consensus tree from the 1,000 replicates. The dendrogram generated in PHYLIP was plotted as an unrooted radial tree using TREEVIEW (version 1.6.6, Page 1996).

## Effective Number of Breeders

The effective number of breeders $\left(\mathrm{N}_{\mathrm{b}}\right)$ was estimated for pre- and postsupplementation program collections (where possible) to investigate whether hatchery programs had affected that genetic metric over the operational period. Wang (2009) derived an equation for effective size $\left(\mathrm{N}_{\mathrm{e}}\right)$ as a function of the frequency of nested full-sib and half-sib families in a random collection of individuals.

$$
\begin{equation*}
\frac{1}{N_{e}}=\frac{1+3 \alpha}{4}\left(Q_{1}+Q_{2}+2 Q_{3}\right)-\frac{\alpha}{2}\left(\frac{1}{N_{1}}+\frac{1}{N_{2}}\right) \tag{equation10}
\end{equation*}
$$

Where $\alpha$ is a measure of the deviation of genotype frequencies from HardyWeinberg expectation (equivalent to Wright's (1969) $\mathrm{F}_{\text {IS }}$ ), $Q_{i}$ are the probabilities that a pair of offspring are paternal half sibs, maternal half sibs, or full sibs, respectively, and $\mathrm{N}_{1}$ and $\mathrm{N}_{2}$ are the number of male and female parents that generation, respectively. Genetic parameters (i.e., sibship distributions) were estimated for summer Chinook collections using algorithms implemented in COLONY (Jones and Wang 2009). To be clear, Wang's (2009) method as implemented here will estimate $N_{b}$, given multi-locus genotypes from each collection were partitioned by brood year for this analysis. To obtain an estimate of $N_{e}$ each $N_{b}$ value must be multiplied by the mean generation time of that population.

## Results

## Collections

A total of 2,350 individuals from 32 collections of temporally replicated samples (six locations) were analyzed (Table 1). Temporally replicated collections of hatchery and natural-origin samples were from the Wenatchee, Methow, and Okanogan Rivers. Temporally replicated hatchery-origin summer Chinook were from Wells Hatchery, Eastbank Hatchery - Wenatchee stock, and Eastbank Hatchery - Methow/Okanogan (MEOK) stock. A total of 232 of those individuals were excluded from any analyses because they failed to amplify at nine or more loci. Data for remaining 2,118 individuals were analyzed to assess differences between temporally replicated natural- and hatchery-origin summer Chinook for each location and to compare the differences among the different collection locations. Summer Chinook data from the temporally replicated collection locations were then combined and compared to fall Chinook data from the GAPS v.3.0 dataset.

## Statistical Analyses

The population statistics (Hardy-Weinberg equilibrium and $\mathrm{F}_{\text {IS }}$ ) calculated for each of the 32 temporally replicated collection locations were consistent with neutral expectations (i.e., no associations among alleles). Three collections did have a single locus that did not meet expectations (Wenatchee hatchery-origin 2006, Wells hatchery 2006, and Okanogan hatchery-origin 2009). Based on these results we suggest the collections represented randomly breeding groups and were not comprised of mixtures of individuals from different genetic source populations.

Population differentiation was assessed for each of the temporally replicated collections from within each location (Table 3). This analysis revealed the only significant difference observed within a collection location pertained to the collection from 1993 Okanogan River natural-origin samples. Because of the significant difference of this collection to the other temporal replicates it was not included in further analyses.

Given the absence of genetic differentiation observed among the temporally replicated collections, the 32 collections from the Wenatchee, Methow, and Okanogan River were combined to form three location-specific collections for analysis. Population differentiation metrics were compared among the composite Wenatchee, Methow, and Okanogan collections and eight other location-specific collections ( 11 locations total). Comparing all collections, there were a total of 39 significant genic test comparisons out of a total 496 (Table 4). Thirty-eight of the 39 statistically significant pairwise differences pertained to the Okanogan River and 2006 Wells Hatchery collections (Table 4). Fst results are described further below.

Within-collection genetic metrics were estimated for the 11 location-specific collections of summer Chinook from the upper Columbia River, in addition to eight collections of fall Chinook (Table 1). The population statistics (HardyWeinberg equilibrium and $\mathrm{F}_{\text {Is }}$ ) calculated for these collections of summer and fall

Chinook were also consistent with neutral expectations. The collection from Lyons Ferry Hatchery had one locus that did not meet expectations and the collections from Crab Creek and Marion Drain both had three loci that did not meet expectations.

The hatchery collections in general had a higher percentage of significantly linked loci; however the observed genetic diversity were similar for the natural and hatchery-origin collections. Analysis of allelic richness was based on 11 individuals per collection, the minimum number of individuals across all collections with complete multilocus genotypes. The largest number of linked loci occurred in the Crab Creek, Entiat River, and Okanogan natural-origin collections. Allelic richness was on average lower in the collections of summer Chinook (10.7) collections in comparison to the collections of fall Chinook (11.0).

Pairwise $\mathrm{F}_{\text {ST }}$ (Table 4) estimates revealed low levels of differentiation, where all observed $\mathrm{F}_{\text {ST }}$ values between the collections of summer Chinook were lower than 0.0096 . There were 15 out of 28 comparisons between collections of summer Chinook that were significantly different from zero and occurred primarily from comparisons of the Okanogan River (hatchery and natural-origin) and Wells Hatchery to all other collections. The collection of Eastbank Hatchery - MEOK stock was differentiated from the Wenatchee River natural-origin and Entiat River collections. The collection from the Chelan River had a small sample size of 23 individuals and only differentiated from the Eastbank Hatchery - MEOK stock. $\mathrm{F}_{\text {ST }}$ estimates regarding pairwise comparisons between each of four fall Chinook collection locations (Crab Creek, Lyons Ferry Hatchery, Marion Drain, and Snake River) to all other collections were significantly different from zero (Table 5). Pairwise comparisons for three other fall Chinook collections (Hanford Reach, lower Yakima River, and Umatilla River) to the collections of summer Chinook were significantly different from zero (Table 6). The only fall Chinook collection that was not significantly differentiated from all of the summer Chinook was Priest Rapids.

The relative genetic relationships among the test groups were assessed using the consensus clustering analysis (Figure 1). Statistical support for the dendrogram topology (i.e., tree shape) was low regarding the branching that separated the collections of summer Chinook from the upper Columbia River. The collections of fall Chinook; however were supported with bootstrap support over $76 \%$ with the exception of three collections (lower Yakima River, Crab Creek, and Umatilla River). In other words, 760 of the 1000 bootstrap replicates supported the placement of the node separating summer and fall collections. The collection from the Chelan River had bootstrap support of 68\%; however the sample size for that collections was small $(\mathrm{N}=23)$. Even though the bootstrap support was low among the collections of summer Chinook there was concordance between geography and genetic distance.

Where comparisons were possible between pre- and post-supplementation program collections, the effective number of breeders $\left(N_{b}\right)$ estimated to have comprised those collections were slightly lower for contemporary (2008) collections; however in all cases the 95\% confidence intervals overlapped between historical and contemporary collections, suggesting statistical equivalency. Regarding Wenatchee River collections, the point estimates of $\mathrm{N}_{\mathrm{b}}$ ranged from 134 (08FU) to 190 (93DD), where all collections had overlapping confidence intervals (Table 7). The upper bound of the 1989 brood year for collection 93DD was very large, suggesting the sample size was insufficient for properly inferring the sibship distribution within the collection. Comparing the Okanogan natural collections 93ED and 08GA, the estimated $\mathrm{N}_{\mathrm{b}}$ were 142 (CI 102 - 203) and 127 (CI 92 - 180), respectively. For the Eastbank Hatchery MEOK stock comparisons, the $\mathrm{N}_{\mathrm{b}}$ estimated for the 93DF collection was 171 (CI 129 - 229), as compared to the $166(\mathrm{Cl} 126-226)$ estimated for collection 08 MO . In all cases, the estimated $\mathrm{N}_{\mathrm{b}}$ can be converted to effective population size $\left(\mathrm{N}_{\mathrm{e}}\right)$ by multiplying the estimate by the mean generation time.

## Discussion

The collections of summer Chinook populations from the upper Columbia River are of interest because census sizes are reduced below historic levels and are the subject of mitigation and supplementation hatchery programs. Concern over the impacts of hatchery supplementation programs on the genetic integrity of natural-origin populations led to our primary objective, which was to evaluate genetic metrics for temporally replicated collections of summer Chinook in the upper Columbia River pre and post hatchery supplementation. A similar analysis by Kassler and Dean (2010) was conducted on spring Chinook in the Tucannon River to evaluate the effects of a supplementation and captive brood program on natural-origin stocks. Additionally, upper Columbia River spring Chinook supplementation programs (Blankenship et al. 2007; Small et al. 2007), spring and fall Chinook populations in the Yakima Basin (Kassler et al. 2008), and a potentially unique population of fall Chinook in Crab Creek (Small et al. 2010) have been evaluated. In the present analysis of summer Chinook populations, collections of pre- and post- supplementation summer Chinook were collected from the Wenatchee River, Methow River, and Okanogan River Basins and analyzed to determine if the genetic profile has changed as a result of the supplementation program. Analysis was then conducted on the collections of summer run to compare the fall run Chinook collections in the upper Columbia River basin.

Allozyme analyses of these three summer run Chinook stocks in the upper Columbia River have identified that each stock was distinct, with a closer relationship detected between the Wenatchee and Methow Rivers (WDF and WDW 1993, Marshall 2002). Wenatchee summer Chinook are thought to be a mixture of native summer Chinook and Chinook from the Grand Coulee Fish Maintenance Project (GCFMP). The goal of the GCFMP project between 1939 and 1943 was to trap migrating Chinook salmon at Rock Island dam ( 75 miles below Grand Coulee) and homogenize the populations, which reduced the
genetic uniqueness of the distinct tributary populations present in the upper Columbia River.

We found allele frequencies for individual temporally replicated hatchery- and natural-origin collection locations of adult summer Chinook were not significantly different from that expected of a single underlying population, except for one collection (1993 Okanogan natural-origin; Table 3). This collection was differentiated to the Okanogan collections in 2006 and 2008; however it was not differentiated from the collection in 1992. The Okanogan collection from 1992 was also not differentiated to any other collection; therefore the difference in the collection from Okanogan 1993 was likely not an indication of genetic change from pre supplementation to post supplementation. The collection was however dropped from further analyses so as to not confuse interpretation of results. The lack of allelic differentiation observed among the temporally replicated collections was interpreted as the genetic metrics from each location in the early 1990's did not differ from the samples collected in 2008. Spanning a few generations, allele frequencies are not expected to change for large populations at genetic equilibrium. In contrast, changes in allele frequencies of small populations may occur due to the stochastic sampling of genes from one generation to the next (i.e., genetic drift).

A second round of analyses was conducted to evaluate the genetic relationships of the summer run collections (temporal collections were combined) with data from the Entiat River, Chelan River, and eight collections of fall Chinook. Assessment of the relationship between the summer run collections in comparison to each other provided very little evidence of genetic differentiation between these collections. While population differentiation did show some significant differences between the Okanogan River and Wells Hatchery collections, all of the pairwise $F_{S T}$ values were below 0.003 . Meaning that a very small proportion of the observed genetic variation could be attributed to restrictions in gene flow (i.e., population structure)

The comparison of the hatchery-origin collections revealed a lack of differentiation between the Eastbank Hatchery - Wenatchee stock, Eastbank Hatchery - MEOK stock, and the Wells Hatchery (with exception of the 2006 collection). The genetic similarity or low level of genetic differentiation among these stocks suggests that there has been an integration of natural- and hatchery-origin summer Chinook in the upper Columbia River or a lack of ancestral genetic difference. The difference of the 2006 Wells Hatchery collection to the other collections is most likely a result of sampling effect because of the lack of differentiation among the stocks in the basin. If the 2006 collection had been mixed from different sources of summer Chinook there would not be a detectable level of differentiation as was seen with the 2006 sample.

The analyses to compare summer and fall Chinook collections provided some understanding on the genetic relationships of Chinook with different run timings in the upper Columbia River basin. Historically, the hatchery programs in the upper Columbia River were separated into groups of the early returning fish (i.e., stream-type) and later returning fish (i.e., ocean-type), but the programs did not sort individuals identified as "summer" or "fall" stocks (Waknitz et al. 1995). Now all Chinook salmon that are migrating above Rock Island Dam descend from a mixture of different stocks from the upper Columbia River basin, but also a mixture between the endemic summer and fall life histories.

Small et al. (2010) conducted an analysis on summer run and fall run Chinook in the upper Columbia River and concluded that Crab Creek Chinook in the upper Columbia River were genetically distinct to all other fall and summer run Chinook stocks that were analyzed. They did note a departure from Hardy Weinberg expectation as a result of a null allele at the microsatellite locus Ogo-4 and a higher linkage disequilibrium value due to the inclusion of family groups in one of their samples. Kassler et al. (2008) found differentiation among spring and fall Chinook populations in the Yakima River.

The tests of pairwise Fst $_{\text {indicated a very low level of genetic differentiation (less }}$ than one percent difference) between collections of summer-run Chinook and fall-run Chinook. The range of pairwise $\mathrm{F}_{\text {St }}$ values for comparisons between the summer run and fall run collections was $0.0016-0.0248$. The larger values from the range were associated to the collections from Crab Creek, Lyons Ferry Hatchery, and Marion Drain. Studies by Kassler et al. (2008) and Small et al. (2010) have documented differences among the populations of these collections to others within the upper Columbia River basin. The low pairwise $\mathrm{F}_{\text {ST }}$ values between Priest Rapids and Hanford Reach collections and the summer run collections were not surprising because summer-run Chinook originating from above Rock Island Dam were believed to have spawned extensively with Hanford Reach and Priest Rapids Hatchery fish during the 1970's and 80's (Chapman 1994). The lack of differentiation among the summer and fall stocks in the Columbia River was also identified by Utter et al. (1995) and the HGMP where they state physical evidence and genetic data suggests that summer and fall Chinook may have become homogenized.

Despite low levels of statistical bootstrap support for dendrogram topology (i.e., tree shape), there was concordance observed between geographic location and the genetic relationships among the summer and fall Chinook populations. The collections from the Okanogan (hatchery and natural-origin) did separate out with collections from Wells Dam Hatchery, Entiat River, and Eastbank Hatchery MEOK stock, and were next to a group of the Methow and Wenatchee collections. The fall Chinook populations are also separated to the summer collections and the position of all but three of these collections (lower Yakima River, Crab Creek, and Umatilla River) were statistically supported. The geographic proximity of the fall collections seemed to follow the observed pattern in this dendrogram. The relationship of the Snake River and Lyons Ferry Hatchery in proximity to the collection from Marion Drain was not surprising while
the relationship between Priest Rapids and Hanford Reach was easily a result of the stocking practices of fall Chinook in the 1970 and 1980's.

A secondary objective of this study was to determine if the effective population size of upper Columbia River summer Chinook populations had changed over time due to supplementation efforts. We observed that the number of effective breeders in the collections from 1993 and 2008 has not changed thus providing reason to believe that the genetic diversity of summer Chinook in the upper Columbia River has not been altered through the supplementation program.

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## References

Banks, M.A., M.S. Blouin, B.A. Baldwin, V.K. Rashbrook, H.A. Fitzgerald, S.M. Blankenship, and D. Hedgecock. 1999. Isolation and inheritance of novel microsatellites in Chinook salmon (Oncorhynchus tshawytscha). Journal of Heredity 90:281-288.

Belkhir, K., P. Borsa, L. Chikhi, N. Raufaste, and F. Bonhomme. 2001. Genetix, logiciel sous Windows TM pour la genetique des populations. Laboratoire Genome, Populations, Interactions: CNRS UMR 5000, Universite de Montpellier II, Montpellier, France.

Blankenship, S.M., J.F. VonBargen, K.I. Warheit, and A.R. Murdoch. 2007. Assessing the Genetic Diversity of Natural Chiwawa River Spring Chinook Salmon and Evaluating the Effectiveness of its Supportive Hatchery Supplementation Program. . Final Report. Unpublished Washington Department of Fish and Wildlife Molecular Genetics Laboratory Report submitted to Chelan County PUD.

Cairney, M., J.B. Taggart, and B. Hoyheim. 2000. Characterization of microsatellite and minisatellite loci in Atlantic salmon (Salmo salar L.) and cross-species amplification in other salmonids. Molecular Ecology 9:21752178.

Cavalli-Sforza, L.L. and A.W.F. Edwards. 1967. Phylogenetic analysis: models and estimation procedures. Evolution 32:550-570.

Chapman, D., A. Giorgi, T. Hillman, D. Deppert, M. Erho, S. Hays, C. Peven, B. Suzumoto, and R. Klinge. 1994. Status of summer/fall chinook salmon in the mid-Columbia region. Report for Chelan, Douglas, and Grant County PUDs. 412 p. + app. (Available from Don Chapman Consultants, 3653 Rickenbacker, Ste. 200, Boise, ID 83705.)

DCPUD. 2005. Conceptual approach for monitoring and evaluating the Douglas County Public Utility District hatchery programs. Douglas County Public Utility District, Wenatchee, Washington. 105 p.

Felsenstein, J. 1993. PHYLIP (Phylogeny Inference Package) version 3.5c. Distributed by the author. Department of Genetics, University of Washington, Seattle, WA.

Glaubitz, J.C. 2003. CONVERT (version 1.2): A user-friendly program to reformat diploid genotypic data for commonly used population genetic software packages.
http://www.agriculture.purdue.edu/fnr/html/faculty/Rhodes/Students\ an d\%20Staff/glaubitz/software.htm.

Goudet, J. 2001. FSTAT, a program to estimate and test gene diversities and fixation indices (version 293). Updated from Goudet (1995). Available from http://www.unilch/izea/softwares/fstat.html.

Greig, C., J.P. Jacobson, and M.A. Banks. 2003. New tetranucleotide microsatellites for fine-scale discrimination among endangered Chinook salmon (Oncorhynchus tshawytscha). Molecular Ecology Notes 3:376379.

Hays, S., T. Hillman, T. Kahler, R. Klinge, R. Langshaw, B. Lenz, A. Murdoch, K. Murdoch, and C. Peven. 2006. Decision rules for monitoring and evaluating district hatchery programs. Draft study plan. 27 p.

HGMP. Draft Hatchery and Genetic Management Plans (2005) for Wenatchee, Methow, and Okanogan River summer Chinook. Available at Washington Department of Fish and Wildlife, 600 Capitol Way N, Olympia, WA, 98501.

Kassler, T.W., J.F. VonBargen, and D. Hawkins. 2008. DNA-based population of-Origin Assignments of Chinook Salmon Smolts Outmigrating Past Chandler Trap at Prosser Dam (Yakima River) in 2007. Final Report. Unpublished Washington Department of Fish and Wildlife Molecular Genetics Laboratory Report submitted to Bonneville Power Administration (BPA).

Kassler, T.W. and C.A. Dean. 2010. Genetic analysis of natural-origin spring Chinook and comparison to spring Chinook from an integrated supplementation program and captive broodstock program in the Tucannon River. Final Report. Unpublished Washington Department of Fish and Wildlife Molecular Genetics Laboratory Report submitted to Mike Gallinat, WDFW - Snake River Laboratory, Dayton, WA.

Jones, O. and J. Wang. 2009. COLONY: a program for parentage and sibship inference from multilocus genotype data. Molecular Ecology Resources 10: 551-555.

Lewis, P. O. and D. Zaykin. 2001. Genetic Data Analysis: Computer program for the analysis of allelic data. Version 1.0 (d16c). Free program distributed by the authors over the internet from http://lewis.eeb.uconn.edu/lewishome/software.html

Marshall, A.R., C. Smith, R. Brix, W. Dammers, J. Hymer, and L. LaVoy. 1995. Genetic diversity units and major ancestral lineages for chinook salmon in Washington. In C. Busack and J. B. Shaklee (eds.), Genetic diversity units and major ancestral lineages of salmonid fishes in Washington, p. 111173. Wash. Dep. Fish Wildl. Tech. Rep. RAD 95-02. (Available from Washington Department of Fish and Wildlife, 600 Capital Way N., Olympia WA 98501-1091.)

Marshall, A. 2002. 16 August memo to Ann Blakley (WDFW) and Amilee Wilson (WDFW) regarding genetic analyses of selected Washington Chinook stocks. WDFW. Olympia.

Mobrand, L. (chair), J. Barr, L. Blankenship, D. Campton, T. Evelyn, C. Mahnken, P. Seidel, L. Seeb and B. Smoker. 2004. Hatchery Scientific Review Group (HSRG) - March 2004. Hatchery Reform Recommendations for the Puget Sound and Coastal Washington Hatchery Reform Project. Long Live the Kings, 1305 Fourth Avenue, Suite 810, Seattle, WA 98101 (available from www.hatcheryreform.org).

Murdoch, A. and C. Peven. 2005. Conceptual approach for monitoring and evaluating the Chelan County Public Utility District hatchery programs. Chelan County Public Utility District, Wenatchee, Washington. 105 p.

Myers, J.M., R.G. Kope, G.J. Bryant, D. Teel, L.J. Lierheimer, T.C. Wainwright, W.S. Grand, F.W. Waknitz, K. Neely, S.T. Lindley, and R.S. Waples. 1998. Status review of chinook salmon from Washington, Idaho, Oregon, and California. U.S. Dept. Commer., NOAA Tech. Memo. NMFS-NWFSC35, 443 p.

Olsen, J.B., P. Bentzen, and J.E. Seeb. 1998. Characterization of seven microsatellite loci derived from pink salmon. Molecular Ecology 7(8):10871089.

Page, R.D.M. 1996. TREEVIEW: an application to display phylogenetic trees on personal computers. Computer Application Biosciences 12:351-358.

Raymond, M. and F. Rousset. 1995. GENEPOP (Version 3.3): Population genetics software for exact tests and ecumenicism. Journal of Heredity 86:248-249.

Rexroad, C.E., III, R.L. Coleman, A.M. Martin, W.K. Hershberger, and J. Killefer. 2001. Thirty-five polymorphic microsatellite markers for rainbow trout (Oncorhynchus mykiss). Animal Genetics 32:317-319.

Rice, W.R. 1989. Analyzing tables of statistical tests. Evolution 43:223-225.

Saitou, N. and M. Nei. 1987. The neighbor-joining method: A new method for reconstructing phylogenetic trees. Molecular Biology and Evolution 4:406425.

Seeb, L.W., A. Antonovich, M.A. Banks, et al. 2007. Development of a standardized DNA database for Chinook salmon. Fisheries 32:11.

Small, M.P., K.I. Warheit, C.A. Dean, and A.R. Murdoch. 2007. Methow spring Chinook genetic monitoring. Final Report. Unpublished Washington Department of Fish and Wildlife Molecular Genetics Laboratory Report.

Small, M.P., D. Burgess, C. Dean, and K. Warheit. 2010. Does Lower Crab Creek in the Eastern WA desert have a native population of Chinook salmon? Submitted to special edition of American Fisheries Society, Proceedings from the Coastwide Salmonid Genetics Meeting, Boise, ID.

Stuehrenberg, L.C., G.A. Swan, L.K. Timme, P.A. Ocker, M.B. Eppard, R.N. Iwamoto, B.L. Iverson, and B.P. Sanford. 1995. Migrational characteristics of adult spring, summer, and fall chinook salmon passing through reservoirs and dams of the mid-Columbia River. Final report. CZES Division, NWFSC, NMFS, Seattle, WA, 115 p.

Utter, F.M., D.W. Chapman, and A.R. Marshall. 1995. Genetic population structure and history of chinook salmon of the Upper Columbia River. American Fisheries Society Symposium 17:149-165.

Waknitz, F.W., G.M. Matthews, T. Wainwright, and G.A. Winans. 1995. Status review for Mid-Columbia River summer chinook salmon. NOAA Tech. Mem. NMFS-NWFSC-22, 80 p. (Available from Natl. Mar. Fish. Serv., Northwest Fisheries Science Center, Coastal Zone and Estuarine Studies Division, 2725 Montlake Blvd. E., Seattle, WA 98112-2097.)

Wang, J. 2009. A new method for estimating effective population sizes from a single sample of multilocus genotypes. Molecular Ecology 18:2148-2164

Wang, J. and A.W. Santure. 2009. Parentage and sibship inference from multilocus genotype data under polygamy. Genetics 181: 1579-1594.

Waples R.S. 1990. Conservation genetics of Pacific salmon. III. Estimating effective population size. Journal of Heredity 81:277-289

Waples R.S., M. Masuda, and J. Pella. 2007. SALMONNb: a program for computing cohort-specific effective population sizes ( $\mathrm{N}_{\mathrm{b}}$ ) in Pacific salmon and other semelparous species using the temporal method. Molecular Ecology Notes 7, 21-24

WDF (Washington Department of Fisheries) and WDW (Washington Department of Wildlife). 1993. 1992 Washington state salmon and steelhead stock inventory. Appendix Three. Columbia River stocks. WDF. Olympia, WA.

Weir, B.S. and C.C. Cockerham. 1984. Estimating F-statistics for the analysis of population structure. Evolution 38:1358-1370.

Williamson, K.S., J.F. Cordes, and B.P. May. 2002. Characterization of microsatellite loci in chinook salmon (Oncorhynchus tshawytscha) and cross-species amplification in other salmonids. Molecular Ecology Notes 2 (1):17-19.

Wright, S. 1969. Evolution and the Genetics of Populations, Vol. 2, The Theory of Gene Frequencies. The University of Chicago Press, Chicago, Illinois.

Table 1. Samples of adult hatchery- and natural-origin summer and fall Chinook that were analyzed from the upper Columbia River. Total number of individuals that were analyzed / individuals with data for 9 or more loci that were included in the analysis. Collection statistics (allelic richness, linkage disequilibrium (before and after Bonferroni correction), $\mathrm{F}_{\text {IS }}$, heterozygosity $\left(H_{O}\right.$ and $\left.H_{E}\right)$ ) and $p$-values for deviations from Hardy-Weinberg equilibrium (HWE). P-values were defined as significant after implementation of Bonferroni correction for multiple tests (Rice 1989).

| WDFW GSI code ${ }^{\text {a }}$ | Collection location | $\mathrm{N}=$ | Allelic Richness ${ }^{\text {b }}$ | Linkage Disequilibrium ${ }^{\text {c }}$ | $F_{\text {IS }}(\mathrm{p} \text {-value })^{\text {d }}$ | $\mathrm{H}_{\mathrm{O}}$ | $\mathrm{H}_{\mathrm{E}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 93DD | Wenatchee River upstream of Tumwater Dam - natural origin | $51 / 45$ |  |  |  |  |  |
| 93DE | Wenatchee River downstream of Tumwater Dam - natural origin | 88 / 88 |  |  |  |  |  |
| 06CQ | Wenatchee River upstream of Tumwater Dam - natural origin | 95 / 86 |  |  |  |  |  |
| 06CR | Wenatchee River downstream of Tumwater Dam - natural origin | 95 / 82 |  |  |  |  |  |
| 08FV | Wenatchee River upstream of Tumwater Dam - natural origin | 95 / 82 |  |  |  |  |  |
| 08FW | Wenatchee River downstream of Tumwater Dam - natural origin | 95 / 87 |  |  |  |  |  |
|  | Wenatchee River - Natural origin combined | 519/470 | 10.7 | 17 / 4 | 0.001 (0.403) | 0.8504 | 0.8513 |
|  |  |  |  |  |  |  |  |
| 06CP | Wenatchee River - hatchery origin | $95 / 70$ |  |  |  |  |  |
| 08FU | Wenatchee River - hatchery origin | 95 / 83 |  |  |  |  |  |
|  | Wenatchee River - Hatchery origin combined | 190 / 153 | 10.6 | 18 / 6 | 0.018 (0.013) | 0.8409 | 0.8561 |
|  |  |  |  |  |  |  |  |
| 93EC | Methow River - natural origin | 27 / 27 |  |  |  |  |  |
| 06CT | Methow River - natural origin | 95 / 90 |  |  |  |  |  |
| 08FY | Methow River - natural origin | 95 / 88 |  |  |  |  |  |
| 09CO | Methow River - natural origin | 91/80 |  |  |  |  |  |
|  | Methow River - Natural origin combined | 308/285 | 10.7 | 4 / 1 | 0.006 (0.160) | 0.8506 | 0.8554 |
|  |  |  |  |  |  |  |  |
| 06CS | Methow River - hatchery origin | 14 / 8 |  |  |  |  |  |
| 08FX | Methow River - hatchery origin | 21/18 |  |  |  |  |  |
| 09CP | Methow River - hatchery origin | $19 / 18$ |  |  |  |  |  |
|  | Methow River - Hatchery origin combined | 54 / 44 | 10.8 | 11 / 2 | -0.003 (0.593) | 0.8553 | 0.8523 |


| Table 1 continued. |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 92FM | Okanogan River - natural origin | 49 / 46 |  |  |  |  |  |
| 93ED* | Okanogan River - natural origin | 103 / 87 |  |  |  |  |  |
| 06CV | Okanogan River - natural origin | 95 / 88 |  |  |  |  |  |
| 08GA | Okanogan River - natural origin | 95 / 92 |  |  |  |  |  |
| 09CN | Okanogan River - natural origin | $133 / 126$ |  |  |  |  |  |
|  | Okanogan River - Natural origin combined | 475 / 439 | 10.8 | 9 / 4 | 0.003 (0.304) | 0.8563 | 0.8596 |
| * - not included in the combined dataset |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |
| 06CU | Okanogan River - hatchery origin | $58 / 49$ |  |  |  |  |  |
| 08FZ | Okanogan River - hatchery origin | 19 / 18 |  |  |  |  |  |
| 09CM | Okanogan River - hatchery origin | 117 / 107 |  |  |  |  |  |
|  | Okanogan River - hatchery origin combined | 194 / 174 | 10.8 | $31 / 10$ | -0.011 (0.920) | 0.8678 | 0.8586 |
|  |  |  |  |  |  |  |  |
| 91FL | Wells Hatchery | 68 / 42 |  |  |  |  |  |
| 92FK | Wells Hatchery | $25 / 23$ |  |  |  |  |  |
| 93DG | Wells Hatchery | 11/9 |  |  |  |  |  |
| 06DM | Wells Hatchery | 95 / 91 |  |  |  |  |  |
| 08HY | Wells Hatchery | 95 / 91 |  |  |  |  |  |
|  | Wells Hatchery combined | 294 / 256 | 10.7 | 8 / 3 | -0.001 (0.529) | 0.8670 | 0.8665 |
|  |  |  |  |  |  |  |  |
| 08MN | Eastbank Hatchery - Wenatchee River stock | 95 / 90 | 10.7 | 6 / 1 | 0.020 (0.024) | 0.8326 | 0.8498 |
|  |  |  |  |  |  |  |  |
| 92FO | Eastbank Hatchery - Methow / Okanogan (MEOK) stock | $36 / 33$ |  |  |  |  |  |
| 93DF | Eastbank Hatchery - Methow / Okanogan (MEOK) stock | 90 / 86 |  |  |  |  |  |
| 08MO | Eastbank Hatchery - Methow / Okanogan (MEOK) stock | 95 / 88 |  |  |  |  |  |
|  | Eastbank Hatchery - MEOK stock combined | 221 / 207 | 10.7 | $2 / 0$ | -0.005 (0.782) | 0.8647 | 0.8604 |
|  |  |  |  |  |  |  |  |
|  |  | 2,350 / 2,118 |  |  |  |  |  |


| Table 1 continued. |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 06KN | Chelan River | 70 / 23 | 10.3 | 11 / 0 | 0.027 (0.118) | 0.8334 | 0.8556 |
| Data provided by USFWS |  |  |  |  |  |  |  |
|  | Entiat River - summer Chinook | 190 | 10.9 | $33 / 10$ | 0.008 (0.119) | 0.8553 | 0.8625 |
| Data from Small et al. (2010) |  |  |  |  |  |  |  |
| 08EH | Crab Creek | 108 |  |  |  |  |  |
| 09AZ | Crab Creek | 291 |  |  |  |  |  |
|  | Crab Creek | 399 | 10.5 | $35 / 14$ | 0.018 (0.000) | 0.8519 | 0.8676 |
| GAPS v.3.0 data |  |  |  |  |  |  |  |
|  | Priest Rapids Hatchery - fall Chinook | 81 | 11.1 | $3 / 2$ | 0.015 (0.079) | 0.8591 | 0.8723 |
|  | Hanford Reach - fall Chinook | 220 | 11.3 | $4 / 0$ | 0.010 (0.068) | 0.8661 | 0.8746 |
|  | Umatilla - fall Chinook | 96 | 11.2 | 17 / 6 | -0.003 (0.623) | 0.8719 | 0.8693 |
|  | lower Yakima River - fall Chinook | 103 | 11.0 | $3 / 1$ | 0.000 (0.511) | 0.8724 | 0.8721 |
|  | Marion Drain - fall Chinook | 190 | 10.8 | $9 / 4$ | 0.022 (0.001) | 0.8586 | 0.8782 |
|  | Lyons Ferry Hatchery - fall Chinook | 186 | 10.6 | $7 / 4$ | 0.013 (0.033) | 0.8527 | 0.8641 |
|  | Snake River - fall Chinook | 521 | 11.1 | $0 / 0$ | -0.001 (0.634) | 0.8720 | 0.8708 |
|  |  | NA / 2,00 |  |  |  |  |  |
| ${ }^{\text {a }}$ - Year that samples were collected is identifed by the two numbers in the WDFW GSI code |  |  |  |  |  |  |  |
| ${ }^{\text {b }}$ - based on a minimum of 11 diploid individuals |  |  |  |  |  |  |  |
| ${ }^{\text {c }}$ - adjusted alpha $p$-value $=0.0006$ |  |  |  |  |  |  |  |
| ${ }^{\text {d }}$ - adjusted alpha $p$-value $=0.0002$ |  |  |  |  |  |  |  |

Table 2. PCR conditions and microsatellite locus information (number alleles/locus and allele size range) for multiplexed loci used for the analysis of Chinook. Also included are the observed and expected heterozygosity $\left(H_{0}\right.$ and $\left.H_{e}\right)$ for each locus.

| PCR Conditions |  |  | Locus statistics |  | Heterozygosity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Poolplex | Locus | Dye Label | \# Alleles/ Locus | Allele Size Range (bp) | $\mathrm{H}_{0}$ | $\mathrm{H}_{\mathrm{e}}$ | References |
| Ots-M | Ots-201b | blue | 49 | 137-334 | 0.9474 | 0.9544 | Unpublished |
|  | Ots-208b | yellow | 56 | 154-378 | 0.9523 | 0.9672 | Greig et al. 2003 |
|  | Ssa-408 | red | 32 | 184-308 | 0.9177 | 0.9214 | Cairney et al. 2000 |
| Ots-N | Ogo-2 | red | 22 | 206-260 | 0.8526 | 0.8673 | Olsen et al. 1998 |
| Ots-O | Ogo-4 | blue | 20 | 128-170 | 0.6694 | 0.7028 | Olsen et al. 1998 |
|  | Ots-213 | yellow | 45 | 178-370 | 0.9430 | 0.9525 | Greig et al. 2003 |
|  | Ots-G474 | red | 16 | 152-212 | 0.6816 | 0.6838 | Williamson et al. 2002 |
| Ots-R | Ots-3M | blue | 15 | 128-158 | 0.7854 | 0.7938 | Banks et al. 1999 |
|  | Omm-1080 | green | 54 | 162-374 | 0.9517 | 0.9670 | Rexroad et al. 2001 |
| Ots-S | Ots-9 | red | 9 | 99-115 | 0.6531 | 0.6543 | Banks et al. 1999 |
|  | Ots-212 | blue | 33 | 123-251 | 0.9205 | 0.9360 | Greig et al. 2003 |
| Ots-T | Oki-100 | blue | 50 | 164-361 | 0.9500 | 0.9567 | Unpublished |
|  | Ots-211 | red | 34 | 188-327 | 0.9325 | 0.9414 | Greig et al. 2003 |

Table 3. Tests of population differentiation for temporal collections of summer Chinook from natural and hatchery-origin populations in the upper Columbia River. P-values that are highlighted grey are significantly different after Bonferroni correction (Rice 1989). Adjusted alpha $p$-value was 0.0001 . The H and W in the collection identifier is for wild or hatchery-origin and the two digit number identifes the year samples were collected.

## Wenatchee River

|  | WenW93U | WenW93D | WenH06 | WenW06U | WenW06D | WenH08 | WenW08U WenW08D |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WenW93U | **** |  |  |  |  |  |  |  |  |
| WenW93D | 0.0162 | $* * * *$ |  |  |  |  |  |  |  |
| WenH06 | 0.0033 | 0.0102 | $* * * *$ |  |  |  |  |  |  |
| WenW06U | 0.3039 | 0.1642 | 0.4795 | $* * * *$ |  |  |  |  |  |
| WenW06D | 0.0261 | 0.0160 | 0.0678 | 0.5300 | $* * * *$ |  |  |  |  |
| WenH08 | 0.1126 | 0.0708 | 0.0073 | 0.4359 | 0.0893 | $* * * *$ |  |  |  |
| WenW08U | 0.2115 | 0.1148 | 0.4191 | 0.7243 | 0.3830 | 0.8856 | $* * *$ |  |  |
| WenW08D | 0.1915 | 0.0014 | 0.7047 | 0.4928 | 0.1671 | 0.7755 | 0.7665 | $* * * *$ |  |

D - collection was downstream of Tumwater Dam; U-collection was upstream of Tumwater Dam

| Methow River |  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MetW93 | MetH06 | MetW06 | MetH08 | MetW08 | MetW09 | MetH09 |  |  |
| MetW93 | **** |  |  |  |  |  |  |  |  |
| MetH06 | 0.3962 | $* * * *$ |  |  |  |  |  |  |  |
| MetW06 | 0.5481 | 0.4688 | $* * * *$ |  |  |  |  |  |  |
| MetH08 | 0.1408 | 0.1192 | 0.2052 | $* * * *$ |  |  |  |  |  |
| MetW08 | 0.8219 | 0.8937 | 0.6156 | 0.3779 | $* * * *$ |  |  |  |  |
| MetW09 | 0.2564 | 0.4282 | 0.2502 | 0.0328 | 0.7309 | $* * * *$ |  |  |  |
| MetH09 | 0.1543 | 0.5678 | 0.0547 | 0.0017 | 0.0098 | 0.0073 | $* * * *$ |  |  |
|  |  |  |  |  |  |  |  |  |  |
| Okanogan River |  |  |  |  |  |  |  |  |  |
|  | OkanW92 | OkanW93 | OkanH06 | OkanW06 | OkanH08 | OkanW08 | OkanH09 | OkanW09 |  |
| OkanW92 | $* * * *$ |  |  |  |  |  |  |  |  |
| OkanW93 | 0.0066 | $* * * *$ |  |  |  |  |  |  |  |
| OkanH06 | 0.0193 | 0.0000 | $* * * *$ |  |  |  |  |  |  |
| OkanW06 | 0.2843 | 0.0082 | 0.0031 | $* * * *$ |  |  |  |  |  |
| OkanH08 | 0.1290 | 0.1106 | 0.0652 | 0.7329 | $* * * *$ |  |  |  |  |
| OkanW08 | 0.0106 | 0.0029 | 0.0082 | 0.4075 | 0.7396 | $* * * *$ |  |  |  |
| OkanH09 | 0.0187 | 0.0001 | 0.0094 | 0.0551 | 0.2214 | 0.0281 | $* * * *$ |  |  |
| OkanW09 | 0.0527 | 0.0000 | 0.0024 | 0.7130 | 0.0262 | 0.0065 | 0.0002 | $* * * *$ |  |

Table 3 continued. $\qquad$

| Wells Dam Hatchery |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | Wells91 | Wells92 | Wells93 | Wells06 | Wells08 |
| Wells91 | $* * * *$ |  |  |  |  |
| Wells92 | 0.5863 | $* * * *$ |  |  |  |
| Wells93 | 0.0490 | 0.0784 | $* * * *$ |  |  |
| Wells06 | 0.0089 | 0.0100 | 0.0542 | $* * * *$ |  |
| Wells08 | 0.0819 | 0.1088 | 0.2552 | 0.0256 | $* * * *$ |
|  |  |  |  |  |  |

Eastbank Hatchery - Wenatchee and MEOK stocks

|  | EBHWen08 | EBHME92 | EBHME93 | EBHME08 |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| EBHWen08 | $* * * *$ |  |  |  |  |
| EBHME92 | 0.8681 | $* * * *$ |  |  |  |
| EBHME93 | 0.0251 | 0.8661 | $* * * *$ |  |  |
| EBHME08 | 0.0086 | 0.9563 | 0.1895 | $* * * *$ |  |

Table 4. $\mathrm{F}_{\text {ST }}$ pairwise comparisons and genotypic tests of differentiation for hatchery- and natural-origin summer Chinook from the upper Columbia River. Above the diagonol are the $F_{\text {ST }}$ values and below are $p$-values for the test of genotypic differentiation. Nonsignificant $p$-values for the result of the genotypic differentiation test are in bold type and $\mathrm{F}_{\mathrm{ST}}$ values that are not significantly different from zero are in bold type.

|  | Wenatchee Hatchery | Wenatchee Natural | Methow Hatchery | Methow Natural | Okanogan Hatchery | Okanogan Natural | Wells Hatchery | Eastbank Wenatchee stock | $\begin{gathered} \text { Eastbank } \\ \text { MEOK } \\ \text { stock } \\ \hline \end{gathered}$ | Entiat <br> River | Chelan River |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wenatchee Hatchery | **** | 0.0000 | 0.0011 | 0.0000 | 0.0013 | 0.0010 | 0.0015 | 0.0004 | 0.0007 | 0.0004 | 0.0072 |
| Wenatchee Natural | 0.4351 | **** | 0.0016 | 0.0000 | 0.0014 | 0.0016 | 0.0024 | 0.0006 | 0.0012 | 0.0009 | 0.0068 |
| Methow Hatchery | 0.3800 | 0.0205 | **** | 0.0012 | 0.0029 | 0.0008 | 0.0027 | 0.0014 | 0.0022 | 0.0019 | 0.0078 |
| Methow Natural | 0.2237 | 0.6566 | 0.1502 | **** | 0.0011 | 0.0011 | 0.0013 | 0.0007 | 0.0007 | 0.0008 | 0.0053 |
| Okanogan Hatchery | 0.0001 | 0.0000 | 0.0364 | 0.0008 | **** | 0.0010 | 0.0014 | 0.0029 | 0.0000 | 0.0007 | 0.0055 |
| Okanogan Natural | 0.0000 | 0.0000 | 0.1755 | 0.0000 | 0.0003 | **** | 0.0016 | 0.0023 | 0.0005 | 0.0008 | 0.0049 |
| Wells Hatchery | 0.0000 | 0.0000 | 0.0129 | 0.0000 | 0.0000 | 0.0000 | **** | 0.0036 | 0.0006 | 0.0008 | 0.0041 |
| Eastbank Wenatchee | 0.5261 | 0.4102 | 0.1215 | 0.8404 | 0.0015 | 0.0000 | 0.0000 | **** | 0.0018 | 0.0030 | 0.0096 |
| Eastbank MEOK stock | 0.0485 | 0.0000 | 0.4246 | 0.0009 | 0.5786 | 0.0051 | 0.0000 | 0.0065 | **** | 0.0005 | 0.0039 |
| Entiat River | 0.0565 | 0.0000 | 0.1795 | 0.0044 | 0.0005 | 0.0000 | 0.0032 | 0.0039 | 0.0042 | **** | 0.0052 |
| Chelan River | 0.0091 | 0.0026 | 0.0182 | 0.0156 | 0.0048 | 0.0030 | 0.0066 | 0.0059 | 0.0493 | 0.0617 | **** |

Table 5. $\mathrm{F}_{\text {ST }}$ pairwise comparisons and genotypic tests of differentiation for fall Chinook. Above the diagonol are the $\mathrm{F}_{\text {ST }}$ values and below are p-values for the test of genotypic differentiation. Non-significant $p$-values for the result of the genotypic differentiation test are in bold type and $F_{S T}$ values that are not significantly different from zero are in bold type.

|  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |

Table 6. $F_{S T}$ pairwise comparisons and genotypic tests of differentiation for hatchery- and natural-origin summer Chinook from the upper Columbia River and fall Chinook. Above the diagonol are the $F_{S T}$ values and below are $p$-values for the test of genotypic differentiation. Non-significant $p$-values for the result of the genotypic differentiation test are in bold type and $F_{S T}$ values that are not significantly different from zero are in bold type.

| Population Differentiation |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Wenatchee Hatchery | Wenatchee Natural | Methow Hatchery | Methow Natural | Okanogan Hatchery | Okanogan Natural | Wells Hatchery | Wenatchee stock | MEOK <br> stock | Entiat River | Chelan River |
| Crab Creek | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Hanford Reach Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0349 |
| Lyons Ferry Hatchery Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| lower Yakima River Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0074 |
| Marion Drain Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Priest Rapids Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0642 |
| Umatilla River Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0579 |
| Snake River Fall | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |


| Table 6 continued. |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pairwise $\mathrm{F}_{\text {ST }}$ |  |  |  |  |  |  |  |  |
|  | Crab Creek | Hanford Reach Fall | Ferry Hatchery | Yakima River | Marion Drain Fall | Priest Rapids Fall | Umatilla River Fall | Snake River Fall |
| Wenatchee Hatchery | 0.0158 | 0.0054 | 0.0180 | 0.0056 | 0.0153 | 0.0025 | 0.0053 | 0.0103 |
| Wenatchee Natural | 0.0162 | 0.0059 | 0.0185 | 0.0063 | 0.0157 | 0.0030 | 0.0059 | 0.0102 |
| Methow Hatchery | 0.0191 | 0.0104 | 0.0248 | 0.0095 | 0.0220 | 0.0069 | 0.0107 | 0.0165 |
| Methow Natural | 0.0148 | 0.0057 | 0.0182 | 0.0051 | 0.0148 | 0.0033 | 0.0055 | 0.0101 |
| Okanogan Hatchery | 0.0146 | 0.0041 | 0.0166 | 0.0042 | 0.0151 | 0.0016 | 0.0041 | 0.0082 |
| Okanogan Natural | 0.0163 | 0.0064 | 0.0187 | 0.0062 | 0.0170 | 0.0035 | 0.0068 | 0.0113 |
| Wells Hatchery | 0.0120 | 0.0051 | 0.0135 | 0.0044 | 0.0120 | 0.0028 | 0.0046 | 0.0077 |
| Wenatchee stock | 0.0184 | 0.0073 | 0.0203 | 0.0074 | 0.0167 | 0.0047 | 0.0084 | 0.0128 |
| Eastbank MEOK stock | 0.0128 | 0.0036 | 0.0143 | 0.0038 | 0.0135 | 0.0019 | 0.0038 | 0.0079 |
| Entiat River | 0.0147 | 0.0059 | 0.0176 | 0.0057 | 0.0156 | 0.0028 | 0.0056 | 0.0100 |
| Chelan River | 0.0074 | 0.0046 | 0.0110 | 0.0040 | 0.0160 | 0.0047 | 0.0035 | 0.0072 |

Table 7. Effective number of breeders per brood year with the largest number of samples of summer Chinook in the upper Columbia River. Brood years with sample size less than 19 individuals (shown in bold type) were not analyzed with exception of the 2008 Wells Hatchery collection. A comparison could not be made between an early and late collection from Wells Hatchery.

| WDFW <br> Code | Collection Location | Sample Size | $\mathrm{Nb}=$ | C195(L) = | CI95(U) = |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 93DD ${ }^{\text {A }}$ | Wenatchee Natural - upstream | 23/19 | 152 / 190 | 77 / 87 | 616 / 2,147,483,647 |
| 08FV | Wenatchee Natural - upstream | 56 | 162 | 112 | 249 |
| 93DE ${ }^{\text {A }}$ | Wenatchee Natural - downstream | $39 / 34$ | 145/152 | 94 / 95 | 256 / 302 |
| 08FW | Wenatchee Natural - downstream | 67 | 140 | 105 | 199 |
| 08FU | Wenatchee Hatchery | 60 | 134 | 90 | 213 |
| 93EC ${ }^{\text {A }}$ | Methow Natural | 10 / 15 | --- | --- | --- |
| 08FY | Methow Natural | 62 | 150 | 106 | 218 |
| 08FX | Methow Hatchery | 9 | --- | --- | --- |
| 93ED | Okanogan Natural | 69 | 142 | 102 | 203 |
| 08GA | Okanogan Natural | 59 | 127 | 92 | 180 |
| 08FZ | Okanogan Hatchery | 16 | --- | --- | --- |
| 93DG | Wells Hatchery | 6 | --- | --- | --- |
| $08 \mathrm{HY}{ }^{\text {B }}$ | Wells Hatchery | 24 / 39 | --- | --- | --- |
| 08MN | Eastbank Hatchery - Wenatchee | 88 | 190 | 144 | 263 |
| 93DF | Eastbank Hatchery - MEOK | 84 | 171 | 129 | 229 |
| 08MO | Eastbank Hatchery - MEOK | 88 | 166 | 126 | 226 |
| A - calculations were made for samples from brood year 1988 / brood year 1989 |  |  |  |  |  |
| B - samples were collected from brood year 2003 / brood year 2004 |  |  |  |  |  |



Figure 1. Relationship of natural- and hatchery-origin Chinook collections from the upper Columbia River basin using Cavalli-Sforza and Edwards (1967) chord distance. Bootstrap values are shown at each node.

Appendix $\mathbf{N}$

Summer Chinook Spawning Ground Surveys in the Methow River Basin and Chelan River, 2015


March 10, 2016

## To: Chelan and Grant Public Utility Districts

From: Denny Snyder, Keith Watson, and Mark Miller
Re: 2015 Summer Chinook spawning ground surveys in the Methow Basin and Chelan River.
The purpose of this memo is to provide information on the supplemented natural spawning population of summer Chinook in the Methow and Chelan River basins. This work is part of a larger effort focused on monitoring and evaluating Grant and Chelan PUDs' hatchery supplementation program. The tasks and objectives associated with implementing Grant and Chelan PUDs’ Hatchery M\&E Plan for 2015 are outlined in Hillman et al. (2013). Figures and tables are presented at the end of this memo. In 2015, The Okanogan Basin was surveyed by the Colville Confederated Tribes (CCT).

## METHODS

Spawning ground surveys were conducted by foot and raft beginning the last week of September and ending late-November. We did not use aerial surveys on the Methow River because past work has demonstrated that ground counts were more accurate than aerial surveys (Miller and Hillman 1997). Ground surveys were used to provide more accurate counts and a complete census of Chinook redds within their spawning distribution. Observers floated through sampling reaches and recorded the location and numbers of redds each week. Observers recorded the date, water temperature, river mile, and constructed a drawing of the area where redds were located. A different symbol was used each week to record the number of new and incomplete redds.
To maintain consistency, at least one observer surveyed the same stream reach on successive dates. In areas where numerous summer Chinook spawn, we constructed detailed maps of the river and used the cell-area-method (Hamilton and Bergersen 1984) to identify the number of redds within each cell. Cells were bound by noticeable landmarks along the banks (e.g., bridges or trees) or at stream habitat boundaries (e.g., transitions between pools and riffles). The number of redds were then recorded in the corresponding grid on the map. When possible, observers estimated the number of redds in a large disturbed area by counting females that defended redds. We assumed that the area or territory defended by a female was one redd.
Carcasses of summer Chinook were sampled to describe the spawning population. Biological data collection included: scale samples for age analysis, length measurements (POH and FKL), sex, egg voidance, marks, and PIT tag detection. These data will be used to assess length-at-age,
size-at-age, egg voidance, origin (hatchery or naturally produced), and stray rates. No DNA samples were collected on summer Chinook this year. We only report the escapement and number of redds for the Okanogan Basin.

## RESULTS

## Methow

There were 1,231 summer Chinook redds counted within seven reaches of the Methow River (Table 1). No redds were counted in the Chewuch and Twisp Rivers this year. This was the fifth highest redd count observed in the last 25 years for the Methow River (Table 3). Spawning began the last week of September, peaked in early October, and ended the third week of November (Figure 1). Spawning may have started the third week of September given the unusually large number of Chinook on spawning grounds. Stream temperatures in the Methow River when spawning began varied from $9.0-10.0^{\circ} \mathrm{C}$ in late September. Spawning peaked the last week of September in reaches M6 and M7, while peak spawning occurred in reaches M3-M5 the first week of October. Spawning peaked the second week of October in reaches M1and M2.

Most redds (78\%) were located in reaches from the mouth to the town of Twisp (M1-M3) (Table 1). In 2015, reach M1 experienced a dramatic increase in spawning with 350 redds compared to 9 redds observed in 2014. This increase is most likely because the fine sediments that covered spawning areas in 2014, as a result of the Carlton Complex Fires and landslides, were flushed from the system during high spring flows in 2015. Estimated escapement based on expansion of redd counts from the sex-ratio observed at Wells Dam during broodstock collection indicates that 3,952 summer Chinook ( 1,231 redds x 3.21 fish/redd) escaped to the Methow River.

There were 839 summer Chinook salmon carcasses sampled within the seven reaches of the Methow River (Table 2). The presence or absence of an adipose fin could not be determined on one fish. Twenty-one percent of the fish returning to the Methow River were sampled based on the estimated escapement of 3,952 summer Chinook. Ad-clipped hatchery fish made up $19 \%$ and naturally produced fish (adipose fin present) made up $81 \%$ of the fish sampled (Table 2). Most ( $94 \%$ ) of the ad-clipped hatchery fish were located in reaches M1-M3, while naturally produced fish were more evenly distributed among survey reaches (Figure 2). Naturally produced fish made up $100 \%$ of the fish sampled in upper reaches (M6 and M7). Females made up 49\% of the carcasses examined. Based on sampling 413 female carcasses, average egg voidance was $99 \%$. Seven females ( $2 \%$ ) died before spawning (i.e., they retained all their eggs).

## Chelan River

There were 448 redds counted in the Chelan River. Spawning activity began the first week of October and peaked two weeks later (Figure 3). Spawning continued into the last week of November. As more information is collected on time of spawning, the average spawn time will likely not appear bimodal. The majority of spawning occurred in the Powerhouse tailrace (48\%), Columbia River tailrace (24\%), and in the Habitat channel (20\%) (Table 1). Estimated escapement based on expansion of redd counts from the sex-ratio observed at Wells Dam during broodstock collection indicates that 1,438 summer Chinook (448 redds x 3.21 fish/redd) escaped to the Chelan River.

There were 363 summer Chinook carcasses sampled in the Chelan River (Table 2). Twenty-five percent of the summer Chinook returning to the Chelan River were sampled based on the estimated escapement of 1,438 fish. Ad-clipped hatchery fish made up $47 \%$ and naturally produced fish were $53 \%$ of the fish examined. The distribution of ad-clipped hatchery fish and naturally produced fish was similar, except in the pool upstream of the habitat channel where only hatchery fish were recovered (Figure 4). A disproportionate number of fish (compared to redds counts) were sampled in the Columbia River tailrace, because carcasses drifted from upstream spawning areas and settled in the Columbia River tailrace. Females made up $77 \%$ of the carcasses examined (Table 2). Mean egg voidance from 281 female carcasses was $84 \%$. Twenty females ( $7 \%$ ) died before spawning. Five Coho were sampled within the Chelan River and these data were submitted to the Yakima Nation (Peshastin Office).

## Okanogan Basin

In 2015, CCT conducted summer Chinook surveys in the Okanogan Basin. A total of 4,128 redds were counted in the Okanogan Basin. Based on expanded redd counts, the estimated escapement for the Okanogan basin was 13,272 summer Chinook (Personal Communication, Andrea Pearl, CCT).

## REFERENCES

Hamilton, K. and E. P. Bergersen. 1984. Methods to estimate aquatic habitat variables. Report for Bureau of Reclamation, Division of Planning and Technical Services, Denver, Colorado. Colorado Cooperative Fishery Research Unit, Colorado State University, Fort Collins, CO.

Hillman, T., T. Kahler, G. Mackey, J. Murauskas, A. Murdoch, K. Murdoch, T. Pearsons, and M. Tonseth. 2013. Updated monitoring and evaluation plan for PUD hatchery programs. Report to the Hatchery Committees, Wenatchee, East Wenatchee, and Ephrata, WA.

Hillman, T., M. Miller, M. Johnson, C. Moran, M. Tonseth, A. Murdoch, C. Willard, L. Keller, B. Ishida, C. Kamphaus, T. Pearsons, and P. Graf. 2015. Monitoring and evaluation of the Chelan and Grant County PUDs hatchery programs: 2014 annual report. Report to the HCP and PRCC Hatchery Committees, Wenatchee and Ephrata, WA.

Miller, M. D. and T. W. Hillman. 1997. Summer/fall Chinook salmon spawning ground surveys in the Methow and Okanogan river basins, 1997. Report to Chelan County PUD. Don Chapman Consultants, Inc. Boise, ID.


Figure 1. Number of new redds counted each week from late September to mid-November. The figure displays the beginning, peak and end of spawning for summer Chinook in the Methow River in 2015 compared to a 24 -year average (1991-2014).

## Carcass Distribution <br> Methow River



Figure 2. Percent distribution of ad-clipped hatchery and naturally produced fish plotted against the percent distribution of redds observed in reaches on the Methow River, 2015.

Time of Spawning


Figure 3. Number of new redds counted each week from late September to mid-November. The figure displays the beginning, peak and end of spawning for summer Chinook in the Chelan River in 2015 compared to a 3-year average (2012-2014).

## Carcass Distribution <br> Chelan River



Figure 4. Percent distribution of ad-clipped hatchery and naturally produced fish plotted against the percent distribution of redds observed in reaches on the Chelan River, 2015.

Table 1. Number of summer Chinook redds observed each week within the Methow and Chelan rivers, 2015. Dashes indicate that no survey occurred.

| Reach | Location (Rkm) | Sep |  | Oct |  |  |  |  |  |  |  |  | Total | Percent |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 27-3 | $\begin{gathered} \hline 4-10 \\ \hline 41 \end{gathered}$ | $\begin{gathered} 11-17 \\ \hline 42 \\ \hline \end{gathered}$ | $\begin{array}{\|c\|} \hline 18-24 \\ \hline 43 \\ \hline \end{array}$ | $\frac{25-31}{44}$ | $\frac{1-7}{45}$ | $\begin{gathered} \hline 8-14 \\ \hline 46 \end{gathered}$ | $\begin{array}{\|c\|} \hline 15-21 \\ \hline 47 \\ \hline \end{array}$ | $\begin{gathered} 22-28 \\ \hline 48 \end{gathered}$ | $\begin{array}{\|c\|} \hline 29-5 \\ \hline 49 \end{array}$ |  |  |
|  |  | 39 | 40 |  |  |  |  |  |  |  |  |  |  |  |
| Methow River |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| M1 | 0.0-23.8 | -- | 0 | 85 | 108 | 37 | 85 | 21 | 13 | 1 | -- | -- | 350 | 28 |
| M2 | 23.8-43.8 | -- | 30 | 120 | 124 | 23 | 8 | 4 | -- | -- | -- | -- | 309 | 25 |
| M3 | 43.8-63.7 | -- | 42 | 150 | 104 | 11 | 0 | 0 | -- | -- | -- | -- | 307 | 25 |
| M4 | 63.7-72.3 | -- | 9 | 41 | 20 | 2 | 0 | 0 | -- | -- | -- | -- | 72 | 6 |
| M5 | 72.3-80.1 | -- | 31 | 74 | 32 | 9 | 0 | -- | -- | -- | -- | -- | 146 | 12 |
| M6 | 80.1-83.0 | -- | 11 | 0 | 2 | 0 | 0 | -- | -- | -- | -- | -- | 13 | 1 |
| M7 | 83.0-96.1 | -- | 21 | 10 | 3 | 0 | -- | -- | -- | -- | -- | -- | 34 | 3 |
| Total: |  | -- | 144 | 480 | 393 | 82 | 93 | 25 | 13 | 1 | -- | -- | 1,231 | 100 |
| Chelan River |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Powerhouse Tailrace |  | -- | 0 | 2 | 21 | 98 | 48 | 25 | 18 | 1 | 4 | 0 | 217 | 48 |
| Columbia R. Tailrace |  | -- | 0 | 1 | 7 | 25 | 32 | 37 | 3 | 1 | 0 | 0 | 106 | 24 |
| Pool |  | -- | 0 | 0 | 13 | 15 | 3 | 3 | 0 | 0 | 0 | 0 | 34 | 8 |
| Habitat Channel |  | -- | 0 | 2 | 19 | 39 | 9 | 18 | 3 | 0 | 1 | 0 | 91 | 20 |
| Total: |  | -- | 0 | 5 | 60 | 177 | 92 | 83 | 24 | 2 | 5 | 0 | 448 | 100 |

Table 2. Number and percent of hatchery (ad-clipped) and naturally produced (not ad-clipped) summer Chinook collected in Methow and Chelan rivers, 2015. The origin of three fish sampled could not be determined in the Methow River.

| Reach | Location (Rkm) | Ad-Clipped Hatchery |  |  |  | Naturally Produced |  |  |  | Reach Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Male | Female | Total | Percent | Male | Female | Total | Percent |  |
| Methow River |  |  |  |  |  |  |  |  |  |  |
| M1 | 0.0-23.8 | 29 | 31 | 60 | 26 | 71 | 97 | 168 | 74 | $229{ }^{1}$ |
| M2 | 23.8-43.8 | 43 | 12 | 55 | 28 | 78 | 61 | 139 | 72 | 194 |
| M3 | 43.8-63.7 | 22 | 16 | 38 | 17 | 98 | 85 | 183 | 83 | 221 |
| M4 | 63.7-72.3 | 4 | 0 | 4 | 7 | 20 | 32 | 52 | 93 | 56 |
| M5 | 72.3-80.1 | 3 | 2 | 5 | 5 | 38 | 52 | 90 | 95 | 95 |
| M6 | 80.1-83.0 | 0 | 0 | 0 | 0 | 9 | 10 | 19 | 100 | 19 |
| M7 | 83.0-96.1 | 0 | 0 | 0 | 0 | 11 | 14 | 25 | 100 | 25 |
| Total |  | 101 | 61 | 162 | 19 | 325 | 351 | 676 | 81 | 839 |
| Chelan River |  |  |  |  |  |  |  |  |  |  |
| Powerho | se Tailrace | 2 | 15 | 17 | 35 | 3 | 29 | 32 | 65 | 49 |
| Columbi | R. Tailrace | 24 | 89 | 113 | 44 | 40 | 102 | 142 | 56 | 255 |
|  | ool | 5 | 12 | 17 | 94 | 1 | 0 | 1 | 6 | 18 |
| Habit | Channel | 3 | 20 | 23 | 56 | 4 | 14 | 18 | 44 | 41 |
|  | tal | 34 | 136 | 170 | 47 | 48 | 145 | 193 | 53 | 363 |

${ }^{1 .}$ Origin of one female carcass in Reach 1 could not assigned.

Table 3. Historical aerial and ground redd counts of summer Chinook in the Methow, Chelan, Okanogan, and Similkameen rivers, 1956-2015.

| Year | Methow |  | Okanogan |  | Similkameen |  | Chelan |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Aerial | Ground | Aerial | Ground | Aerial | Ground | Aerial | Ground |
| 1956 | 109 | -- | 37 | -- | 30 | -- | -- | -- |
| 1957 | 451 | -- | 53 | -- | 30 | -- | -- | -- |
| 1958 | 335 | -- | 94 | -- | 31 | -- | -- | -- |
| 1959 | 130 | -- | 50 | -- | 23 | -- | -- | -- |
| 1960 | 194 | -- | 29 | -- | -- | -- | -- | -- |
| 1961 | 120 | -- | -- | -- | -- | -- | -- | -- |
| 1962 | 678 | -- | -- | -- | 17 | -- | -- | -- |
| 1963 | 298 | -- | 9 | -- | 51 | -- | -- | -- |
| 1964 | 795 | -- | 112 | -- | 67 | -- | -- | -- |
| 1965 | 562 | -- | 109 | -- | 154 | -- | -- | -- |
| 1966 | 1,275 | -- | 389 | -- | 77 | -- | -- | -- |
| 1967 | 733 | -- | 149 | -- | 107 | -- | -- | -- |
| 1968 | 659 | -- | 232 | -- | 83 | -- | -- | -- |
| 1969 | 329 | -- | 103 | -- | 357 | -- | -- | -- |
| 1970 | 705 | -- | 656 | -- | 210 | -- | -- | -- |
| 1971 | 562 | -- | 310 | -- | 55 | -- | -- | -- |
| 1972 | 325 | -- | 182 | -- | 64 | -- | -- | -- |
| 1973 | 366 | -- | 138 | -- | 130 | -- | -- | -- |
| 1974 | 223 | -- | 112 | -- | 201 | -- | -- | -- |
| 1975 | 432 | -- | 273 | -- | 184 | -- | -- | -- |
| 1976 | 191 | -- | 107 | -- | 139 | -- | -- | -- |
| 1977 | 365 | -- | 276 | -- | 268 | -- | -- | -- |
| 1978 | 507 | -- | 195 | -- | 268 | -- | -- | -- |
| 1979 | 622 | -- | 173 | -- | 138 | -- | -- | -- |
| 1980 | 345 | -- | 118 | -- | 172 | -- | -- | -- |
| 1981 | 195 | -- | 55 | -- | 121 | -- | -- | -- |
| 1982 | 142 | -- | 23 | -- | 56 | -- | -- | -- |
| 1983 | 65 | -- | 36 | -- | 57 | -- | -- | -- |
| 1984 | 162 | -- | 235 | -- | 301 | -- | -- | -- |
| 1985 | 164 | -- | 138 | -- | 309 | -- | -- | -- |
| 1986 | 169 | -- | 197 | -- | 300 | -- | -- | -- |
| 1987 | 211 | -- | 201 | -- | 164 | -- | -- | -- |
| 1988 | 123 | -- | 113 | -- | 191 | -- | -- | -- |
| 1989 | 126 | -- | 134 | -- | 221 | 370 | -- | -- |
| 1990 | 229 | -- | 88 | 47 | 94 | 147 | -- | -- |
| 1991 | -- | 153 | 55 | 64 | 68 | 91 | -- | -- |
| 1992 | -- | 107 | 35 | 53 | 48 | 57 | -- | -- |
| 1993 | -- | 154 | 144 | 162 | 152 | 288 | -- | -- |
| 1994 | -- | 310 | 372 | 375 | 463 | 777 | -- | -- |
| 1995 | -- | 357 | 260 | 267 | 337 | 616 | -- | -- |


| Year | Methow |  | Okanogan |  | Similkameen |  | Chelan |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Aerial | Ground | Aerial | Ground | Aerial | Ground | Aerial | Ground |
| 1996 | -- | 181 | 100 | 116 | 252 | 419 | -- | -- |
| 1997 | -- | 205 | 149 | 158 | 297 | 486 | -- | -- |
| 1998 | -- | 225 | 75 | 88 | 238 | 276 | -- | -- |
| 1999 | -- | 448 | 222 | 369 | 903 | 1,275 | -- | -- |
| 2000 | -- | 500 | 384 | 549 | 549 | 993 | -- | 196 |
| 2001 | -- | 675 | 883 | 1,108 | 865 | 1,540 | -- | 240 |
| 2002 | -- | 2,013 | 1,958 | 2,667 | 2,000 | 3,358 | -- | 253 |
| 2003 | -- | 1,624 | 1,099 | 1,035 | 103 | 378 | -- | 173 |
| 2004 | -- | 973 | 1,310 | 1,327 | 2,127 | 1,660 | -- | 185 |
| 2005 | -- | 874 | 1,084 | 1,611 | 1,111 | 1,423 | -- | 179 |
| 2006 | -- | 1,353 | 1,857 | 2,592 | 1,337 | 1,666 | -- | 208 |
| 2007 | -- | 620 | 1,265 | 1,301 | 523 | 707 | -- | 86 |
| 2008 | -- | 599 | 1,019 | 1,146 | 673 | 1,000 | -- | 153 |
| 2009 | -- | 692 | 1,109 | 1,672 | 907 | 1,298 | -- | 246 |
| 2010 | -- | 887 | 688 | 1,011 | 642 | 1,107 | -- | 398 |
| 2011 | -- | 941 | 1,203 | 1,714 | 1,047 | 1,409 | -- | 413 |
| 2012 | -- | 960 | 1,170 | 1,613 | 762 | 1,066 | -- | 426 |
| 2013 | -- | 1,551 | NA | 2,267 | NA | 1,280 | -- | 729 |
| 2014 | -- | 591 | NA | 2,231 | NA | 2,022 | -- | 400 |
| 2015 | -- | 1,231 | NA | $4,276^{1}$ | NA | -- | -- | 448 |

${ }^{1 .}$ The redd count here is for the entire Okanogan Basin (Similkameen + Okanogan rivers).


[^0]:    ${ }^{1}$ Unnamed tributary that drains the eastside of Chiwawa Ridge. Its confluence with the Chiwawa River is about 1 mile ( 1.6 km ) downstream from the mouth of Phelps Creek.

[^1]:    ${ }^{2}$ The study period 1992-2015 includes only 23 years of sampling because there was no sampling in 2000.

[^2]:    ${ }^{3}$ The habitat use index was calculated as follows: Multiple channel use $=\left(\operatorname{parr}_{m c} / \operatorname{parr}_{t}\right) /\left(\operatorname{area}_{m c} /\right.$ area $\left._{t}\right)$, where parr $m c$ $=$ the number of parr counted in multiple channel habitat, $\operatorname{parr}_{t}=$ the total number of parr counted within all habitat types, $\operatorname{area}_{m c}=$ the area of multiple channel habitat within the sampling frame, and area ${ }_{t}=$ the total area of the sampling frame. A multiple channel use index value of 1 would indicate that parr were uniformly distributed among habitat types and exhibited no preference for multiple habitat types. Values of the use index greater than 1 indicate use of multiple channels to a greater extent than the average, while scores between 0 and 1 indicate below-average use of multiple channel habitat.

[^3]:    ${ }^{4}$ The $\gamma$ parameter in the Gamma model was greater than 0 , which means that this model is nearly identical to the Ricker model.

[^4]:    ${ }^{5}$ Because there are no estimates for probability of detecting bull trout with daytime underwater observation methods in the Chiwawa River basin, we could not adjust bull trout numbers based on detectability. Therefore, the numbers reported in this report likely underestimate the "true" number of bull trout in the survey area.

[^5]:    ${ }^{1}$ Includes the lower 0.2 miles of Minnow Creek.

[^6]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^7]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^8]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^9]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^10]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^11]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^12]:    ${ }^{1}$ Includes lower 0.2 miles of Minnow Creek.

[^13]:    ** PA pond - No discharge. PA pond system down during hatchery rebuild.

[^14]:    ${ }^{1}$ Defined as "above Wells Dam" because hatchery origin, adipose-clipped steelhead release into the Methow and Okanogan rivers from the Wells FH and Winthrop NFH have the same marks and are indistinguishable from one another.

[^15]:    ${ }^{1}$ Also includes fish detected downstream of release point (fallbacks).
    ${ }^{2}$ Detection efficiency $p_{\text {all }}=0.406$ in 2009 was assigned from 2010 data.
    ${ }^{3}$ Technical difficulties with the White River PIT array prevented the calculation of detection efficiency and a markrecapture based escapement estimate.
    ${ }^{4}$ In 2015, 45 sockeye salmon were detected in Chiwaukum Creek.

[^16]:    ${ }^{1}$ Samples taken from scale cards provided by Jeff Fryer (CRITFC)

[^17]:    ${ }^{\text {a }}$ Does not include subyearling fry prior to July 1.
    ${ }^{\mathrm{b}}$ 10-year average of complete brood data, BY2004-2013.

[^18]:    *Majority occurring during incident detailed in Appendix D.

[^19]:    ${ }^{\text {a }}$ Number of complete redds in White River (Hillman et al. 2015)
    ${ }^{\mathrm{b}}$ Mean annual fecundity of spring Chinook broodstock at Chiwawa River Hatchery
    ${ }^{\text {c }}$ Estimate is based on capture of parr collected during summer/fall and does not include fry captured prior to July1.
    ${ }^{\mathrm{d}}$ Did not operate trap; no production estimates were made.

